

# Semiparametric robust mean estimations based on the orderliness of quantile averages

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## Hodges–Lehmann inequality and $\gamma$ - $U$ -orderliness

The Hodges–Lehmann estimator stands out as a unique robust location estimator due to its definition being substantially dissimilar from conventional  $L$ -estimators,  $R$ -estimators, and  $M$ -estimators. In their landmark paper, *Estimates of location based on rank tests*, Hodges and Lehmann (1) proposed two methods for computing the H-L estimator: the Wilcoxon score  $R$ -estimator and the median of pairwise means. The Wilcoxon score  $R$ -estimator is a location estimator based on signed-rank test, or  $R$ -estimator, (1) and was later independently discovered by Sen (1963) (2, 3). However, the median of pairwise means is a generalized  $L$ -statistic and a trimmed  $U$ -statistic, as classified by Serfling in his novel conceptualized study in 1984 (4). Serfling further advanced the understanding by generalizing the H-L kernel as  $hl_k(x_1, \dots, x_k) = \frac{1}{k} \sum_{i=1}^k x_i$ , where  $k \in \mathbb{N}$  (4). Here, the weighted H-L kernel is defined as  $whl_k(x_1, \dots, x_k) = \frac{\sum_{i=1}^k x_i \mathbf{w}_i}{\sum_{i=1}^k \mathbf{w}_i}$ , where  $\mathbf{w}_i$ s are the weights applied to each element.

By using the  $whl_k$  kernel and the  $L$ -estimator, it is now clear that the Hodges–Lehmann estimator is an  $LL$ -statistic, the definition of which is provided as follows:

$$LL_{k,\epsilon,\gamma,n} := L_{\epsilon_0,\gamma,n} \left( \text{sort} \left( (whl_k(X_{N_1}, \dots, X_{N_k}))_{N=1}^{\binom{n}{k}} \right) \right),$$

where  $L_{\epsilon_0,\gamma,n}(Y)$  represents the  $L$ -estimator that uses the sorted sequence,  $\text{sort}(whl_k(X_{N_1}, \dots, X_{N_k}))_{N=1}^{\binom{n}{k}}$ , as input, the upper asymptotic breakdown point of the  $L$ -estimator is  $\epsilon_0$ , the lower asymptotic breakdown point is  $\gamma\epsilon_0$ . The upper asymptotic breakdown point of  $LL_{k,\epsilon,\gamma}$  is  $\epsilon = 1 - (1 - \epsilon_0)^{\frac{1}{k}}$ , as proven in another relevant paper. There are two ways to adjust the breakdown point: either by setting  $k$  as a constant and adjusting  $\epsilon_0$ , or by setting  $\epsilon_0$  as a constant and adjusting  $k$ . In the above definition,  $k$  is discrete, but the bootstrap method can be applied to ensure the continuity of  $k$ , also making the breakdown point continuous. Specifically, if  $k \in \mathbb{R}$ , let the bootstrap size be denoted by  $b$ , then first sampling the original sample  $(1 - k + \lfloor k \rfloor)b$  times with each sample size of  $\lfloor k \rfloor$ , and then subsequently sampling  $(1 - \lceil k \rceil + k)b$  times with each sample size of  $\lceil k \rceil$ ,  $(1 - k + \lfloor k \rfloor)b \in \mathbb{N}$ ,  $(1 - \lceil k \rceil + k)b \in \mathbb{N}$ . The corresponding kernels are computed separately, and the pooled sorted sequence is used as the input for the  $L$ -estimator. Indeed, for any finite sample,  $X$ , when  $k = n$ , the sorted sequence becomes a single point,  $whl_{k=n}(X_1, \dots, X_n)$ .

**Data Availability.** Data for Figure ?? are given in SI Dataset S1. All codes have been deposited in [GitHub](#).

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T.L. designed research, performed research, analyzed data, and wrote the paper.

The author declares no competing interest.

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