Penalized Regression

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STAT 7200-Multivariate Statistics

Objectives

- Introduce ridge regression and discuss the bias-variance trade-off
- · Introduce Lasso regression and discuss variable selection
- · Discuss cross-validation for parameter tuning

Recall: Least Squares Estimation i

- · Let $\mathbf{Y}_1, \dots, \mathbf{Y}_n$ be a random sample of size n, and let $\mathbf{X}_1, \dots, \mathbf{X}_n$ be the corresponding sample of covariates.
 - · \mathbf{Y}_i and \mathbf{X}_i are of dimension p and q, respectively.
- We will write $\mathbb Y$ and $\mathbb X$ for the matrices whose i-th row is $\mathbf Y_i$ and $\mathbf X_i$, respectively.
- From the linear model assumption, we can then write $E(\mathbb{Y}\mid \mathbb{X})=\mathbb{X}B.$
- · The least-squares criterion is given by

$$LS(B) = \operatorname{tr}\left[(\mathbb{Y} - \mathbb{X}B)^T (\mathbb{Y} - \mathbb{X}B) \right].$$

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Recall: Least Squares Estimation ii

· The minimum is attained at at

$$\hat{B} = (\mathbb{X}^T \mathbb{X})^{-1} \mathbb{X}^T \mathbb{Y}.$$

• The least-squares estimator is *unbiased*:

$$E(\hat{B} \mid \mathbb{X}) = B.$$

· If we let $\hat{\beta}_i$ be the i-th column of \hat{B} , we have

$$\operatorname{Cov}(\hat{\beta}_i, \hat{\beta}_j) = \sigma_{ij}(\mathbb{X}^T \mathbb{X})^{-1},$$

where σ_{ij} is the (i, j)-th entry of $\Sigma = \operatorname{Cov}(\mathbf{Y}_i \mid \mathbf{X}_i)$.

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Multicollinearity

- As we can see, the variance of the regression coefficients depend on the inverse of $\mathbb{X}^T\mathbb{X}$.
- Multicollinearity is when the columns of $\mathbb X$ are almost linearly dependent.
 - · Note: This can happen when a covariate is almost constant.
- As a consequence, $\mathbb{X}^T\mathbb{X}$ is nearly singular, and therefore the variance of the regression coefficients can blow up.

Ridge regression

• Solution: Add a small positive quantity along the diagonal of $\mathbb{X}^T\mathbb{X}$.

$$\cdot \ \mathbb{X}^T \mathbb{X} \to \mathbb{X}^T \mathbb{X} + \lambda I$$

 \cdot The **Ridge estimator** of B is given by

$$\hat{B}_R = (\mathbb{X}^T \mathbb{X} + \lambda I_q)^{-1} \mathbb{X}^T \mathbb{Y}.$$

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Example i

```
library(tidyverse)
url <- "https://maxturgeon.ca/w20-stat7200/prostate.csv"</pre>
prostate <- read csv(url)</pre>
# Separate into training and testing sets
data_train <- filter(prostate, train == TRUE) %>%
  dplvr::select(-train)
data_test <- filter(prostate, train == FALSE) %>%
  dplyr::select(-train)
```

Example ii

```
# OLS
model1 <- lm(lpsa ~ .,</pre>
              data = data_train)
pred1 <- predict(model1, data_test)</pre>
mean((data_test$lpsa - pred1)^2)
## [1] 0.521274
```

Example iii

Example iv

```
X test <- model.matrix(lpsa ~ .,</pre>
             data = data test)
pred2 <- X test %*% B ridge
mean((data_test$lpsa - pred2)^2)
## [1] 0.5180924
# Compare both estimates
head(cbind(coef(model1), B_ridge))
```

Example v

```
[,2]
                      [,1]
##
## (Intercept)
               0.42917013
                            0.1323063
## lcavol
                0.57654319
                            0.5709660
## lweight
                0.61402000
                            0.6160020
               -0.01900102 -0.0173843
## age
                            0.1395858
## lbph
                0.14484808
## svi
                0.73720864
                            0.6683160
```

