

cursor imp. → select.

cursor par. → ese turnee cu id.

cursor imbres. → premiu.

tipul
 \sum

~~Curs 14~~

Jinego Cifare

1) $E[|xy|] \leq \sqrt{E[x^2]E[y^2]}$: Cauchy - Schwartz

2) f convexă $E[\varphi(x)] \geq \varphi(E[x])$ Jensen
 f concavă $E[\varphi(x)] \leq \varphi(E[x])$

3) $x > 0, a > 0 \quad P(X > a) \leq \frac{E[X]}{a}$, Markov

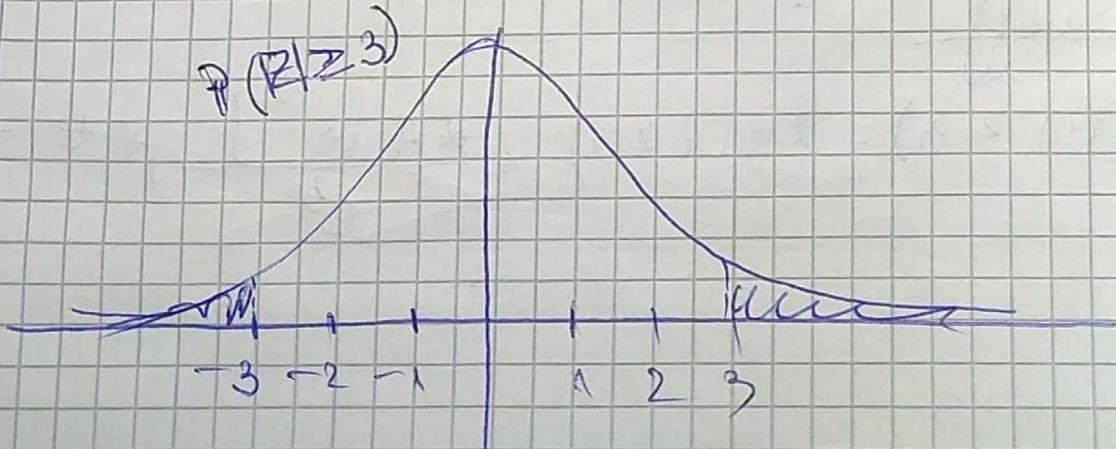
4) $X \sim \text{univ.} \quad E[X] = \mu, \text{Var}(X) = \sigma^2$
 $P(|X - \mu| \geq a) \leq \frac{\text{Var}(X)}{a^2}, \forall a > 0$ Chebyshev.

5) $X \sim \text{univ.}, a > 0, t > 0$ altăori

$$P(X \geq a) \leq \frac{E[e^{tX}]}{e^{ta}}, \forall t \text{ Chernoff}$$

Ex.: $Z \sim N(0,1)$. Vom să mărginim superior

$P(|Z| \geq 3)$ folosind Markov, Chebyshev, Chernoff



Anew prop. 68-95-99,7

$$P(|z| \leq 1) \approx 0,68$$

$$P(|z| \leq 2) \approx 0,95$$

$$P(|z| \leq 3) = 0,997$$

Aufg., aware $P(|z| \geq 3) \approx 0,003$

a) Markow

$$P(|z| \geq 3) \leq \frac{E|z|}{3} \leq \frac{0}{3} = 0,26$$

$$\begin{aligned} E[z] &= \int |z| \varphi(z) dz \\ &= \int |z| \cdot \frac{1}{\sqrt{2\pi}} \cdot e^{-\frac{z^2}{2}} dz \end{aligned}$$

$$= 2 \int_0^\infty \frac{z}{\sqrt{2\pi}} \cdot e^{-\frac{z^2}{2}} dz = \sqrt{\frac{2}{\pi}} \int_0^\infty z e^{-\frac{z^2}{2}} dz$$

$$= \sqrt{\frac{2}{\pi}} \left(-e^{-\frac{z^2}{2}} \right) \Big|_0^\infty = \sqrt{\frac{2}{\pi}}$$

b) Chebyshev

$$P(|z| \geq 3) = P(|z - 0| \geq 3) \leq \frac{\text{Var}(z)}{9} = \frac{1}{9} = 0,11$$

