A Continuous Paradoxical Colouring Rule Using Group Action

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Abstract

Given a probability space $(X, \mathcal{B}, m)$, measure preserving transformations $g_1, \ldots, g_k$ of X, and a colour set C, a colouring rule is a way to colour the space with C such that the colours allowed for a point x are determined by that point's location and the colours of the finitely $g_1(x), \ldots, g_k(x)$ with $g_i(x) \neq x$ for all i and almost all x. We represent a colouring rule as a correspondence F defined on $X \times C^k$ with values in C. A function $f : X \to C$ satisfies the rule at x if $f(x) \in F(x, f(g_1 x), \ldots, f(g_k x))$. A colouring rule is paradoxical if it can be satisfied in some way almost everywhere with respect to m, but not in **any** way that is measurable with respect to a finitely additive measure that extends the probability measure m and for which the finitely many transformations $g_1, \ldots, g_k$ remain measure preserving. We show that a colouring rule can be paradoxical when the $g_1, \ldots, g_k$ are members of a group G, the probability space X and the colour set C are compact sets, C is convex and finite dimensional, and the colouring rule says if $c : X \to C$ is the colouring function then the colour $c(x)$ must lie (m a.e.) in $F(x, c(g_1(x)), \ldots, c(g_k(x)))$ for a non-empty upper-semi-continuous convex-valued correspondence F defined on $X \times C^k$. We show that any colouring that approximates the correspondence by ϵ for small enough positive ϵ cannot be measurable in the same finitely additive way. Furthermore any function satisfying the colouring rule illustrates a paradox through finitely many measure preserving shifts defining injective maps from the whole space to subsets of measure summing up to less than one.

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1 Introduction

A common belief is that measure theoretic paradoxes, like the Banach-Tarski paradox, are not relevant to real events. The reasoning is that such paradoxes require the Axiom of Choice (AC), and therefore the exhibition of paradoxical behaviour would negate the fact that we can remain logically consistent after rejecting AC (Note however that there is a class of theorems called shadows of AC which proofs use AC and exist even after rejecting AC; see [3] for the details.)

There are two parts to showing that a decomposition is paradoxical, a part showing that it cannot be done in a measurable way and a part showing its existence. It is the second part that requires some variation of AC. We don't employ AC for the first part.

In [6], we considered colouring rules such that the allowed colours for a point are determined by location and the colours of finitely many of the point's neighbours in a graph. The adjacency relation is defined by measure preserving transformations and the finitely many transformations of a point x are called its *descendants*. A colouring rule is paradoxical if it can be satisfied in some way almost everywhere, but not in any way that is measurable with respect to a finitely additive measure for which the transformations defining the descendants remain measure preserving. We demonstrated several paradoxical colouring rules and proved that if the measure preserving transformations belong to a group and there are finitely many colour classes then any colouring of a paradoxical colouring rule has colour classes that jointly with the measure preserving transformations and the Borel sets define a *measurably G-paradoxical decomposition* of the probability space, by which we mean the existence of two measurable sets of different measures that are G-equidecomposable (see [6], Thm. 1).

In the conclusion of [6], we asked whether a colouring rule could be paradoxical if the colour classes belonged to a finite dimensional convex set and the colouring rule was defined by an upper-semi-continuous convex-valued non-empty correspondence. We call such colouring rules *probabilistic* colouring rules. In [7], we demonstrated a paradoxical probabilistic colouring rule where some of the measure preserving transformations were non-invertible, hence do not belong to a group. In the present paper, we give an example where they do belong to a group. Invertibility of the transformations defining the descendants presents special challenges. To show that a colouring rule is paradoxical it is advantageous to show that satisfaction of the rule is done with extremal points (of colours) almost everywhere. It is easier to do this if the transformations are uncountable-to-one. We do it below with transformations that are one-to-one.

One interpretation of the upper-semi-continuous convex valued correspondence is that the choosing of colours is according to a maximisation or minimisation of a continuous and affine evaluation of options, with indifference between two options implying indifference between all of their convex combinations. This is such an example.

Throughout this paper, by a *proper* finitely additive extension we mean a finitely additive measure that extends the Borel measure and is invariant with respect to the measure preserving actions used.

We use an inclusive concept of approximation. A solution is ϵ-stable if the expected gains from deviation, evaluated at each point individually and then integrated over the space, are no more than ϵ. We show that every ϵ-stable solution for small enough ϵ cannot be measurable with respect to any proper finitely additive extension.

We were inspired optimistically by an example from [6]. A brief description of that example follows.

Consider $G = \mathbf{F}_2$ (the group generated freely by two generators) and the space $X = \{-1, +1\}^G$ acted on by G in the canonical way (through shifting). Let g_1, g_2 be the free generators of G. Where the e coordinate of an $x \in X$ lies in $\{-1, +1\}$ determines whether an arrow from x should be directed toward the choice of either $g_1 x$ or $g_2 x$ (if $x^e = +1$) or rather toward the choice of either $g_1^{-1} x$ or $g_2^{-1} x$ (if $x^e = -1$). If a point in the space X has two or more arrows directed to it, it is *congested*; if not, it is *uncongested*. The rule is to direct an arrow, if possible, toward a point that is uncongested (if not possible or possible in both directions, the rule allows the arrow to be directed in either direction). We showed that if almost all points follow this rule, the set of congested points is a subset of measure zero. Let the degree of x be the total number of potential inward arrows toward x, as determined by the choice of -1 or $+1$ for the e coordinate of its four neighbours by the four directions $g_1, g_2, g_1^{-1}, g_2^{-1}$. The argument, that the set of congested points is of measure zero when the rule is satisfied, was combinatorial. The structure supporting a congested point requires a continuous repetition of three or four degree points in an infinite chain of vertices. Because the average degree of a vertex is two, such a structure is restricted to a subset of measure zero. This implied that the rule was paradoxical, because in at least $\frac{1}{16}$ of the space there was no possibility of any arrow moving inward (and that probability is slightly higher than $\frac{1}{16}$, due to some configurations of two or more points with no possibility of arrows coming in toward that configuration). From every point in the space there was one arrow going outwards, but going inwards on the average there could be at most $\frac{15}{16}$ (assuming a measurable structure). This approach followed an inspiration from combinatorics, inclusion-exclusion. It is an example of a probability space $(X, \mathcal{B}, \mu)$ partitioned into n parts $A_1, \ldots, A_n$ such that after moving these parts by invertible measurable preserving transformations $\sigma_1, \ldots, \sigma_n$, we have $\mu(X \setminus (\cup_{i=1}^n \sigma_i A_i)) > 0$ and for every $i \neq j$ $\mu(\sigma_i A_i \cap \sigma_j A_j) = 0$. By inclusion-exclusion this is a contradiction to measurability.

One could try to repeat the argument with a weighted choice between the two options (meaning some distribution (p_1, p_2) with $0 \leq p_1 \leq 1$ and $p_1 + p_2 = 1$), and devise some definition of what is means to be congested and how congested. With a rule that requires pointing the arrow to the less congested option, one could hope to reproduce a situation where the average weight of received arrows would be strictly less than 1. With discrete choices there is no option between 1 arrow coming in and 2 arrows coming in. With weighted choices, however, using the critical weight of $1 + \frac{1}{16}$ to defined crowded, there are too many ways to distribute weights so that crowdedness appears throughout typical infinite chains of connected vertices. We kept the same idea that the differences between the degrees of vertices is a useful stochastic process, but we had to look for more sophisticated ways to determine

how weights should be distributed.

Our idea was that there should be two kinds of crowdedness, a crowdedness at a point receiving weights, what we call *passive pain*, and another kind of crowdedness at the location from where weights come, what we call *active pain*. The idea was that if x could direct some weight to y it is not the crowdedness at y, the passive pain, that counts at x but the total weight sent from x to y times that passive pain. This has the effect of distributing the weights more evenly. It also relates the pain levels to entropy type inequalities, critical to the proof. Note that the colouring rule for some x cannot be determined in any way by the colour of x, however we incorporate into the colour of y a variable that reflects the weight coming from x. Rather than a combinatorial argument as before, we show that pain, both active and passive, almost everywhere has to increase on the average along an infinite chain. As there is a finite upper limit to the level of both kinds of pain, that would mean that pain can be sustained only in a set of measure zero.

We discovered that two choices for distributing weights was not enough. Though the group generated freely by two independent elements has arbitrarily many independent choices, to keep the structure simple we equated generators with choices. With n generators, as before, -1 for the e coordinate of x means that weight can be sent only in the negative directions (to $g_1^{-1}x, \ldots, g_n^{-1}x$) and $+1$ means that weight can be sent only in the positive directions (to $g_1^{+1}x, \ldots, g_n^{+1}x$). Define the degree of a point y in X to be the number of directions from which weight can be be sent toward y. With n generators, the degree ranges from 0 to $2n$ with an average of n. The distribution of degrees is determined by the binomial expansion.

No matter how many generators were used, there is always a possibility for both passive and active pain to decrease. As a general rule, smaller degrees in a chain means an increase in pain, larger degrees a decrease of pain. The break-even degree is exactly n. Conditioned on having reached a point with some edge, adding one for that connection, the average degree is $n+\frac{1}{2}$. The product rule for determining active pain means that when the degree is $n-1$ the resulting increase of pain overweighs the resulting decrease of pain when the degree is $n+1$. The is because there is a reciprocal relationship between the the weight sent to a vertex and the active pain level, and these weights relate to the degree of that vertex. If a weight of p is sent from x to a vertex y with a passive pain level of r, the active pain level at x by choosing p is rp. On the other hand, if a weight of q is sent to a different vertex z with passive pain level w, and the active pain levels equate, we have $rp = qw$ or $w = \frac{rp}{q}$. A closely related analogy is the fact that $\frac{1}{n-1} + \frac{1}{n+1} > \frac{2}{n}$ and $\frac{1}{n-1} \cdot \frac{1}{n+1} > \frac{1}{n^2}$; the effect is stronger for $\frac{1}{n-k}$ and $\frac{1}{n+k}$ when $1 < k < n$. But to exploit this influence sufficiently we needed a variety of possible degrees below the average. We found success at $n = 5$, after failing at $n = 2$ and $n = 3$. At $n = 4$, some preliminary work suggested some difficulty in formulating a proof. We suspect that it can be done with $n = 4$, but not as nicely as with $n = 5$.

In the next section we describe the colouring rule. In the third section we show that this colouring rule is paradoxical, given a stochastic structure and analysis. In the fourth section, we present a computer program confirming that stochastic analysis. In the fifth section we apply the colouring rule to a problem of local optimisation and show that solutions for small enough approximations of the colouring

rule cannot be measurable with respect to any proper finitely additive extension. In conclusion we consider related problems.

2 A Probabilistic Colouring Rule

Let G be the group freely generated by T_1, T_2, T_3, T_4, T_5, and let $X = \{-1, 1\}^G$. For any $x \in X$ and $g \in G$, x^g stands for the g coordinate in x. With e the identity in G, the e coordinate of x is x^e. There is a canonical right group action on X, namely $g(x)^h = x^{gh}$ for every $g, h \in G$. We use the canonical product topology and probability measure, that giving 2^{-n} for every cylinder determined by particular choices of -1 or $+1$ for any n distinct group elements. With this Borel probability measure the group G is measure preserving.

Definition 1. *The* graph *of X is the directed subgraph of the orbit graph of the action of G on X induced by the edge subset*

$$\{(x, T_i x) \mid x^e = +1\} \cup \{(x, T_i^{-1} x) \mid x^e = -1\}.$$

We orient the graph of X by placing arrows from x to all five of the $T_i(x)$ if $x^e = +1$ and arrows from x pointed to all five of the $T_i^{-1}(x)$ if $x^e = -1$.

The subset of X where G does not act freely has measure zero. Without loss of generality, we will be interested only in those orbits of G and connected components of the graph of X where G acts freely.

Definition 2. *We define $S(y) := \{x \mid y = T_i^{x^e}(x)\}$ and define $|S(y)|$ to be the degree of y (the number of neighbours in the graph with arrows pointed to y).*

The graph of X involves two independent structures of arrows. Each point x has a passive and active role, an active role in one structure and a passive role in the other. The active and passive roles alternate. We are interested in that alternation, moving from a point in its passive role to its neighbours in their active roles, and from a point in its active role to its neighbours in their passive roles. Every point has an active role, namely a connection to five different points in their passive roles. The degree of a point concerns its passive role. Not every point has a passive role, meaning that they are of degree zero. The points of degree zero play indirectly a key role in the main argument.

Every point $x \in X$ has a colour in $\Delta(\{1, 2, 3, 4, 5\}) \times ([0, 1] \times_{z \in S(x)} [0, 1]_z)$ where the dimension of the last part of the colouring is equal to the degree of x. The first part, $\Delta(\{1, 2, 3, 4, 5\})$, a four-dimensional simplex, we call the *active* part of the colour. The $[0, 1] \times_{z \in S(x)} [0, 1]_z$ part we call the *passive* part of the colour.

What is the colouring rule, which we call **Q**?

Usually the word "cost" is used to describe a function that should be minimised. With this example, we prefer the word "pain", because it represents a situation that should be avoided. For every direction $i \in \{1, 2, 3, 4, 5\}$ we define the *active pain* for x to be $v_i \cdot r_x$ where v_i is the first coordinate of the passive colour of $y = T_i^{x^e}(x)$ and r_x is the coordinate corresponding to $x \in S(y)$ in the passive colour of $y = T_i^{x^e}(x)$. The rule for the active colour of x is to choose those directions

where the active pain is minimised. If more than one are minimal, then any convex combination of the minimal directions is allowed. The quantity of the active colour of x given to the i coordinate (in the direction of T_i or T_i^{-1}) is called the *weight* given in the direction i or toward $y = T_i^{x^e}(x)$.

The first part of the passive colouring is called the *passive pain*. The rule for the first part of the passive colouring is as follows. Whenever the sum of the active colours in $S(y)$ moving toward y is less than $1 + 2^{-11}$, then $v = 0$ is required by the rule $\mathbf{Q}$ for the first passive coordinate. If that sum is more than $1 + 2^{-11}$, then $v = 1$ is required by the rule $\mathbf{Q}$. And if the sum is exactly $1 + 2^{-11}$ then any value in $[0, 1]$ is acceptable for v.

The rule for the $x \in S(y)$ coordinate of the passive colour is very simple, it is the copy of the i coordinate of the active part of the point x pointed toward y such that $y = T_i^{x^e}(x)$.

It is now clear from the colouring rule $\mathbf{Q}$, why the T_i and their inverses should remain measure preserving with any finitely additive extension. A critical aspect of the colouring rule uses that from any y all the x such that $x \to y$ are treated equally, e.g. their weights are summed without prejudice. The same holds for the active part of the colour, that each of the five directions are treated equally.

