Lecture 16: Sept 27

Last time

- Transformations of Random Variables
- Expected Values
- Moments

Today

- One-page one-sided letter-size hand-written cheat sheet for midterm 1
- Expected Values

Expected Values

Definition The expected value or mean of a random variable g(X), denoted by Eg(X), is

$$Eg(X) = \begin{cases} \int_{-\infty}^{\infty} g(x)f(x)dx & \text{if } X \text{ is continuous} \\ \sum_{x \in \mathcal{X}} g(x)\Pr(X = x) & \text{if } X \text{ is discrete} \end{cases}$$

Provided the integral or summation exists.

If we let g(X) = X, then we get

$$EX = \begin{cases} \int_{-\infty}^{\infty} x f(x) dx & \text{if } X \text{ is continuous} \\ \sum_{x \in \mathcal{X}} x \Pr(X = x) & \text{if } X \text{ is discrete} \end{cases}$$

Example (Exponential mean) Suppose X has an exponential (λ) distribution, $X \sim Exp(\lambda)$, that is, it has pdf given by

$$f_X(x) = \frac{1}{\lambda} e^{-x/\lambda}, \quad 0 \le x < \infty, \lambda > 0.$$

Find out EX.

Solution:

Example (Binomial mean) if X has a binomial distribution, $X \sim Binomial(n, p)$, its pmf is given by

$$\Pr(X = x) = \binom{n}{x} p^x (1-p)^{n-x}, \quad x = 0, 1, \dots, n,$$

where n is a positive integer, $0 \le p \le 1$, and for every fixed pair n and p the pmf sums to 1. Find out EX.

Solution:

The process of taking expectations is a linear operation, which means that the expectation of a linear function of X can be easily evaluated by noting that for any constants a and b, such that

$$E(aX + b) = aEX + b$$

Theorem Let X be a random variable and let a, b, and c be constants. Then for any functions $g_1(x)$ and $g_2(x)$ whose expectations exist,

- 1. $E(ag_1(X) + bg_2(X) + c) = aEg_1(X) + bEg_2(X) + c$.
- 2. If $q_1(x) \ge 0$ for all x, then $Eq_1(X) \ge 0$.
- 3. If $g_1(x) \ge g_2(x)$ for all x, then $Eg_1(X) \ge Eg_2(X)$.
- 4. If $a \leq g_1(x) \leq b$ for all x, then $a \leq Eg_1(X) \leq b$.

Proof:

Example (Method of indicators) An example of how the above properties are useful. Let $X \sim Binomial(n, p)$ for n positive integer and $0 \le p \le 1$ (n is the number of independent identical binary trials and p is the probability of success). We can write

$$X = \sum_{i=1}^{n} I_i$$

where I_i is the indicator that i^{th} trial is a success (i.e. $I_i \stackrel{\text{i.i.d.}}{\sim} Bernoulli(p)$). We have

$$EI_i = 1 \cdot p + 0 \cdot (1 - p) = p.$$

Therefore,

$$EX = \sum_{i=1}^{n} EI_i = \sum_{i=1}^{n} p = np.$$

Theorem For a non-negative random variable X (i.e. f(x) = 0 for x < 0).

$$EX = \begin{cases} \int_0^\infty (1 - F(x)) dx, & X \text{ continuous} \\ \sum_{x=0}^\infty (1 - F(x)), & X \in \mathbb{Z} \end{cases}$$

Proof:

Moments

Example (Minimizing distance) The expected value of a random variable has another property, one that we can think of as relating to the interpretation of EX as a good guess at a value of X.

Suppose we measure the distance between a random variable X and a constant b by $(X-b)^2$. The closer b is to X, the smaller this quantity is. We can now determine the value of b that minimizes $E[(X-b)^2]$ and, hence, will provide us with a good predictor of X. (Note that it does no good to look for a value of b that minimizes $(X-b)^2$, since the answer would depend on X, making it a useless predictor of X.)

