Lecture 26: Oct 27

Last time

• Common Continuous Distribution

Today

• Common Continuous Distribution

Common continuous distributions

Exponential Distribution Denoted $X \sim Exp(\lambda)$:

• sample space: $x \ge 0$

• pdf:

$$f(x) = \begin{cases} \lambda e^{-\lambda y} & \text{for } y \geqslant 0\\ 0 & \text{otherwise} \end{cases}$$

• cdf:

$$F(x) = \int_{-\infty}^{x} f(y)dy = \begin{cases} 1 - e^{-\lambda x} & \text{for } x \ge 0\\ 0 & \text{for } x < 0 \end{cases}$$

• moments:

$$E(X) = 1/\lambda$$

$$Var(X) = 1/\lambda^{2}$$

$$M_{X}(t) = \lambda/(\lambda - t), \quad t < \lambda$$

Interpretation The exponential can be derived as the waiting time between Poisson events. Suppose that the number of events in a unit interval of time follows a $Poisson(\lambda)$ distribution. Then, let Y be the time until the first event.

$$Pr(Y > t) = Pr(0 \text{ events in } [0, t])$$

and the number of events in [0, t] follows a Poisson distribution with parameter λt . Therefore,

$$\Pr(Y > t) = e^{-\lambda t}$$
.

The cdf of Y is

$$F(t) = 1 - \Pr(Y > t) = 1 - e^{-\lambda t}$$

and hence the density is $f(t) = \lambda e^{-\lambda t}$.

Alternative parameterization Many books write the density as

$$f(y) = \begin{cases} \frac{1}{\theta} e^{-y/\theta} & \text{for } y \ge 0\\ 0 & \text{otherwise} \end{cases}$$

so that $E(Y) = \theta$ and $Var(Y) = \theta^2$. In this case $\theta = 1/\lambda$ is called the *mean parameter*, while $\lambda = 1/\theta$ is called the *rate parameter*.

Memoryless property The exponential has a memoryless property, just like the geometric.

$$\Pr(Y > s + t | Y > t) = \Pr(Y > s)$$

Same interpretation as the geometric for continuous time:

- The probability of an event in a time interval depends only on the length of the interval, not the absolute time of the interval.
- The underlying Poisson process is stationary: the rate λ is constant. (In the geometric case, the probability, p of getting an event in every discrete time unit is constant).

Shifted exponential Let $X \sim Exp(\lambda)$ and $Y = X + v, v \in \mathbb{R}$. Then, Y has the shifted exponential distribution with pdf:

$$f(y) = \begin{cases} \lambda e^{-(y-v)\lambda} & \text{for } y \geqslant v \\ 0 & \text{otherwise} \end{cases}$$

Interpretation:

- v > 0: Event is delayed
- v < 0: The news of the event is delayed

Does the shifted exponential maintain the memoryless property?

Double exponential The double exponential distribution is formed by reflecting an exponential distribution around zero. It has pdf:

$$f(x) = \frac{1}{2}\lambda e^{-\lambda|x|}, \quad x \in \mathbb{R}$$

Laplace distribution Suppose X has the above distribution with $\lambda = 1$. Now let $Y = \sigma X + \mu, \mu \in \mathbb{R}$ (shifting) and $\sigma > 0$ (scaling). Then Y has the Laplace distribution with pdf:

$$f_Y(y) = \frac{1}{2\sigma} \exp\left(-\frac{|y-\mu|}{\sigma}\right)$$

with moments

$$EY = \mu, \quad Var(Y) = 2\sigma^2$$

The Laplace distribution provides an alternative to the normal for centered data with fatter tails but all finite moments.

Normal Distribution Introduced by De Moivre (1667 - 1754) in 1733 as an approximation to the binomial. Later studied by Laplace and others as part of the Central Limit Theorem. Gauss derived the normal as a suitable distribution for outcomes that could be thought of as sums of many small deviations.

• Sample space: $\mathbb{R} = (-\infty, \infty)$

• pdf: For $Y \sim N(\mu, \sigma^2)$,

$$f(y) = \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{(y-\mu)^2}{2\sigma^2}} - \infty < y < \infty$$

- cdf: There is no closed form.
- When $\mu = 0$ and $\sigma = 1$, the distribution is called *standard normal*:

$$\Phi(y) = \Pr(Y \leqslant y), \quad \Phi(-y) = 1 - \Phi(y)$$

• Mean:

$$EY = \mu$$

• Variance:

$$Var(Y) = E(Y - \mu)^2 = \sigma^2$$

• Higher central moments:

$$E(Y - \mu)^m = \begin{cases} \frac{m!}{2^{m/2}(m/2)!} \sigma^m & m \text{ is even} \\ 0 & m \text{ is odd} \end{cases}$$

• In particular:

$$\mu_3 = E(Y - \mu)^3 = 0$$
(Skewness)
 $\mu_4 = E(Y - \mu)^4 = 3\sigma^4$

• Moment generating function:

$$M_Y(t) = \exp(\mu t + \sigma^2 t^2 / 2)$$

Standardization

$$Y \sim N(\mu, \sigma^2) \iff Z = \frac{Y - \mu}{\sigma} \sim N(0, 1)$$

Shifting and scaling:

$$Z \sim N(0,1) \iff Y = \sigma Z + \mu \sim N(\mu, \sigma^2)$$

Notes

- Normal distribution is useful in many practical settings. E.g. measurement error.
- Plays an important role in *sampling distributions* in *large samples*, since the Central Limit Theorem syas that the sums of independent identically distributed random variables are approximately normal
- There are many important distributions that can be derived from functions of normal random variables (e.g. χ^2 , t, F). We will briefly present the pdf's and sample spaces of these distributions.