

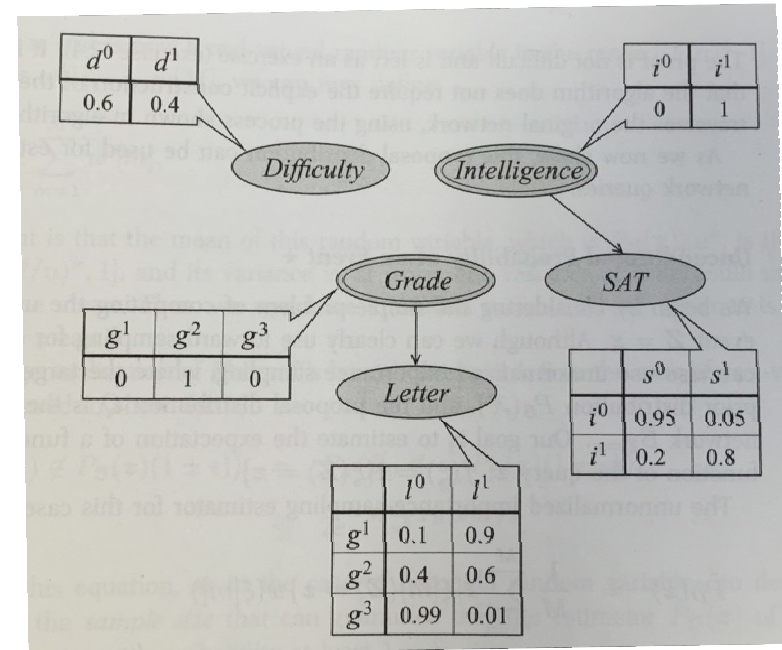
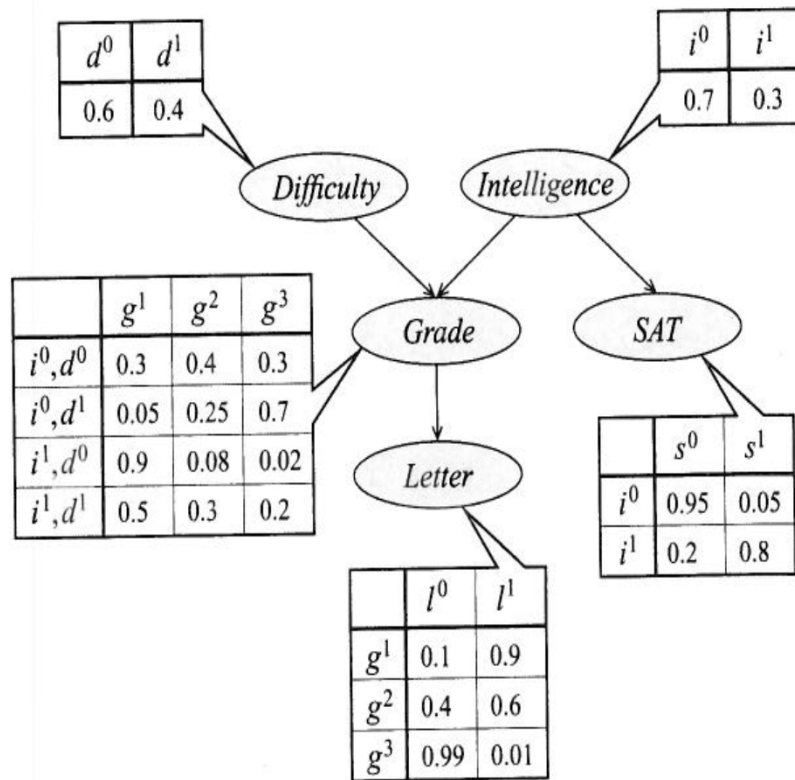
Lecture 18: March 30, 2015
cs 573: Probabilistic Reasoning
Professor Nevatia
Spring 2015

Review

- HW #5 due today
- HW #6 in two parts, one to be posted today, other part later in the week, both due 4/8/15
- Previous Lecture
 - Various sampling approaches
 - Likelihood weighting
 - Unnormalized and normalized importance sampling
 - MCMC Intro
- Today's objective
 - Markov Chain Monte Carlo (MCMC) methods
 - Intro to temporal models