We could proceed with the minimization of $E(X - b)^2$ by using calculus, but there is a simpler method:

$$E(X - b)^{2} = E(X - EX + EX - b)^{2}$$

$$= E[(X - EX) + (EX - b)]^{2}$$

$$= E(X - EX)^{2} + (EX - b)^{2} + 2E[(X - EX)(EX - b)],$$

where we have expanded the square. Note that E[(X - EX)(EX - b)] = (EX - b)E(X - EX) = 0, since EX - b is constant and comes out of the expectation, E(X - EX) = EX - EX = 0. This means

$$E(X - b)^{2} = E(X - EX)^{2} + (EX - b)^{2}.$$

Such that $E(X - b)^2$ is minimized at b = EX. And $E(X - EX)^2$ is actually the variance of X $(VarX = E(X - EX)^2)$.

The various moments of a distribution are an important class of expectations.

Definition For each integer n, the nth moment of X (or $F_X(x)$), μ'_n , is

$$\mu'_n = EX^n.$$

The nth central moment of X, μ_n , is

$$\mu_n = E(X - \mu)^n,$$

where $\mu = \mu'_1 = EX$.

Notes:

- $\mu'_0 = EX^0 = 1$
- μ'_1 is the *mean*, usually denoted by μ .
- $\mu_0 = E(X \mu)^0 = 1$
- $\mu_1 = 0$
- $\mu_2 = E(X EX)^2$ is the variance
- $\mu_3 = E(X EX)^3$ is related to the skewness.
- $\mu_4 = E(X EX)^4$ is related to the kurtosis.

Definition The variance of a random variable X is its second central moment, $Var(X) = E[(X - EX)^2]$. The positive square root of Var(X) is the standard deviation of X.

The variance gives a measure of the degree of spread of a distribution around its mean. Figure 15.1 shows a plot of two samples, one sample draws 100 numbers from a normal distribution with mean 0 and variance 1, N(0,1). The other sample draws 100 numbers from a normal distribution with mean 0 and variance 100, N(0,100).

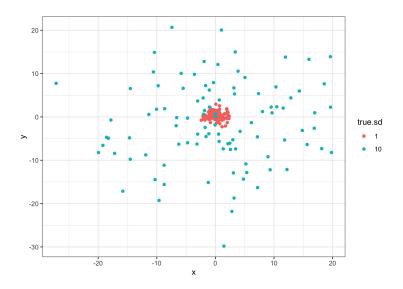


Figure 15.1: Figure 2.1.2. Two samples of 100 numbers drawn from N(0, 1) and N(0, 100).

Example (Exponential variance) Let X have the exponential (λ) distribution. We can calculate the variance of X now. Solution:

$$Var(X) = E(X - \lambda)^{2}$$
$$= \int_{0}^{\infty} (x - \lambda)^{2} \frac{1}{\lambda} e^{-x/\lambda} dx$$

Theorem If X is a random variable with finite variance, then for any constants a and b,

$$\operatorname{Var}(aX + b) = a^{2}\operatorname{Var}(X)$$
.

Proof:

From the definition, we have

$$Var (aX + b) = E [(aX + b) - E(aX + b)]^{2}$$

$$= E(aX - aEX)^{2}$$

$$= a^{2}E(X - EX)^{2}$$

$$= a^{2}Var (X).$$

It is sometimes to use an alternative formula for the variance, given by

$$Var(X) = E(X^2) - (EX)^2,$$

which is easily established by

$$Var(X) = E(X - EX)^{2} = E[X^{2} - 2XEX + (EX)^{2}]$$
$$= EX^{2} - 2(EX)^{2} + (EX)^{2}$$
$$= EX^{2} - (EX)^{2}.$$

Example (Binomial variance) Let $X \sim Binomial(n, p)$, that is ,

$$\Pr(X = x) = \binom{n}{x} p^x (1 - p)^{n - x}.$$

What is the variance of X? Solutions: