## Semiparametric robust mean estimations based on the orderliness of quantile averages

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## Hodges–Lehmann inequality and $\gamma$ -U-orderliness

The Hodges–Lehmann estimator stands out as a unique robust location estimator due to its definition being substantially dissimilar from conventional L-estimators, R-estimators, and M-estimators. In their landmark paper, Estimates of location based on rank tests, Hodges and Lehmann (1) proposed two methods for computing the H-L estimator: the Wilcoxon score R-estimator and the median of pairwise means. The Wilcoxon score R-estimator is a location estimator based on signed-rank test, or R-estimator, (1) and was later independently discovered by Sen (1963) (2, 3). However, the median of pairwise means is a generalized L-statistic and a trimmed U-statistic, as classified by Serfling in his novel conceptualized study in 1984 (4). Serfling further advanced the understanding by generalizing the H-L kernel as  $hl_k(x_1,\ldots,x_k) = \frac{1}{k} \sum_{i=1}^k x_i$ , where  $k \in \mathbb{N}$  (4). Here, the weighted H-L kernel is defined as  $whl_k(x_1,\ldots,x_k) = \frac{\sum_{i=1}^k x_i \mathbf{w}_i}{\sum_{i=1}^k \mathbf{w}_i}$ , where  $\mathbf{w}_i$ s are the weights applied to each element.

By using the weighted H-L kernel and the L-estimator, it is now clear that the Hodges-Lehmann estimator is an LL-statistic, the definition of which is provided as follows:

$$LL_{k,\epsilon,\gamma,n} := L_{\epsilon_0,\gamma,n} \left( \operatorname{sort} \left( \left( whl_k \left( X_{N_1}, \cdots, X_{N_k} \right) \right)_{N=1}^{\binom{n}{k}} \right) \right),$$

where  $L_{\epsilon_0,\gamma,n}(Y)$  represents the  $\epsilon_0,\gamma$ -L-estimator that uses the sorted sequence, sort  $(whl_k(X_{N_1}, \dots, X_{N_k}))_{N=1}^{\binom{n}{k}}$ , as input. The upper asymptotic breakdown point of  $LL_{k,\epsilon,\gamma}$  is  $\epsilon =$  $1-(1-\epsilon_0)^{\frac{1}{k}}$ , as proven in DSSM II. There are two ways to adjust the breakdown point: either by setting k as a constant and adjusting  $\epsilon_0$ , or by setting  $\epsilon_0$  as a constant and adjusting k. In the above definition, k is discrete, but the bootstrap method can be applied to ensure the continuity of k, also making the breakdown point continuous. Specifically, if  $k \in \mathbb{R}$ , let the bootstrap size be denoted by b, then first sampling the original sample (1 - k + |k|)b times with each sample size of |k|, and then subsequently sampling  $(1-\lceil k \rceil + k)b$  times with each sample size of  $\lceil k \rceil$ ,  $(1-k+|k|)b \in \mathbb{N}$ ,  $(1-\lceil k \rceil+k)b \in \mathbb{N}$ . The corresponding kernels are computed separately, and the pooled sorted sequence is used as the input for the L-estimator. Indeed, for any finite sample, X, when k = n, the sorted sequence becomes a single point,  $whl_{k=n}(X_1,\ldots,X_n)$ .

**Data Availability.** Data for Figure ?? are given in SI Dataset S1. All codes have been deposited in GitHub.

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The author declares no competing interest.

T.L. designed research, performed research, analyzed data, and wrote the paper

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