First, it is easy to show, with AC (for uncountable families of sets as there are uncountably many group orbits), that there is some non-measurable solution to the rule $\mathbf{Q}$ valid almost everywhere. By AC we can choose in each orbit where G acts freely a special point x to correspond to $e \in G$. Classify each point y in the orbit containing x according to the length of the word in G needed to move from x to y. Choose a direction from y that involves a word of one length greater and allowed by the coordinate y^e. As there is only one possibility for a direction from y corresponding to a word of one length less (and no such possibility if $y = x$), there will be always an option to satisfy this requirement. Because one always chooses an arrow from a point with a shorter word to one with a longer word, it is not possible for two chosen arrows to be aimed toward the same point. The end result will be a colouring satisfying the rule $\mathbf{Q}$ where there is no pain, passive or active.

Later, we show that if the rule $\mathbf{Q}$ is satisfied then all points y where the weight sent to y is greater than $1 + \frac{1}{10^{11}}$ is contained in a Borel set of measure zero. In this way a kind of paradox is witnessed by the active part of **any** colouring satisfying the rule without any additional application of theory. The active part of the colouring can be seen as a distribution of 1 in ten different directions such that at least $\frac{1}{2^{10}}$ of the space receives no weight at all. We show, however, that outside of a set of Borel measure zero, no point receives weight more than $1 + \frac{1}{2^{11}}$. This can be seen as a kind of paradox, a measure preserving flow where the flow out (1) is greater than the flow in (no more than $(1 - \frac{1}{2^{10}})(1 + \frac{1}{2^{11}})$). Furthermore we will show that the same kind of paradoxical behaviour holds if the rule $\mathbf{Q}$ is followed to a sufficiently small approximation.

Given what will be proven later, we show that any such colouring generates a measurable G-paradoxical decomposition. From [6], we proved it suffices to have a finite partition of X (generated from the colouring, actions of the group, and the Borel sets) for which no proper finitely additive extension can make all partition members measurable. First approximate the weights in all directions by integer

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multiples of $\frac{1}{N}$ that add up to 1 (according to the N different intervals of values between 0 and 1 and the ten different directions) so that the whole space is broken into finitely many parts according to the values given to the ten different directions of the colouring (toward the T_i if $x^e = +1$ and toward the T_i^{-1} if $x^e = -1$). If N is large enough, the N copies are shifted in this way (and invariant measurability is assumed), the total weight coming into the vertices will remain less than the total weight coming out of the vertices. After defining N different partitions from this, at least one of them cannot have a proper finitely additive extension for which all partition members are measurable. Hence by [6] we can generate from this partition two Borel measurable sets of different measure that are G-equi-decomposable.

3 Paradoxical Colouring

The goal of this section is to complete the proof of the following theorem:

Theorem 1. *The colouring rule* **Q** *is paradoxical.*

We have shown already that there is some way to satisfy the rule. To complete the proof of Theorem 1, we will show that satisfaction of the colouring rule implies that the set where the passive and active pain is positive is contained in a Borel subset of measure zero. To show that **Q** is paradoxical, it suffices to show that the set of points where the passive pain is equal to 1 is contained in a set of measure strictly less than $\frac{2^{-11}}{9}$. That would be enough to show that the average weight moving inward toward all points is strictly less than

$$(1 - \frac{2^{-11}}{9} - 2^{-10})(1 + 2^{-11}) + 10\frac{2^{-11}}{9} = 1 - 2^{-21} - \frac{2^{-22}}{9}.$$

3.1 Chains

Definition 3. *Given that* $x \to y$, *meaning that* $y = T_i x$ *if* $x^e = +1$ *or* $y = T_i^{-1}x$ *if* $x^e = -1$, *the* chain *generated by* $x \to y$ *are all the points* z *in the graph of* X *that can be reached from* x *without going through* y *and involve alternating arrows, meaning that if* z *is an odd distance from* x *then the passive role* $z \leftarrow$ *connects* z *to* x *and if* z *is an even distance from* x *then the active role* $z \to$ *connects* z *to* x.

A chain involves an alternating process of using the active and passive colouring functions of the points. If $x \to y$ then x distributes weights to four other points z_1, z_2, z_3, z_4. In turn, depending on their degrees, there are further directed edges $x^* \to z_i$ (or in the rare case that the degree of each z_i is 1, no further directed edges). One could see a chain as a quarter of a G orbit; the choice of seeing y as passive or active, and the choice of moving in the x direction rather than in the direction of the potentially other x^* with $x^* \to y$.

Definition 4. *A* 0-level terminating point *of a chain generated by* $x \to y$ *is any vertex of degree* 1 *of positive even distance from* y. *A* 1st level terminating point *of the chain generated by* $x \to y$ *is some vertex* x^* *of the chain such that* $x^* \to z$ *is a directed edge of the chain,* $z \neq y$, *and* z *is terminating of level* 0 *(equivalently*

has degree 1). If $i \geq 2$ is even then an i-level terminating point of the chain is some vertex z such x^* is a terminating point of $i - 1$ or less for every x^* such that $x^* \to z$, x^* is further from y than z, and furthermore there is at least one such x^* that is a $(i - 1)$-level terminating point. If $i \geq 3$ is odd then an ith level terminating point of a chain is some vertex x^* of the chain such that $x^* \to z$ is the last step in the alternating path from y to z and z is a terminating point of level $i - 1$. A terminating point is a vertex that is a terminating point of some level. A chain generated by $x \to y$ is terminating if x is a terminating point, and its terminating level is the terminating level of x. If the chain generated by $x \to y$ is not terminating, then we say that the chain and the edge $x \to y$ is non-terminating. The non-terminating part of a non-terminating chain is the non-terminating chain with its terminating points removed.

Remark: A subchain of a terminating chain may not be a terminating chain. A subchain of a non-terminating chain may be a terminating chain. We could have a terminating chain generated by $x \to y$ with $x \to z_1$, x terminating of level 1, z_1 terminating of level 0, and $x \to z_2$ with $x^* \to z_2$ generating a non-terminating chain for some $x^* \neq x$. Likewise $x \to y$ could be non-terminating, $x \to z$ with z of degree three or more, with $x^* \to z$ generating a terminating chain for some $x^* \neq x$. Likewise if $x \to y$ is terminating it does not imply that y is a terminating point for all chains it may belong to, as there could be at least two edges $x^* \to y$ with $x^* \neq x$ such that $x^* \to y$ is not a terminating chain.

Lemma 1. *If $x \to y$ generates a terminating chain, then in any colouring of X that satisfies the colouring rule **Q**, the active pain level at x is zero, meaning that if the passive pain level at y is positive then no weight is given by x toward y.*

Proof. We prove the lemma by induction on the level of the terminating points; we claim that any terminating point of even level experiences no passive pain and any terminating point of odd level experiences no active pain. Suppose that $x^* \to y^*$ is part of the chain and y^* is terminating of level 0. As y^* is degree one, it is not possible for y^* to experience passive pain, since the maximal weight sent to y^* is at most 1. By choosing all weight to y^* there is no resulting active pain, and since the rule **Q** requires that x^* minimise the active pain, x^* cannot experience active pain in any direction. We notice that such an x^* is a terminating point of level 1. We continue with the induction assumption. Suppose y^* is a terminating point of even level i with x^* the vertex such that $x^* \to y^*$ and the path from x to y^* passes through x^*. If y^* was experiencing any passive pain, all the points $\hat{x} \to y^*$ such that $\hat{x} \neq x^*$ would give zero weight to y^*, since by induction they all experience no active pain. And with only one vertex x^* possibly giving weight to y^*, it is impossible for y^* to experience any passive pain, a contradiction. But then x^* does not experience any active pain either, because it could put all weight toward y^*, with the result of no active pain. ■

We describe a structure essential to our following stochastic arguments. Instead of looking at some x according to its topological location in X, we think of x as a member of a chain. Given that x sends the weight $p > 0$ to y with $t > 0$ the passive pain level at y, we consider what the passive pain levels t_i must be at the

four z_i with $x \to z_i$ for all $i = 1, 2, 3, 4$ so that the active pain levels for each of the five choices are equal (through satisfying the rule **Q**). If p_i is the weight sent from x to z_i, we must have the equations $p_i t_i = pt$ for each $i = 1, 2, 3, 4$. With the equations $t_i = \frac{pt}{p_i}$, we also consider the degrees of the z_i and how these quantities continue in further branches in the chain generated by $x \to y$. We want to show that almost everywhere, given $t > 0$, the pain values in further stages must be unbounded. Since these values cannot exceed 1, we have shown that positive passive pain happens only in a subset of measure zero. It is a kind of reverse engineering, determining what pain values must exist as implied by the rule **Q**. In this analysis we do not focus on directional choices determined by the e coordinate; we look instead on the degrees of the vertices of odd distance to x (even distance from y). We use that the probability of degree k is $\binom{9}{k-1} 2^{-9}$. It does not follow the formula $\binom{10}{k} 2^{-10}$ because we condition on the existence of a particular edge coming toward the vertex.

An important first step toward the main argument is to eliminate all terminating chains from the analysis and look at only non-terminating chains and their non-terminating parts. The following Lemma does this.

Lemma 2. *The probability that $x \to y$ generates a non-terminating chain is approximately $\hat{q} = .991603$.*

Proof. Because of the homogeneous structure to the space, there is a recursive formula for the value of $\hat{q}$. Given that $x \to z$ and $z \neq y$ and $x^* \to z$ with $x^* \neq x$, $\hat{q}$ is also the probability that x^* is not a terminating point of the chain. The probability that z is a terminating point is $(1 - \frac{\hat{q}}{2})^9$, hence the probability of it not being a terminating point is $1 - (1 - \frac{\hat{q}}{2})^9$. For x to be a non-terminating point each of the four such z must fail to be terminating points. Therefore $\hat{q}$ is the root of the polynomial $q = (1 - (1 - \frac{q}{2})^9)^4$. Applying Wolfram Alpha, the largest root of this polynomial strictly less than 1 is approximately $\hat{q} = .991603$. ∎

By its definition, we notice that $\hat{q} = \sum_{1 \leq j_i \leq 9} \prod_{i=1}^4 \binom{9}{j_i} (\frac{\hat{q}}{2})^{j_1} (1 - \frac{\hat{q}}{2})^{9-j_i}$.

In what follows, we will assume that all terminating points are removed so that the probability distribution on the degrees in a chain follow the binomial expansion applied to $\frac{\hat{q}}{2}$ and $1 - \frac{\hat{q}}{2}$ (instead of $\frac{1}{2}$ and $\frac{1}{2}$) and then conditioned to the probability $\hat{q}$. Because a terminating point is a terminating point of some finite level, and the terminating points of a fixed finite level define an open set, the non-terminating points form a closed subset of X. Likewise the space of non-terminating chains is a compact space with a probability distribution determined by the special value of $\hat{q}$. Along with a choice for a weight of p given by some x to y at the start of the chain, this topology defines a collection of Borel sets on the space of chains.

Definition 5. *Given a non-terminating chain generated by $x \to y$, any $p \in (0, 1)$, and any colouring c of that chain satisfying the rule **Q** with p the weight of x given to y and 1 the passive pain at y, define $t(x \to y, c, p)$ to be the supremum of the active pain in the colouring c in that chain.*

Because we are concerned with the ratio to the quantity 1, e.g., the passive pain at y could be some very small positive v, we allow for values above 1, although

strictly speaking there can never be pain, passive or active, above the level of 1. By definition, $t(x \to y, c, p)$ is no less than the active pain at x, which is p.

Definition 6. *Define $u(x \to y, p)$ to be the infimum of $t(x \to y, c, p)$ over all the colourings c satisfying the rule* **Q**. *The value $u(x \to y, p)$ is called the* chain minimum *with respect to p. This value can be infinite and we will show that for all positive p it is almost everywhere infinite.*

It is straightforward that $u(x \to y, p)$ is increasing in p.

Proposition 1. *The function $u(x \to y, p)$ is infinite for all $p > 0$ and almost all $x \to y$.*

Proposition 1 implies Theorem 1. As both passive and active pain cannot exceed 1, Proposition 1 implies that the places where the pain, passive or active, is positive is a set of measure zero. To prove Theorem 1, it suffices to show this for all $p > \frac{1}{10}$, as any point with passive pain 1 must be next to some vertex pointed toward it with weight more than $\frac{1}{10}$.

Let's look at the chain generated by $x \to y$, where $x \to z_i$ for $i = 1, 2, 3, 4$ and x^* is another vertex where $x^* \to z_i \neq y$ for some i. Assume that p is the weight given by x toward y. The function $u(x \to y, p)$ is discovered by equalising the $u(x^* \to z_i, p^*) \cdot v_i$ over all the choices for weights q_i from the $x^* \neq x$ to the various z_i and by the weights p^* sent to the z_i from the x^* and the corresponding induced passive pain levels v_i for the z_i following the equality $v_i q_i = p$. Notice that some $u(x^* \to z_i, p^*)$ could be infinite, in which case the p^* could be zero. Either p is strictly more than this common equal value, in which case $p = u(x \to y, p)$, or p is less than or equal to this common equal value, in which case $u(x \to y, p)$ is that common value. This follows by monotonicity, that dividing by q is strictly decreasing in q and that $u(x \to y, p)$ is strictly increasing in p. The same holds true for all the $x^* \to y^*$ further in the chain. This means that the $u(x^* \to y^*, p^*)$ times the induced passive pain levels, when these values are finite, form a super-martingale through the minimising process (with potentially decreasing future values).

