# Sim\_setting\_binary

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For dichotomous outcomes, we use a logit link i.e.,  $log \frac{E[Y|X]}{1-E[Y|X]} = \sum_{i=0}^{p} \theta_i X_i$  and the true average treatment effect is  $E[Y^1 = 1] - E[Y^0 = 1]$ .

In this project, we want to simulate the binary outcome so that the treatment (i.e. smoking status) causes an increases in the probability of the occurrence of the outcome (i.e. lung cancer) by 15 percent. For our population without treatment (smoking) history, we assume that the probability of getting lung cancer is 0.2. Namely,  $E[Y^0 = 1] = 0.2$ . From this fact, we know that our true log odds ratio  $(\beta_1)$  for our binary logistic regression is approximately 0.767.

We set our desired proportion of treatment is 0.14 given that 14% of U.S. adults were current smokers in 2019.

Based on this simulation design, we can obtain our  $\alpha_0$  coefficient for our propensity score model and  $\beta_0$  and  $\beta_1$  for our logistic model with some algebra computation.

With logit model and standardization method (creating pseudo population),

$$E[Y^1 = 1] - E[Y^0 = 1]$$

$$\Leftrightarrow E[Y^1 = 1|A = 1] - E[Y^0 = 1|A = 0]$$
 (by exchangeabiltiy)

$$\Leftrightarrow \ E[Y=1|A=1] - E[Y=1|A=0]$$

$$\Leftrightarrow \frac{e^{\beta_0 + \beta_1 A + \beta_2 L_1 + \beta_3 L_2}}{1 + e^{\beta_0 + \beta_1 A + \beta_2 L_1 + \beta_3 L_2}} - 0.2 = 0.15$$

$$\Leftrightarrow e^{\beta_0 + \beta_1 A + \beta_2 L_1 + \beta_3 L_2} = 0.35 + 0.35e^{\beta_0 + \beta_1 A + \beta_2 L_1 + \beta_3 L_2}$$

$$\Leftrightarrow \beta_0 + \beta_1 A + \beta_2 L_1 + \beta_3 L_2 = log \frac{0.35}{0.65}$$

Now we use,

$$E[Y^0 = 1] = e^{\beta_0 + \beta_2 L_1 + \beta_3 L_2} = 0.2$$

$$\Leftrightarrow \beta_0 + \beta_2 L_1 + \beta_3 L_2 = log \frac{0.2}{0.8}$$

$$\Leftrightarrow \beta_2 L_1 + \beta_3 L_2 = log \frac{0.2}{0.8} + \beta_0$$

Using this equation, we get the following:

$$\Rightarrow \beta_0 + \beta_1 A + \log \frac{0.2}{0.8} + \beta_0 = \log \frac{0.35}{0.65}$$

$$\Leftrightarrow \beta_0 \approx (-0.619 + 1.386 + 0.767A)/2$$

$$\Rightarrow E[\hat{\beta}_0] \approx E[-0.619 + 1.386 + 0.767A]/2$$

$$\Rightarrow \hat{\beta}_0 \approx (-0.619 + 1.386 + 0.767E[A])/2$$

$$\Rightarrow \hat{\beta_0} \approx (-0.619 + 1.386 + 0.767 * 0.14)/2$$

$$\Rightarrow \hat{\beta_0} \approx 0.437$$

From this computation, we can see that no matter what other covariates are in the model, we can obtain  $\hat{\beta}_0$ .

## Finding $\alpha_0$ and coefficients for L1, L2 ,L3 in propensity score model

This process allows us to find the values of coefficients for our propensity score model in order to achieve the expected proportion of the treated in our data simulation.

```
pre_data <- defData(varname = "L1", formula = "0", variance = 1,</pre>
                 dist = "normal")
pre_data <- defData(pre_data, varname = "L2", formula = "0", variance = 1,</pre>
                 dist = "normal")
pre_data <- defData(pre_data, varname = "L3", formula = "0", variance = 1,</pre>
                 dist = "normal")
set.seed(77)
df <- genData(5000, pre_data)</pre>
data_gen <- function(desired_prop) {</pre>
   alpha 0 = \log(\text{desired prop}/(1 - \text{desired prop}))
   coef_L1 <- sample(df$L1, 1000, replace = TRUE)</pre>
   coef_L2 <- sample(df$L2, 1000, replace = TRUE)</pre>
   coef_L3 <- sample(df$L3, 1000, replace = TRUE)</pre>
   A_logit <- vector(mode = "list",length = length(coef_L2))
   p_A <- vector(mode = "numeric",length = length(coef_L2))</pre>
   u \leftarrow alpha_0 + coef_L1[1]*df$L1 + coef_L2[1]*df$L2 - coef_L3[1]*df$L3
   p_A[1] \leftarrow mean(exp(u)/(1 + exp(u)))
   tol <- 0.005
   i = 1
   while (abs(p_A[i] - 0.14) > tol) {
     i = i + 1
     A_logit[[i]] <-
       alpha_0 + coef_L2[i]*df$L1 + coef_L2[i]*df$L2 - coef_L3[i]*df$L3
     p_A[i] <- mean(exp(A_logit[[i]])/(1 + exp(A_logit[[i]])))</pre>
     if (abs(p_A[i] - 0.14) < tol) {
       mean_treated_proportion <- p_A[i]</pre>
       desired_coef_L1 <- coef_L1[i]</pre>
       desired_coef_L2 <- coef_L2[i]</pre>
        desired_coef_L3 <- coef_L3[i]</pre>
     if (i > length(coef_L2)) {
```

```
stop("You need better coverage, mate.")
    }
   }
   return(tibble(alpha_0, mean_treated_proportion,
            desired_coef_L1, desired_coef_L2, desired_coef_L3))
}
set.seed(77)
data_gen(0.14)
## # A tibble: 1 x 5
     alpha_0 mean_treated_proporti~ desired_coef_L1 desired_coef_L2 desired_coef_L3
##
       <dbl>
                               <dbl>
                                               <dbl>
                                                                <dbl>
                                                                                <dbl>
       -1.82
                                              -0.978
## 1
                               0.145
                                                                0.127
                                                                               -0.285
```