c) Chebyscheff

$$P(|z| \geq 3) = 2P(z \geq 3) \leq \frac{2E[e^{tz}]}{e^{6t}}, t > 0$$

$$\begin{aligned}
 \text{suchet } \{e^{tx}\} &= \int_{-\infty}^{\infty} e^{tx} \cdot \frac{1}{\sqrt{2\pi}} \cdot e^{-\frac{x^2}{2}} dx \\
 &= \frac{1}{\sqrt{2\pi}} \int e^{-\frac{x^2}{2} + tx} dx \\
 &= \frac{1}{\sqrt{2\pi}} \int e^{-\frac{1}{2}(x^2 - 2tx + t^2 + t^2)} dx \\
 &\sim \frac{1}{\sqrt{2\pi}} \int e^{-\frac{(x-t)^2}{2} + \frac{t^2}{2}} dx \\
 &= \frac{e^{\frac{t^2}{2}}}{\sqrt{2\pi}} \int e^{-\frac{(x-t)^2}{2}} dt = 1, \quad X \sim N(t, 1) \\
 &= e^{\frac{t^2}{2}}
 \end{aligned}$$

$$\rightarrow P(|Z| \geq 3) \leq \frac{2e^{\frac{t^2}{2}}}{e^{\frac{9t}{2}}}, \forall t$$

pt. t \geq 3 \text{ annehmen} \quad P(|Z| \geq 3) \leq 2e^{-\frac{9}{2}} = 0,02

Exp. X u.a; x \in [a, b]; E[x] = \mu; V(x) = \sigma^2

$$P(|x-\mu| \geq t) \leq \frac{\sigma^2}{t^2}, \forall t > 0$$

Wahrscheinlichkeit

$$\text{der approx. } \leq \frac{(b-a)^2}{4t^2}$$

var. e wäng. Intervall [a, b]

Vrem să arătăm:

$$\cancel{Q^2 \geq \frac{(b-a)^2}{4}}$$

$$Q^2 \leq \frac{(b-a)^2}{4}$$

$$g(\bar{x}) = E\{(x-\bar{x})^2\}$$

$$\bar{x} = E[x]$$

$$E\{(x-\bar{x})^2\} \geq E\{(x-E[x])^2\}; \forall \bar{x} \in \mathbb{R}$$

$$E\{(x-\bar{x})^2\} = E\left[\underbrace{(x-E[x])}_{=0} + \underbrace{\underbrace{(E[x]-\bar{x})}_{=0}}_{=0}\right]^2$$

$$= E\{(x-E[x])^2\} + 2 \underbrace{E[(x-E[x])(E[x]-\bar{x})]}_{=0} +$$

$$+ (E[x]-\bar{x})^2$$

$$! \geq E\{(x-E[x])^2\}; \forall \bar{x} \in \mathbb{R}$$

$$Q^2 = \text{Var}(x) = E\{(x-E[x])^2\} \leq E\{(x-\bar{x})^2\}, \forall \bar{x}$$

$$\text{Pf. } \bar{x} = \frac{a+b}{2}$$

$$! E\{(x-\frac{a+b}{2})^2\} = E\left[(x-a)(x-b) + \frac{(b-a)^2}{4}\right]$$
$$= E\{(x-a)(x-b)\} + \underbrace{\frac{(b-a)^2}{4}}_{\leq 0} \leq \underbrace{\frac{(b-a)^2}{4}}_{\leq 0}$$

Cum arăta?

$$(x-\bar{x})^2 = A + \frac{(b-a)^2}{4}$$

≤ 0
gasim ceea ce

pe \bar{x} ca să fie îndeplinită condiția

Teoreme Cunîjă. Xogea M., măr. (jumătate)

Def.: Fie $(X_n)_{n \geq 1}$ un sir de v.a. si X_0 v.a. peste (Ω, \mathcal{F}, P) .

În același mod, dacă $\{X_n\}_{n \in \mathbb{N}}$ este o succesiune de variabile aleatori și convergentă la X în sensul probabilistic, atunci există un număr natural N astfel încât pentru orice $n > N$, avem $P(X_n = X) = 1$.

$$\boxed{\mathbb{P} \left(\lim_n X_n = X \right) = 1}$$

exercice 1) $A = \{ w \in \Sigma^* \mid \lim_n X_n(w) = X(w) \}$

→ exp. pt. P \ O → com. apr. sign pt că prob. cui O e O (?)

Fie (X_n) un sucesión u.a. si X es u.a def.

