

$$X|Y=0 \sim \begin{pmatrix} 1 & 2 & 3 \\ \frac{3}{18} & 0 & \frac{4}{18} \\ \frac{7}{18} & & \frac{7}{18} \end{pmatrix}$$

$$X|X>\underline{2} \sim \begin{pmatrix} -1 & 0 & 2 \\ \frac{1}{18} & \frac{3}{18} & \frac{2}{18} \\ \frac{6}{18} & \frac{6}{18} & \frac{6}{18} \end{pmatrix}$$

$X, Y$  nu sunt indep  $\Rightarrow$   $\text{fara } 0 \dots$  (\*\*\*\*)

$$\Rightarrow E[XY] = \frac{1}{18}(-1) + 2 \cdot \frac{2}{18} + (-2) \cdot \frac{2}{18} + 4 \cdot \frac{3}{18} + 6 \cdot \frac{3}{18} = \dots$$

$$\begin{aligned} E[2X+3Y] &= 2E[X] + 3E[Y] \\ &\stackrel{\text{sum}}{=} \sum (2x+3y) P(X=x, Y=y) \end{aligned}$$

$$E[X|Y=y] = \sum_x x P(X=x|Y=y)$$

media  
condiționată  
a lui  $X$  la  $Y=y$

$$* E[X|Y] = g(Y)$$

$\rightarrow$   $x$  e condit. la  $Y$  în val.  $g(y_i)$ .  
 $=$   
 din  $Y$

A

MAIL LAB

F

9 feb ora 9

Cursul 13 : 09.02.

$f_{xy}(x,y)$  densitatea comună

$$f_x(x) = \int f_{xy}(x,y) dy$$

$$f_y(y) = \int f_{xy}(x,y) dx$$

$$f_{X|A}(x) = \frac{f_X(x)}{P(X \in A)}$$

$$P(x \in B | A) = \int_B f_{X|A}(x) dx$$

$$f_{X|Y}(x|y) = \frac{f_{X,Y}(x|y)}{f_Y(y)}$$

Formula probab. totale

$$f_X(x) = \sum_{i=1}^n f_{X|A_i} P(A_i)$$

a)  $y$  este o variabilă a discretei  $\{y_1, \dots, y_m\}$   
 $x$  este o v.a. cont.  $f_x$

$$f_X(x) = \sum_{i=1}^m f_{X|Y}(x|y_i) P(Y=y_i)$$

b)  $y$  v.a. cont.  
 $x$  v.a. cont cu dens.  $f_x$

$$f_X(x) = \int f_{X|Y}(x|y) f_Y(y) dy$$

Independență c. a.

$$\underline{x \perp\!\!\!\perp y} \Leftrightarrow P(x \in A, y \in B) = P(x \in A) P(y \in B), \forall A, B \subseteq \mathbb{R}$$

$$\text{pt. } A = (-\infty, x] ; B = (-\infty, y] \Rightarrow P(X \leq x, Y \leq y) = P(X \leq x) \cdot P(Y \leq y)$$

$$\int_{-\infty}^x \int_{-\infty}^y f_{x,y}(u,v) du dv = \int_{-\infty}^x f_x(u) du \int_{-\infty}^y f_y(v) dv$$

$$\frac{\partial}{\partial x} \int_{-\infty}^y f_{x,y}(u,v) du =$$

derivație  
după  $x$  și  $y$

$$= \frac{\partial}{\partial x} \int_{-\infty}^x f_x(u) du \frac{\partial}{\partial y} \int_{-\infty}^y f_y(v) dv$$

$$f_{x,y}(x,y) = f_x(x) f_y(y)$$

! **P** Fie  $x, y$  v.a. cu densități  $f_x$  respectiv  $f_y$ .  
 Atunci  $x \perp\!\!\!\perp y \Leftrightarrow f_{x,y}(x,y) = f_x(x) \cdot f_y(y)$ .