Definition 7. *Given that $u(x \to y, p)$ is finite, we define the colouring resulting from finding equality at each step (as described above) the* chain minimiser. *This holds if the minimum is reached with p at the start or any other location on the chain, with the chain minimiser finding the minimum for the following part of the chain. If $u(x \to y, p)$ is infinite, then the chain minimiser is the result of the same process of finding equality, but with p^* replacing $u(x^* \to y^*, p^*)$ in the calculations.*

Lemma 3. *The value of $u(x \to y, p)$ and the chain minimiser used to define it are Borel measurable functions of the future degrees and p.*

Proof. For every i let $t_i(x \to y, c, p)$ be the maximum active pain of any $x^* \to y^*$ of distance i or less from $x \to y$ with the colouring c and p the weight given by x to y. Let $u_i(x \to y, p)$ be the corresponding minimum over the various c. As a function of p, the degrees, and c, the t_i is continuous. Notice that $u_i(x \to y, p)$ are non-decreasing functions and is always less than or equal to $u(x \to p)$ for every i. Notice also that if there were a sequence of p_i converging to p and the sequence

of $u_i(x \to y, p_i)$ were converging to a value strictly less than $u(x \to y, p)$ then the colourings c_i associated with these solutions would have a subsequence converging pointwise to some colouring c with $t(x \to y, c, p)$ strictly less than $u(x \to y, p)$, a contradiction. Hence where it is finite, $u(x \to y, p)$ is the pointwise limit of increasing continuous functions, hence it is Borel measurable. Return to definition of the chain minimiser at the start. If $u(x \to y, p)$ is infinite, then because it is defined only on the p^* of the next stage, the chain minimiser there is Borel measurable. Returning to $x \to y$ and the functions $u(x^* \to y^*, p^*)$ from the next stage continuations, because they are Borel measurable the function where the equality holds for the different weights is Borel measurable and the set where this equality is obtained is a Borel measurable set. Since the equalities are unique solutions and the inverse image of a Borel measurable set is a Borel measurable set, the colouring that define these equalities are also Borel measurable. We proceed by induction on the stages. ∎

Definition 8. *Define $u(p)$ to be the greatest lower bound of all r such that r is greater than $u(p, x \to y)$ for some set of $x \to y$ of positive measure.*

In the above process of equalising, we would like to minimise using the function $u(p^*)$ instead of the $u(x^* \to y^*, p^*)$. Of course the actual process of pain minimisation could look very different because $u(x^* \to y^*, p^*)$ may be much larger than $u(p^*)$. But if we minimise with this assumption, we obtain a result which is not higher than the proper result almost everywhere. Our goal is to show that with this assumption, by applying the rule of minimising the $u(p^*) \cdot v^*$ level on the next stage, and applying this evaluation stage after stage, the resulting limit superior of the product is infinite almost everywhere. It follows that the real value must be infinite almost everywhere also.

Unfortunately we do not know the function u explicitly, and ultimately we will show that it is infinite for all positive values; therefore the function u cannot be of practical use. So instead of minimising according to u, we introduce some new function w to replace the function u, with w finite everywhere.

There are two useful facts from the replacement of $u(x^* \to z_i, p^*)$ by the uniform functions $u(p^*)$ or $w(p^*)$ (should the former be finite). We explain this with the function w, which ultimately we will use.

First, if one aims to minimise active pain levels according to $w(p^*)$ times the passive pain level of z_i on the next stage (where p^* is the weight given by some x^* to z_i), because w is a increasing function, it is sufficient to uses equal weights p^* from all the $x^* \neq x$ such that $x^* \to z$.

The second fact, following from the equal weights consequence, is by minimising we do not have to consider any weight p given to some z_i by some $x^* \neq x$ such that $p < \frac{1}{100}$. As we already assumed that all the weights on the other side of z_i were equal, if $p < \frac{1}{100}$ were these common weights it means that the weight from x to z_i is at least .91. This means that the weights from x to the other z_j with $z_j \neq y$ add up to no more than 0.09. Therefore the passive pain level at these other $z_j \neq z_i, y$ are at least 10 times that of z_i. Assuming that this is not a terminating chain, there is a $\overline{x}$ giving weight of at least $\frac{1}{10}$ to one of these z_j. From the monotonicity of w the chain $\overline{x} \to z_j$ receives at least ten times the level of the chain $\hat{x} \to z_i$. This means

that we could redistribute the weights coming from x, giving more to the other z_j and less to z_i with a reduction in the level as determined by the function w. We discovered from our computer calculations that this bound of 0.01 could be replaced by 0.055.

Definition 9. *Let $\mathcal{C}$ be the collection of non-terminating chains with their roots. Let $\mathcal{P}$ be the set of infinite paths in the chains starting at their roots.*

The set $\mathcal{C}$ should be understood as the choices of degrees at each stage and in each position. The set $\mathcal{C}$ is a compact set, as the removal of terminating points is the removal of an open set. The probability $\hat{q}$ (approximately .991603) determines whether or not a potential edge exists, and non-terminating implying that after the removal of the terminating points there is an infinite continuation in each of the four directions $x^* \to y^*$ where x^* is closer to the root than y^*. Based on the choices of degrees, we have a canonical conditional probability distribution on the collection $\mathcal{C}$ as determined by $\hat{q}$.

Whether we use the function u, some other function w, or leave it as a sequence of multiplying by ps and dividing by qs, we should notice the connection to entropy. Assuming that we are in the chain generated by $x_0 \to y_0$, a member ω of $\mathcal{P}$ is a sequence of $(y_0, x_0, y_1, x_1, \dots)$ such that $x_0 \to y_0, x_0 \to y_1, x_1 \to y_1, x_1 \to y_2, \dots$. If ω is such an infinite sequence inside a chain $C \in \mathcal{C}$, and weights p_i are given to the various $x_i \to y_i$ and weights q_i are given to the various $x_{i-1} \to y_i$, and assuming a normalised passive pain of 1 at y_0, the active pain at some $x_N \to y_N$ in the sequence is $\frac{\prod_{i=0}^{N} p_i(\omega)}{\prod_{i=1}^{N} q_i(\omega)}$. This will be combined with a probability distribution on these sequences not dissimilar to that used to define entropy.

Definition 10. *For any path $\omega \in P$ and i define $\phi_i(\omega)$ as*

$$\phi_i(\omega) := \sum_{k=0}^{i} \log(p_k(\omega)) - \sum_{k=1}^{i} \log(q_k(\omega)).$$

We say that $\omega \in C$ if the path ω is contained in C.

Notice that we want to show that, for almost all $C \in \mathcal{C}$,

$$\limsup_{i \to \infty} \max_{\omega \in C} \phi_i(\omega) = \infty.$$

It doesn't really matter if we replace the last p_i with a finite $u(p_i)$ or $w(p_i)$, as long as there is fixed ratio by which p_i cannot differ from either $u(p_i)$ or $w(p_i)$. This leads to the following definition.

Definition 11. *For any $\omega \in P$, we define $\phi_i^w(\omega)$ as*

$$\phi_i^w(\omega) := \log(w(p_i(\omega))) + \sum_{k=0}^{i-1} \log(p_k(\omega)) - \sum_{k=1}^{i} \log(q_k(\omega)).$$

Given that the sequences $p_i(\omega)$ and the $q_i(\omega)$ are from the chain minimisers of some $C \in \mathcal{C}$, we will show that the ratio of $\log(w(p_{i+1}) - \log(q_i) + \log(p_i)$ to $\log(w(p_i))$ tends to increase in expectation. To make sense of this tendency, we need to define a probability distribution on the $\mathcal{P}$. We have already a probability distribution on $\mathcal{C}$, the set of chains, determined by the probabilities for the future degrees. But to get a probability distribution on $\mathcal{P}$, we need to extend it to a conditional probability defined on each chain. We define that probability distribution using the function w and the process of equalisation, as described above.

3.2 The Function w

We have to define a function $w : [0, 1) \rightarrow [1, \infty)$. The function $w(x)$ is not far away from $\frac{x}{1-x}$, so it is easiest to represent it as $w(x) = \frac{x}{1-x} f(x)$ with a function $f : [0, 1] \rightarrow [\frac{4}{5}, \frac{8}{5}]$ such that f is in C^2.

Where does the function w come from? We consider a non-terminating chain generated by $x \rightarrow y$. We assume that the passive pain at some point y is normalised at 1, the probability coming from x to y is p, and all vertices are non-terminating of degree 5. This means that, if $p = \frac{1}{5}$, the resulting passive pain of the entire chain has the value of 1 and the active pain $\frac{1}{5}$. Following from x to the four other vertices on the other side from y, we assume that the weights are distributed evenly. This means that from x to the four points z_1, z_2, z_3, z_4, the weight is $\frac{1-p}{4}$. The passive pain v at each of these four places z_i has to satisfy $\frac{1-p}{4} v = p$ or $v = \frac{4p}{(1-p)}$. In the definition of the function w, we will ignore the initial multiple of 4 and also drop the 4 from other places where it is superfluous. We continue with weights $\frac{3+p}{16}$ in the further directions from each of the z_i. The weights at the next stage are $\frac{13-p}{64}$. The recursive calculation, a generating function, for the active pain in the limit becomes

$$\frac{p}{1-p} \cdot \frac{4(3+p)}{13-p} \cdot \frac{4(51+p)}{205-p} \cdot \frac{4(819+p)}{3277-p} = \frac{p}{1-p} \prod_{i=0}^{\infty} \frac{4(\frac{16^{i+1}-1}{5} + p)}{\frac{4 \cdot 16^{i+1}+1}{5} - p}.$$

There is a problem with this definition for the function w. The above infinite product is based on the idea that continuation with all points of degree 5 is the proper way to represent what happens in general when the pain levels converge to finite levels. However we claim a slight tendency for these pain levels to approach infinity, meaning that the function we need should be a slight distortion of the above. The above function works well for the values p that are used most commonly in the equalisation process. As stated above, all such p are greater than .55, however usually they are the p between $\frac{1}{10}$ and $\frac{1}{2}$. The values outside of this range need to be altered with only minimal change for the values within this range. For the function f with $w(p) = \frac{p}{1-p} f(p)$ we define

$$f(p) = e^{-1_{[0, \frac{1}{5}]}(p) \frac{20}{27}(\frac{1}{5} - p)^3} \cdot e^{-1_{[\frac{1}{2}, 1]}(p) \frac{1}{4}(p - \frac{1}{2})^3} \cdot \frac{4(3+p)}{13-p} \cdot \frac{4(51+p)}{205-p} \cdot \frac{4(819+p)}{3277-p} \cdots,$$

meaning that the function is reduced for the extremes of $p \in [0, \frac{1}{5}]$ and $p \in [\frac{1}{2}, 1]$. The function is altered so that it remains in C^2.

3.3 The Equalisation Process

Let's look at the chain generated by $x \to y$, where $x \to z_i$ for $i = 1, 2, 3, 4$ and x^* is another vertex where $x^* \to z_i \neq y$ for some i. Assume that p is the weight given by x toward y and $j_i + 1$ is the degree of z_i. Instead of equalising the $u(x^* \to z_i, p^*)^{\frac{1}{q_i}}$ over all the choices, we equalise the $\frac{w(\frac{1-q_i}{j_i})}{q_i}$. This can be performed, as $w(p)$ can be calculated easily, while the $u(x^* \to z_i, p^*)$ is known only through understanding the infinite structure. In what follows, this is what we call the equalisation process, not the process described above of equalising the $u(x^* \to z_i, p^*)$.

We have to relate the equalisation process to the chain minimisers. This is accomplished by the following lemmas.

Lemma 4. *(a) For every choice of $k = 1, \ldots, 9$ and x, the second derivative of $-\log(1 - \frac{1-x}{k}) + \log(f(\frac{1-x}{k}))$ is negative. (b) For every choice of x, $\frac{1}{6} \leq \frac{d(\log(f(x)))}{dx} < \frac{5}{9}$. (c) The function $xw(x)$ is convex in x.*

Proof. (a) We group the terms by $\log(\frac{k-1+x}{k}) - 1_{[1-\frac{k}{5},1]}\frac{20}{27}(\frac{1}{5} - \frac{1-x}{k})^3 - 1_{[0,\frac{k}{2}]}\frac{1}{4}(\frac{1-x}{k} - \frac{1}{2})^3 + \log(4(3 + \frac{1-x}{k}))$ and then the rest, followed by the pair $-\log 13 - \frac{1-x}{k} + \log(4(51 + \frac{1-x}{k}))$ and so on. Taking the first derivative of the first three gives

$$-\frac{1}{k-1+x} - \frac{1}{k}1_{[1-\frac{k}{5},1]}(x)\frac{20}{9}(\frac{1}{5} - \frac{1-x}{k})^2 + \frac{1}{k}1_{[0,\frac{k}{2}]}\frac{3}{4}(\frac{1-x}{k} - \frac{1}{2})^2 - \frac{1}{k}\frac{1}{3 + \frac{1-x}{k}}.$$

Taking the second derivative gives

$$\frac{1}{(k-1+x)^2} - \frac{1}{k^2}1_{[1-\frac{k}{5},1]}(x)\frac{40}{9}(\frac{k}{5} - \frac{1-x}{k}) - \frac{1}{k^2}1_{[0,\frac{k}{2}]}\frac{3}{2}(\frac{1-x}{k} - \frac{1}{2}) - \frac{1}{k^2}\frac{1}{(3 + \frac{1-x}{k})^2}.$$

Where the two special restrictions apply are mutually exclusive. In the region $x \notin [0, \frac{k}{2}]$ the minimum occurs at $x = 1$, for the quantity $\frac{1}{k^2}(1 - \frac{8}{9} - \frac{1}{9}) = 0$. Where $x \in [0, \frac{k}{2}]$ the result is at least $\frac{1}{k^2}(1 - \frac{3}{4} - \frac{1}{9}) > 0$. Moving to the first pair of following terms, taking the first derivative gives $\frac{1}{13-x} + \frac{1}{x+52}$. The second derivative is $\frac{1}{(13-x)^2} - \frac{1}{(51+x)^2}$, which is also positive. And the same follows for the rest of the pairs.

(b) The second derivative of $\log(f(x))$ is negative throughout and hence the derivative reaches its maximum at $x = 0$ for the quantity $\frac{4}{45} + \frac{1}{3} + \frac{1}{13} + \frac{1}{51} + \cdots < \frac{5}{9}$ and its minimum at 1 for $-\frac{3}{16} + \frac{1}{4} + \frac{1}{12} + \frac{1}{52} + \cdots > \frac{1}{6}$.