Suntem că suntem X_n conv. în probabilitate, și
 notăm $X_n \xrightarrow{P} X$ dacă $\forall \varepsilon > 0$

$$\lim_{n \rightarrow \infty} P(|X_n - X| \geq \varepsilon) \geq 0$$

Obs., $\forall \varepsilon > 0$, $\exists \delta > 0$, $\exists n_0 \in \mathbb{N}$ s.t. $n \geq n_0$

$$\mathbb{P}(|X_n - x| \geq \varepsilon) \leq d$$

accuracy approximation

Ex. : $X_n \sim U([0,1])$ indep,

$$Y_{\text{min}} = \min \{x_1, x_2, \dots, x_n\}$$

$$A\overset{P}{\longrightarrow} 0$$

\rightarrow you converge in probability if

Pf. $\varepsilon > 0$

$$P(|Y_m - 0| \geq \varepsilon) \xrightarrow[n \rightarrow \infty]{} 0$$

$$P(|Y_m| \geq \varepsilon) = P(Y_m \geq \varepsilon) =$$

\downarrow
 ε pozitiv \Rightarrow surj. form. din var. poz
 $Y_m \geq 0$

$$= P(x_1 \geq \varepsilon, x_2 \geq \varepsilon, \dots, x_m \geq \varepsilon)$$

$$\stackrel{\text{indep.}}{=} P(x_1 \geq \varepsilon) P(x_2 \geq \varepsilon) \cdots P(x_m \geq \varepsilon)$$

$$= \underbrace{(1 - P(x_1 < \varepsilon))}_{\varepsilon} \underbrace{(1 - P(x_2 < \varepsilon))}_{\varepsilon} \cdots \underbrace{(1 - P(x_m < \varepsilon))}_{\varepsilon}$$

v. uniforme

$$(1 - \varepsilon)(1 - \varepsilon) \cdots (1 - \varepsilon) = (1 - \varepsilon)^m$$

$$P(|Y_m| \geq \varepsilon) = \underbrace{(1 - \varepsilon)^m}_{\downarrow n \rightarrow \infty} \text{ pt. } \varepsilon \in (0, 1) \quad \text{pt. } \varepsilon \in \mathbb{C}^*, \mathbb{D}$$

\downarrow
0.

Def. Numirea esantionului volumul unui pătrat

Q, v.a. X_1, X_2, \dots, X_n indep. și identic repartizat
nd. (i.i.d.) cu $P(X_i^{-1}) = Q$

Media esantionului $\Rightarrow \bar{X}_n = \frac{X_1 + X_2 + \dots + X_n}{n}$

Pp. X_1, X_2, \dots, X_n esantion de medie μ , și varianta σ^2
($E[X_i] = \mu$, $V(X) = \sigma^2$)

$$E[\bar{X}_n] = E\left\{\frac{X_1 + X_2 + \dots + X_n}{n}\right\} = \frac{1}{n}(E[X_1] +$$

$$+ E[X_2] + \dots + E[X_n]) = \mu$$

$$V_2(\bar{x}_n) = V_2\left(\frac{x_1 + x_2 + \dots + x_n}{n}\right) = \frac{1}{n^2} \left(V_2(x_1) + V_2(x_2) + \dots + V_2(x_n) \right) = \frac{\sigma^2}{n}$$

$\hookrightarrow n \cdot \sigma^2$

Abgeleitete Nr. 2:

(LNM) (stabsä)

Für (x_n) n unabh. d. v.a. i.i.d. en $E[x_1] = \mu < \infty$,
 $V_2(x_1) = \sigma^2 < \infty$

