

## Lecture 18: Oct 7

### Last time

- Midterm Exam 1

### Today

- Midterm 1 review
- Moments and moment generating function

### Moments

**Example** (Minimizing distance) The expected value of a random variable has another property, one that we can think of as relating to the interpretation of  $EX$  as a good guess at a value of  $X$ .

Suppose we measure the distance between a random variable  $X$  and a constant  $b$  by  $(X - b)^2$ . The closer  $b$  is to  $X$ , the smaller this quantity is. We can now determine the value of  $b$  that minimizes  $E[(X - b)^2]$  and, hence, will provide us with a good predictor of  $X$ . (Note that it does no good to look for a value of  $b$  that minimizes  $(X - b)^2$ , since the answer would depend on  $X$ , making it a useless predictor of  $X$ .)

We could proceed with the minimization of  $E(X - b)^2$  by using calculus, but there is a simpler method:

$$\begin{aligned} E(X - b)^2 &= E(X - EX + EX - b)^2 \\ &= E[(X - EX) + (EX - b)]^2 \\ &= E(X - EX)^2 + (EX - b)^2 + 2E[(X - EX)(EX - b)], \end{aligned}$$

where we have expanded the square. Note that  $E[(X - EX)(EX - b)] = (EX - b)E(X - EX) = 0$ , since  $EX - b$  is constant and comes out of the expectation,  $E(X - EX) = EX - EX = 0$ . This means

$$E(X - b)^2 = E(X - EX)^2 + (EX - b)^2.$$

Such that  $E(X - b)^2$  is minimized at  $b = EX$ . And  $E(X - EX)^2$  is actually the variance of  $X$  ( $Var X = E(X - EX)^2$ ).

The various moments of a distribution are an important class of expectations.

**Definition** For each integer  $n$ , the  $n$ th *moment* of  $X$  (or  $F_X(x)$ ),  $\mu'_n$ , is

$$\mu'_n = EX^n.$$

The  $n$ th *central moment* of  $X$ ,  $\mu_n$ , is

$$\mu_n = E(X - \mu)^n,$$

where  $\mu = \mu'_1 = EX$ .

Notes:

- $\mu'_0 = EX^0 = 1$
- $\mu'_1$  is the *mean*, usually denoted by  $\mu$ .
- $\mu_0 = E(X - \mu)^0 = 1$
- $\mu_1 = 0$
- $\mu_2 = E(X - EX)^2$  is the *variance*
- $\mu_3 = E(X - EX)^3$  is related to the *skewness*.
- $\mu_4 = E(X - EX)^4$  is related to the *kurtosis*.

**Definition** The *variance* of a random variable  $X$  is its second central moment,  $\text{Var}(X) = E[(X - EX)^2]$ . The positive square root of  $\text{Var}(X)$  is the *standard deviation* of  $X$ .

The variance gives a measure of the degree of spread of a distribution around its mean. Figure 18.1 shows a plot of two samples, one sample draws 100 numbers from a normal distribution with mean 0 and variance 1,  $N(0, 1)$ . The other sample draws 100 numbers from a normal distribution with mean 0 and variance 100,  $N(0, 100)$ .

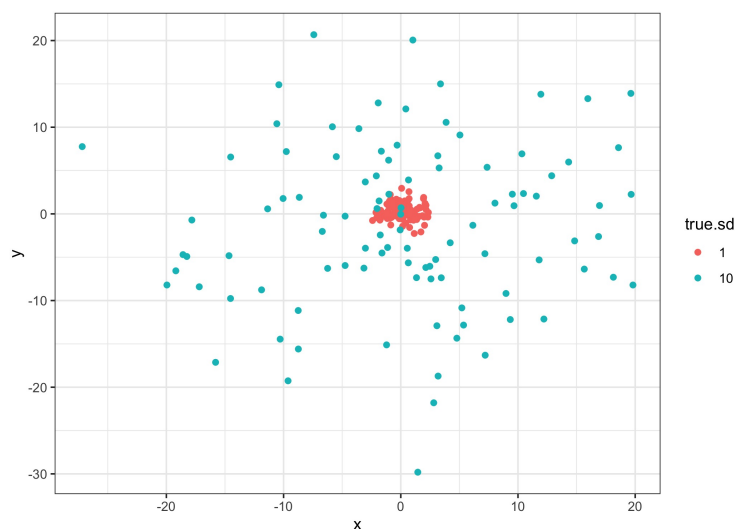


Figure 18.1: Figure 2.1.2. Two samples of 100 numbers drawn from  $N(0, 1)$  and  $N(0, 100)$ .

**Example** (Exponential variance) Let  $X$  have the exponential( $\lambda$ ) distribution. We can calculate the variance of  $X$  now.

*Solution:*

$$\begin{aligned}\text{Var}(X) &= E(X - \lambda)^2 \\ &= \int_0^\infty (x - \lambda)^2 \frac{1}{\lambda} e^{-x/\lambda} dx\end{aligned}$$

**Theorem** If  $X$  is a random variable with finite variance, then for any constants  $a$  and  $b$ ,

$$\text{Var}(aX + b) = a^2 \text{Var}(X).$$

*Proof:*

From the definition, we have

$$\begin{aligned}\text{Var}(aX + b) &= E[(aX + b) - E(aX + b)]^2 \\ &= E(aX - aEX)^2 \\ &= a^2 E(X - EX)^2 \\ &= a^2 \text{Var}(X).\end{aligned}$$

It is sometimes to use an alternative formula for the variance, given by

$$\text{Var}(X) = E(X^2) - (EX)^2,$$

which is easily established by

$$\begin{aligned}\text{Var}(X) &= E(X - EX)^2 = E[X^2 - 2XEX + (EX)^2] \\ &= EX^2 - 2(EX)^2 + (EX)^2 \\ &= EX^2 - (EX)^2.\end{aligned}$$

**Example** (Binomial variance) Let  $X \sim \text{Binomial}(n, p)$ , that is ,

$$\Pr(X = x) = \binom{n}{x} p^x (1 - p)^{n-x}.$$

What is the variance of  $X$ ?

*Solutions:*

**Definition** Let  $X$  be a random variable with cdf  $F_X$ . The *moment generating function (mgf)* of  $X$  (or  $F_X$ ), denoted by  $M_X(t)$ , is

$$M_X(t) = Ee^{tX},$$

provided that the expectation exists for  $t$  in some neighborhood of 0. That is, there is an  $h > 0$  such that, for all  $t$  in  $-h < t < h$ ,  $Ee^{tX}$  exists. If the expectation does not exist in a neighborhood of 0, we say that the moment generating function does not exist.

More explicitly, we can write the mgf of  $X$  as

$$M_X(t) = \int_{-\infty}^{\infty} e^{tx} f_X(x) dx, \quad \text{if } X \text{ is continuous,}$$

or

$$M_X(t) = \sum_x e^{tx} \Pr(X = x), \quad \text{if } X \text{ is discrete.}$$

It is easy to see how the mgf generates moments as in the following theorem.

**Theorem** If  $X$  has mgf  $M_X(t)$ , then

$$EX^n = M_X^{(n)}(0),$$

where we define

$$M_X^{(0)} = \frac{d^n}{dt^n} M_X(t) \Big|_{t=0}.$$

That is, the  $n^{th}$  moment is equal to the  $n^{th}$  derivative of  $M_X(t)$  evaluated at  $t = 0$ .

*Proof:*