#### Generating the treatment and outcome data

Here, our intention is to introduce confounders in the data generation process in order to see how using propensity matching and standardization can tackle the problems caused by confounders in predicting the average treatment effect.

## [1] 0.1864

## Nearest neighbor propensity score matching

```
summary(matched)[2]
```

```
## $nn
## Control Treated
## All (ESS) 4068 932
## All 4068 932
```

```
## Matched (ESS) 932 932
## Matched 932 932
## Unmatched 3136 0
## Discarded 0 0

matched_df <-
   match.data(matched, distance = "ps")</pre>
```

## PS matched model variance-covariance matrix

```
matched_ps_model <- glm(Y ~ A + ps, data = matched_df, family = "binomial")

vcov(matched_ps_model)

## (Intercept) A ps
## (Intercept) 0.013001011 -0.0046680524 -0.0313372204

## A -0.004668052 0.0100425924 0.0004670379

## ps -0.031337220 0.0004670379 0.1160915103</pre>
```

### simple bootstrap

```
nboot <- 100
# set up a matrix to store results
boots <- data.frame(i = 1:nboot,</pre>
                     se_ATE = NA,
                     se_{OR} = NA,
                     log_OR = NA,
                     mean1 = NA,
                     mean0 = NA,
                     difference = NA
# loop to perform the bootstrapping
for (i in 1:nboot) {
  # sample with replacement
  sampl <- matched_df %>% filter(subclass %in% sample(levels(subclass),500, replace = TRUE))
  bootmod \leftarrow glm(Y \sim A + ps, data = sampl,
                  weights = weights, family = binomial)
  # create new data sets
  sampl.treated <- sampl %>%
    mutate(A = 1)
  sampl.untreated <- sampl %>%
    mutate(A = 0)
  # predict values
  sampl.treated$pred.y <-</pre>
```

```
predict(bootmod, sampl.treated, type = "response")
sampl.untreated$pred.y <-</pre>
  predict(bootmod, sampl.untreated, type = "response")
 # output results
boots[i, "log OR"] <- summary(bootmod)$coeff[2,1]</pre>
boots[i, "se OR"] <- summary(bootmod)$coeff[2,2]</pre>
boots[i, "se_ATE"] <-</pre>
  sqrt((summary(bootmod)$coeff[2,2]*mean(sampl.treated$pred.y) *
     (1 - mean(sampl.treated$pred.y)))^2 +
  (summary(bootmod)$coeff[2,2]*mean(sampl.untreated$pred.y) *
     (1 - mean(sampl.untreated$pred.y)))^2)
boots[i, "mean1"] <- mean(sampl.treated$pred.y)</pre>
boots[i, "mean0"] <- mean(sampl.untreated$pred.y)</pre>
boots[i, "difference"] <- boots[i, "mean1"] - boots[i, "mean0"]</pre>
mean_log_OR <- mean(boots$log_OR)</pre>
Empirical_se_ATE <- sd(boots$difference)</pre>
mean_se_ATE <- mean(boots$se_ATE)</pre>
Empirical_se_log_OR <- sd(boots$log_OR)</pre>
mean_se_log_OR <- mean(boots$se_OR)</pre>
ATE <- mean(boots$difference)</pre>
# once loop is done, print the results
if (i == nboot) {
  cat("ATE:")
  cat(ATE)
  cat("\n")
  cat("\n")
  cat("Empirical_se_ATE:")
  cat(Empirical_se_ATE)
  cat("\n")
  cat("\n")
  cat("mean se ATE:")
  cat(mean_se_ATE)
  cat("\n")
  cat("\n")
  cat("95% CI for ATE:")
  cat(ATE - 1.96*Empirical_se_ATE,
      ATE + 1.96*Empirical_se_ATE)
  cat("\n")
  cat("\n")
  cat("mean_log_OR:")
```

```
cat(mean_log_OR)
    cat("\n")
    cat("\n")
    cat("Empirical_se_log_OR:")
    cat(Empirical_se_log_OR)
    cat("\n")
    cat("\n")
    cat("mean_se_log_OR:")
    cat(mean_se_log_OR)
    cat("\n")
    cat("\n")
    cat("95% CI for log odds ratio:")
    cat(mean_log_OR - 1.96*mean_se_log_OR,
        ",",
        mean_log_OR + 1.96*mean_se_log_OR)
  }
}
## ATE:0.1264747
##
## Empirical_se_ATE:0.02339231
##
## mean_se_ATE:0.04828172
##
## 95% CI for ATE:0.08062574 , 0.1723236
##
## mean_log_OR:0.5823668
##
## Empirical_se_log_OR:0.1091926
##
## mean_se_log_OR:0.1558736
## 95% CI for log odds ratio:0.2768546 , 0.887879
True ATE: 0.15
True log OR: 0.767
```