**P** Fie  $x, y$  două v.a. și  $g$  și  $h$  două funcții.

Dacă  $f_{x,y}(x,y) = g(x) h(y)$ , atunci  $x \perp\!\!\!\perp y$

**P** Dacă  $x$  și  $y$  sunt 2 variabile aleatoare independente, atunci

$$! E[g(x) h(y)] = E[g(x)] E[h(y)]$$

Dacă  $g(x) = x$  și  $h(y) = y \Rightarrow$  fol. identitate.  
 atunci  $E[x,y] = E[x] \cdot E[y]$

$$! \text{Var. } (X+Y) = \text{Var. } (X) + \text{Var. } (Y) \quad \text{nu e reciproc}$$

Formula lui Bayes

$$f_{x|y}(x|y) = f_{x,y}(x|y) \cdot f_y(y)$$

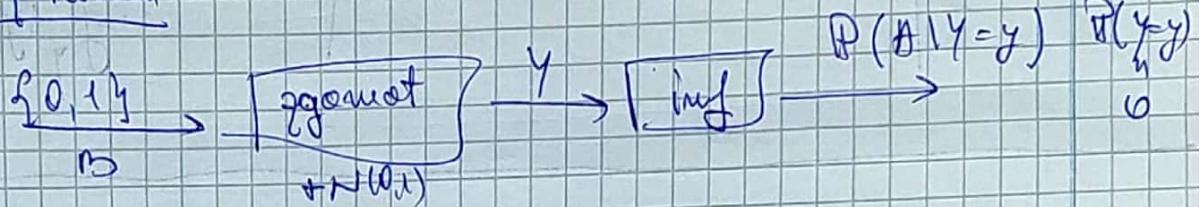
$$= f_{y|x}(y|x) f_x(x)$$

$$\Rightarrow f_{x|y}(x|y) = \frac{f_{x,y}(x|y)}{f_y(y)} = \frac{f_{y|x} \cdot f_x(x)}{f_y(y)} =$$

$$= \frac{f_{y|x} \cdot f_x(x)}{\int f_{y|x}(y|x') \cdot f_x(x') dx'}$$

Für ein Bayes

Cay hilft



$$P(A|Y=y) = \lim_{dy \rightarrow 0} P(A \cap Y \in (y, y+dy))$$

$$(y - \frac{dy}{2}, y + \frac{dy}{2})$$

$$= \lim_{dy \rightarrow 0} \frac{P(A \cap Y \in (y, y+dy))}{P(Y \in (y, y+dy))} =$$

$$= \lim_{dy \rightarrow 0} \frac{P(A) \cdot P(Y \in (y, y+dy) | A)}{P(Y \in (y, y+dy))}$$

$$= \lim_{dy \rightarrow 0} \frac{P(A) \cdot \int_y^{y+dy} f_{y|A}(u) du}{\int_y^{y+dy} f_y(u) du}$$

$$= \lim_{dy \rightarrow 0} \frac{P(A) \cdot \int_y^{y+dy} f_{y|A}(u) dy}{\int_y^{y+dy} f_y(u) dy}$$

$$P(A|Y=y) = \frac{P(A) \cdot \int_y^{y+dy} f_{y|A}(u) dy}{\int_y^{y+dy} f_y(u) dy}$$

$$\begin{aligned} & \text{durch marginale f. ein Bayes} \\ & = f_{y|A}(y) P(A) + \\ & + f_{y|A^c}(y) \cdot P(A^c) \end{aligned}$$

## Formule probab. totale

$X \setminus Y$	discret	cont
discret	$P(X=x) = \sum_y P(X=x \cap Y=y) \cdot P(Y=y)$	$P(X=x) = \int P(X=x \mid Y=y) \cdot f_Y(y) dy$
cont.	$f_X(x) = \sum_y f_{X Y}(x \mid y) P(Y=y)$	$f_X(x) = \int f_{X Y}(x \mid y) \cdot f_Y(y) dy$

