### 9 Lecture 10: Feb 5

#### Last time

• Sum of squares

### Today

- R-square
- Statistical model of SLR

### Sample correlation coefficient

Definition: The sample correlation coefficient  $r_{xy}$  of the paired data  $(x_1, y_1), (x_2, y_2), ..., (x_n, y_n)$  is defined by

$$r_{xy} = \frac{\sum (x_i - \bar{x})(y_i - \bar{y})/(n-1)}{\sqrt{\sum (x_i - \bar{x})^2/(n-1) \times \sum (y_i - \bar{y})^2/(n-1)}} = \frac{s_{xy}}{s_x s_y}$$

 $s_{xy}$  is called the sample covariance of x and y:

$$s_{xy} = \frac{\sum (x_i - \bar{x})(y_i - \bar{y})}{n - 1}$$

 $s_x = \sqrt{\sum (x_i - \bar{x})^2/(n-1)}$  and  $s_y = \sqrt{\sum (y_i - \bar{y})^2/(n-1)}$  are, respectively, the sample standard deviations of X and Y.

Some properties of  $r_{xy}$ :

- $r_{xy}$  is a measure of the linear association between x and y in a dataset.
- correlation coefficients are always between -1 and 1:

$$-1 \leqslant r_{xy} \leqslant 1$$

• The closer  $r_{xy}$  is to 1, the stronger the positive linear association between x and y

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- $\bullet$  The closer  $r_{xy}$  is to -1, the stronger the negative linear association between x and y
- The bigger  $|r_{xy}|$ , the stronger the linear association
- If  $|r_{xy}| = 1$ , then x and y are said to be perfectly correlated.
- $\hat{\beta}_1 = \frac{\sum (x_i \bar{x})(y_i \bar{y})}{\sum (x_i \bar{x})^2} = \frac{s_{xy}}{s_x^2} = r_{xy} \frac{s_y}{s_x}$

### R-square

The ratio of RegSS to TSS is called the *coefficient of determination*, or sometimes, simply "r-square". it represents the proportion of variation observed in the response variable y which can be "explained" by its linear association with x.

- In simple linear regression, "r-square" is in fact equal to  $r_{xy}^2$ . (But this isn't the case in multiple regression.)
- It is also equal to the squared correlation between  $y_i$  and  $\hat{y}_i$ . (This is the case in multiple regression.)

For Davis's regression of measured on reported weight:

$$TSS = 4753.8$$
  
 $RSS = 418.87$   
 $RegSS = 4334.9$ 

Thus,

$$r^2 = \frac{4334.9}{4753.8} = 1 - \frac{418.87}{4753.8} = 0.9119$$

# The statistical model of Simple Linear Regression

Standard statistical inference in simple regression is based on a *statistical model* that describes the population or process that is sampled:

$$y_i = \beta_0 + \beta_1 x_i + \epsilon_i$$

where the coefficients  $\beta_0$  and  $\beta_1$  are the population regression parameters. The data are randomly sampled from some population of interest.

- $y_i$  is the value of the response variable
- $x_i$  is the explanatory variable
- $\epsilon_i$  represents the aggregated omitted causes of y (i.e., the causes of y beyond the explanatory variable), other explanatory variables that could have been included in the regression model, measurement error in y, and whatever component of y is inherently random.

# Key assumptions of SLR

The key assumptions of the SLR model concern the behavior of the errors, equivalently, the distribution of y conditional on x:

• Linearity. The expectation of the error given the value of x is 0:  $\mathbf{E}(\epsilon) \equiv \mathbf{E}(\epsilon|x_i) = 0$ . And equivalently, the expected value of the response variable is a linear function of the explanatory variable:  $\mu_i \equiv \mathbf{E}(y_i) \equiv \mathbf{E}(y_i|x_i) = \mathbf{E}(\beta_0 + \beta_1 x_i + \epsilon_i|x_i) = \beta_0 + \beta_1 x_i$ .