8) Aufgabe $x_n \xrightarrow{P} \mu$

$\sigma^2 < \infty$
 e. g. \lim

Versicherungstheorie

x_n v.a. iid. $E[|x_1|] < \infty$; $E[x_1] = \mu$

$\bar{x}_n \xrightarrow{a.s.} \mu$

Beweis: Pf. $\varepsilon > 0$

$$\Pr(|\bar{x}_n - \mu| \geq \varepsilon) \xrightarrow{n \rightarrow \infty} 0$$

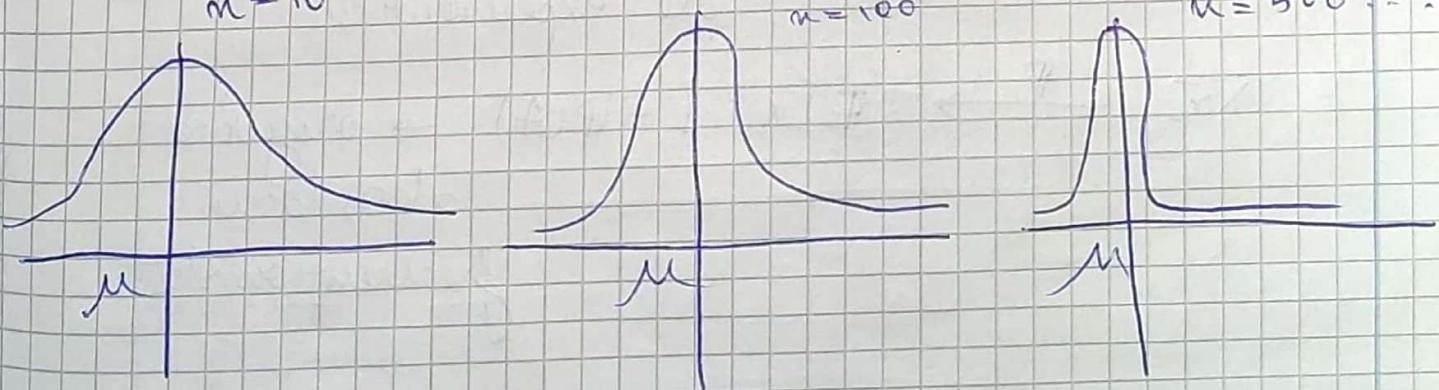
Frage: Chibyshev:

$$\Pr(|\bar{x}_n - \mu| \geq \varepsilon) \leq \frac{V_2(\bar{x}_n)}{\varepsilon^2} = \frac{\sigma^2}{n\varepsilon^2} \xrightarrow{n \rightarrow \infty} 0$$

$n = 10$

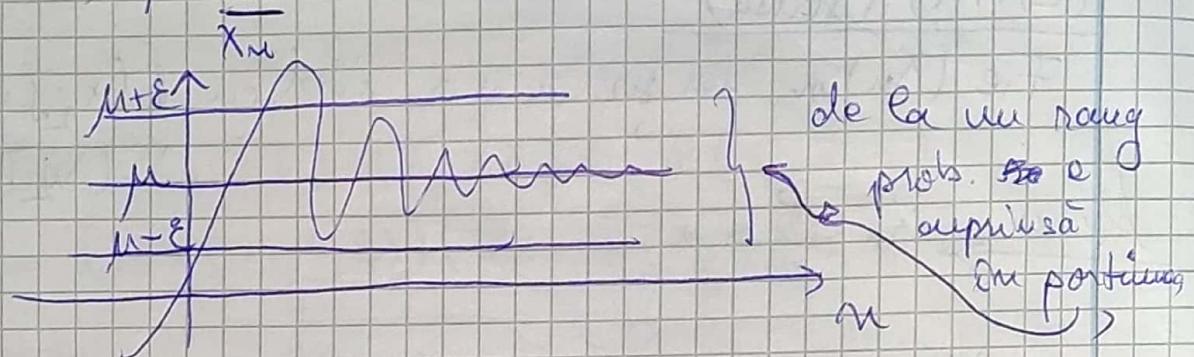
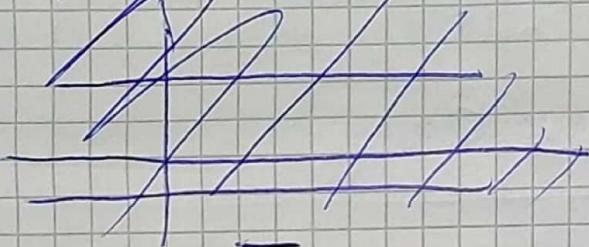
$n = 100$

$n = 500 \dots$



Pentru un nivel de acuratete slab

$$\overline{P} \left(\frac{(\mu - \epsilon, \mu + \epsilon)}{\cancel{\text{X}_n}} \rightarrow \overline{X_n} \right) \xrightarrow{n \rightarrow \infty} 1$$



$\exists p \cdot (\Sigma, f, P) \text{ c.f.}; A \in \mathbb{N}$

$$\text{Fie } x_i = \begin{cases} 1, & w^i \in A \\ 0, & \text{alif} \end{cases}$$

$$x_i \in B(p)$$

$$p = P(X_i = 1) = P(A)$$

$$\bar{x}_m = \frac{x_1 + x_2 + \dots + x_n}{m}$$

- frecvența relativă de apariție a unui A în n repetițiile experimentului

$$\overline{x_n} \xrightarrow{P} E[x_i] = P(A) \rightarrow \text{origina}$$

frequentist

Ex. Fie p procentul din populație care votăză pe A.