(c) We need the second derivate of $x \log(w(x)) = x \log(x) - x \log(1-x) - 1_{[0,\frac{1}{5}]}(x)\frac{20}{27}(\frac{1}{5} - x)^3 x - 1_{[\frac{1}{2},1]}(x)\frac{1}{4}(x - \frac{1}{2})^3 x + x \log(4(3+x)) - x \log(13-x) + x \log(51+x) - x \log(205-x) + \ldots$. The first derivative is

$$-1_{[0,\frac{1}{5}]}(x)\frac{20}{27}(\frac{1}{5} - x)^3 + 1_{[0,\frac{1}{5}]}(x)\frac{20}{9}(\frac{1}{5} - x)^2 x - 1_{[\frac{1}{2},1]}(x)\frac{1}{4}(x - \frac{1}{2})^3 1_{[\frac{1}{2},1]}(x)\frac{3}{4}(x - \frac{1}{2})^2$$

$$\log(x) - \log(1-x) + \log(4(3+x)) - \log(13-x) + \log(51+x) - \log(205-x) \cdots +$$

$$1 + \frac{x}{1-x} + \frac{x}{3+x} + \frac{x}{13-x} + \frac{x}{51+x} + \frac{x}{205-x} + \ldots.$$

The second derivative is

$$1_{[0,\frac{1}{5}]}(x)\frac{40}{9}(\frac{1}{5}-x)^2 - 1_{[0,\frac{1}{5}]}(x)\frac{40}{9}(\frac{1}{5}-x)x - 1_{[\frac{1}{2},1]}\frac{3}{2}(x-\frac{1}{2})^2 - 1_{[\frac{1}{2},1]}\frac{3}{2}(x-\frac{1}{2})x$$

$$+\frac{1}{x} + \frac{1}{1-x} + \frac{1}{3+x} + \frac{1}{13-x} + \frac{1}{51+x} + \frac{1}{205-x} + \cdots$$

$$\frac{1}{(1-x)^2} + \frac{3}{(3+x)^2} + \frac{13}{(13-x)^2} + \frac{51}{(51+x)^2} + \frac{205}{(205-x)^2} + \cdots .$$

The negative terms are dominated by $\frac{1}{x}$ when $x \leq \frac{1}{5}$ and by $\frac{1}{(1-x)^2}$ when $x \geq \frac{1}{2}$. ∎

Definition 12. *Define $g_k(x)$ to be $\frac{1}{x+k-1} + \frac{1}{k}\frac{f'(\frac{1-x}{k})}{f(\frac{1-x}{k})}$, which is the negative of the first derivative of $-\log(\frac{1-x}{k}) + \log(f(\frac{1-x}{k}))$. From Lemma 4 we know that $g_k(x)$ in decreasing in x.*

3.4   The Stochastic Process

To define a stochastic process on $\mathcal{P}$, we need a probability distribution on paths. Before we do that, to use the computer analysis effectively we need to perform a reduction of the system to discretely many values. We alter slightly the chain minimiser of every chain. We start by rounding down the initial weight p_0, the weight from x_0 to y_0, to the highest value of $\frac{k}{20,000}$ less than or equal to it (for $k = 0$ or k a positive integer). From x_0 to the four different potential y_1 we can increase the weights, so that they add up exactly to $1 - p_0$. For the weights from various x_1 to some y_1, with the weight from x_0 to y_1, they add up to at least $1 + 2^{-11}$. We round down these various numbers p_1 to the highest value of $\frac{k}{20,000}$ less than or equal to p_1 for k a non-negative integer. Also to simplify the analysis, those rounded down numbers, which always add up to at least 1, are normalised to add up to 1. They must add up to at least 1 because by rounding down one can reduce the quantity by no more than $\frac{1}{20,000}$, there are at most 9 such numbers, and $\frac{9}{20,000}$ is less than 2^{-11}. We require only that this rounding down process is done in a Borel measurable way. We can continue this process to all stages, always reducing the value of the products. After this alteration, if we show that the limit superior of of the maximum of ϕ_i^w or ϕ_i within a chain is infinite almost everywhere, the same is true before this alteration. Indeed, we show that any colouring function obeying the rule **Q** can be altered in this way so that the limit superior is infinite almost everywhere.

The following can occur in that analysis: there is a p_i with $p_i = \frac{19,999}{20,000}$ followed by some passive point z of degree 2, meaning that there is only one p_{i+1} following in that direction. As less than $\frac{1}{20,000}$ can distributed to z, it follows that this p_{i+1} must be more than $\frac{19,999}{20,000}$. As there are no other such p_{i+1}, it will be rounded down to exactly $\frac{19,999}{20,000}$. The sum of $\frac{19,999}{20,000}$ with the weight to z will be strictly less than 1. Notice that this situation cannot occur, because with less than $\frac{1}{20,000}$ sent to z in one direction and at least 1 in the other direction there is significantly less than $1 + \frac{1}{2^{11}}$ weight toward z. The result would be that p_{i+1} should be exactly 1, all the next q_{i+1} should be zero, and the process jumps immediately to an infinite value. But we will not exclude this possibility from the analysis. Instead of letting

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the process jump immediately to infinity we will give it some very high but finite value. This is because the following analysis involves a Markov chain with bounds for the variance on each stage. In the effort to show that a process should be infinite almost everywhere, we ignore such situations where it may go to infinity in one step. To include this in the analysis would greatly complicate it.

Next there are some facts about any q_1, q_2, q_3, q_4 chosen to satisfy the equalisation process: namely, that

$$\frac{w(\frac{1-q_i}{j_i})}{q_i} = \frac{w(\frac{1-q_j}{j_j})}{q_j},$$

for all i and $p + q_1 + q_2 + q_3 + q_4 = 1$. The obvious fact is that $j_i = j_k$ implies that $r_i = r_k$ and $j_i < j_k$ implies $r_i > r_k$. The not so obvious, based on the many calculations, is the following lemma.

Lemma 5. *The following holds for all $p = \frac{n}{20,000}$, where n is a positive integer less than 20,000:*
(a) if $j_i > j_k$ for some i, k, then $q_i < \frac{1}{3}$;
(b) if $j_i = j_k + 1$, then $|q_i - q_k| \leq \frac{1}{4}$;
(c) if $j_i > j_k > 1$, then $q_l + q_m + p > \frac{2}{9}$ for $\{l, m\} \cap \{i, k\} = \emptyset$.

The proof of Lemma 5 is confirmed by the computer. The discussion of the nymerical calculations is deferred to Section 4.

Lemma 6. *Let $0 < p < 1$ and let $q = (q_1, \ldots, q_4)$ be a solution to the equalities*

$$\frac{w(\frac{1-q_i}{j_i})}{q_i} = \frac{w(\frac{1-q_j}{j_j})}{q_j},$$

for all i and $p + q_1 + q_2 + q_3 + q_4 = 1$ with $p = \frac{k}{20,000}$ for some integer k with $0 < k < 20,000$ and $j_1, \ldots, j_k$ positive integers between 1 and 9 inclusive. Then, there exists some $s \in \Delta(\{1, \ldots, k\})$ such that for all $q^* = (q_1^*, \ldots, q_k^*)$ with all q_i^* positive such that $p + q_1^* + \cdots + q_k^* = 1$ it follows that

$$\phi_s(q^*) := \sum_i s_i \, \log\left(\frac{w(\frac{1-q_i^*}{j_i})}{q_i^*}\right) \geq \phi_s(q) = \sum_i s_i \, \log\left(\frac{w(\frac{1-q_i}{j_i})}{q_i}\right).$$

Furthermore for every ϵ there are finitely many values $q_1^*, \ldots, q_n^*$ for q^* such that after rounding up to the nearest q_k^* the inequality holds up to ϵ.

Proof. With Lagrangian multipliers $L := \phi_s(q^*) - \lambda(1 - p - \sum_i q_i^*)$, we define s and create a critical point for ϕ_s at $q^* = q$ by setting

$$\frac{1}{t_i} := -\frac{\partial \log\left(\frac{w(\frac{1-q_i^*}{j})}{q_i^*}\right)}{\partial q_i^*}(q_i)$$

and defining $s_i := \frac{t_i}{\sum_j t_j}$. From Lemma 4, all the t_i are positive. As convergence to the boundary of the domain of q^* (where $q_i^* = 0$ for some i) gives convergence

of ϕ_s to positive infinity, it suffices to show that there can be only one unique critical point of ϕ_s so defined, namely the $q = (q_1, \ldots, q_k)$. Another critical point $r = (r_1, \ldots, r_k) \neq q$ would have to satisfy the equalities for some other λ'.

According to Lemma 4, the second derivative of $-\log(\frac{w(\frac{1-x}{j})}{x})$ is equal to $\frac{1}{x^2} - \frac{1}{(1-x)^2}$ plus a positive term, meaning that the second derivative of $\frac{1}{2} + x$ plus the second derivative of $\frac{1}{2} - x$ is always positive. As we could switch q and r, without loss of generality, assume that $\lambda' \geq \lambda$. As the q_i and the r_i sum up to the same quantity $1 - p$, a second critical point is possible only if there is some $r_i \geq \frac{1}{2}$ with $q_i \leq r_i$ with $q_i + r_i > 1$ and $r_k \leq q_k$ for all other $k \neq i$. We can now assume that the difference between q and r implies that $\lambda' > \lambda$, $r_i > q_i$ and $r_j < q_j$ for all $j \neq i$, and of course that $p < \frac{1}{2}$.

A second critical point implies that

$$\frac{\frac{1}{r_i} + \frac{1}{1-r_i} + \frac{1}{j_i - 1 + r_i} + \frac{1}{j} \frac{f'(\frac{1-r_i}{j_i})}{f(\frac{1-r_i}{j_i})}}{\frac{1}{q_i} + \frac{1}{1-q_i} + \frac{1}{j_i - 1 + q_i} + \frac{1}{j_i} \frac{f'(\frac{1-q_i}{j_i})}{f(\frac{1-q_i}{j_i})}}$$

is equal to

$$\frac{\frac{1}{r_k} + \frac{1}{1-r_k} + \frac{1}{j_k - 1 + r_k} + \frac{1}{j_k} \frac{f'(\frac{1-r_k}{j_k})}{f(\frac{1-r_k}{j_k})}}{\frac{1}{q_k} - \frac{1}{1-q_k} - \frac{1}{j_k - 1 + q_k} + \frac{1}{j_k} \frac{f'(\frac{1-q_k}{j_k})}{f(\frac{1-q_k}{j_k})}},$$

for all $i \neq k$ and all of these ratios are equal to $\frac{\lambda'}{\lambda}$. We show that a second critical point is not possible by showing that this is not possible.

We start with the assumption that the $\{x_l, y_l\}$ are equal to the $\{q_l, r_l\}$ for all l with $y_i > x_i$ for only one i and choose any $k \neq i$ such that $\frac{(1-y_i)x_k}{(1-x_i)y_k}$ is maximal. Without loss of generality let $i = 0$ and $k = 1$ (with x_2, x_3 and y_2, y_3 the other variables). With their sums equal to $1 - p$ and $y_0 > x_0$ it follows that $\frac{(1-y_i)x_k}{(1-x_i)y_k} > 1$, meaning that $x_1(1 - y_0) > y_1(1 - x_0)$.

We will take two approaches to proving the above equalities are impossible. Either we will show directly that the equality is not possible or we will demonstrate that the equality implies that

$$\frac{\frac{1}{y_1} + \frac{1}{1-y_1}}{\frac{1}{y_0} + \frac{1}{1-y_0}} \leq \frac{\frac{1}{x_1} + \frac{1}{1-x_1}}{\frac{1}{x_0} + \frac{1}{1-x_0}}.$$

With $\frac{y_0(1-y_0)}{y_1(1-x_1)} > 1$, we get the two inequalities:

$$x_1(1 - y_0) > y_1(1 - x_0) \quad x_0 y_1(1 - y_1)(1 - x_0) \geq y_1 x_1(1 - x_1)(1 - y_0).$$

Multiplying together gives $x_0(1 - y_1) > y_0(1 - x_1)$ and adding $y_1(1 - y_0) > x_1(1 - x_0)$ to this inequality we get $x_0 + x_1 > y_0 + y_1$. This implies that there must be some $n = 2$ or $n = 3$ with $y_n > x_n$, a contradiction to $y_j < x_j$ for all $j \neq 0$.

Let $k = j_0$ and $l = j_1$. We use that $g_j(x)$ is decreasing in x for any choice of j.

Case 1, $k < l$: Notice that the situation where $y_i = q_i$ is included in this case, since $y_0 > \frac{1}{2}$ implies that $k < l$. If $k + 1 = l$ we get $|x_0 - x_1| \leq \frac{1}{4}$ in both cases of $x_0 = q_0$ or $x_0 = r_0$ from Lemma 5 , since $x_0 > x_1$, $y_0 > x_0$ and $y_1 < x_1$. So regardless of the values of k and l, we conclude that $b := g_k(x_0) > d := g_l(x_1)$. Likewise we define $a := g_k(y_0)$ and $c := g_l(y_1)$ with the the result $b > a$ and $c > d$. To demonstrate that the above inequality implies the impossibility of a second critical point, and using $bc > ad$, it is sufficient to show that

$$d(\frac{1}{y_0} + \frac{1}{1 - y_0}) + a(\frac{1}{x_1} + \frac{1}{1 - x_1}) \leq c(\frac{1}{x_0} + \frac{1}{1 - x_0}) + b(\frac{1}{y_1} + \frac{1}{1 - y_1}).$$

By the inequalities $b > a$ and $c > d$ this is implied by

$$d(\frac{1}{y_0} + \frac{1}{1 - y_0}) + b(\frac{1}{x_1} + \frac{1}{1 - x_1}) \leq d(\frac{1}{x_0} + \frac{1}{1 - x_0}) + b(\frac{1}{y_1} + \frac{1}{1 - y_1}).$$

Now using $b > d$ and that $y_1 < x_1 < \frac{1}{2}$ implies $\frac{1}{y_1} + \frac{1}{1 - y_1} > \frac{1}{x_1} + \frac{1}{1 - x_1}$ and likewise $y_0 + x_0 > 1$, $y_0 > x_0$ and $y_0 > \frac{1}{2}$ implies $\frac{1}{y_0} + \frac{1}{1 - y_0} > \frac{1}{x_0} + \frac{1}{1 - x_0}$, it suffices to prove that

$$\frac{1}{y_0} + \frac{1}{1 - y_0} + \frac{1}{x_1} + \frac{1}{1 - x_1} \leq \frac{1}{x_0} + \frac{1}{1 - x_0} + \frac{1}{y_1} + \frac{1}{1 - y_1}.$$

We separate into two parts, to show that $\frac{1}{1 - x_0} + \frac{1}{y_1} \geq \frac{1}{1 - y_0} + \frac{1}{x_1}$ and $\frac{1}{x_0} + \frac{1}{1 - y_1} \geq \frac{1}{y_0} + \frac{1}{1 - x_1}$. To deal with the first part, after clearing the dominators one gets equivalence to $(1 - y_1)(1 - x_1)(y_0 - x_0) \geq x_0 y_0 (x_1 - y_1)$. This inequality follows from $y_0 + y_1 < 1$, $x_0 + x_1 < 1$, and $y_0 + y_1 > x_0 + x_1$. The other part reduces to the same inequality, after clearing the dominators.

Case 2, $k = l$: This is broken down into two cases: **Case 2A,** $x_0 \geq x_1$ and **Case 2B,** $x_0 > x_1$. In the former case, we have something of the form $\frac{A}{B} = \frac{C}{D}$ where $C > A$ and $B > D$, which is impossible. In the latter case, we have the same inequalities of Case 1.

Case 3, $k > l$: This is broken down into two cases. In both cases, since $g_k(x_0) > g_k(y_0)$, they bring down the fraction on that side. Therefore, for the sake of contradiction, we assume that

$$\frac{\frac{1}{y_0} + \frac{1}{1 - y_0}}{\frac{1}{x_0} + \frac{1}{1 - x_0}} \geq \frac{\frac{1}{y_1} + \frac{1}{1 - y_1} + \frac{1}{y_1 - 1 + l} + e}{\frac{1}{x_1} + \frac{1}{1 - x_1} + \frac{1}{x_1 - 1 + l} + e},$$

where e is an upper limit for the negative of the derivative of $\log(f(\frac{1 - y_1}{l}))$ (which is larger than when y_1 is replaced by x_1).