Bias-Variance tradeoff i

• The ridge estimator is biased:

$$E(\hat{B}_R \mid \mathbb{X}) = (\mathbb{X}^T \mathbb{X} + \lambda I_q)^{-1} \mathbb{X} E(\mathbb{Y} \mid \mathbb{X})$$
$$= (\mathbb{X}^T \mathbb{X} + \lambda I_q)^{-1} \mathbb{X}^T \mathbb{X} B$$
$$\neq B.$$

• But the variance is potentially smaller:

$$\mathrm{Cov}(\hat{\beta}_i, \hat{\beta}_j) = \sigma_{ij} (\mathbb{X}^T \mathbb{X} + \lambda I_q)^{-1} \mathbb{X}^T \mathbb{X} (\mathbb{X}^T \mathbb{X} + \lambda I_q)^{-1}.$$

- This is an example of the classical bias-variance tradeoff:
 - · We increase bias and decrease variance.

Bias-Variance tradeoff ii

 Ideally, this is done in such a way to reduce the mean squared error:

$$MSE = \frac{1}{n} \operatorname{tr} \left[(\mathbb{Y} - \hat{\mathbb{Y}})^T (\mathbb{Y} - \hat{\mathbb{Y}}) \right].$$

· Should we compute the MSE with the training of the test data?

Example (cont'd) i

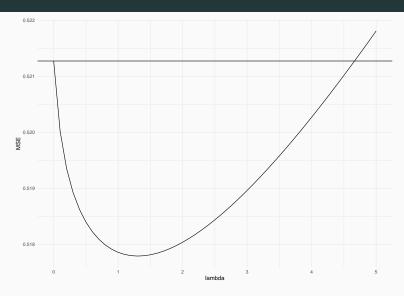
```
mse df <- purrr::map df(seq(0, 5, by = 0.1),
                         function(lambda) {
  B_ridge <- solve(crossprod(X_train) + diag(lambda, 9),</pre>
                  t(X train)) %*% Y train
  pred2 <- X test %*% B ridge
  mse <- mean((data test$lpsa - pred2)^2)</pre>
  return(data.frame(MSE = mse,
                     lambda = lambda))
  })
```

Example (cont'd) ii

```
ols_mse <- mean((data_test$lpsa - pred1)^2)

ggplot(mse_df, aes(lambda, MSE)) +
  geom_line() + theme_minimal() +
  geom_hline(yintercept = ols_mse)</pre>
```

Example (cont'd) iii



Regularized regression

 The ridge estimator can also be defined as a solution to a regularized least squares problem:

$$LS_R(B;\lambda) = \operatorname{tr}\left[(\mathbb{Y} - \mathbb{X}B)^T(\mathbb{Y} - \mathbb{X}B)\right] + \lambda \operatorname{tr}\left(B^TB\right).$$

 Yet another way to define the ridge estimator is as a solution to a constrained least squares problem:

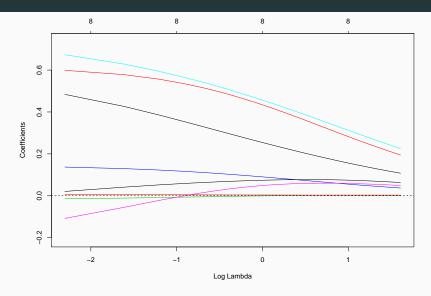
$$\min_{B} \operatorname{tr} \left[(\mathbb{Y} - \mathbb{X}B)^{T} (\mathbb{Y} - \mathbb{X}B) \right], \quad \operatorname{tr} \left(B^{T}B \right) \leq c.$$

Solution path i

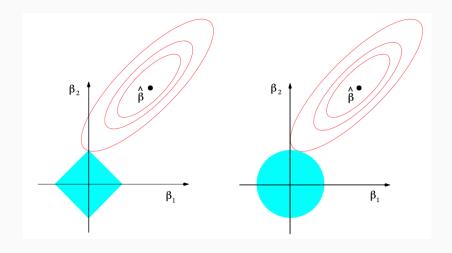
Solution path ii

```
# Plot the value of the coefficients
# as a function of lambda
plot(ridge_fit, xvar = "lambda")
abline(h = 0, lty = 2)
```