des esantionare stratificată (at. când alegeră persoane care le întră în pt. sondaj după diverse criterii)

acei → alegeră aleatoriu
 $X_1, X_2, \dots, X_m \sim B(p)$ indep.

$$\bar{X}_m = \frac{X_1 + X_2 + \dots + X_m}{m} = \text{procentul persoanelor care votăză pe A din totalul de } m.$$

→ Reg. Chebyshev și

$$P(|\bar{X}_m - p| \geq \varepsilon) \leq \frac{\text{Var}(\bar{X}_m)}{\varepsilon^2} = \frac{p(1-p)}{m\varepsilon^2}$$

Dar

p - necunoscut

dar putem spune că

$$\Rightarrow P(|\bar{X}_m - p| \geq \varepsilon) \leq \frac{1}{4m\varepsilon^2} = \frac{1}{4m\varepsilon^2} = p(1-p) \leq \frac{1}{4}$$

$\varepsilon = 0,01$ (acurătatea)

$\frac{1}{4m\varepsilon^2} = 0,05$ (nivel de incertitudine)

$\hookrightarrow m \approx 50000$

trebui să petreac multă lume

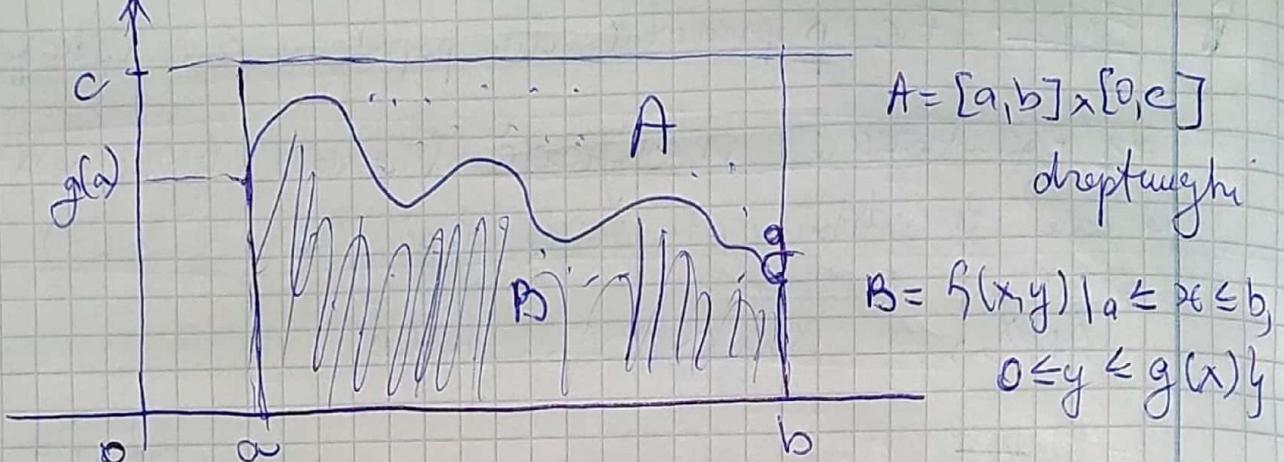
Integrarea Monte-Carlo

→ metode de aprox. a unor valori, integrările
fol. v.a. și legea lui Markov

Pp că avem o fct. g și neavă să calculăm

$$\int_a^b g(x) dx$$

Pp că pe $[a, b]$ avem $0 \leq g(x) \leq c$



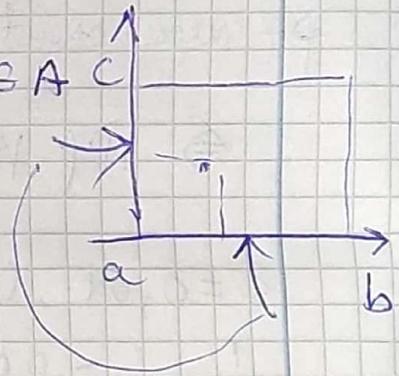
Generarea patr. Unif (A)

uniform repartitate ~~pe A~~
pe A

$$(x_1, y_1), \dots, (x_n, y_n) \sim U(A)$$

$$(x, y) \sim U(A)$$

$$f_{(x,y)}(x,y) = \begin{cases} \frac{1}{c(b-a)} & ; (x,y) \in A \\ 0 & \text{altfel} \end{cases}$$



$$A(A) = c(b-a)$$

$$x \sim U([a, b]) \quad \text{indep.}$$

$$y \sim U([0, c])$$

periferia e unif. pe
supr. desusă

$$f_x(x) = \frac{1}{b-a} \mathbb{1}_{[a,b]}(x)$$

$$f_y(y) = \frac{1}{c} \cdot \mathbb{1}_{[0,c]}(y)$$

$$f_{(x,y)}(x,y) = \text{indep. } f_x(x) f_y(y) = \frac{1}{c(b-a)} \mathbb{1}_A(x,y)$$

