

Semiparametric robust mean estimations based on the orderliness of quantile averages

Tuban Lee

This manuscript was compiled on June 10, 2023

semiparametric | mean-median-mode inequality | asymptotic | unimodal
| Hodges–Lehmann estimator

Hodges–Lehmann inequality and γ - U -orderliness

The Hodges–Lehmann estimator stands out as a unique robust location estimator due to its definition being substantially dissimilar from conventional L -estimators, R -estimators, and M -estimators. In their landmark paper, *Estimates of location based on rank tests*, Hodges and Lehmann (1) proposed two methods for computing the H-L estimator: the Wilcoxon score R -estimator and the median of pairwise means. The Wilcoxon score R -estimator is a location estimator based on signed-rank test, or R -estimator, (1) and was later independently discovered by Sen (1963) (2, 3). However, the median of pairwise means is a generalized L -statistic and a trimmed U -statistic, as classified by Serfling in his novel conceptualized study in 1984 (4). Serfling further advanced the understanding by generalizing the H-L kernel as $hl_k(x_1, \dots, x_k) = \frac{1}{k} \sum_{i=1}^k x_i$, where $k \in \mathbb{N}$ (4). Here, the weighted H-L kernel is defined as $whl_k(x_1, \dots, x_k) = \frac{\sum_{i=1}^k x_i \mathbf{w}_i}{\sum_{i=1}^k \mathbf{w}_i}$, where \mathbf{w}_i s are the weights applied to each element.

By using the weighted H-L kernel and the L -estimator, it is now clear that the Hodges–Lehmann estimator is an LL -statistic, the definition of which is provided as follows:

$$LL_{k,\epsilon,\gamma,n} := L_{\epsilon_0,\gamma,n} \left(\text{sort} \left((whl_k(X_{N_1}, \dots, X_{N_k}))_{N=1}^{\binom{n}{k}} \right) \right),$$

where $L_{\epsilon_0,\gamma,n}(Y)$ represents the ϵ_0, γ - L -estimator that uses the sorted sequence, $\text{sort}(whl_k(X_{N_1}, \dots, X_{N_k}))_{N=1}^{\binom{n}{k}}$, as input. The upper asymptotic breakdown point of $LL_{k,\epsilon,\gamma}$ is $\epsilon = 1 - (1 - \epsilon_0)^{\frac{1}{k}}$, as proven in DSSM II. There are two ways to adjust the breakdown point: either by setting k as a constant and adjusting ϵ_0 , or by setting ϵ_0 as a constant and adjusting k . In the above definition, k is discrete, but the bootstrap method can be applied to ensure the continuity of k , also making the breakdown point continuous. Specifically, if $k \in \mathbb{R}$, let the bootstrap size be denoted by b , then first sampling the original sample $(1 - k + \lfloor k \rfloor)b$ times with each sample size of $\lfloor k \rfloor$, and then subsequently sampling $(1 - \lceil k \rceil + k)b$ times with each sample size of $\lceil k \rceil$, $(1 - k + \lfloor k \rfloor)b \in \mathbb{N}$, $(1 - \lceil k \rceil + k)b \in \mathbb{N}$. The corresponding kernels are computed separately, and the pooled sorted sequence is used as the input for the L -estimator. Let \mathbf{S}_k represent the sorted sequence. Indeed, for any finite sample, X , when $k = n$, \mathbf{S}_k becomes a single point, $whl_{k=n}(X_1, \dots, X_n)$. When $\mathbf{w}_i = 1$, the minimum of \mathbf{S}_k is $\frac{1}{k} \sum_{i=1}^k X_i$, due to the property of order statistics. The maximum of \mathbf{S}_k is $\frac{1}{k} \sum_{i=1}^k X_{n-i+1}$. The monotonicity of the order statistics implies the monotonicity of the extrema with respect to k , i.e., the support of \mathbf{S}_k shrinks monotonically. For

unequal \mathbf{w}_i s, the shrinkage of the support of \mathbf{S}_k might not be strictly monotonic, but the general trend remains, since all LL -statistics converge to the same point, as $k \rightarrow n$. Therefore, if $\frac{\sum_{i=1}^n X_i \mathbf{w}_i}{\sum_{i=1}^n \mathbf{w}_i}$ approaches the population mean when $n \rightarrow \infty$, all LL -statistics based on such consistent kernel function approach the population mean as $k \rightarrow \infty$. For example, if $whl_k = \text{BM}_{\nu, \epsilon_k, n=k}$, $\nu \ll \epsilon_k^{-1}$, $\epsilon_k \rightarrow 0$, such kernel function is consistent. These cases are termed the LL -mean ($LLM_{k,\epsilon,\gamma,n}$). By substituting the $WA_{\epsilon_0,\gamma,n}$ for the $L_{\epsilon_0,\gamma,n}$ in LL -statistic, the resulting statistic is referred to as the weighted L -statistic ($WL_{k,\epsilon,\gamma,n}$).

Data Availability. Data for Figure ?? are given in SI Dataset S1. All codes have been deposited in [GitHub](#).

ACKNOWLEDGMENTS. I sincerely acknowledge the insightful comments from the editor which considerably elevated the lucidity and merit of this paper.

1. J Hodges Jr, E Lehmann, Estimates of location based on rank tests. *The Annals Math. Stat.* **34**, 598–611 (1963).
2. PK Sen, On the estimation of relative potency in dilution (-direct) assays by distribution-free methods. *Biometrics* pp. 532–552 (1963).
3. M Ghosh, MJ Schell, PK Sen, A conversation with pranab kumar sen. *Stat. Sci.* pp. 548–564 (2008).
4. RJ Serfling, Generalized l-, m-, and r-statistics. *The Annals Stat.* **12**, 76–86 (1984).

T.L. designed research, performed research, analyzed data, and wrote the paper.

The author declares no competing interest.

¹To whom correspondence should be addressed. E-mail: tl@biomathematics.org