## Formule lui Bayes

$X \setminus Y$	discret	cont
discret	$P(Y=y \mid X=x) = \frac{P(X=x \mid Y=y) \cdot P(Y=y)}{P(X=x)}$	$f_{Y X}(y \mid x) = \frac{P(X=x \mid Y=y) \cdot f_Y(y)}{P(X=x)}$
cont.	$P(Y=y \mid X=x) = \frac{f_{X Y}(x \mid y) \cdot P(Y=y)}{f_X(x)}$	$f_{Y X}(y \mid x) = \frac{f_{X Y}(x \mid y) \cdot f_Y(y)}{f_X(x)}$

Exp: A și B : durata de viață A Exp ( $\lambda_0$ )  
                       $\sim \text{Exp} (\lambda_1)$   
 Ap. că primul urmăruștele de la  $\lambda_0 < \lambda_1$

A are prob.  $p_0$  să fie de la B cu prob.  $p_1 = 1 - p_0$

Fie  $T$  durata de viață a telefornicului pe linii.

- a) Fct. de rep. și densitatea lui  $T$
- b) Vrem să găsim prob. ca telefonul să fie prezent de la B stîncoală  $T=t$ .

$\exists$

$T$  v. q. acut

Fie  $I$  v. a.

$$\begin{cases} 0, \text{ dacă } I=0 \\ 1, \text{ dacă } I=1 \end{cases}$$

$$\text{Exp}(\lambda) \cdot f(x) = \lambda e^{-\lambda x}$$

$$F(x) = 1 - e^{-\lambda x}$$

$$P(I=0) = p_0$$

$$P(I=1) = p_1 = 1 - p_0$$

$$T | I=0 \sim \text{Exp}(\lambda_0)$$

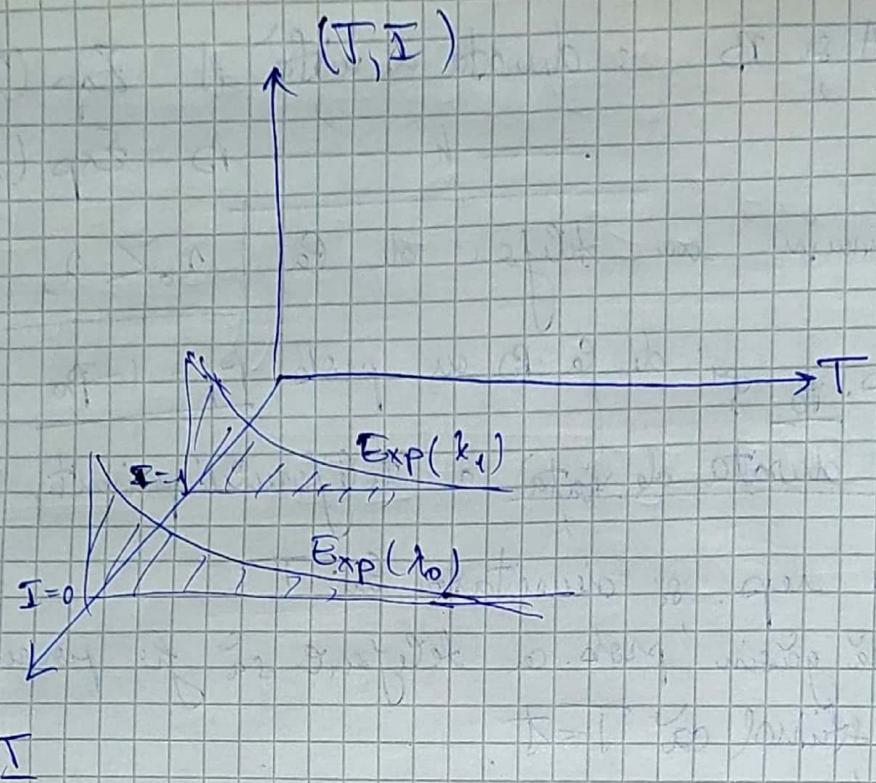
$$T | I=1 \sim \text{Exp}(\lambda_1)$$

$$P(T \leq t) = P(T \leq t | I=0) \cdot P(I=0) +$$

$$+ P(T \leq t | I=1) \cdot P(I=1)$$

$$= (1 - e^{-\lambda_0 t}) p_0 + (1 - e^{-\lambda_1 t}) (1 - p_0)$$

$$f_T(t) = \frac{d}{dt} F_T(t) = \lambda_0 e^{-\lambda_0 t} p_0 + \lambda_1 e^{-\lambda_1 t} (1 - p_0) \quad t > 0$$