~~Z $\sim P(p)$~~

~~P(Z=1)~~

$$\text{Fie } z_i = \begin{cases} 1, & (x_i, y_i) \in B \\ 0, & \text{a. g. f.}, z_i \sim P(p) \end{cases}$$

$$p = P(z_i = 1) = P((x_i, y_i) \in B)$$

$$= \iint_B f_{(x,y)}(x,y) dx dy = \frac{A(B)}{A(A)}$$

$$\text{f in LNM; } z_i = \frac{z_1 + \dots + z_n}{n} \xrightarrow{P} p = \frac{A(B)}{c(b-a)}$$

$$= \frac{\int_a^b g(x) dx}{c(b-a)}$$

$$\boxed{\int_a^b g(x) dx = c(b-a) \cdot \bar{z}_n}$$

in 2:

Fie $U_1, U_2, \dots, U_n \sim U[a, b]$ i.i.d.

$X_1 = g(U_1); X_2 = g(U_2) \dots X_n = g(U_n)$ i.i.d.

LNM: $\bar{X}_n \xrightarrow{P} E[X_1] = E[g(U_1)]$

$$= \int g(x) f_{U_1}(x) dx$$

$$= \int g(x) \frac{1}{b-a} \mathbb{1}_{[a,b]}(x) dx$$

$$= \frac{1}{b-a} \int_a^b g(x) dx$$

$$\int_a^b g(x) dx = (b-a) \cdot \frac{x_1 + x_2 + \dots + x_n}{n}$$

[Klausur]

$$\int_a^b g(x) dx = (b-a) \cdot \frac{x_1 + x_2 + \dots + x_n}{m}$$

$$\int_a^b g(x) dx = (b-a) \cdot \frac{g(u_1) + \dots + g(u_m)}{m}$$

$$u_i \sim U(a, b)$$

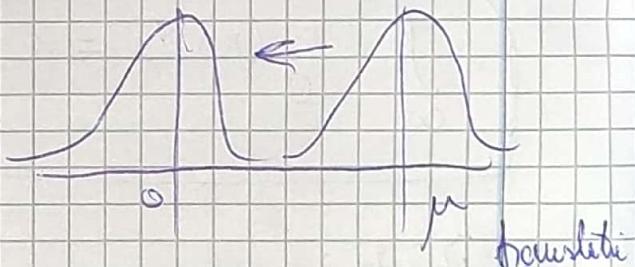
Th. limite Centrală

(TLC)

Din LNM: $\bar{x}_n \xrightarrow{P} E[x_i]$

Fie x_1, x_2, \dots, x_n iid $E[x_i] = \mu$, $V_n(x) = \sigma^2$

$$V_n\left(\frac{x - E[x]}{\sqrt{Var(x)}}\right) = 1$$



$$z = \frac{x - E[x]}{\sqrt{V_n(x)}} \quad \text{s.m. z-scor}$$

sau variabilă normalizată

$$z_n = \frac{x_1 + x_2 + \dots + x_n - E[x_1 + \dots + x_n]}{\sqrt{Var(x_1 + \dots + x_n)}}$$

$$= \frac{x_1 + \dots + x_n - n\mu}{\sqrt{n\sigma^2}}$$

variabilă de
foc / variabilă
normalizată

$$Z_m = \sqrt{m} \left(\frac{\bar{X}_m - \mu}{\sigma} \right)$$

T (Teorema limită Centrală)

Fie $(X_m)_{m \in \mathbb{N}}$ un sir de variabile aleatorii iid. cu $E[X_1] = \mu < \infty$ și $V[X_1] = \sigma^2 < \infty$. Atunci

$$\lim_{m \rightarrow \infty} P(Z_m \leq x) = \Phi(x), \forall x$$

$$\text{unde } Z_m = \sqrt{m} \cdot \frac{\bar{X}_m - \mu}{\sigma}, \quad \Phi(x) = \int_{-\infty}^x \frac{1}{\sqrt{2\pi}} e^{-\frac{t^2}{2}} dt$$