Case 3A, $l = 1$: We show that

$$\frac{\frac{1}{y_0} + \frac{1}{1 - y_0}}{\frac{1}{x_0} + \frac{1}{1 - x_0}} \geq \frac{\frac{1}{y_0} + \frac{1}{1 - y_0}}{\frac{1}{x_0} + \frac{1}{1 - x_0}},$$

leading to the above contradiction. From $e \leq \frac{5}{9}$ it suffices to show that

$$(\frac{1}{y_0} + \frac{1}{1 - y_0})(\frac{1}{x_1} + \frac{5}{9}) \leq (\frac{1}{x_0} + \frac{1}{1 - x_0})(\frac{1}{y_1} + \frac{5}{9}).$$

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That $\frac{1}{1-x_0}\frac{1}{y_1} < \frac{1}{1-y_0}\frac{1}{x_1}$ follows from the choice of x_1 and y_1. We use that $x_0 < \frac{1}{3}$ implies that $y_0 > \frac{2}{3}$. It also implies that $\frac{1}{y_1 x_0}$ is greater than $\frac{1}{x_1 y_0} + \frac{5}{9}\frac{1}{1-y_0}$ and of course that $\frac{5}{9x_0}$ is greater than $\frac{5}{9y_0}$.

Case 3B, $l \geq 2$: We use that $\frac{1}{y_1+l-1} < 1$ and $e \leq \frac{5}{18}$, so that their sum is no more than $\frac{23}{18}$. We show that

$$\frac{\frac{1}{y_0} + \frac{1}{1-y_0}}{\frac{1}{x_0} + \frac{1}{1-x_0}} \geq \frac{\frac{1}{y_1} + \frac{1}{1-y_1} + \frac{23}{18}}{\frac{1}{x_1} + \frac{1}{1-x_1} + \frac{23}{18}}$$

is impossible, or with cross multiplication that

$$\left(\frac{1}{y_0} + \frac{1}{1-y_0}\right)\left(\frac{1}{x_1} + \frac{1}{1-x_1} + \frac{23}{18}\right) \geq \left(\frac{1}{x_0} + \frac{1}{1-x_0}\right)\left(\frac{1}{y_1} + \frac{1}{1-y_1} + \frac{23}{18}\right)$$

is impossible. From the choice of x_1 and y_1 we get $\frac{1}{y_1(1-x_0)} > \frac{1}{x_1(1-y_0)}$. From Lemma 5 we have $x_0 + x_1 < \frac{7}{9}$, with of course $x_0 \leq \frac{1}{3}$ and $x_1 \leq \frac{7}{9}$. From the choice of x_1 and y_1 we have $\frac{1}{x_0 y_1} > \frac{1-x_0}{x_1(1-y_0)x_0}$. From $x_0 < \frac{1}{3}$ and $1 - x_1 = x_0 + p + x_2 + x_3$ and $p + x_2 + x_3 \geq \frac{2}{9}$ we have $\frac{1-x_1}{x_0} > \frac{5}{3}$. From $1 - x_0 = p + x_1 + x_2 + x_3$ we have $\frac{1-x_0}{x_1} > \frac{9}{7}$. Together we get $\frac{21}{45}\frac{1}{x_0 y_1} > \frac{1}{(1-y_0)(1-x_1)}$. From the choice of x_1 and y_1 we get that $\frac{1-y_0}{y_1} > \frac{1-x_1}{x_1} > \frac{9}{7}$. So we can write $\frac{1}{3}\frac{7}{9}\frac{23}{18}\frac{1}{x_0 y_1} > \frac{23}{18}\frac{1}{1-y_0}$. Notice that $\frac{1}{3}\frac{7}{9}\frac{23}{18} = \frac{161}{486}$. With $\frac{161}{486} + \frac{21}{45} < 1$ we can conclude that $\frac{1}{x_0 y_1} > \frac{1}{(1-y_0)(1-x_1)} + \frac{23}{18}\frac{1}{1-y_0}$. It is only left to show that

$$\frac{23}{18}\frac{1}{x_0} + \frac{1}{(1-x_0)}\left(\frac{1}{y_1} + \frac{1}{1-y_1}\right) + \frac{1}{x_0(1-y_1)} > \frac{23}{18}\frac{1}{y_0} + \frac{1}{x_1 y_0} + \frac{1}{y_0(1-x_1)}.$$

Case 3Bi, $l \geq 2$, $x_1 \geq \frac{1}{2}$: From Lemma 5 we get $x_0 < 1 - \frac{1}{2} - \frac{2}{9} = \frac{5}{18}$ and therefore $y_0 > \frac{13}{18}$, so $\frac{23}{18}(\frac{1}{x_0} - \frac{1}{y_0}) > \frac{23}{18}\frac{144}{65}$. From the definition of y_1 and x_1 it holds that $\frac{y_1}{x_1} < \frac{5}{13}$. With $x_1 < \frac{7}{9}$ and $1 - x_0 < y_0$ we get $\frac{1}{(1-x_0)}\frac{1}{y_1} - \frac{1}{x_1 y_0} > \frac{144}{65}\frac{9}{7}$. We have $\frac{1}{x_0(1-y_1)} > \frac{18}{5}$ and with x_1 no more than $\frac{7}{9}$ and $y_0 > \frac{18}{23}$ we have $\frac{1}{y_0(1-x_1)} < \frac{23}{4}$. The quantity $\frac{1}{(1-x_0)}\frac{1}{1-y_1}$ is at least 1. With $(\frac{23}{18} + \frac{9}{7})\frac{144}{65} + \frac{18}{5} + 1 > \frac{23}{4}$, the case is settled.

Case 3Bii, $l \geq 2$, $x_1 \leq \frac{1}{2}$: With $x_0 < \frac{1}{3}$ and $y_0 > \frac{2}{3}$ we have $\frac{23}{18}(\frac{1}{x_0} - \frac{1}{y_0}) > \frac{69}{36}$. From the definition of y_1 and x_1, it holds that of $\frac{y_1}{x_1}$ is less than $\frac{1}{2}$. With $x_1 \leq \frac{1}{2}$ and with $1 - x_0 < y_0$, we get $\frac{1}{(1-x_0)}\frac{1}{y_1} - \frac{1}{x_1 y_0} > 2$. The quantity $\frac{1}{(1-x_0)}\frac{1}{1-y_1}$ is at least 1. With x_1 no more than $\frac{1}{2}$, we have $\frac{1}{x_0(1-y_1)} > 3$ and $\frac{1}{y_0(1-x_1)} < 3$, and the case is settled.

Finally, notice that there are finitely many possibilities for the choice of degrees (j_1, j_2, j_3, j_4) and $p = \frac{k}{20,000}$. For each such choice, any q_k^* small enough so that $s_k\frac{1-q_k^*}{q_k^* j_k}$ alone exceeds the total expectation of $\sum_{i=1}^4 s_i\frac{1-q_i}{q_i j_i}$ suffices for the lowest value needed. The approximation by ϵ follows from the fact that $\frac{1}{x}$ is uniformly continuous when positive x is bounded from below. ∎

Given any fixed $C \in \mathcal{C}$, we need to determine a conditional probability distribution $P(\cdot \mid C)$ on the paths ω that belong to C. We start with the root $x \to y$ of the chain C, and call $x = x_0$ and $y = y_0$. Let z_1, z_2, z_3, z_4 be the points such

that $x_0 \to z_i$, with j_1, j_2, j_3, j_4 the positive integers between 1 and 9 such that $j_i + 1$ is the degree of z_i. Let q_1, q_2, q_3, q_4 be the weights from x_0 to the z_i that solve the equalisation process. The probability of moving in the direction from x_0 to z_i is the quantity s_i as determined by Lemma 6. The probability of moving from z_i to $x_{i,k}$ is $\frac{p_{i,k}}{\sum_{l=1}^{4} p_{i,l}}$, where $p_{i,l}$ is the weight given to z_i by $x_{i,k}$. We continue in this way defining the probability in terms of these products.

With a conditional probability distribution defined on each chain, and a probability distribution defined on the chains, we need to extend this to a probability distribution defined on $\mathcal{P}$. To do this, we use the expectations on the $P(\cdot \mid C)$. In order for this to make any sense, the conditional values we get on the chains must be Borel measurable. We get that from Lemma 3 It could be noticed that if there is symmetry to the way the quantities are rounded down, for any given choices for j_1, j_2, j_3, j_4 and i the expectation for the $p_{i,k}$ will be equal. We do not use this in the proof.

Corollary 1. *According to the above probability distribution, if the q_i and $p_{i,k}$ are from the chain minimiser, the expectation of $\log(w(p_{i,k}) - \log(w(p)) + \log(p) - \log(q_i)$ is positive.*

Proof. It follows from Lemma 6 and Lemma 4 part c (from the fact that equal quantities of $p_{i,k}$ for a given i defines a critical point and from the convexity there is a unique minimiser). ∎

Recall the definition of $\phi_i^w(\omega)$. We can break this sum into two parts. We can perform the equalisation process at each step, and for any sequence $p_0, q_1, p_1, q_2, p_2, \ldots, p_{i-1}, q_i, p_i$ corresponding to a path ω define a sequence of triples (p_0, q_1', p_1'), $(p_1, q_2', p_2'), \ldots, (p_{i-1}, q_i', p_i')$ where the q_i' are defined by the equalisation process and the p_i' are defined by equality for each weight going to the same point z in the chain. We can break down the expression of $\phi_i^w(\omega)$ into two parts, that involving the triples and the difference. Call $\overline{\phi_i^w}$ the sum of the part involving the triples and $\tilde{\phi}_i^w$ the difference $\phi_i^w - \overline{\phi_i^w}$. By the above corollary, we have shown that, conditioned on any chain, the expectation of $\tilde{\phi}_i^w$ is positive. Now, we turn to the other part, the $\overline{\phi_i^w}$.

We need to show that the ϕ_i^w functions are unbounded on almost every chain in $\mathcal{C}$. As we use only $20,000$ many values for the p_i, it suffices to do the same for the ϕ_i^w. As it does not matter where on the chain the value of ϕ_i^w is maximal, it suffices to show that the expectation of ϕ_i^w goes to infinity on almost all chains. As the expectation of $\tilde{\phi}_i^w$ is always positive, attention is drawn to the $\overline{\phi_i^w}$, the part of the process from the sequence of triples.

We want to define a Markov chain from the triples that define $\overline{\phi_i^w}$ and show that it defines a submartingale on this Markov chain that approaches infinity almost everywhere. However strictly speaking the triples do not define a Markov chain. The problem is that each system of weights is determined by the membership of some chain C in $\mathcal{C}$, and therefore those weights are determined by the future. However we can relate this process to a Markov chain through an inequality.

Definition 13. *For every $p = \frac{k}{20,000}$ and every (j_1, j_2, j_3, j_4) (choice of $1 \le j_i \le 9$) $i = 1, 2, 3, 4$ such that $p + \sum_{i=1}^{4} q_i = 1$, define $w_1(p, j_1, j_2, j_3, j_4)$ to be the common value for $\frac{w(\frac{1-q_i}{j_i})}{q_i}$ from the equalisation process. Define $r(p, j_1, j_2, j_3, j_4)$ to be*

$\log(w_1(p, j_1, j_2, j_3, j_4)) - \log(w(p))$. *For every choice* (j_1, j_2, j_3, j_4) *and every* $p = \frac{k}{20,000}$ *for all* $k = 1, \ldots, 19,999$, *we define*

$$\overline{r}(j_1, j_2, j_3, j_4) := \min_{k=1,\ldots,19,999} w_1(\frac{k}{20,000}, j_1, j_2, j_3, j_4).$$

As before, $\hat{q}$ is the probability that a chain is terminating, which we approximated at $\hat{q} = .991603$. For each choice of $1 \leq j_1, j_2, j_3, j_4 \leq 9$, we sum up the logarithm of $\overline{r}(j_1, j_2, j_3, j_4)$ times the probability $\prod_{i=1}^{4} \binom{9}{j_i}(\frac{\hat{q}}{2})^{j_1}(1 - \frac{\hat{q}}{2})^{9-j_i}$ and divide by $\hat{q}$ (to condition on the event that the chain is not terminating) to get the rate of increase s.

Lemma 7. *The rate s, the conditional expectation of $\overline{r}$, is at least $\frac{1}{1,000}$.*

The proof of Lemma 7 is done with the help of the computer. A discussion of these numerical calculations can be found in Section 4.

Proposition 2. *The process $\phi_i^{\overline{w}}$ converges to positive infinity almost everywhere.*

Proof. The Markov chain from the $\overline{r}(j_1, j_2, j_3, j_4)$ is well defined. The Kolmogorov inequality states that, if $X_1, X_2, \ldots$ is a martingale starting at X_0, then, for $\epsilon > 0$, the probability that $\max_{0 < i \leq n} |X_i - X_0| > \epsilon$ is no more than the sum of the variances of the $X_i - X_{i-1}$ divided by ϵ^2. As only finitely many values for p and (j_1, j_2, j_3, j_4) are used, the variances at each stage have a uniform bound $B > 0$ (determined by the two extremes of $j_1 = j_2 = j_3 = j_4 = 1$ and $p = \frac{19,999}{20,000}$ and $j_1 = j_2 = j_3 = j_4 = 9$ and $p = \frac{1}{20,000}$). After subtracting the $s > 0$ a martingale x_i is defined with $X_n = \sum_{i=1}^{n} x_i$. The cumulative variance of the process to the nth stage is the sum of the variances at each stage, which is no more than nB. If the subset where the limit superior before removing the s is not infinite has positive measure, there must be an $\epsilon > 0$ such that for every n the probability that $|X_n|$ is greater than $\frac{ns}{2}$ is at least ϵ. But this is not true, since the Kolmogorov inequality says that this probability is not greater than $\frac{4nB}{n^2 s^2}$ for every n. ∎

Now we can prove the main result.

Proposition 3. *The limit superior of the maximal values of the ϕ_i is infinite for almost all chains $C \in \mathcal{C}$.*

Proof. Suppose there is a bound M and a subset A of chains of positive measure is such that the highest value of ϕ_i for all i in the subset A is M. Because the expectation of $\tilde{\phi}_i^w$ is non-negative, this means that in this subset A the expectation of $\overline{\phi_i^w}$ must be less than M. But this is impossible, since $\overline{\phi_i^w}$ approaches infinity almost everywhere. ∎

4 The Numerical Calculations

The first problem is to calculate the function w. As it is defined above, it is difficult to calculate with great precision because the infinite product doesn't converge quickly. The influence of each term is approximately one-fourth of the previous

term, and that means to gain accuracy to less than one-millionth requires the use of around ten terms. After ignoring the exponential part, we want to convert the infinite product into a few products followed by a power series. However we notice that the coefficients of 3,13, 52, . . . are not easy to work with. We make a simple substitution, $t = p - \frac{1}{5}$, with t now standing for the difference from the norm of $\frac{1}{5}$.

