Semiparametric robust mean estimations based on the orderliness of quantile averages

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semiparametric | mean-median-mode inequality | asymptotic | unimodal | Hodges—Lehmann estimator

Inequalities related to weighted averages

So far, it is quite natural to hypothesize that the value of ϵ, γ -trimmed mean should be monotonically related to the breakdown point in a semiparametric distribution, since it is a linear combination of quantile averages as shown in Section ??. Analogous to the γ -orderliness, the γ -trimming inequality for a right-skewed distribution is defined as $\forall 0 \leq \epsilon_1 \leq \epsilon_2 \leq \frac{1}{1+\gamma}$, $TM_{\epsilon_1,\gamma} \geq TM_{\epsilon_2,\gamma}$. γ -orderliness is a sufficient condition for the γ -trimming inequality, as proven in the SI Text. The next theorem shows a relation between the ϵ, γ -quantile average and the ϵ, γ -trimmed mean under the γ -trimming inequality, suggesting the γ -orderliness is not a necessary condition for the γ -trimming inequality.

Theorem .1. For a distribution that is right-skewed and follows the γ -trimming inequality, it is asymptotically true that the quantile average is always greater or equal to the corresponding trimmed mean with the same ϵ and γ , $0 \le \epsilon \le \frac{1}{1+\gamma}$.

Proof. According to the definition of the γ -trimming inequality: $\forall 0 \leq \epsilon \leq \frac{1}{1+\gamma}, \ \frac{1}{1-\epsilon-\gamma\epsilon+2\delta} \int_{\gamma\epsilon-\delta}^{1-\epsilon+\delta} Q\left(u\right) du \geq \frac{1}{1-\epsilon-\gamma\epsilon} \int_{\gamma\epsilon}^{1-\epsilon} Q\left(u\right) du$, where δ is an infinitesimal positive quantity. Subsequently, rewriting the inequality gives $\int_{\gamma\epsilon-\delta}^{1-\epsilon+\delta} Q\left(u\right) du - \frac{1-\epsilon-\gamma\epsilon+2\delta}{1-\epsilon-\gamma\epsilon} \int_{\gamma\epsilon}^{1-\epsilon} Q\left(u\right) du \geq 0 \Leftrightarrow \int_{1-\epsilon}^{1-\epsilon+\delta} Q\left(u\right) du + \int_{\gamma\epsilon-\delta}^{\gamma\epsilon} Q\left(u\right) du - \frac{2\delta}{1-\epsilon-\gamma\epsilon} \int_{\gamma\epsilon}^{1-\epsilon} Q\left(u\right) du \geq 0$ 25 0. Since $\delta \to 0^+$, $\frac{1}{2\delta} \left(\int_{1-\epsilon}^{1-\epsilon+\delta} Q\left(u\right) du + \int_{\gamma\epsilon-\delta}^{\gamma\epsilon} Q\left(u\right) du \right) = \frac{Q(\gamma\epsilon) + Q(1-\epsilon)}{2} \geq \frac{1}{1-\epsilon-\gamma\epsilon} \int_{\gamma\epsilon}^{1-\epsilon} Q\left(u\right) du$, the proof is complete.

An analogous result about the relation between the ϵ, γ -trimmed mean and the ϵ, γ -Winsorized mean can be obtained in the following theorem.

Theorem .2. For a right-skewed distribution following the γ -trimming inequality, asymptotically, the Winsorized mean is always greater or equal to the corresponding trimmed mean with the same ϵ and γ , provided that $0 \le \gamma \le 1$. If assuming γ -orderliness, the inequality is valid for any non-negative γ .

of the first assertion is complete. The second assertion is established in Theorem 0.3. in the SI Text. \Box

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Replacing the TM in the γ -trimming inequality with WA forms the definition of the γ -weighted inequality. The γ -orderliness also implies the γ -Winsorization inequality when $0 \le \gamma \le 1$, as proven in the SI Text. To construct weighted averages based on the ν th γ -orderliness and satisfying the corresponding weighted inequality, let $\mathcal{B}_i = \int_{i\epsilon}^{(i+1)\epsilon} \mathrm{QA}\left(u,\gamma\right) du$, $ka = k\epsilon + c$.

Data Availability. Data for Figure ?? are given in SI Dataset S1. All codes have been deposited in GitHub.

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