b)  $P(I=1 | T=t) = \frac{f_{T|I}(t|1) \cdot P(I=1)}{f_T(t)}$

$$= \frac{\lambda_1 \cdot e^{-\lambda_1 t} (1-p_0)}{\lambda_0 \cdot e^{-\lambda_0 t} \cdot p_0 + \lambda_1 \cdot e^{-\lambda_1 t} (1-p_0)}$$

Media unei fct. de v.a

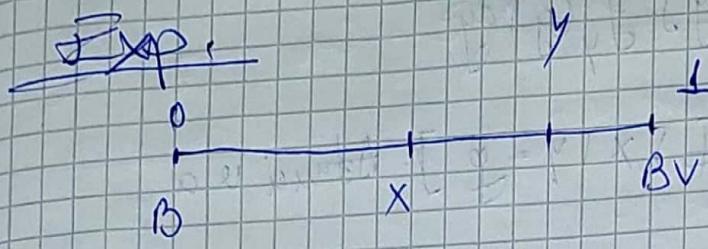
$x, y$  două v.a.  $f_{x,y}(x,y)$  și  $g: \mathbb{R}^2 \rightarrow \mathbb{R}$

$$\boxed{E[g(x,y)] = \iint g(x,y) \cdot f_{x,y}(x,y) dx dy}$$

In particular

~~E<sup>xy</sup>~~

$$E[xy] = \iint xy f_{x,y}(x,y) dx dy$$



$$x, y \sim U[0,1] \text{ indep.}$$

$$\mathbb{E}[|x-y|] = \iint |x-y| f_{x,y}(x,y) dx dy$$

$$\begin{aligned} \mathbb{E}[|x-y|] &= \iint |x-y| \cdot \mathbf{1}_{[0,1]}(x) \cdot \mathbf{1}_{[0,1]}(y) dx dy \\ &= \int_0^1 \int_0^1 (x-y) dx dy + \int_0^1 \int_0^y (y-x) dx dy \\ &= \left. \int_0^1 \frac{x^2}{2} - yx \right|_y^1 dy + \left. \int_0^1 yx - \frac{x^2}{2} \right|_0^y dy = \frac{1}{3} \end{aligned}$$

Media conditioanala

Xv.a and si A un ev.  $P(A) > 0$

$$\mathbb{E}[x|A] = \int_x f_{x|A}(x) dx$$

$$\text{Dacă } A = \{y = y\}$$

$$\mathbb{E}[x|y=y] = \int_x f_{x|y}(x|y) dx$$

Formule probabile totale

$$f_x(x) = \sum_{i=1}^m f_{x|A_i} \cdot P(A_i)$$

| apoi integrz.