We start with $\frac{p}{1-p}$. After the substitution $p = t + \frac{1}{5}$ we get $\frac{t+\frac{1}{5}}{\frac{4}{5}-t} = \frac{1+5t}{4-5t}$. As multiplying one time by 4 doesn't change anything, we get $\frac{1+5t}{1-\frac{5}{4}t}$.

Next comes $\frac{3+p}{13-p}$. After the substitution $p = t + \frac{1}{5}$ we get $4\frac{t+3+\frac{1}{5}}{\frac{64}{5}-t} = \frac{5t+16}{16-\frac{5}{4}t}$. We recognise the pattern

$$\tilde{w}(t) := \frac{1 + 5t}{1 - \frac{5}{4}t} \frac{16 + 5t}{16 - \frac{5}{4}t} \frac{16^2 + 5t}{16^2 - \frac{5}{4}t} \cdots .$$

To make accurate calculations of $\tilde{w}$, we use the first three products and then change the rest into the geometric power series. We get

$$\tilde{w}(t) := \frac{1 + 5t}{1 - \frac{5}{4}t} \frac{16 + 5t}{16 - \frac{5}{4}t} \frac{16^2 + 5t}{16^2 - \frac{5}{4}t}(1 + \frac{5t}{16^3})(1 + \frac{5t}{16^4}) \cdots$$

$$(1+\frac{5}{16^3 \cdot 4}t + \frac{5^2}{16^6 \cdot 4^2}t^2 + \frac{5^3}{16^9 4^3}t^3 + \ldots)(1 + \frac{5}{16^4 \cdot 4}t + \frac{5^2}{16^8 \cdot 4^2}t^2 + \frac{5^3}{16^{12}4^3}t^3 + \ldots) \cdots$$

Collecting the t and t^2 terms via the geometric series and including the first t^3 term gives a very good approximation:

$$\frac{1 + 5t}{1 - \frac{5}{4}t} \frac{16 + 5t}{16 - \frac{5}{4}t} \frac{16^2 + 5t}{16^2 - \frac{5}{4}t}\left(1 + \frac{5}{3 \cdot 4 \cdot 16^2}t + \frac{100}{16^5 \cdot 9 \cdot 17}t^2 + \frac{125}{4^3 \cdot 16^9}t^3\right).$$

The first t^3 term dominates the rest (and true also of the higher powers of t) and so the error is less than $\frac{1}{10^{10}}$. Even dropping the second power term puts one within $\frac{1}{10^6}$, which is good enough, considering that the rate of expansion is slightly greater than $\frac{1}{1,000}$.

The first term is easy to calculate with a geometric series: it is $\frac{5t}{16^3} \frac{4}{3} = \frac{5t}{16^2 \cdot 12}$. The second term comes in two parts. First there are the terms that come directly from the t^2, or $\frac{5^2 t^2}{16^7} \frac{16^2}{255} = \frac{5^2 t^2}{16^5 \cdot 255}$. The rest are products of single powers of t. We use that if $a_1, \ldots, a_n$ are numbers and we want to calculate $\sum_{i<j} a_i a_j$ we could calculated instead $\frac{1}{2}((a_1 + \cdots + a_n)^2 - a_1^2 - \ldots a_n^2)$. If we want to calculate $\sum_{0 \le i < j} ab^i$ for some positive b less than 1 we get $\frac{1}{2}((\frac{a}{1-b})^2 - \frac{a^2}{(1-b^2)}) = \frac{(1+b)a^2 - (1-b)a^2}{2(1-b)^2(1+b)} = \frac{a^2 b}{(1-b)^2(1+b)}$. In our case it is $a = \frac{5}{16^3}$, $b = \frac{1}{4}$ and we get $\frac{25t^2 \frac{1}{4}}{16^6 \frac{9}{16} \frac{5}{4}} = \frac{5t^2}{16^5 \cdot 9}$. For the second term we get the sum $\frac{5t^2}{16^5 \cdot 9} + \frac{5^2 t^2}{16^5 \cdot 255} = \frac{100t^2}{16^5 \cdot 17 \cdot 9}$.

The function $\tilde{w}(t)$ gets converted back to $w(p)$ with the substitution $t = p - \frac{1}{5}$ and the inclusion of the exponentials at the two ends, which appear on lines 99–110 on Page 32.

The function generic_thread(j1,j2,j3,j4) (on Page 31, starting from line 35) solves for the equalisation process, for each j1,j2,j3,j4, and does the bookkeeping for keeping track of the various quantities with which the statment of Lemma 5 is concerned.

In the final nested loop of the programme (lines 142–156 on Page 33), the global bounds required by Lemma 5 and the calculation of rate s from Lemma 7.

The output of the code confirms the statements of Lemma 5 and Lemma 7.

```
1   The value of \hat{q} in Lemma 2: 0.9916
2
3   Proof of Lemma 5.
4   The maximum difference |q_k − q_l| for j_k=j_l+1: 0.22955
5   The maximum q_k for j_k>j_l: 0.32546
6   The minimum value of p+q_3+q_4 for j_2>j_1>1: 0.23163
7
8   Proof of Lemma 7.
9   The expectation of the ratios: 0.0010956
```

5   Approximation

We can define the colouring rule in terms of a problem of local optimisation. At every point choices are made according to an objective function, which will be the sum total of three variables corresponding to the three types of choices that are made, the choice of five weights, the copying of those weights by adjacent points, and the choice of a passive pain level. We use the term *solution* for a function from X to the colour set C obeying the rule approximately, so as not to confuse it with "objective function".

The rule for the active colouring is already phrased in terms of an optimisation, the minimisation of active pain. As for the passive colourings, it is easy to make it the result of a minimisation. Let $c(y)$ be the sum total of weights directed at y. Choosing a level of $0 \leq b \leq 1$ at y results in a cost of $(1 - b) \cdot c(y) + b \cdot (1 + \frac{1}{2^{11}})$, with preference for $b = 0$ if $c(y) < 1 + \frac{1}{2^{11}}$, preference for $b = 1$ if $c(y) > 1 + \frac{1}{2^{11}}$, and any value for b if $c(y) = 1 + \frac{1}{2^{11}}$. The copying of the weight of an adjacent point is done easily by taking the absolute value of the difference between the weight and the choice. Approximate copying will be done later in an affine way with finitely many options when we present the local optimisation again as a Bayesian game.

We can see from its formulation that the invariance of the group G for any finitely additive extension is necessary for this optimisation problem. At any point, the weights toward it from different directions are given equal consideration for determining the passive pain. The same is true for the five directions involved in the choice of minimal active pain.

Of course for every positive ϵ there will be a measurable ϵ-optimal solution where optimality is understood with respect to all the measurable options. On the other hand, given a measurable solution, we can integrate the objective function over the whole space and from the need for the weights inward to equal the weights outward it follows that expectation of the objective function will not go below $\frac{1}{9}(2^{-11}$ (from the passive pain alone). This does not come close to the 0 result almost everywhere when using some non-measurable solutions. Both of these options for understanding ϵ-optimality are not interesting.

5.1 Stability

We are interested in a special kind of ϵ-optimality, which we call ϵ-stability. For each $x \in X$, let $t(x)$ be the possible improvement in the objective function at x, keeping the solution for all other $y \neq x$ fixed. Let μ be a proper finitely additive extension. A solution is ϵ-*stable* (w.r.t. μ) if the μ-expectation of $t(x)$ is no more than $\epsilon \geq 0$, meaning that there is no finite disjoint collection $A_1, \ldots, A_n$ of μ measurable sets such that the objection function can be improved by at least t_i at all points in A_i and $\sum_{i=1} \mu(A_i)t_i$ is greater than ϵ. Another way of understanding ϵ-stability is that X is a uncountable space of human society or molecules, and the solution is ϵ-stable if the gains from the individual deviations do not add up to an expectation of ϵ. Our claim is that there is a positive ϵ such that no solution that is measurable with respect to any proper finitely additive extension is ϵ-stable (and likewise for any $\epsilon^* < \epsilon$). This does not mean that if the deviations happened simultaneously there would be such an improvement for all concerned; indeed the result may be worse for all concerned.

There are two ways that a measurable solution must obey ϵ-stability. First, the set where there is significant divergence from optimality must be small. Second, where divergence from optimality exists in a subset of large measure, that divergence must be small. That can be formalised in the following way: if a solution is $\epsilon \cdot \delta$-stable, then the subset where it diverges from optimality by more than δ cannot be of measure more than ϵ.

For any $\delta > 0$ the rule $\mathbf{Q}^\delta$ applies to a point x if all three aspects of the colour at x (choosing weights, copying weights for each direction separately, and responding with passive pain) are δ-optimal at x with respect to the rule $\mathbf{Q}$ and furthermore in its passive role there is no terminating point x^* of odd level such that $x^* \to x$ and the weight given by x^* to x is more than $\frac{1}{10}2^{-12}$. The condition on non-terminating points is a way to ignore the terminating points and reduce our analysis to the non-terminating points. By δ-optimal we mean that an improvement by δ in each aspect is allowed, but no more. In this way the rule becomes a closed relation. That the δ applies to each aspect of the colouring separately greatly simplifies the following analysis.

Assume that there is an option to choose a or b, a gives a cost of 0, b a cost of $(1 - \delta)$-optimality for a positive δ means that there cannot be more than δ weight given to b, since otherwise by switching one could gain by more than δ. A choice of exactly δ for b and $1 - \delta$ for a is δ-optimal, because by switching to a only a gain of δ can be accomplished.

Again we introduce the concept of the stochastic process on non-terminating points, except that the rule $\mathbf{Q}$ is replaced by the approximate rule $\mathbf{Q}^\delta$. The stochastic process is defined only for the non-terminating points, so that we retain the analysis using the probability $\hat{q}$ for non-terminating points. As before, minimising of the future pain levels is done with the function w used at every stage. And as before, the analysis is almost identical, showing that with near certainty the pain, both passive and active, must reach unobtainable levels and therefore the assumption of a significant probability of passive pain at level 1 is not possible. There are two main differences however. First, we cannot make this claim for all positive passive

pain levels, as we did for the **Q** rule. If the passive pain level is small compared to δ, one could slip away from the logic of the rule. Second, we have to re-introduce the influence of the terminating points, for the same reason, that extremely small pain levels could be involved. The w-process is defined only on non-terminating points, but to make it apply properly we have to assume that the contributions from terminating points are sufficiently small.

With δ sufficiently small, the $\mathbf{Q}^\delta$ rule implies that the quantities directed to a passive point with passive pain of at least δ must be at least $1 + \frac{9}{20,000}$. After rounding down to quantities of the form $\frac{k}{20,000}$ for positive integers k, we have the same structure as before, that the weights toward each passive point add up to 1 (with those rare exceptions already discussed above).

We must still deal with the terminating points and a subset where the $\mathbf{Q}^\delta$ might not apply.

Lemma 8. *Let y have passive pain of level v, x be a terminating point of odd level n in the chain generated by $x \to y$ with x giving y weight of at least $\frac{1}{10 \cdot 2^{12}}$. Furthermore assume that each point between y and the terminating point of level 0 satisfies the rule $\mathbf{Q}^\delta$. It follows that $\delta \geq v(\frac{2^{-12}}{10})^{n+1}$.*

Proof. Let x be terminating of level $n \geq 3$. Because x gives weight of at least $\frac{1}{10 \cdot 2^{12}}$ to y, its active pain is at least $v10 \cdot 2^{12} - \delta 10 \cdot 2^{12})$ in all directions and therefore the terminating point of level $n-1$ next to x has passive pain of at least $v10 \cdot 2^{12} - \delta 10 \cdot 2^{12}$ (as the weight given to any other point cannot exceed $1 - \frac{1}{10 \cdot 2^{12}}$). The result follows by induction, after noticing that a point of 0 terminating level (degree 1) cannot have a passive pain level of more than $\delta 2^{-11}$. ∎

Lemma 9. *Let i be odd and let q_i be the probability of a chain $x \to y$ having terminating level i (meaning that x is a terminating point of level i). Then the probability q_1 is less than $\frac{1}{128}$ and the probability of $q_{\frac{i-1}{2}}$ is less than $\frac{1}{128}(\frac{1}{6})^{\frac{i-1}{2}}$.*

Proof. The probability that $x \to z$ and $z \neq y$ is terminating of level 0 is exactly 2^{-9}. Since there are four such z, q_1 is no more than 4×2^{-9}. Now assume that $x \to z$ and $z \neq y$ is a terminating point of level $i-1$. There is at least one x^* that is a terminating point of level $i-2$ with $x^* \to z$ and no other $\hat{x} \to z$ that is non-terminating. The probability is no more than $9 \cdot (1 - \frac{\hat{q}}{2})^8 q_{i-2}$, where $\hat{q}$ is approximately $.991603$. Since this could happen in any one of four places, q_i is no more than $4 \cdot 9 \cdot (1 - \frac{\hat{q}}{2})^8 q_{i-2}$. The conclusion holds by induction and that $4 \cdot 9 \cdot (1 - \frac{\hat{q}}{2})^8 < \frac{1}{6}$. ∎

The argument that the expectation over the paths ω in $\mathcal{P}$ of the sequences $\log(p_0(\omega)) - \log(q_1(\omega)) + \log(p_1(\omega)) - \log(q_2(\omega)) + \cdots - \log(q_i(\omega)) + \log(p_i(\omega))$ approaches positive infinity does not use the **Q** rule, rather holds for any choice of the sequences $p_0, q_1, \ldots, p_i$. All that was required to define the stochastic process, and the Markov chain lying within it, is that expectations for the p_i and q_i values are well defined at each stage. We could do this in at least one of two ways. One way would be to define a unique chain minimiser with the $\mathcal{Q}^\delta$ rule, show that it is Borel measurable, and proceed in the same way as before. Another way would be to work directly with any finitely additive G-invariant measure. We choose the latter way.

To do it the latter way, we prefer to reformulate the stochastic process with only finitely many possibilities at each stage. These choices for the finitely many values must be independent of the distributions implies by the finitely additive measure, otherwise we may run into trouble due to the lack of countable additivity. We are justified in this by Lemma 6. In what follows, we assume that there are finitely many values for the q_i and p_i and that with this assumption the expectation of $\overline{r}$ is at least $\frac{s}{2}$. To define the stochastic process, we use that the colouring function is measurable according to any finitely additive process. But implicit in the probability calculations following the binomial expansion is that the finitely additive measure is proper.

