

## Lecture 30: Nov 6

Last time

- Common Continuous Distribution

Today

- HW4 Q1.3
- HW4 Q6
- Common Continuous Distribution

### Common continuous distributions

Gamma distribution   Notation:  $Y \sim \text{Gamma}(a, \lambda)$ .

- pdf:

$$f(y) = \frac{\lambda e^{-\lambda y} (\lambda y)^{a-1}}{\Gamma(a)}, \quad y \geq 0$$

where  $\Gamma(a)$  is the gamma function,

$$\Gamma(a) = \int_0^{\infty} x^{a-1} e^{-x} dx$$

- cdf: In general, there is no closed form, unless  $a$  is an integer.
- moments:

$$\begin{aligned} E(Y) &= a/\lambda \\ \text{Var}(Y) &= a/\lambda^2 \end{aligned}$$

- MGF:

$$M_Y(t) = \left( \frac{1}{1 - t/\lambda} \right)^a, \quad t < \lambda$$

Another parameterization   Same as the exponential distribution, we can let  $\beta = \frac{1}{\lambda}$ , then we have

- pdf:

$$f(y) = \frac{y^{a-1} e^{-y/\beta}}{\Gamma(a) \beta^a}, \quad y \geq 0$$

- moments:

$$\begin{aligned} EX &= \alpha\beta \\ \text{Var}(X) &= \alpha\beta^2 \end{aligned}$$

- MGF:

$$M_Y(t) = \left( \frac{1}{1 - t\beta} \right)^a, \quad t < \frac{1}{\beta}$$

Notes:

- The special case  $a = 1$  corresponds to an *exponential*( $\lambda$ )
- The parameter  $a$  is known as the *shape parameter*, since it most influences the peakedness of the distribution.
- The parameter  $\beta$  is called the *scale parameter* since most of its influence is on the spread of the distribution.
- The special case  $\text{Gamma}(a = n/2, \lambda = 1/2)$ , for integer  $n$ , corresponds to the  $\chi_n^2$  distribution with  $n$  degrees of freedom.
- The gamma distribution can be derived as the sum of  $a$  independent *exponential*( $\lambda$ ) distributions.

**Beta distribution** Notation:  $Y \sim \text{Beta}(a, b)$ .

- Sample space:  $[0, 1]$
- pdf:

$$f(y) = \frac{y^{a-1}(1-y)^{b-1}}{B(a, b)}, \quad 0 \leq y \leq 1$$

where  $B(a, b)$  is the Beta function,

$$B(a, b) = \int_0^1 x^{a-1}(1-x)^{b-1} dx = \frac{\Gamma(a)\Gamma(b)}{\Gamma(a+b)},$$

and  $\Gamma(a)$  is the gamma function. Note that if  $a$  and  $b$  are integers, then  $B(a, b)$  can be calculated in closed form.

- cdf: In general, there is no closed form, except if  $a$  and  $b$  are integers.
- moments:

$$EY = \frac{a}{a+b}$$

$$\text{Var}(Y) = \frac{ab}{(a+b)^2(a+b+1)}$$

The beta distribution is very flexible, and can take a wide variety of shapes by varying its parameters.

- Special case:  $\text{Beta}(1, 1) = U(0, 1)$ .

Omitted distributions: Weibull distribution, and Cauchy distribution.

**Exponential Families** A family of pdfs or pmfs with vector parameter  $\boldsymbol{\theta}$  is called an *exponential family* if it can be expressed as

$$f(x|\boldsymbol{\theta}) = h(x)c(\boldsymbol{\theta})\exp\left(\sum_{j=1}^k w_j(\boldsymbol{\theta})t_j(x)\right), \quad x \in S \subset \mathbb{R} \quad (1)$$

where  $S$  is not defined in terms of  $\boldsymbol{\theta}$ ,  $h(x)$ ,  $c(\boldsymbol{\theta}) \geq 0$  and the functions are just functions of the parameters specified; i.e.  $h$  is free of  $\boldsymbol{\theta}$ ,  $c(\boldsymbol{\theta})$  is free of  $x$ , etc...

Examples:

- One-dimensional: Exponential, Poisson
- Two-dimensional: Gaussian

Exponential family parameterizations are unique except for multiplying constant factors.

**Example: Gaussian** Let  $f(x|\mu, \sigma^2)$  be the  $n(\mu, \sigma^2)$  family of pdfs, where  $\boldsymbol{\theta} = (\mu, \sigma^2)$ . Then

$$\begin{aligned} f(x|\mu, \sigma^2) &= \frac{1}{\sqrt{2\pi}\sigma} \exp\left(-\frac{(x-\mu)^2}{2\sigma^2}\right) \\ &= \frac{1}{\sqrt{2\pi}\sigma} \exp\left(-\frac{\mu^2}{2\sigma^2}\right) \exp\left(-\frac{x^2}{2\sigma^2} + \frac{\mu x}{\sigma^2}\right) \end{aligned}$$

Thus

$$\begin{aligned} h(x) &= \frac{1}{\sqrt{2\pi}} & c(\mu, \sigma) &= \frac{1}{\sigma} \exp\left(-\frac{\mu^2}{2\sigma^2}\right) \\ w_1(\mu, \sigma) &= -\frac{1}{2\sigma^2} & w_2(\mu, \sigma) &= \frac{\mu}{\sigma^2} \\ t_1(x) &= x^2 & t_2(x) &= x \end{aligned}$$

The parameter space is  $(\mu, \sigma^2) \in \mathbb{R} \times (0, \infty)$ .