Solution path iii



Constrained regression



Lasso regression

- Lasso regression puts a different constraint on the size of the regression coefficients B:
 - · Ridge regression: $\operatorname{tr}\left(B^TB\right) = \sum_{ij} B_{ij}^2 \leq c$
 - · Lasso regression: $\|B\|_1 = \sum_{ij} |B_{ij}| \le c$
- Just as with ridge regression, this is also equivalent to a regularized least squares problem:

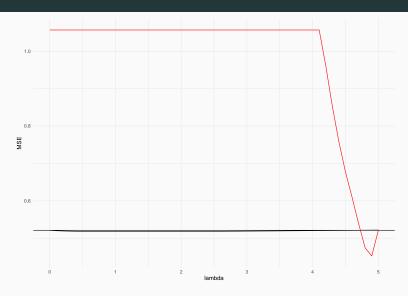
$$LS_L(B;\lambda) = \operatorname{tr}\left[(\mathbb{Y} - \mathbb{X}B)^T (\mathbb{Y} - \mathbb{X}B) \right] + \lambda ||B||_1.$$

• Major difference: Lasso regression performs variable selection.

Example (cont'd) i

Example (cont'd) ii

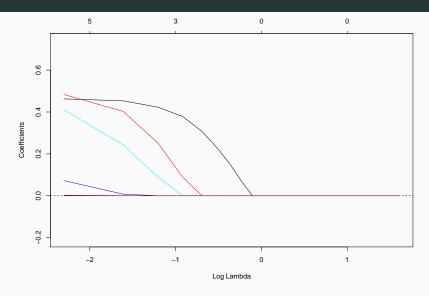
Example (cont'd) iii



Example (cont'd) iv

```
# Plot the value of the coefficients
# as a function of lambda
plot(lasso_fit, xvar = "lambda")
abline(h = 0, lty = 2)
```

Example (cont'd) v



Example (cont'd) vi

```
# Where is the min MSE?
filter(lasso mse df, MSE == min(MSE))
##
          MSF lambda
## 1 0.4526232 4.9
# What are the estimates?
coef(lasso fit, s = 4.9)
```

Example (cont'd) vii

```
## 9 x 1 sparse Matrix of class "dgCMatrix"
##
                       1
## (Intercept) 2.452345
## lcavol
## lweight
## age
## lbph
## svi
## lcp
## gleason
## pgg45
```

Comments

- · There are other forms of penalized regression:
 - · Elastic net, SCAD, adaptive lasso, group lasso, etc.
- They each have different asymptotic and finite sample properties.
 - E.g. Lasso is asymptotically biased; Elastic net and SCAD are asymptotically unbiased.
- · In general, how do we select λ when we don't have a test set?
 - · Answer: Cross-validation.

K-fold cross-validation

- Goal: Find the value of λ that minimises the MSE on test data.
- \cdot K-fold cross-validation (CV) is a resampling technique that estimates the test error from the training data.
- It is also an efficient way to use all your data, as opposed to separating your data into a training and a testing subset.

Algorithm

Let K>1 be a positive integer.

- 1. Separate your data into K subsets of (approximately) equal size.
- 2. For $k=1,\ldots,K$, put aside the k-th subset and use the remaining K-1 subsets to train your algorithm.
- 3. Using the trained algorithm, predict the values for the held out data.
- 4. Calculate MSE_k as the Mean Squared Error for these predictions.
- 5. The overall MSE estimate is given by

$$MSE = \frac{1}{K} \sum_{k=1}^{K} MSE_k.$$

Example i

```
# Take all the data
dataset <- dplyr::select(prostate, -train)</pre>
dim(dataset)
## [1] 97 9
set.seed(7200)
library(caret)
```

Loading required package: lattice

Example ii

dotPlot

##

```
##
## Attaching package: 'lattice'
## The following object is masked from 'package:openintro
##
##
       lsegments
##
## Attaching package: 'caret'
## The following object is masked from 'package:openintro
##
```

