Semiparametric robust mean estimations based on the orderliness of quantile averages

Tuban Lee

This manuscript was compiled on June 10, 2023

semiparametric | mean-median-mode inequality | asymptotic | unimodal | Hodges-Lehmann estimator

Hodges–Lehmann inequality and γ -U-orderliness

- The Hodges-Lehmann estimator stands out as a unique robust location estimator due to its definition being substantially
- dissimilar from conventional L-estimators, R-estimators, and
- M-estimators. In their landmark paper, Estimates of location
- based on rank tests, Hodges and Lehmann (1) proposed two
- methods for computing the H-L estimator: the Wilcoxon score
- R-estimator and the median of pairwise means. The Wilcoxon
- score R-estimator is a location estimator based on signed-
- rank test, or R-estimator, (1) and was later independently
- discovered by Sen (1963) (2, 3). However, the median of
- pairwise means is a generalized L-statistic and a trimmed
- U-statistic, as classified by Serfling in his novel conceptualized
- study in 1984 (4). Serfling further advanced the understanding

- by generalizing the H-L kernel as $hl_k(x_1, ..., x_k) = \frac{1}{k} \sum_{i=1}^k x_i$, where $k \in \mathbb{N}$ (4). Here, the weighted H-L kernel is defined as $whl_k(x_1, ..., x_k) = \frac{\sum_{i=1}^k x_i \mathbf{w}_i}{\sum_{i=1}^k \mathbf{w}_i}$, where \mathbf{w}_i s are the weights applied to each element
- applied to each element.

By using the weighted H-L kernel and the L-estimator, it is now clear that the Hodges-Lehmann estimator is an LLstatistic, the definition of which is provided as follows:

$$LL_{k,\epsilon,\gamma,n} \coloneqq L_{\epsilon_0,\gamma,n}\left(\operatorname{sort}\left(\left(whl_k\left(X_{N_1},\cdots,X_{N_k}\right)\right)_{N=1}^{\binom{n}{k}}\right)\right),$$

- where $L_{\epsilon_0,\gamma,n}(Y)$ represents the ϵ_0,γ -L-estimator that uses
- the sorted sequence, sort $(whl_k(X_{N_1}, \dots, X_{N_k}))_{N=1}^{\binom{k}{k}}$, as input. The upper asymptotic breakdown point of $LL_{k,\epsilon,\gamma}$ is $\epsilon =$
- 21
- $1 (1 \epsilon_0)^{\frac{1}{k}}$, as proven in DSSM II.
- Data Availability. Data for Figure ?? are given in SI Dataset
- S1. All codes have been deposited in GitHub.
- **ACKNOWLEDGMENTS.** I sincerely acknowledge the insightful comments from the editor which considerably elevated the lucidity 26
- and merit of this paper.
- 1. J Hodges Jr, E Lehmann, Estimates of location based on rank tests. The Annals Math. Stat. 34, 598-611 (1963).
- 2. PK Sen, On the estimation of relative potency in dilution (-direct) assays by distribution-free methods. Biometrics pp. 532-552 (1963).
- 3. M Ghosh, MJ Schell, PK Sen, A conversation with pranab kumar sen. Stat. Sci. pp. 548-564
- 4. RJ Serfling, Generalized I-, m-, and r-statistics. The Annals Stat. 12, 76-86 (1984)