$$\mathbb{E}[x] = \sum_{i=1}^m \mathbb{E}[x|A_i] \cdot P(A_i)$$

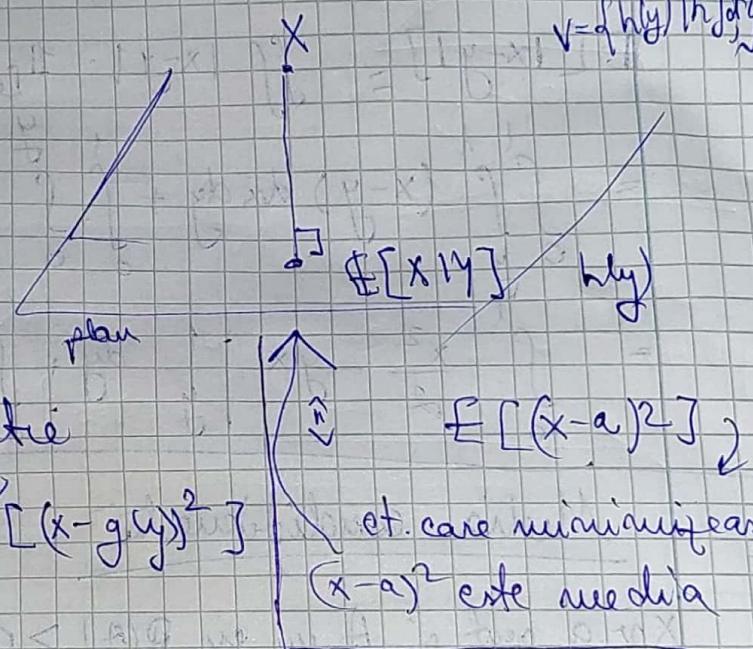
$$\mathbb{E}[X] = \int \mathbb{E}[X|Y=y] f_y(y) dy$$

Def. Fie  $g(y) = \mathbb{E}[X|Y=y]$ . Astăzi vea

$$\mathbb{E}[X|Y] = g(Y)$$

Prop.

$$a) \mathbb{E}[\mathbb{E}[X|Y]] = \mathbb{E}[X]$$



sea mai bună predicție

$$\mathbb{E}[X|Y] = \underset{g}{\operatorname{arg\,min}} \mathbb{E}[(x-g(y))^2]$$

$$\text{Var}(X) = \text{Var}(\mathbb{E}[X|Y]) + \mathbb{E}[\text{Var}(X|Y)]$$

$N$

$X_1, X_2, \dots$

$$T = X_1 + X_2 + \dots + X_N$$

$$\mathbb{E}[T]$$

$$w_1 \quad N(w_1) = 10$$

$$T(w_1) = X_1^{w_1} + \dots + X_{10}^{w_1}$$

## Covarianta și corelație

Def. Fie  $X$  și  $Y$  două v.a. S.m. covariantele dintre  $X$  și  $Y$ .

$$\text{Cov}(X, Y) = E[(X - E[X])(Y - E[Y])]$$

În particular,  $X = Y \Rightarrow \text{Cov}(X, X) = \text{Var}(X)$

(P)  $\boxed{\text{Cov}(X, Y) = E[XY] = E[X]E[Y]}$  !!

Def. Spunem că v.a.  $X$  și  $Y$  sunt necorelate dacă  $\text{Cov}(X, Y) = 0$ .

Astfel, deși  $E[XY] = E[X]E[Y]$ ,

~~Obs~~ Dacă  $X \perp\!\!\! \perp Y \Rightarrow X$  și  $Y$  sunt necorelate

Raciocinare  
falsa

Ex:  $X \sim N(0, 1) \Rightarrow E[X]E[Y] = 0$

$$Y = X^2 \Rightarrow E[XY] = E[X^3] = 0$$

$\Rightarrow X$  și  $Y$  sunt necorelate

$X$  și  $Y$  nu sunt indep!

Prop a)  $\text{Cov}(X, X) = \text{Var}(X)$

b)  $\text{Cov}(X, a) = 0$ ;  $a$  const.