Lemma 10. *Let $\delta > 0$ be smaller than $\frac{1}{20,000}$, let $v > 0$ be the passive pain level at y, let $p > \frac{1}{20,000}$ be a weight from x to y satisfying the $\mathbf{Q}^\delta$ rule, let v_i be the passive pain level at z_i satisfying the $\mathbf{Q}^\delta$ rule with $x \to z_i$, and let q_i be a weight satisfying the $\mathbf{Q}^\delta$. It follows that $\log(v_i) \geq \log(p) - \log(q_i) - (30,000)^2 \cdot \delta \cdot \log(\frac{1}{v})$.*

Proof. The copying of the weight p at y must be within δ of p. Hence the active pain at x in the direction of y must be at least $vp - \delta$. If the active pain in the direction of z_i were not at least $vp - 2001\delta$, there would be a gain of at least δ by replacing all the weight in the y direction over to the z_i direction. As the copying of the weight q_i in the z_i direction is within δ, it follows that by choosing q_i in that direction the active pain at x is also within δv_i of $q_i v_i$. We conclude that v_i is at least $\frac{vp - 2002\delta}{q_i}$. The rest follows by taking the log of both sides and that $p \geq \frac{1}{20,000}$ and $\delta \leq \frac{1}{20,000}$. $\blacksquare$

Theorem 2. *There is a positive γ small enough so that there is no γ-stable solution to the rule $\mathbf{Q}$ that is measurable with respect to any proper finitely additive extension.*

Proof. We assume that the p values have been rounded down to integer multiples of $\frac{1}{20,000}$ and that there are finitely many q values that preserve the property that the expectation of the $\log(p_0(\omega)) - \log(q_1(\omega)) + \cdots - \log(q_i(\omega)) + \log(p_i(\omega))$ goes to infinity as i goes to infinity. We will prove that, with sufficiently small positive ϵ and δ, the subset where $\mathbf{Q}^\delta$ does not hold must exceed ϵ, given that the solution is properly measurable. We start with a hypothetical chain generated by $x \to y$ where the passive pain at y is at least $\frac{3}{4}$ and show that this can happen with only a very small probability.

There is a positive integer N such that the probability is at least $1 - \frac{1}{100}2^{-12}$ that there is some path ω with $\log(p_0(\omega)) - \log(q_1(\omega)) + \cdots + \log(-q_N)$ greater than 3. As there is a lower bound on all the $\log(p_i)$, from Lemma 10, in a non-terminating chain generated by $x \to y$ where y has a passive pain level of at least $\frac{1}{2}$ there is a δ^* such that if the $\mathcal{Q}^*$ rule is followed, then after a distance of N the probability is at least $1 - \frac{1}{100}2^{-12}$ that a passive pain level of 2 is reached, (which is impossible).

The number of vertices of distance N away from a point in a chain of length N does not exceed 50^N. So we make positive ϵ smaller than $\frac{2^{-12}}{100 \cdot 50^N}$ and make positive $\hat{\delta}$ smaller than $\delta^* \frac{1}{2} e^{-NM}$. All that is left is to control for the probability that a terminating point of odd level sends more than a weight of $\frac{1}{10}2^{-12}$ to a non-terminating point in the chain of distance no more than N from the initial y.

By Lemma 9, there is some odd i such that the chances of some terminating point of level i or more sending more than $\frac{1}{100}2^{-12}$ to any of these vertices, at most 50^N of them, is less than ϵ. So, we define our δ to be $(\frac{2^{-12}}{100})^{i+1}\hat{\delta}$ and use Lemma 8 to have our pair δ and ϵ such that a measurable $\delta\epsilon$ stable solution is not possible. ∎

Notice that the last part of the proof incorporates both possibilities of $x \to y$ being either terminating or non-terminating. This proof is far from optimal in choosing a δ and ϵ, and we are sure that this choice can be done much better.

5.2 A Bayesian Game

Our interest in paradoxical colouring rules came originally from game theory, from the desire to show that **all**, not just some, equilibria of a game are not measurable. R. Simon [4] showed that there is a Bayesian game which had no Borel measurable equilibria, though it had non-measurable equilibria. The infinite dihedral group, an amenable group, acted on the equilibria in a way that prevented any equilibrium from being measurable.

R. Simon and G. Tomkowicz [5] showed that there is a Bayesian game with non-measurable equilibria but no Borel measurable ϵ-equilibrium for small enough positive ϵ and later [7] that there is a Bayesian game with non-measurable equilibria but no measurable ϵ-equilibria for small enough positive ϵ where measurable in the above means with respect to any finitely additive measure that extends the Borel measure and respects the probability distributions of the players. These constructions involved the action of a non-amenable semi-group.

Some background to Bayesian games can be found in [7] and [2]. Of particular importance is the relationship to countable Borel equivalence relations.

Let $G = \mathbf{F}_5$ be the group generated freely by five generators, T_1, T_2, T_3, T_4, T_5 and let X be the Cantor set $\{-1,1\}^G$. Let A be the set $\{a_i^+, a_i^- \mid i = 1,2,3,4,5\}$ of cardinality 10 and let B be the set $\{b_i^+, b_i^- \mid i = 1,2,3,4,5\}$ of cardinality 10. We assume that A and B are disjoint. Let C be the set $A \cup B$ of cardinality 20 and let Ω be the space $X \times C$. Let m be the canonical probability distribution such that the measure of a cylinder set defined by

$$\{x \mid x^{g_1} = f_1, \ldots, x^{g_l} = f_l\}$$

is equal to 2^{-l} for every sequence $f_1, \ldots, f_l$ of choices in $\{-1,1\}$ and $g_1, \ldots, g_l$ are mutually distinct. Define the Borel measure μ on Ω by

$$\mu(A \times \{c\}) = \frac{m(A)}{20},$$

for every Borel measurable set A in X and any choice of c in C.

There are two players, the active player, called the green player, and the passive player, called the red player. An information set for a player is another term for a member of that player's partition. For every $x \in X$, the green player has the information set

$$(\{x\} \times A) \bigcup_{i=1,2,3,4,5} \{(T_i(x), b_i^-), (T_i^{-1}(x), b_i^+)\}.$$

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For every $y \in X$, the red player has the information set

$$(\{y\} \times B) \cup_{i=1,2,3,4,5} \{(T_i(y), a_i^-), T_i^{-1}(y), a_i^+)\}.$$

Notice that each information set is of cardinality 20 and for both players these sets partition the space. To identify the information set of the player, the green player is **centred at** x if $\{x\} \times A$ is half of its information set and the red player is **centred at** y if $\{y\} \times B$ is half of its information set, meaning that if nature chooses some (y, b) with $b \in B$ then the green player is centred at some neighbouring point while the red player is centred at y (and a symmetric statement can be made if nature chooses some (y, a) with $a \in A$). We will also refer to (x, a_i^+) as (x, a_y) where $y = T_i(x)$, (x, a_i^-) as (x, a_y) where $y = T_i^{-1}(x)$, (y, b_i^+) as (y, b_x) where $x = T_i^{-1}(y)$, and (y, b_i^-) as (y, b_x) where $x = T_i(y)$.

The green player has the choice of 5 actions, t_1, t_2, t_3, t_4, t_5. A strategy for the green player at any x is a point in the four-dimensional simplex $\Delta(\{1, 2, 3, 4, 5\})$.

The red player centred at y has the choice of $2 \cdot M^{d(y)}$ actions, where M is a very large positive integer, size to be determined later. The set of actions is

$$\{c, u\} \times \prod_{x \in S(y)} \{m_x \mid 0 \leq m_x \leq M - 1\}.$$

The symbol c stands for "crowded" and u for "uncrowded". The choice of a mixed strategy for the red player is for some point in the $2 \cdot M^{d(y)} - 1$ dimensional simplex.

The payoffs for the green player centred at x take place only in $\{x\} \times A$, meaning that in the other ten locations the payoff is uniformly zero. The payoffs for the red player centred at y take place only in $\{y\} \times B$. It is more restrictive than this. The payoffs for the green player centred at x take place only in the five locations $\{(x, a_i^+) \mid i = 1, 2, 4, 5\}$ if $x^e = +1$ or only in the five locations $\{(x, a_i^-) \mid i = 1, 2, 4, 5\}$ if $x^e = -1$. The payoffs for the red player centred at y take place only in that subset of B corresponding to the subset $S(y)$ (meaning only at the b_x with $x \in S(y)$). With both players, as each gives the probability $\frac{1}{20}$ to each point in its information set, the payoff is determined by summing over all the points giving equal weight to each. The key to understanding is that whatever is played by the green player centred at x is done uniformly throughout its information set $(\{x\} \times A) \cup_{i=1,2,3,4,5} \{(T_i(x), b_i^-), (T_i^{-1}(x), b_i^+)\}$, and the same is true for the red player centred at y and its information set.

First we define the payoffs for the green player. We consider what happens to the green player centred at x when choosing the action t_i. The action t_i has a payoff consequence only at the point (x, a_y) where $y = T_i^{x^e}$. Given that the red player centred at y chooses $(c, m_x, *)$, where $*$ stands for any choices of $m_{x'}$ for other $x' \in S(y)$, the payoff to the green player centred at (x, a_y) is $-\frac{m_x}{M}$. . Otherwise for all combination with u instead of c the payoff is 0.

Now we define the payoffs for the red player. For any $x \in S(y)$, meaning $y = T_i^{x^e}(x)$, let t_y be the action t_i. First consider a piece-wise linear convex function $f : [0, 1] \to \mathbf{R}$, where $f = \max_k f_k$ for some affine functions $f_0, \ldots, f_{M-1}$ where f is equal to f_k on $[\frac{k}{M}, \frac{k+1}{M}]$. Let s_k^+ and s_k^- be defined by $f_k(0) = s_k^-$ and $f_k(1) = s_k^+$, and the difference in slopes between consecutive f_i and f_{i+1} is always at least 1.

Define the value of the actions $(c, m_x, *)$ played against t_y at (y, b_x) to be $s^+_{m_x} + 1$, the value of the actions $(u, m_x, *)$ played against t_y at (y, b_x) to be $s^+_{m_x} + 1 + r$, for any $z \neq x$ the value of the actions $(c, m_x, *)$ played against t_z at (y, b_x) to be $s^-_{m_x}$, and for any $z \neq x$ the value of the actions $(u, m_x, *)$ played against t_z at (y, b_x) to be $s^-_{m_x} + 1 + r$.

Because the consequence for the red player centred at y by choosing some m_x for $x \in S(y)$ lies entirely at the point (y, b_x) and is also independent of the choice for c or u, the red player will chose the marginal probabilities for m_x according to $s^+_{m_x}$ and $s^-_{m_x}$ and the probability for t_y performed by the green player centred at x. By the structure of those values, no more than two m_x will be chosen in equilibrium, and only two adjacent $m_x - 1$ and m_x if the probability for t_y is exactly $\frac{m_x}{M}$. When the probability for t_y lies strictly between $\frac{m_x}{M}$ and $\frac{m_x + 1}{M}$ then only m_x will be chosen in equilibrium.

Notice that in equilibrium this game not only approximates the colouring rule of the previous sections, and it can be done so in a way for which the computer calculations also apply. If the green player centred at x chooses the action t_y with probability q, the red player centred at y will mimic with various combinations of $(c, m_{[qM]}, *), (u, m_{[qM]}, *)$ and possibly with some $(c, m_{qM-1}, *), (u, m_{qM-1}, *)$ if qM is an integer. The cost for the green player centred at x and with the action t_i will be $\frac{1}{20}$ times the red player's total probability of playing c centred at y times some quantity that is between $\frac{[qM-1]}{M}$ and $\frac{[qM]}{M}$.

To show a lack of an ϵ-equilibrium (measurable with respect to any proper finitely additive extension) using our previous argument for the lack of an ϵ-stable solution, we require that the process of copying weights is done with sufficient precision. Whatever δ worked for the $\epsilon\delta$-stability argument above, we divide by 3 and declare this to be the quantity needed for the lack of finitely additive measurable $\epsilon\delta/3$-equilibria for this Bayesian game. We make M be larger than $\frac{3}{\delta}$ to insure that there is no inaccuracy up to $\frac{\delta}{3}$ resulting from the intervals used. But lastly, we need to know that there is no relevant distortion from the mixture of the c and the u coordinates with the occasional choice of a level m_j that is not a good copy of the actual weight sent from the relevant point. We need to know that the summation of the probabilities given to the actions $(c, m_j, *)$ is sufficiently close to the average value for m_j times the average proportion for c (the product of expectations from the marginals). Lets suppose that the level m_j is incorrect when m_i is the choice closest to the correct choice on the same side as m_j. Due to the slopes of the lines defining the payoffs, we know that the cost of this mistake is at least $\frac{|j-i|(|j-i|-1)}{2} q_j$ where q_j is the probability of using m_j. We have that the summation over j of the $\frac{|j-i|(|j-i|-1)}{2} q_j$ cannot exceed $\delta/3$. It follows that $\frac{1}{M} \sum_j q_j |i-j|$ cannot exceed $2\delta/M$. By choosing M greater than $\frac{3}{\delta}$, we have the needed accuracy.

6   Conclusion

What interested us initially about paradoxical colouring rules was the connection to the Banach-Tarski Paradox.

Question 1. *For all colourings c satisfying the rule* **Q** *is there a finite partition of*

the colour space into Borel sets such that the inverse images of this finite partition along with the Borel sets $\mathcal{F}$ and shifts in G generate a finite partition of X with the Banach-Tarski property, e.g. they create two copies of X after shifting by members of G?

A further issue is raised by the expected value of the non-measurable solutions. With rule **Q** there exists non-measurable solutions where optimality is perfect, meaning the pain level of 0 almost everywhere. And with all measurable solutions there is an average passive pain level above $\frac{2^{-11}}{9}$.

Question 2. *Does there exist a problem of local optimisation or a Bayesian game such that the optimisation can be accomplished locally or the values can be measured globally, but not both simultaneously?*

Theorem 1 uses a free non-abelian group of rank 5. Given the existence of non-amenable groups without free non-abelian subgroups, demonstrated by Olshanskii and Grigorchuk, (see [TW], Chapter 12 for the details) it is natural to ask the following:

Question 3. *Does there exists a probabilistic paradoxical colouring rule that uses a non-amenable group without free non-abelian subgroups?*

The idea behind Question 3 is related to the complexity behind the proof of Theorem 1. Recall that two or three free choices were not enough to obtain a paradoxical rule. So it is natural to investigate and describe if the required complexity can be forced by generators that are not independent.