- Constant variance. The variance of the errors is the same regardless of the value of x:  $\mathbf{Var}(\epsilon|x_i) = \sigma_{\epsilon}^2$ . The constant error variance implies constant conditional variance of y on given x:  $\mathbf{Var}(y|x_i) = \mathbf{E}((y_i \mu_i)^2) = \mathbf{E}((y_i \beta_0 \beta_1 x_i)^2) = \mathbf{E}(\epsilon_i^2) = \sigma_{\epsilon}^2$ . (Question: why the last equal sign?)
- Normality. The errors are independent identically distributed with Normal distribution with mean 0 and variance  $\sigma_{\epsilon}^2$ . Write as  $\epsilon_i \stackrel{iid}{\sim} N(0, \sigma_{\epsilon}^2)$ . Equivalently, the conditional distribution of the response variable is normal:  $y_i \stackrel{iid}{\sim} N(\beta_0 + \beta_1 x_i, \sigma_{\epsilon}^2)$ .
- Independence. The observations are sampled independently.
- Fixed X, or X measured without error and independent of the error.
  - For experimental research where X values are under direct control of the researcher (i.e. X's are fixed). If the experiment were replicated, then the values of X would remain the same.
  - For research where X values are sampled, we assume the explanatory variable is measured without error and the explanatory variable and the error are independent in the population from which the sample is drawn.
- X is not invariant. X's can not be all the same.

Figure 9.1 shows the assumptions of linearity, constant variance, and normality in SLR model.



Figure 9.1: The assumptions of linearity, constant variance, and normality in simple regression. The graph shows the conditional population distributions Pr(Y|x) of Y for several values of the explanatory variable X, labeled as  $x_1, x_2, \ldots, x_5$ . The conditional means of Y given x are denoted  $\mu_1, \ldots, \mu_5$ .

## Properties of the Least-Squares estimator

Under the strong assumptions of the simple regression model, the sample least squares coefficients  $\hat{\beta}_{ls}$  have several desirable properties as estimators of the population regression coefficients  $\beta_0$  and  $\beta_1$ :

- The least-squares intercept and slope are *linear estimators*, in the sense that they are linear functions of the observations  $y_i$ .

  Proof:
- The sample least-squares coefficients are *unbiased estimators* of the population regression coefficients:

$$\mathbf{E}\left(\hat{\beta}_{0}\right) = \beta_{0}$$

$$\mathbf{E}\left(\hat{\beta}_{1}\right) = \beta_{1}$$

Proof:

• Both  $\hat{\beta}_0$  and  $\hat{\beta}_1$  have simple sampling variances:

$$\operatorname{Var}(\hat{\beta}_0) = \frac{\sigma_{\epsilon}^2 \sum x_i^2}{n \sum (x_i - \bar{x})^2}$$
$$\operatorname{Var}(\hat{\beta}_1) = \frac{\sigma_{\epsilon}^2}{\sum (x_i - \bar{x})^2}$$

Proof:

- Rewrite the formula for  $Var(\hat{\beta}_1) = \frac{\sigma_{\epsilon}^2}{(n-1)S_X^2}$ , we see that the sampling variance of the slope estimate will be small when
  - The error variance  $\sigma_{\epsilon}^2$  is small
  - The sample size n is large
  - The explanatory-variable values are spread out (i.e. have a large variance,  $S_X^2)$
- (Gauss-Markov theorem) Under the assumptions of linearity, constant variance, and independence, the least-squares estimators are BLUE (Best Linear Unbiased Estimator), that is they have the smallest sampling variance and are unbiased. (show this) *Proof:*
- Under the full suite of assumptions, the least-squares coefficients  $\hat{\beta}_0$  and  $\hat{\beta}_1$  are the maximum-likelihood estimators of  $\beta_0$  and  $\beta_1$ . (show this) *Proof:*

• Under the assumption of normality, the least-squares coefficients are themselves normally distributed. Summing up,

$$\hat{\beta}_0 \sim N(\beta_0, \frac{\sigma_{\epsilon}^2 \sum x_i^2}{n \sum (x_i - \bar{x})^2})$$

$$\hat{\beta}_1 \sim N(\beta_1, \frac{\sigma_{\epsilon}^2}{\sum (x_i - \bar{x})^2})$$