c)  $\text{Cov}(ax + by, Y) = b \text{Cov}(X, Y)$

d)  $\text{Cov}(X, Y) = \text{Cov}(Y, X)$

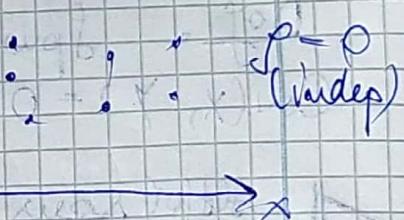
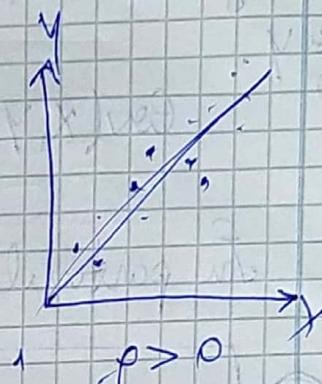
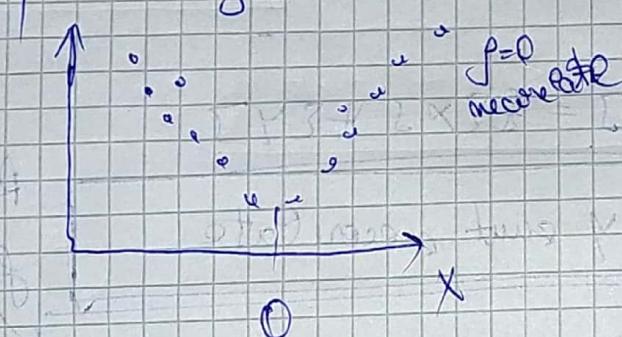
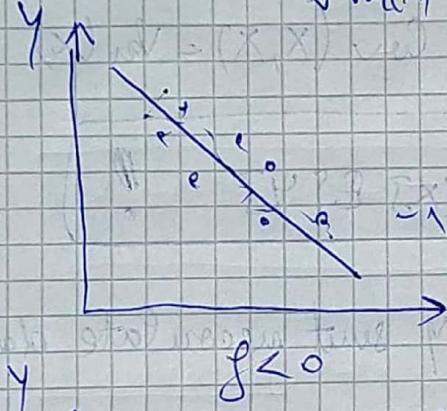
e)  $\text{Var}(X+Y) = \text{Var}(X) + \text{Var}(Y) + 2 \text{Cov}(X, Y)$

$\text{Var}(X_1 + X_2 + \dots + X_n) = \sum_{i=1}^n \text{Var}(X_i) + 2 \sum_{i < j} \text{Cov}(X_i, X_j)$

$$g) \text{Cov}(X+Y, Z) = \text{Cov}(X, Z) + \text{Cov}(Y, Z)$$

Def. Fie  $X$  și  $Y$  două v.a. și definim coefficientul de corelație dintre  $X$  și  $Y$ .

$$f(X, Y) = \frac{\text{Cov}(X, Y)}{\sqrt{\text{Var}(X)} \cdot \sqrt{\text{Var}(Y)}}$$



Prop.  $f \in [-1, 1]$ . Dacă  $f = 1$  sau  $-1$  atunci  $X = a + bY$  ( $Y = a + bX$ ) a.s. (aproape sigur).

$$P(X = a + bY) = 1$$

Dem.  $X, Y : E[X] = \mu_X, \text{Var}(X) = \sigma_X^2$

$$E[Y] = \mu_Y, \text{Var}(Y) = \sigma_Y^2$$

P.p. că  $\mu_X = \mu_Y = 0$  și  $\sigma_X^2 = \sigma_Y^2 = 1 \rightarrow f(X, Y) = E[XY]$

$$f(X, Y) = \frac{\text{Cov}(X, Y)}{\sqrt{\text{Var}(X)} \sqrt{\text{Var}(Y)}} = \frac{E[(X - \mu_X)(Y - \mu_Y)]}{\sigma_X \sigma_Y}$$

$$= \mathbb{E} \left[ \left( \frac{x - \mu_x}{\sigma_x} \right) \left( \frac{y - \mu_y}{\sigma_y} \right) \right]$$

v.a. normalizate

obtinute prin translatie  
si scalare

Stim ca  $\mathbb{E}[(x + xy)^2] \geq 0; \forall x \in \mathbb{R}$

$\Leftrightarrow$

$$\mathbb{E}[x^2] + 2x\mathbb{E}[xy] + \mathbb{E}[y^2] \geq 0; \forall x$$

$$\Delta = 4\mathbb{E}[xy]^2 - 4\mathbb{E}[x^2]\mathbb{E}[y^2] \leq 0$$

$$\mathbb{E}[xy]^2 \leq \mathbb{E}[x^2]\mathbb{E}[y^2] !$$

$$\Rightarrow \mathbb{E}[xy]^2 \leq 1 \Rightarrow |f(x, y)|^2 \leq 1 \Rightarrow |f(x, y)| \leq 1$$