The paradox would be more graphic if passive pain began with 1 rather than $1 + \frac{1}{2^{11}}$, meaning that, outside a set of measure zero, a colouring satisfying the rule defines a flow where to every point there is no more than a total of 1 going inward (and yet in 2^{-10} of the space there is no inward flow). Could one find a colouring rule with a much stronger paradoxical effect? Instead of choosing between the five directions with the incoming arrows of variable degree, one could assume that toward any point there are always five arrows coming in but leaving from any point there are anywhere from 0 to 10 arrows. Instead of a rule defined by the avoidance of pain, the goal might be to obtain pleasure by directing weight toward where weight is lacking. If we could show that in general (except for a set of measure zero) the weights directed toward a point add up to at least 1, then the inward flow is at least 1 but the outward flow is no more than $1 - \frac{1}{2^{-10}}$. Initial investigation suggest that this could have a stronger paradoxical effect.

The proof of Theorem 2 seems convoluted. Terminating points and non-terminating points are treated separately, and it would be nice to have a unified approach. The problem is that in the calculations behind Theorem 1, integrating the effect of terminating points into the argument would involve a division by 0 (as we divide by one less than the degree of the vertex). Indeed terminating points are such that they need infinite levels of pain in order to avoid sending all weight toward them. The present approach is not efficient for establishing a good upper bound for the ϵ for which there is no measurable ϵ-stable solution. Again, a colouring rule with a stronger paradoxical effect is desired.

Question 4. *What is the largest positive ϵ such that there is a probabilistic paradoxical colouring rule defined by a local optimisation where the objective function is between 0 and 1 and there is no ϵ-stable solution that is measurable with respect to any proper finitely additive extension?*

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A Code

In this section, we present the entire C++ code used to establish the correctness of Lemma 5. Section A.1 presents the main function, which controls the parallel computation of the quantities required for the proof of Lemma 5. Section A.2 includes the code for the classes and for various helper functions needed for the numerical calculations. The entire code and compilation instructions can accessed at a GitHub repository [1].

A.1 Main Code

```cpp
// Compile command:
// g++ genanalysis_threaded.cpp -o genericanalysis -pthread -std=c++11

#include<iostream>
#include<fstream>
```

```
 6  #include<cmath>
 7  #include<thread>
 8
 9  using namespace std;
10
11  const long double accuracy = 1e−11L;
12  const long double lb=0.0L,ub=1e+8L; // Bounds for the function values
13  const int N=5;          //
14  const int NoFns=N−1;   // no of functions
15  const int types=9;
16  const int jconfigs = 495; // #(a,b,c,d)\in types^4 s.t. a<=b<=c<=d
17
18  const int maxC=19999;
19  const long double M = (long double) maxC+1;
20  const int minC=1;
21
22  long double vcur[maxC+1];
23  // long double vnext[maxC+1];
24
25  long double r;
26
27  long double maxdiff1[types+1][types+1][types+1][types+1];
28  long double maxjumpval[types+1][types+1][types+1][types+1];
29  long double mintail[types+1][types+1][types+1][types+1];
30  long double mingen[types+1][types+1][types+1][types+1];
31
32  #include "Func1.cpp"
33  #include "helpers1.cpp"
34
35  void generic_thread(int j1, int j2, int j3, int j4){
36
37    Func* Fns[NoFns];
38    long double roots[NoFns];
39    Fns[0] = new Func_iter(j1,vcur);
40    Fns[1] = new Func_iter(j2,vcur);
41    Fns[2] = new Func_iter(j3,vcur);
42    Fns[3] = new Func_iter(j4,vcur);
43
44    long double mxd=0.0L, mxj=0.0, mnt=1.0L;
45    mingen[j1][j2][j3][j4]=ub;
46
47    for(int m=1;m<=maxC;m++){
48      long double budget = 1.0 − (m/M);
49
50      SimulSolver(Fns,NoFns,roots,budget,lb,ub,accuracy);
51
52      if (j1==j2−1)
53        mxd = max(mxd,abs(roots[1]−roots[0]));
54      if (j2==j3−1)
55        mxd = max(mxd,abs(roots[2]−roots[1]));
56      if (j3==j4−1)
57        mxd = max(mxd,abs(roots[3]−roots[2]));
58      if (j1<j2)
59        mxj = max(mxj,roots[1]);
60      if (j1<j3)
```

```
61      mxj = max( mxj , roots [ 2 ] ) ;
62    if ( j1<j4 )
63      mxj = max( mxj , roots [ 3 ] ) ;
64    if ( j2<j3 )
65      mxj = max( mxj , roots [ 2 ] ) ;
66    if ( j2<j4 )
67      mxj = max( mxj , roots [ 3 ] ) ;
68    if ( j3<j4 )
69      mxj = max( mxj , roots [ 3 ] ) ;
70    if ( j1>1 && j2>j1 )
71      mnt = min( mnt , (m/M) + roots [ 2 ] + roots [ 3 ] ) ;
72
73      long double t = ( ∗ Fns [ 0 ] ) ( roots [ 0 ] ) ;
74      long double ratio = log (m∗t /M) − log ( vcur [m] ) ;
75      if ( ratio < mingen [ j1 ] [ j2 ] [ j3 ] [ j4 ] )
76        mingen [ j1 ] [ j2 ] [ j3 ] [ j4 ] = ratio ;
77
78    }
79    maxdiff1 [ j1 ] [ j2 ] [ j3 ] [ j4 ] = mxd ;
80    maxjumpval [ j1 ] [ j2 ] [ j3 ] [ j4 ] = mxj ;
81    mintail [ j1 ] [ j2 ] [ j3 ] [ j4 ] =  mnt ;
82 }
83
84 int main ( ) {
85    cout . precision ( 5 ) ;
86
87    binomcoef = new int [ types +1];
88    calcbinomcoef ( types ) ;
89
90    r=rSolver (N,0.0 ,0.9 ,1.0 , accuracy ) ; // r should be very close 1
91    cout << ”The value of r : ” << r << endl ;
92
93    // for calculating v_1 from innerfn ()
94    Func_ini f ( 1 ) ;
95    // Adjustments for the bottom and top end of the range
96    for ( int m=minC ;m<=maxC;m++){
97      vcur [m]=f . innerfn (m/M − 1.0/N) ;
98      //    cout << m << ” : ” << vcur [m] << endl ;
99      long double g = 1.0 ;
100     if (m<4000){
101       g = (m/M − 0.2 ) ;
102       g ∗= (20.0/27.0)∗g∗g ;
103       g = exp ( g ) ;
104     }
105     if (m>10000){
106       g = (m/M − 0.5 ) ;
107       g ∗= −0.25∗g∗g ;
108       g = exp ( g ) ;
109     }
110     vcur [m] ∗= g ;
111   }
112
113   thread ∗myth [ jconfigs ] ;
114   int t_i = 0 ;
115   for ( int j1 =1; j1<=types ; j1++){
```

```
116        for(int j2=j1;j2<=types;j2++){
117          for(int j3=j2;j3<=types;j3++){
118            for(int j4=j3;j4<=types;j4++){
119              myth[t_i] = new thread(generic_thread,j1,j2,j3,j4);
120              t_i++;
121            }
122          }
123        }
124      }
125
126      t_i = 0;
127      for(int j1=1;j1<=types;j1++){
128        for(int j2=j1;j2<=types;j2++){
129          for(int j3=j2;j3<=types;j3++){
130            for(int j4=j3;j4<=types;j4++){
131              if((*myth[t_i]).joinable()){
132                (*myth[t_i]).join();
133                t_i++;
134              }
135            }
136          }
137        }
138      }
139
140      long double mxd=0.0L, mxj=0.0L, mnt=1.0L;
141      long double sum = 0.0L;
142      for(int j1=1;j1<=types;j1++){
143        for(int j2=j1;j2<=types;j2++){
144          for(int j3=j2;j3<=types;j3++){
145            for(int j4=j3;j4<=types;j4++){
146              mxd = max(mxd,maxdiff1[j1][j2][j3][j4]);
147              mxj = max(mxj,maxjumpval[j1][j2][j3][j4]);
148              mnt = min(mnt,mintail[j1][j2][j3][j4]);
149
150              long double p = probfn(j1,j2,j3,j4,r);
151              long double w = mingen[j1][j2][j3][j4];
152              sum += w*p*noperm(j1,j2,j3,j4);
153            }
154          }
155        }
156      }
157
158      cout << endl << "Proof_of_Lemma_5." << endl;
159      cout << "The_maximum_difference_|q_k_-_q_l|_for_j_k=j_l+1:_"
160          << mxd << endl;
161      cout << "The_maximum_q_k_for_j_k>j_l:_" << mxj << endl;
162      cout << "The_minimum_value_of_p+q_3+q_4_for_j_2>j_1>1:_" << mnt
163          << endl << endl;
164
165      sum /= r;
166      cout << "Proof_of_Lemma_7." << endl;
167      cout << "The_expectation_of_the_ratios:_" << sum << endl;
168  }
```

A.2 Helper Functions

```cpp
1   // implements function (1-(1-x/2)^{2n-1})^{n-1} - x
2   // to find the value of r using bisection solver rSolver
3   long double rfun(long double x, long double n){
4
5      double result = pow((1.0-(x/2.0)),2*n-1);
6      result = 1 - result;
7      result = pow(result,n-1)-x;
8      return result;
9   }
10
11  // bisection solver for rfun() function
12  long double rSolver(long double n,long double target,long double lb,
13                      long double ub,long double accuracy){
14     long double mid;
15     do{
16        mid=(lb+ub)/2;
17        //    cout << mid << "    ";
18        long double y=rfun(mid,n);
19        // cout << y << endl;
20        if (y>=target)
21           lb=mid;
22        else
23           ub=mid;
24        }while (ub-lb >= accuracy);
25     return mid;
26  }
27
28  // factorial function for n>=1
29  int factorial(int n){
30     int f = 1;
31     while (n>1)
32        f *= n--;
33     return f;
34  }
35
36  // calculates no. of different permutations of j1,j2,j2,j4
37  // 1 <= j1 <= j2 <= j3 <= j4 <= maxtypes
38  int noperm(int j1, int j2, int j3, int j4){
39     int t[4];
40     t[0]=j1;
41     t[1]=j2;
42     t[2]=j3;
43     t[3]=j4;
44
45     int no = 24; // 4!
46     int i = 1; // index
47     int rep = 1; // number of repetitions of a repeated value
48     while(i<4){
49        if (t[i]==t[i-1]){
50           i++;
51           rep++;
52        }
53        else{
```

```
54        no /= factorial(rep);
55        i++;
56        rep=1;
57      }
58    }
59    no /= factorial(rep);
60    return no;
61  }
62
63  int *binomcoef;
64
65  void calcbinomcoef(const int n)
66  {
67    binomcoef[0]=1;
68    for(int i=1;i<=n;i++){
69      binomcoef[i] = binomcoef[i-1] * (n-i+1) / i;
70    }
71  }
72
73  long double probfn(int i,int j,int k,int l,long double r){
74    long double result=1.0;
75    result *= binomcoef[i];
76    result *= binomcoef[j];
77    result *= binomcoef[k];
78    result *= binomcoef[l];
79    result *= pow(1.0*r/2.0,i+j+k+l);
80    result *= pow((1.0-1.0*r/2.0),36-i-j-k-l);   // 36 = NoFns * types
81
82    return result;
83  }
84
85
86  // BisectionSolver tries to solve for x in [lb,ub] such that f(x)=target
87  // It stops when ub-lb < accuracy
88  // It initially assumes and maintains that f(lb) >= target >= f(ub)
89  long double BisectionSolver(Func* f,long double B,long double target,
90                              long double lb,long double ub,
91                              long double accuracy){
92    long double mid;
93    do{
94      mid=(lb+ub)/2;
95      long double y=(*f)(mid);
96      if (y>=target)
97        lb=mid;
98      else
99        ub=mid;
100      }while (ub-lb >= accuracy);
101    return mid;
102  }
103
104  // SimulSolver tries to solve nofn functions simultaneously such that
105  // fn[1](x_1)=fn[2](x_2)=...=fn[nofn](x_nofn) and x_1+x_2+...+x_nofn=B
106  // The common function value should be in [lb,ub]
107  // fn[i]s are assumed to be non-increasing
108  // The equalities are checked within accuracy
```

```
109   // SimulSolver updates argument array roots with corresponding x_i values
110   void SimulSolver(Func* fn[],int nofn,long double* roots, long double B,
111                    long double lb,long double ub,long double accuracy){
112     do{
113       long double sum=0.0;
114       long double mid=(lb+ub)/2;
115       for(int i=0;i<nofn;i++){
116         roots[i]=BisectionSolver(fn[i],B,mid,0.0,1.0,accuracy);
117         sum+=roots[i];
118       }
119       if (B-sum >= accuracy)
120         ub=mid;
121       else if (sum-B>accuracy)
122         lb=mid;
123       else
124         break;
125     }while (ub-lb >= accuracy);
126     return;
127   }
```

What follows contain the definition of the Func class that is used to represent functions $w(x/j_i)$ from Section 3.4.

```
1   class Func{
2   public:
3     virtual long double operator()(long double x)=0;
4     long double k;
5     virtual long double innerfn(long double x)=0;
6   };
7
8   class Func_iter:public Func{
9   public:
10    Func_iter(long double key, long double *vec);
11    long double operator()(long double x);
12    long double innerfn(long double x);
13
14  private:
15    long double *v;
16  };
17
18  class Func_ini:public Func{
19  public:
20    Func_ini(long double key);
21    long double operator()(long double x);
22    long double innerfn(long double x);
23
24  private:
25    long double *numerator, *denom, *series;
26  };
27
28  Func_iter::Func_iter(long double key, long double *vec){
29    k=key;
30    v=vec;
31  }
32
33  long double Func_iter::innerfn(long double x){
```

```
34     static int minindex=maxC+1,maxindex=0;
35     if (x>=1)
36       return 1e+10;
37     int i = (int) floor( M * x);
38     return v[i];
39   }
40
41   long double Func_iter::operator()(long double x){
42     return innerfn((1.0-x) / k) / x;
43   }
44
45   Func_ini::Func_ini(long double key){
46     k=key;
47     series = new long double[4];
48     series[0] = 1.0;
49     series[1] = 5.0/3072.0;
50     series[2] = 100.0/17.0/9.0/1024.0/1024.0;
51     series[3] = 125.0/1024.0/1024.0/1024.0/4096.0;
52   }
53
54   long double Func_ini::innerfn(long double x){
55     long double result=series[3];
56     for (int i=2;i>=0;i--){
57       result *= x;
58       result += series[i];
59     }
60     result *= (1+5.0*x) * (1+5.0*x/16.0) * (1+5.0*x/256);
61     result /= (1-5.0*x/4.0) * (1-5.0*x/64.0) * (1-5.0*x/1024.0);
62     return result;
63   }
64
65   long double Func_ini::operator()(long double x){
66     return innerfn((1.0-x)/k-0.2) / x;
67   }
```