(fct. de rep. a $N(0,1)$) \rightarrow ajung la o normală

Obs. X_1, X_2, \dots, X_m iid., $E[X_1] = \mu$

$$S_m = X_1 + \dots + X_m$$

$$V[X_1] = \sigma^2$$

$$\begin{aligned} P(S_m \leq c) &= P\left(\frac{S_m - E[S_m]}{\sqrt{V[S_m]}} \leq \frac{c - E[S_m]}{\sqrt{V[S_m]}}\right) \\ &\stackrel{\text{TLC}}{=} \Phi\left(\frac{c - \mu m}{\sigma \sqrt{m}}\right) \end{aligned}$$

Pt. m suficient de mare

$$\begin{cases} S_m \sim N(m\mu, m\sigma^2) \\ \bar{X}_m \sim N(\mu, \frac{\sigma^2}{m}) \end{cases} \leftrightarrow \text{approximativ normal}$$

Ex . . 100 pachete
 $\sim \mathcal{U}([5, 50])$

Care e probabilitatea ca greutatea totală să fie ≥ 3000 kg?

Așa că,

Fie x_1, \dots, x_{100}

$$S_{100} = x_1 + \dots + x_{100}$$

pachete $\sim \mathcal{U}([5, 50])$

greutatea totală

$$\mathbb{P}(S_{100} \geq 3000) = ?$$

Răz.

$$\mathbb{P}(S_{100} \geq 3000) = \mathbb{P}\left(\frac{S_{100} - \mathbb{E}[S_{100}]}{\sqrt{\text{Var}(S_{100})}} \geq \frac{3000 - \mathbb{E}[S_{100}]}{\sqrt{\text{Var}(S_{100})}}\right)$$

$$Z_{100} \sim N(0,1)$$

$$\text{TLC} = 1 - \mathbb{P}\left(\frac{3000 - \mathbb{E}[S_{100}]}{\sqrt{\text{Var}(S_{100})}}\right)$$

$$\mathbb{E}[S_{100}] = 100 \cdot \mathbb{E}[x_i] = 100 \cdot \frac{55}{2} = 27,5 \cdot 100 = 2750$$

$$\text{Var}(S_{100}) = 100 \cdot \text{Var}(S_1) = 100 \cdot \frac{(50-5)^2}{12} = 16875$$

$$X \sim \mathcal{U}[a, b]$$

$$\text{Var}(X) = \frac{(b-a)^2}{12}$$

$$x = a + (b-a)y$$

$$Y \sim \mathcal{U}(0,1)$$

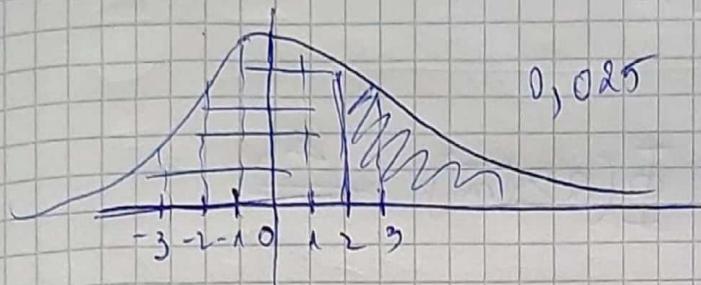
$$\mathbb{E}[Y] = \frac{1}{2}$$

$$\text{Var}(Y) = \frac{1}{3} - \frac{1}{4} = \frac{1}{12}$$

$$\mathbb{E}[Y^2] = \int_0^1 y^2 dy = \frac{1}{3}$$

$$P(L_{100} \geq 3000) \approx 1 - \Phi\left(\frac{3000 - 2750}{\sqrt{16875}}\right) = 0,0274$$

$$\cong 1 - \Phi(1,92)$$



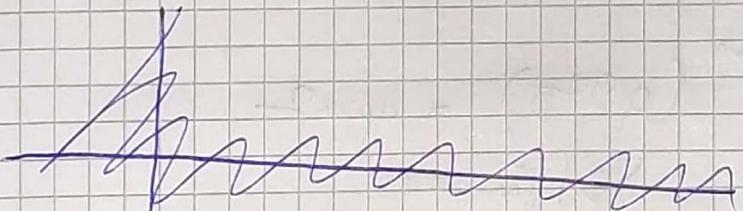
$\stackrel{?}{=} 1 - \text{probam.}(1,92)$