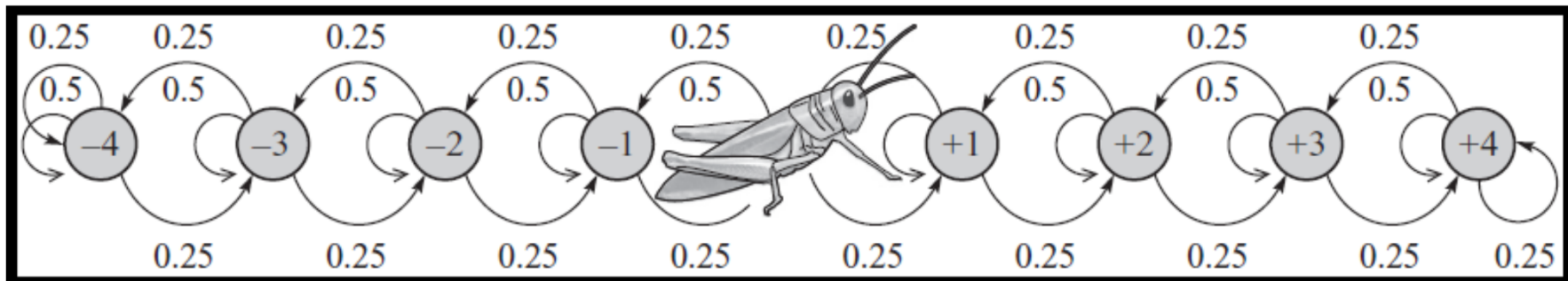
Mutilated Network

- Note: error in figure in last lecture (error in earlier edition of book)



Markov Chains

- Defined by a transition function $T(\mathbf{x} \rightarrow \mathbf{x}')$ between a pair of states $(\mathbf{x}, \mathbf{x}')$ which defines the probability of going from current state \mathbf{x} to new state \mathbf{x}' . A state is given by assignments to variables.
 - Note T will have n^2 entries if \mathbf{X} can take n values
 - Can be viewed as a matrix
- Homogeneous Markov Chain
 - Transition probability does not change over time
- Grasshopper Example
 - State: 9 integers from -4 to +4
 - Initial position: 0
 - At each instance, $T(i \rightarrow i) = .5$, $T(i \rightarrow i-1) = .25$, $T(i \rightarrow i+1) = .25$
 - At two ends, can not jump beyond (stays in the same state)
 - $T(4 \rightarrow 4) = .75$
 - Write as a transition matrix



$$P^{(t+1)}(X^{(t+1)} = x') = \sum_{x \in \text{Val}(X)} P^{(t)}(X^{(t)} = x) T(x \rightarrow x').$$

At $t=0$, $P(X^0=0) = 1$

At $t = 1$, $P(X^1 = 0) = .5$, $P(X^1 = 1) = .25$, $P(X^1 = -1) = .5$

At $t = 2$, $P(X^2 = 0) = .5 \times .5 + .25 \times .25 + .25 \times .25 = .375$

$P(X^2 = 1 \text{ or } -1) = .5 \times .25 + .25 \times .5 = .25$

$P(X^2 = 2 \text{ or } -2) = .25 \times .25 = .0625$

Position probability converges to a nearly uniform distribution
with time for this example

Stationary Distribution

- At convergence, we expect:

$$P^{(t)}(x') \approx P^{(t+1)}(x') = \sum_{x \in \text{Val}(X)} P^{(t)}(x) T(x \rightarrow x')..$$

- Stationary Distribution

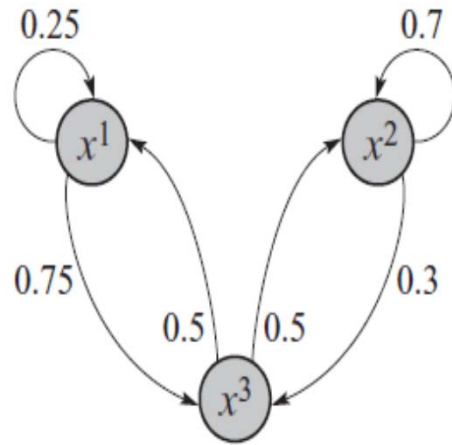
A distribution $\pi(X)$ is a stationary distribution for a Markov chain T if it satisfies:

$$\pi(X = x') = \sum_{x \in \text{Val}(X)} \pi(X = x) T(x \rightarrow x').$$

*A stationary distribution is also called an *invariant distribution*.*

- In linear algebra formulation: $T \pi(x) = \pi(x)$; *i.e.* the stationary distribution is an eigenvector of the transition matrix with eigenvalue = 1

Example 12.7



Stationary distribution must satisfy

$$\pi(x^1) = 0.25\pi(x^1) + 0.5\pi(x^3)$$

$$\pi(x^2) = 0.7\pi(x^2) + 0.5\pi(x^3)$$

$$\pi(x^3) = 0.75\pi(x^1) + 0.3\pi(x^2),$$

Transition
equations

$$\pi(x^1) + \pi(x^2) + \pi(x^3) = 1.$$

Normalize

Soution gives $\pi(x^1) = .2$, $\pi(x^2) = .5$, $\pi(x^3) = .3$

Some chains can oscillate between two distributions: *periodic chains*

In some, there are distinct regions not reachable from others;
stationary distribution depends on choice of first sample:
Reducible Markov Chains

