Semiparametric robust mean estimations based on the orderliness of quantile averages

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This manuscript was compiled on June 11, 2023

semiparametric | mean-median-mode inequality | asymptotic | unimodal | Hodges—Lehmann estimator

Inequalities related to weighted averages

So far, it is quite natural to hypothesize that the value of ϵ, γ -trimmed mean should be monotonically related to the breakdown point in a semiparametric distribution, since it is a linear combination of quantile averages as shown in Section ??. Analogous to the γ -orderliness, the γ -trimming inequality for a right-skewed distribution is defined as $\forall 0 \leq \epsilon_1 \leq \epsilon_2 \leq \frac{1}{1+\gamma}$, $TM_{\epsilon_1,\gamma} \geq TM_{\epsilon_2,\gamma}$. γ -orderliness is a sufficient condition for the γ -trimming inequality, as proven in the SI Text. The next theorem shows a relation between the ϵ, γ -quantile average and the ϵ, γ -trimmed mean under the γ -trimming inequality, suggesting the γ -orderliness is not a necessary condition for the γ -trimming inequality.

Theorem .1. For a distribution that is right-skewed and follows the γ -trimming inequality, it is asymptotically true that the quantile average is always greater or equal to the corresponding trimmed mean with the same ϵ and γ , $0 \le \epsilon \le \frac{1}{1+\gamma}$.

Proof. According to the definition of the γ -trimming inequality: $\forall 0 \leq \epsilon \leq \frac{1}{1+\gamma}, \ \frac{1}{1-\epsilon-\gamma\epsilon} \int_{\gamma\epsilon-\delta}^{1-\epsilon+\delta} Q\left(u\right) du \geq \frac{1}{1-\epsilon-\gamma\epsilon} \int_{\gamma\epsilon}^{1-\epsilon} Q\left(u\right) du, \text{ where } \delta \text{ is an infinitesimal positive quantity.}$ Subsequently, rewriting the inequality gives $\int_{\gamma\epsilon-\delta}^{1-\epsilon+\delta} Q\left(u\right) du - \frac{1-\epsilon-\gamma\epsilon+2\delta}{1-\epsilon-\gamma\epsilon} \int_{\gamma\epsilon}^{1-\epsilon} Q\left(u\right) du \geq 0 \Leftrightarrow \int_{1-\epsilon}^{1-\epsilon+\delta} Q\left(u\right) du + \int_{\gamma\epsilon-\delta}^{\gamma\epsilon} Q\left(u\right) du - \frac{2\delta}{1-\epsilon-\gamma\epsilon} \int_{\gamma\epsilon}^{1-\epsilon} Q\left(u\right) du \geq 0 \Leftrightarrow 0.$ Since $\delta \to 0^+, \ \frac{1}{2\delta} \left(\int_{1-\epsilon}^{1-\epsilon+\delta} Q\left(u\right) du + \int_{\gamma\epsilon-\delta}^{\gamma\epsilon} Q\left(u\right) du\right) = \frac{Q(\gamma\epsilon) + Q(1-\epsilon)}{2} \geq \frac{1}{1-\epsilon-\gamma\epsilon} \int_{\gamma\epsilon}^{1-\epsilon} Q\left(u\right) du, \text{ the proof is complete.}$

An analogous result about the relation between the ϵ, γ -trimmed mean and the ϵ, γ -Winsorized mean can be obtained in the following theorem.

Theorem .2. For a right-skewed distribution following the γ -trimming inequality, asymptotically, the Winsorized mean is always greater or equal to the corresponding trimmed mean with the same ϵ and γ , provided that $0 \le \gamma \le 1$. If assuming γ -orderliness, the inequality is valid for any non-negative γ .

$$\begin{array}{lll} \text{36} & \textit{Proof.} \text{ According} & \text{to} & \text{Theorem} & .1, & \frac{Q(\gamma\epsilon)+Q(1-\epsilon)}{2} & \geq \\ \text{37} & \frac{1}{1-\epsilon-\gamma\epsilon} \int_{\gamma\epsilon}^{1-\epsilon} Q\left(u\right) du & \Leftrightarrow & \gamma\epsilon\left(Q\left(\gamma\epsilon\right)+Q\left(1-\epsilon\right)\right) & \geq \\ \text{38} & \left(\frac{2\gamma\epsilon}{1-\epsilon-\gamma\epsilon}\right) \int_{\gamma\epsilon}^{1-\epsilon} Q\left(u\right) du. & \text{Then,} & \text{if} & 0 & \leq & \gamma & \leq \\ \text{39} & 1, \left(1-\frac{1}{1-\epsilon-\gamma\epsilon}\right) \int_{\gamma\epsilon}^{1-\epsilon} Q\left(u\right) du & + & \gamma\epsilon\left(Q\left(\gamma\epsilon\right)+Q\left(1-\epsilon\right)\right) & \geq \\ \text{40} & 0 \Rightarrow \int_{\gamma\epsilon}^{1-\epsilon} Q\left(u\right) du + \gamma\epsilon Q\left(\gamma\epsilon\right) + \epsilon Q\left(1-\epsilon\right) \geq \int_{\gamma\epsilon}^{1-\epsilon} Q\left(u\right) du + \\ \text{41} & \gamma\epsilon\left(Q\left(\gamma\epsilon\right)+Q\left(1-\epsilon\right)\right) & \geq & \frac{1}{1-\epsilon-\gamma\epsilon} \int_{\gamma\epsilon}^{1-\epsilon} Q\left(u\right) du, & \text{the proof} \end{array}$$

of the first assertion is complete. The second assertion is established in Theorem 0.3. in the SI Text. $\hfill\Box$