Ex. p pe A. în proporție populatie

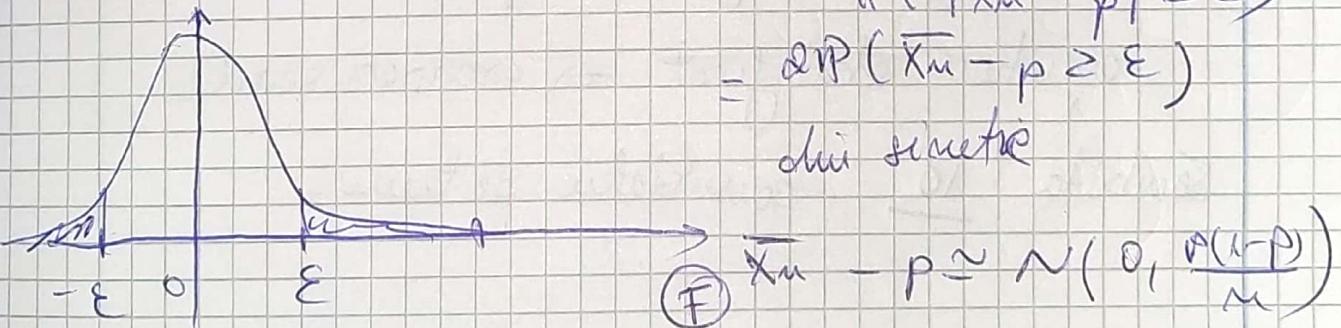
$$x_1, \dots, x_n \sim P(p)$$

$$(1) \bar{x}_n \sim N(\mu, \frac{\sigma^2}{n}) = N(p, \frac{p(1-p)}{n})$$

$$(2) \bar{x}_n - p \sim N(0, \frac{p(1-p)}{n})$$



$$\begin{aligned} & P(|\bar{x}_n - p| \geq \varepsilon) \\ & = 2P(\bar{x}_n - p \geq \varepsilon) \\ & \text{obișnuită} \end{aligned}$$



$$P(|\bar{x}_n - p| \geq \varepsilon) = 2P(\bar{x}_n - p \geq \varepsilon)$$

$$= 2P\left(\frac{\bar{x}_n - p}{\sqrt{\frac{p(1-p)}{n}}} \geq \frac{\varepsilon}{\sqrt{\frac{p(1-p)}{n}}}\right)$$

$$\begin{aligned} & \stackrel{(1)}{=} 2\left(1 - \Phi\left(\frac{\varepsilon}{\sqrt{\frac{p(1-p)}{n}}}\right)\right) \\ & \stackrel{(2)}{\rightarrow} 2\left(1 - \Phi\left(\frac{\varepsilon}{\sqrt{\frac{p(1-p)}{n}}}\right)\right) \leq 2\left(1 - \Phi\left(2\sqrt{\frac{p(1-p)}{n}}\right)\right) \end{aligned}$$

$$p(1-p) = \frac{1}{n} \Rightarrow p\frac{1-p}{n} = \frac{1}{n^2} \Rightarrow \sqrt{\frac{1}{P(1-p)}} \geq \sqrt{\frac{1}{n}}$$

$$\Rightarrow \frac{\varepsilon}{\sqrt{\frac{P(1-p)}{n}}} \geq 2\varepsilon\sqrt{n}$$

$$\Rightarrow \Phi\left(\frac{\varepsilon}{\sqrt{\frac{P(1-p)}{n}}}\right) \geq \Phi(2\varepsilon\sqrt{n}) \stackrel{*}{=}$$

$$P(|\bar{x}_n - p| \geq \varepsilon) \leq \underbrace{(1 - \Phi(2\varepsilon\sqrt{n}))}_{0,05}$$

$\varepsilon = 0,01$

$$\left. \begin{array}{l} 2(1 - \Phi(2\varepsilon\sqrt{n})) = 0,05 \\ \varepsilon = 0,01 \end{array} \right\} \Phi(2 \times 0,01\sqrt{n}) = 0,975$$

$$\Rightarrow 2 \times 0,01\sqrt{n} = \Phi^{-1}(0,975) = \text{approx } 2$$

$\rightarrow \begin{cases} n \approx 100 \\ n \geq 10000 \text{ personen} \end{cases}$

examination stratified \Rightarrow increased sample size

sample size 10 consultation per team