

## Lecture 27: Oct 30

### Last time

- Common Continuous Distribution

### Today

- Show HW4 Q9
- Common Continuous Distribution

### Common continuous distributions

**Shifted exponential** Let  $X \sim \text{Exp}(\lambda)$  and  $Y = X + v, v \in \mathbb{R}$ . Then,  $Y$  has the *shifted exponential distribution* with pdf:

$$f(y) = \begin{cases} \lambda e^{-(y-v)\lambda} & \text{for } y \geq v \\ 0 & \text{otherwise} \end{cases}$$

Interpretation:

- $v > 0$ : Event is delayed
- $v < 0$ : The news of the event is delayed

Does the shifted exponential maintain the memoryless property?

**Double exponential** The *double exponential distribution* is formed by reflecting an exponential distribution around zero. It has pdf:

$$f(x) = \frac{1}{2} \lambda e^{-\lambda|x|}, \quad x \in \mathbb{R}$$

**Laplace distribution** Suppose  $X$  has the above distribution with  $\lambda = 1$ . Now let  $Y = \sigma X + \mu, \mu \in \mathbb{R}$  (shifting) and  $\sigma > 0$  (scaling). Then  $Y$  has the *Laplace distribution* with pdf:

$$f_Y(y) = \frac{1}{2\sigma} \exp\left(-\frac{|y - \mu|}{\sigma}\right)$$

with moments

$$EY = \mu, \quad \text{Var}(Y) = 2\sigma^2$$

The Laplace distribution provides an alternative to the normal for centered data with fatter tails but all finite moments.

**Normal Distribution** Introduced by De Moivre (1667 - 1754) in 1733 as an approximation to the binomial. Later studied by Laplace and others as part of the Central Limit Theorem. Gauss derived the normal as a suitable distribution for outcomes that could be thought of as sums of many small deviations.

- Sample space:  $\mathbb{R} = (-\infty, \infty)$
- pdf: For  $Y \sim N(\mu, \sigma^2)$ ,

$$f(y) = \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{(y-\mu)^2}{2\sigma^2}} \quad -\infty < y < \infty$$

- cdf: There is no closed form.
- When  $\mu = 0$  and  $\sigma = 1$ , the distribution is called *standard normal*:

$$\Phi(y) = \Pr(Y \leq y), \quad \Phi(-y) = 1 - \Phi(y)$$

- Mean:

$$EY = \mu$$

- Variance:

$$\text{Var}(Y) = E(Y - \mu)^2 = \sigma^2$$

- Higher central moments:

$$E(Y - \mu)^m = \begin{cases} \frac{m!}{2^{m/2}(m/2)!} \sigma^m & m \text{ is even} \\ 0 & m \text{ is odd} \end{cases}$$

- In particular:

$$\begin{aligned} \mu_3 &= E(Y - \mu)^3 = 0 \text{ (Skewness)} \\ \mu_4 &= E(Y - \mu)^4 = 3\sigma^4 \end{aligned}$$

- Moment generating function:

$$M_Y(t) = \exp(\mu t + \sigma^2 t^2 / 2)$$

**Standardization**

$$Y \sim N(\mu, \sigma^2) \iff Z = \frac{Y - \mu}{\sigma} \sim N(0, 1)$$

**Shifting and scaling:**

$$Z \sim N(0, 1) \iff Y = \sigma Z + \mu \sim N(\mu, \sigma^2)$$

**Notes**

- Normal distribution is useful in many practical settings. E.g. measurement error.

- Plays an important role in *sampling distributions* in *large samples*, since the Central Limit Theorem says that the sums of independent identically distributed random variables are approximately normal
- There are many important distributions that can be derived from functions of normal random variables (e.g.  $\chi^2$ ,  $t$ ,  $F$ ). We will briefly present the pdf's and sample spaces of these distributions.

**$\chi^2$  distribution** If  $Z \sim N(0,1)$ , then  $X = Z^2$  has the  $\chi^2$  distribution with 1 degree of freedom. More generally, we have the  $\chi^2$  distribution with  $v$  degrees of freedom with pdf:

$$f(x) = \frac{(x/2)^{\frac{v}{2}-1} e^{-x/2}}{2\Gamma(v/2)}, \quad x > 0$$

where  $\Gamma(a)$  is the complete gamma function,

$$\Gamma(a) = \int_0^{\infty} x^{a-1} e^{-x} dx$$

The  $\chi^2(v)$  distribution is a special case of the gamma distribution, so it is easier to derive its properties from the gamma.

Facts about the Gamma function

- $\Gamma(a+1) = a\Gamma(a), a > 0$
- $\Gamma(1) = 1$
- $\Gamma(n) = (n-1)!$
- $\Gamma(1/2) = \sqrt{\pi}$

**Student's  $t$  and  $F$  distributions**  $Y$  has a  $t_k$  distribution ( $t$  with  $k$  degrees of freedom) if its pdf can be written as:

$$f(y) = \frac{\Gamma[(v+1)/2]}{\sqrt{v\pi}\Gamma(v/2)} \frac{1}{(1+y^2/v)^{(v+1)/2}}, \quad -\infty < y < \infty$$

$Y$  has an  $F(v_1, v_2)$  distribution if its pdf can be written as:

$$f(y) = \frac{(v_1/v_2)\Gamma[(v_1+v_2)/2]}{\Gamma(v_1/2)\Gamma(v_2/2)} \frac{(v_1 y/v_2)^{v_1/2-1}}{(1+v_1 y/v_2)^{(v_1+v_2)/2}}, \quad 0 \leq y < \infty$$

There are many important properties and relationships between these three distributions (e.g.  $\chi_k^2$  is the distribution of the sum of the squares of  $k$  independent standard normals).

Gamma distribution   Notation:  $Y \sim \text{Gamma}(a, \lambda)$ .

- pdf:

$$f(y) = \frac{\lambda e^{-\lambda y} (\lambda y)^{a-1}}{\Gamma(a)}, \quad y \geq 0$$

where  $\Gamma(a)$  is the gamma function,

$$\Gamma(a) = \int_0^{\infty} x^{a-1} e^{-x} dx$$

- cdf: In general, there is no closed form, unless  $a$  is an integer.
- moments:

$$\begin{aligned} E(Y) &= a/\lambda \\ \text{Var}(Y) &= a/\lambda^2 \end{aligned}$$

- MGF:

$$M_Y(t) = \left( \frac{1}{1 - t/\lambda} \right)^a, \quad t < \lambda$$

Another parameterization   Same as the exponential distribution, we can let  $\beta = \frac{1}{\lambda}$ , then we have

- pdf:

$$f(y) = \frac{y^{a-1} e^{-y/\beta}}{\Gamma(a) \beta^a}, \quad y \geq 0$$

- moments:

$$\begin{aligned} EX &= \alpha\beta \\ \text{Var}(X) &= \alpha\beta^2 \end{aligned}$$

- MGF:

$$M_Y(t) = \left( \frac{1}{1 - t\beta} \right)^a, \quad t < \frac{1}{\beta}$$

Notes:

- The special case  $a = 1$  corresponds to an *exponential*( $\lambda$ )
- The parameter  $a$  is known as the *shape parameter*, since it most influences the peakedness of the distribution.
- The parameter  $\beta$  is called the *scale parameter* since most of its influence is on the spread of the distribution.
- The special case  $\text{Gamma}(a = n/2, \lambda = 1/2)$ , for integer  $n$ , corresponds to the  $\chi_n^2$  distribution with  $n$  degrees of freedom.
- The gamma distribution can be derived as the sum of  $a$  independent *exponential*( $\lambda$ ) distributions.