Example iii

```
## The following object is masked from 'package:purrr':
##
## lift

# 5-fold CV
trainIndex <- createFolds(dataset$lpsa, k = 5)
str(trainIndex)</pre>
```

Example iv

```
## List of 5
## $ Fold1: int [1:20] 6 8 22 23 25 27 28 32 41 46 ...
## $ Fold2: int [1:19] 5 7 15 18 20 26 29 42 44 45 ...
## $ Fold3: int [1:19] 1 11 19 21 24 30 33 48 49 50 ...
## $ Fold4: int [1:19] 3 4 10 12 16 31 34 35 39 43 ...
## $ Fold5: int [1:20] 2 9 13 14 17 36 37 38 40 47 ...
```

Example v

```
# Define function to compute MSE
compute_mse <- function(prediction, actual) {
    # Recall: the prediction comes in an array
    apply(prediction, 2, function(col) {
        mean((actual - col)^2)
      })
}</pre>
```

Example vi

```
MSEs <- sapply(trainIndex, function(indices){</pre>
  X train <- model.matrix(lpsa ~ . - 1,</pre>
                            data = dataset[-indices,])
  Y train <- dataset$lpsa[-indices]</pre>
  X_test <- model.matrix(lpsa ~ . - 1,</pre>
                            data = dataset[indices,])
  lasso_fit <- glmnet(X_train, Y_train, alpha = 1,</pre>
                        lambda = seq(0, 5, by = 0.1)
  lasso pred <- predict(lasso fit, newx = X test)</pre>
  compute mse(lasso pred, dataset$lpsa[indices])
  })
```

Example vii

```
# Each column is for a different fold
dim(MSEs)
## [1] 51 5
CV_MSE <- colMeans(MSEs)</pre>
seq(0, 5, by = 0.1)[which.min(CV_MSE)]
## [1] 0.4
```

Example viii

```
# What are the estimates?
coef(lasso fit, s = 0.4)
## 9 x 1 sparse Matrix of class "dgCMatrix"
##
## (Intercept) 1.63646053
## lcavol 0.37816202
## lweight 0.08802054
## age
## lbph
## svi
```

Example ix

```
## lcp
## gleason
## pgg45
# Conveniently, glmnet has a function for CV
# It also chooses the lambda sequence for you
X <- model.matrix(lpsa ~ . -1, data = dataset)</pre>
lasso cv fit <- cv.glmnet(X, dataset$lpsa, alpha = 1,
                           nfolds = 5)
c("lambda.min" = lasso cv fit$lambda.min,
```

"lambda.1se" = lasso cv fit\$lambda.1se)

Example x

```
## lambda.min lambda.1se
## 0.03250172 0.14400281
# What are the estimates?
coef(lasso cv fit, s = 'lambda.min')
## 9 x 1 sparse Matrix of class "dgCMatrix"
##
                          1
## (Intercept) 0.161190494
## lcavol
            0.508157223
## lweight
               0.552486889
               -0.009374709
## age
```

Example xi

lbph

```
## svi 0.594693519
## lcp .
## gleason 0.004797184
## pgg45 0.002351087
```

0.064736544

```
# 1 SE rule
coef(lasso_cv_fit, s = 'lambda.1se')
```

Example xii

```
## 9 x 1 sparse Matrix of class "dgCMatrix"
##
## (Intercept) 0.2975535
## lcavol
              0.4725492
## lweight
              0.3989087
## age
## lbph
## svi
              0.4400593
## lcp
## gleason
## pgg45
```

Summary

- Regularized regression can help reduce the mean-squared error, especially in the presence of multicollinearity
 - \cdot Ridge regression: Penalizes the L2 norm of the coefficients
 - \cdot Lasso regression: Penalizes the L1 norm of the coefficients
- Unlike ridge regression, lasso regression also performs variable selection.
 - But this comes at a cost: post-selection inference.
- K-fold cross-validation can be used to find the best value of λ .