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Replacing the TM in the γ -trimming inequality with WA forms the definition of the γ -weighted inequality. The γ -orderliness also implies the γ -Winsorization inequality when $0 \le \gamma \le 1$, as proven in the SI Text. To construct weighted averages based on the ν th γ -orderliness and satisfying the corresponding weighted inequality, when $0 \le \gamma \le 1$, let $\mathcal{B}_i = \int_{i\epsilon}^{(i+1)\epsilon} \mathrm{QA}\left(u,\gamma\right) du$, $ka = k\epsilon + c$. From the γ -orderliness, it follows that, $-\frac{\partial \mathrm{QA}}{\partial \epsilon} \ge 0 \Leftrightarrow \forall 0 \le a \le 2a \le \frac{1}{1+\gamma}, -\frac{(\mathrm{QA}(2a,\gamma)-\mathrm{QA}(a,\gamma))}{a} \ge 0 \Rightarrow \mathcal{B}_i - \mathcal{B}_{i+1} \ge 0$, if $0 \le \gamma \le 1$. Suppose that $\mathcal{B}_i = \mathcal{B}_0$. Then, the ϵ,γ -block Winsorized mean, is defined as

$$\mathrm{BWM}_{\epsilon,\gamma,n} := \frac{1}{n} \left(\sum_{i=n\gamma\epsilon+1}^{(1-\epsilon)n} X_i + \sum_{i=n\gamma\epsilon+1}^{2n\gamma\epsilon+1} X_i + \sum_{i=(1-2\epsilon)n}^{(1-\epsilon)n} X_i \right),$$

which is double weighting the leftest and rightest blocks having sizes of $\gamma \epsilon n$ and ϵn , respectively. As a consequence of $\mathcal{B}_i - \mathcal{B}_{i+1} \geq 0$, the γ -block Winsorization inequality is valid, provided that $0 \leq \gamma \leq 1$. The block Winsorized mean uses two blocks to replace the trimmed parts, not two single quantiles. The subsequent theorem provides an explanation for this difference.

Theorem .3. Asymptotically, for a right-skewed γ -ordered distribution, the Winsorized mean is always greater than or equal to the corresponding block Winsorized mean with the same ϵ and γ , if $0 \le \gamma \le 1$.

Proof. From the definitions of BWM and WM, the statement necessitates $\int_{\gamma_{\epsilon}}^{1-\epsilon}Q\left(u\right)du+\gamma\epsilon Q\left(\gamma\epsilon\right)+\epsilon Q\left(1-\epsilon\right)\geq \int_{\gamma_{\epsilon}}^{1-\epsilon}Q\left(u\right)du+\int_{\gamma_{\epsilon}}^{2\gamma\epsilon}Q\left(u\right)du+\int_{1-2\epsilon}^{1-\epsilon}Q\left(u\right)du\Leftrightarrow \gamma\epsilon Q\left(\gamma\epsilon\right)+\epsilon Q\left(1-\epsilon\right)\geq \int_{\gamma_{\epsilon}}^{2\gamma\epsilon}Q\left(u\right)du+\int_{1-2\epsilon}^{1-\epsilon}Q\left(u\right)du.$ Define WMl(x) = $Q\left(\gamma\epsilon\right)$ and BWMl(x) = $Q\left(x\right)$. In both functions, the interval for x is specified as $[\gamma\epsilon,2\gamma\epsilon]$. Then, define WMu(y) = $Q\left(1-\epsilon\right)$ and BWMu(y) = $Q\left(y\right)$. In both functions, the interval for y is specified as $[1-2\epsilon,1-\epsilon]$. The functions, the interval for y is specified as $[1-2\epsilon,1-\epsilon]$. The function $y:[\gamma\epsilon,2\gamma\epsilon]\to[1-2\epsilon,1-\epsilon]$ defined by $y(x)=1-\frac{x}{\gamma}$ is a bijection. WMl(x) + WMu(y(x)) = $Q\left(\gamma\epsilon\right)+Q\left(1-\epsilon\right)\geq \text{BWMl}(x)+\text{BWMu}(y(x))=Q\left(x\right)+Q\left(1-\frac{x}{\gamma}\right)$ is valid for all $x\in[\gamma\epsilon,2\gamma\epsilon]$, according to the definition of γ -orderliness. Integration of the left side yields, $\int_{\gamma\epsilon}^{2\gamma\epsilon}\left(\text{WMl}\left(u\right)+\text{WMu}\left(y\left(u\right)\right)\right)du=\int_{\gamma\epsilon}^{2\gamma\epsilon}Q\left(\gamma\epsilon\right)du+\int_{y(\gamma\epsilon)}^{y(2\gamma\epsilon)}Q\left(1-\epsilon\right)du=\int_{\gamma\epsilon}^{2\gamma\epsilon}Q\left(\gamma\epsilon\right)du+\int_{1-2\epsilon}^{1-\epsilon}Q\left(1-\epsilon\right)du=\gamma\epsilon Q\left(\gamma\epsilon\right)+\epsilon Q\left(1-\epsilon\right),$ while integration of the right side

T.L. designed research, performed research, analyzed data, and wrote the paper.