Ineg. Cauchy-Schwarz

$$\left( \sum_{i=1}^n a_i b_i \right)^2 \leq \left( \sum a_i^2 \right) \left( \sum b_i^2 \right)$$

$$|\langle x, y \rangle| \leq \|x\| \cdot \|y\|$$

Ineg. si teoreme limita

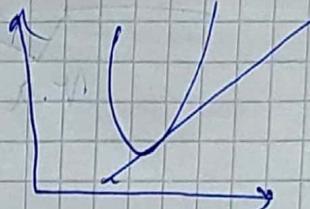
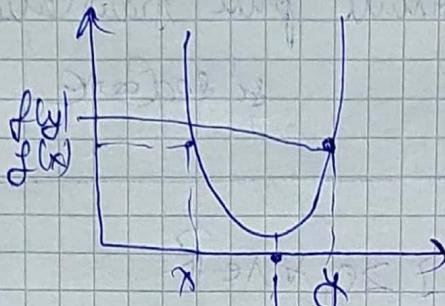
① Ineg. Cauchy-Schwarz

Fie  $x$  si  $y$  v.a. cu  $\text{Var}(x) < \infty$ ,  $\text{Var}(y) < \infty$ , atunci

$$|\mathbb{E}[xy]| \leq \sqrt{\mathbb{E}[x^2]\mathbb{E}[y^2]}$$

## ② Inequalidad de Jensen

### a) Fct. convexa



$$f(tx + (1-t)y) \leq tf(x) + (1-t)f(y)$$

$$f(tx + (1-t)y) \leq tf(x) + (1-t)f(y)$$

Fct. convexa:  $\forall x, y; \forall t \in [0, 1]$

$$f(tx + (1-t)y) \geq tf(x) + (1-t)f(y)$$

### ③ (Ineq. Jensen)

Sea  $X$  v.a. si  $g \circ f$  de convexa. Entonces

$$\mathbb{E}[g(X)] \leq g(\mathbb{E}[X])$$

Dado  $g$  es convexa:  $f[g(X)] \leq g(f(X))$

! Obs.  $\text{Var}(X) \geq 0$

$$\mathbb{E}[X^2] \geq \mathbb{E}[X]^2$$

### ④ (Ineq. Markov)

Sea  $X$  v.a. positiva. Entonces

$$\mathbb{P}(X \geq a) \leq \frac{\mathbb{E}[X]}{a}$$

$$\text{Defn. } y = \begin{cases} 0, & x < a \\ a, & x \geq a \end{cases}$$

$$\mathbb{E}[Y] = a \mathbb{P}(X \geq a)$$

$$Y \leq X \Rightarrow \mathbb{E}[Y] \leq \mathbb{E}[X]$$

Ex.:  $X \sim U([0, 1])$

$$\mathbb{P}(X > 2) \leq \frac{1}{4}$$

$$\mathbb{P}(X \geq 1) \leq \frac{1}{2}$$

$$\mathbb{P}\left(X \geq \frac{1}{2}\right) \leq 1$$

④ (Ineq. Chebychev)

Fie  $X$  v.a.  $\mathbb{E}[X] = \mu < \infty$ ;  $\text{Var}(X) = \sigma^2 < \infty$

$$\mathbb{P}(|X - \mu| \geq a) = \frac{\text{Var}(X)}{a^2}; \quad \forall a > 0$$

$$y = (x - \mu)^2$$

$$\mathbb{P}(Y \geq a^2) \leq \frac{\mathbb{E}[Y]}{a^2} = \frac{\text{Var}(X)}{a^2}$$

Obs.  $a = k\sigma$

$$\mathbb{P}(|X - \mu| \geq k\sigma) \leq \frac{1}{k^2}$$

⑤ (Ineq. Chernoff)

! f.p. (faute poterice)

Fie  $X$  v.a.  $a > 0, t > 0$

$$\mathbb{P}(X \geq a) \leq \frac{\mathbb{E}[e^{tX}]}{e^{ta}} = e^{-ta + at}$$