**Example: Binomial** Let  $f(x|p)$  be the *binomial*( $n, p$ ),  $0 < p < 1$  family of pmfs.

$$\begin{aligned} f(x|p) &= \binom{n}{x} p^x (1-p)^{n-x} = \binom{n}{x} (1-p)^n \left[\frac{p}{1-p}\right]^x \\ &= \binom{n}{x} (1-p)^n \exp\left[\log\left(\frac{p}{1-p}\right)x\right] \end{aligned}$$

Thus,

$$\begin{aligned} h(x) &= \binom{n}{x}, \quad x = 0, \dots, n & w_1(p) &= \log\left(\frac{p}{1-p}\right) \\ c(p) &= (1-p)^n, \quad 0 < p < 1 & t_1(x) &= x \end{aligned}$$

Note that this works when  $p$  is considered the parameter, while  $n$  is fixed. Also,  $p$  cannot be 0 or 1. Otherwise, the range changes.

**More examples** The following distributions belong to Exponential families:

- Continuous: exponential, Gaussian, gamma, beta,  $\chi^2$
- Discrete: Poisson, geometric, binomial (fixed # trials), negative binomial (fixed # successes)

The following distributions not exponential families:

- Continuous:  $t$ ,  $F$ , uniform E.g.:  $X \sim U(0, \theta)$

$$f_X(x) = \theta^{-1} 1(0 < x < \theta)$$

- Discrete: uniform, hypergeometric

**Theorem** If  $X$  is a random variable with pdf or pmf of the form [1](#), then

$$\begin{aligned} E \left( \sum_{i=1}^k \frac{\partial w_i(\boldsymbol{\theta})}{\partial \theta_j} t_i(X) \right) &= -\frac{\partial}{\partial \theta_j} \log c(\boldsymbol{\theta}) \\ \text{Var} \left( \sum_{i=1}^k \frac{\partial w_i(\boldsymbol{\theta})}{\partial \theta_j} t_i(X) \right) &= -\frac{\partial^2}{\partial \theta_j^2} \log c(\boldsymbol{\theta}) - E \left( \sum_{i=1}^k \frac{\partial^2 w_i(\boldsymbol{\theta})}{\partial \theta_j^2} t_i(X) \right). \end{aligned}$$

Although these equations may look formidable, when applied to specific cases they can work out quite nicely. Their advantage is that we can replace integration or summation by differentiation, which is often more straightforward.

**Example (Normal exponential family)** Let  $f(x|\mu, \sigma^2)$  be the  $N(\mu, \sigma^2)$  family of pdfs, where  $\boldsymbol{\theta} = (\mu, \sigma)$ ,  $-\infty < \mu < \infty, \sigma > 0$ . Then

$$\begin{aligned} f(x|\mu, \sigma^2) &= \frac{1}{\sqrt{2\pi}\sigma} \exp \left( -\frac{(x - \mu)^2}{2\sigma^2} \right) \\ &= \frac{1}{\sqrt{2\pi}\sigma} \exp \left( -\frac{\mu^2}{2\sigma^2} \right) \exp \left( -\frac{x^2}{2\sigma^2} + \frac{\mu x}{\sigma^2} \right) \end{aligned}$$

Define

$$\theta_1 = \frac{1}{\sigma^2} > 0, \quad \theta_2 = \frac{\mu}{\sigma^2} \in \mathbb{R}$$

Then

$$f_X(x) = \frac{\sqrt{\theta_1}}{\sqrt{2\pi}} \exp \left( -\frac{\theta_2^2}{2\theta_1} \right) \exp \left( -\theta_1 \frac{x^2}{2} + \theta_2 x \right)$$

and

$$\begin{aligned} h(x) &= 1 \text{ for all } x; \\ c(\boldsymbol{\theta}) &= c(\theta_1, \theta_2) = \exp \left( -\frac{\theta_2^2}{2\theta_1} \right), \quad (\theta_1, \theta_2) \in (0, \infty) \times \mathbb{R} \\ w_1(\boldsymbol{\theta}) &= \theta_1 & t_1(x) &= -x^2/2 \\ w_2(\boldsymbol{\theta}) &= \theta_2 & t_2(x) &= x \end{aligned}$$

Therefore, by the above theorem

$$\begin{aligned} E(X) &= -\frac{\partial}{\partial \theta_2} \log c(\boldsymbol{\theta}) = \frac{\theta_2}{\theta_1} = \mu \\ Var(X) &= -\frac{\partial^2}{\partial \theta_2^2} \log c(\boldsymbol{\theta}) = -\frac{1}{\theta_1} = \sigma^2 \end{aligned} \tag{2}$$