The author declares no competing interest.

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yields $\int_{\gamma_{\epsilon}}^{2\gamma\epsilon} \left(\operatorname{BWM}l\left(x\right) + \operatorname{BWM}u\left(y\left(x\right)\right) \right) dx = \int_{\gamma_{\epsilon}}^{2\gamma\epsilon} Q\left(u\right) du + \int_{\gamma_{\epsilon}}^{2\gamma\epsilon} Q\left(1 - \frac{x}{\gamma}\right) dx = \int_{\gamma_{\epsilon}}^{2\gamma\epsilon} Q\left(u\right) du + \int_{1-2\epsilon}^{1-\epsilon} Q\left(u\right) du$, which are the left and right sides of the desired inequality. Given that the upper limits and lower limits of the integrations are different for each term, the condition $0 \leq \gamma \leq 1$ is necessary for the desired inequality to be valid.

From the second γ -orderliness, $\frac{\partial^2 QA}{\partial^2 \epsilon} \geq 0 \Rightarrow \forall 0 \leq a \leq 2a \leq 3a \leq \frac{1}{1+\gamma}, \frac{1}{a} \left(\frac{(QA(3a,\gamma)-QA(2a,\gamma))}{a} - \frac{(QA(2a,\gamma)-QA(a,\gamma))}{a} \right) \geq 0 \Rightarrow \text{if } 0 \leq \gamma \leq 1, \ \mathcal{B}_i - 2\mathcal{B}_{i+1} + \mathcal{B}_{i+2} \geq 0. \ \text{SM}_{\epsilon} \text{ can thus be interpreted as assuming } \gamma = 1 \ \text{and replacing the two blocks}, \ \mathcal{B}_i + \mathcal{B}_{i+2} \ \text{with one block } 2\mathcal{B}_{i+1}. \ \text{From the } \nu \text{th } \gamma \text{-orderliness}, \ \text{the recurrence relation of the derivatives naturally produces the alternating binomial coefficients,}$

$$(-1)^{\nu} \frac{\partial^{\nu} QA}{\partial \epsilon^{\nu}} \ge 0 \Rightarrow \forall 0 \le a \le \dots \le (\nu+1)a \le \frac{1}{1+\gamma},$$

$$\frac{(-1)^{\nu}}{a} \left(\frac{\frac{QA(\nu a + a, \gamma) \cdot \dots}{a} - \frac{\dots \cdot QA(2a, \gamma)}{a}}{a} - \frac{\frac{QA(\nu a, \gamma) \cdot \dots}{a} - \frac{\dots \cdot QA(a, \gamma)}{a}}{a} \right)$$

$$\ge 0 \Leftrightarrow \frac{(-1)^{\nu}}{a^{\nu}} \left(\sum_{j=0}^{\nu} (-1)^{j} {\nu \choose j} QA((\nu - j + 1) a, \gamma) \right) \ge 0$$

$$\Rightarrow \text{if } 0 \le \gamma \le 1, \sum_{j=0}^{\nu} (-1)^{j} {\nu \choose j} \mathcal{B}_{i+j} \ge 0.$$

Based on the ν th orderliness, the ϵ, γ -binomial mean is introduced as

$$BM_{\nu,\epsilon,\gamma,n} := \frac{1}{n} \left(\sum_{i=1}^{\frac{1}{2}\epsilon^{-1}(\nu+1)^{-1}} \sum_{j=0}^{\nu} \left(1 - (-1)^j {\nu \choose j} \right) \mathfrak{B}_{i_j} \right),$$

where $\mathfrak{B}_{i_j} = \sum_{l=n\gamma\epsilon(j+(i-1)(\nu+1)+1)}^{n\epsilon(j+(i-1)(\nu+1)+1)} (X_l + X_{n-l+1})$. If ν is not indicated, it defaults to $\nu=3$. Since the alternating sum of binomial coefficients equals zero, when $\nu\ll\epsilon^{-1}$ and $\epsilon\to0$, BM $\to\mu$. The solutions for the continuity of the breakdown point is the same as that in SM and not repeated here. The equality $\mathrm{BM}_{\nu=1,\epsilon}=\mathrm{BWM}_{\epsilon}$ holds. Similarly, $\mathrm{BM}_{\nu=2,\epsilon}=\mathrm{SM}_{\epsilon,b=3}$, when $\gamma=1$ and their respective ϵ s are identical.

Data Availability. Data for Figure ?? are given in SI Dataset
 S1. All codes have been deposited in GitHub.

ACKNOWLEDGMENTS. I sincerely acknowledge the insightful comments from the editor which considerably elevated the lucidity and merit of this paper.

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