Regular Chain

- We consider only chains with unique, stationary distributions
- *Regular chain*: There exists a k such that probability of going from x to x' in exactly k steps is > 0
 - True for grasshopper (9 steps) and fig 12.4 (2 steps)
- Thm 12.3: If a chain is regular, it has a unique stationary distribution

Multiple Transition Kernels

- Consider grasshopper to be hopping on a 2-D grid
- Define a separate transition model for each dimension (X, Y)
- Each such model is called a *kernel* .
- Cycle through multiple kernels one at a time or by some stochastic choice
- Multiple kernels have stationary distributions if each kernel has a stationary distribution
- Gibbs sampler is a special case (see next slide)

Gibbs Chain

- Gibbs chain follows the formal transition rule:

$$T_i \{(\mathbf{x}_{-i}, x_i) \rightarrow (\mathbf{x}_{-i}, x_i')\} = P(x_i' | \mathbf{x}_{-i})$$

- Can be shown that posterior distribution $P_\Phi(X | e)$ is a stationary distribution
- Can be shown that only the assignments in the Markov blanket of X_i matter in the equation above

Reversible Chain, Dynamic Balance

- A chain is said to be *reversible* if there exists a unique distribution π such that for all x, x' , it satisfies the following *detailed balance* equation

$$\pi(x) T(x \rightarrow x') = \pi(x') T(x' \rightarrow x)$$

- Pick a random starting state from $\pi(x)$ and a random transition according to transition probability, then probability of a transition from x to x' is same as from x' to x .
- If a chain is regular and satisfies detailed balance according to π then π is the unique stationary distribution of the chain
- Gibbs-chain is a reversible chain
- So is a chain constructed by the Metropolis-Hastings algorithm (next slide)

Metropolis-Hastings Algorithm

- Sample not according to P (may be hard to compute) but some other distribution Q
- Let T^Q define a transition model from x to x'
- We accept this transition according to some probability $A(x \rightarrow x')$; Effectively, the transition model is:

$$\begin{aligned} T(x \rightarrow x') &= T^Q(x \rightarrow x') A(x \rightarrow x') & x \neq x' \\ T(x \rightarrow x) &= T^Q(x \rightarrow x) + \sum_{x' \neq x} T^Q(x \rightarrow x') (1 - A(x \rightarrow x')) \end{aligned}$$

- Choice of Q is rather arbitrary but resulting chain must be regular: for example, we can choose a uniform distribution over values of X_i or a Gaussian over current state x
- To achieve detailed balance, we must have for x not equal to x'

$$\pi(x) T^Q(x \rightarrow x') A(x \rightarrow x') = \pi(x') T^Q(x' \rightarrow x) A(x' \rightarrow x).$$

- See next slide for solution

Metropolis-Hastings Algorithm

- One solution to previous equation is:

$$\mathcal{A}(\mathbf{x} \rightarrow \mathbf{x}') = \min \left[1, \frac{\pi(\mathbf{x}')T^Q(\mathbf{x}' \rightarrow \mathbf{x})}{\pi(\mathbf{x})T^Q(\mathbf{x} \rightarrow \mathbf{x}')} \right]$$

- Metropolis-Hastings algorithm
- Consider case where Q is a uniform distribution, T^Q terms cancel in equation above and we get ratio of $\pi(\mathbf{x}')$ to $\pi(\mathbf{x})$
 - If first is larger, we always transition to \mathbf{x}' (with probability 1), but may also transition when $\pi(\mathbf{x}')$ is smaller.
 - Like stochastic hill climbing
- Thm 12.5: For any proposal distribution Q , the Markov chain defined by previous slide with the above acceptance probability:
 - If the resulting chain is regular, it has a stationary distribution π .

MCMC for Graphical Models

$$\begin{aligned} \mathcal{A}(\mathbf{x}_{-i}, x_i \rightarrow \mathbf{x}_{-i}, x'_i) &= \min \left[1, \frac{\pi(\mathbf{x}_{-i}, x'_i) T_i^{Q_i}(\mathbf{x}_{-i}, x'_i \rightarrow \mathbf{x}_{-i}, x_i)}{\pi(\mathbf{x}_{-i}, x_i) T_i^{Q_i}(\mathbf{x}_{-i}, x_i \rightarrow \mathbf{x}_{-i}, x'_i)} \right] \\ &= \min \left[1, \frac{P_{\Phi}(x'_i, \mathbf{x}_{-i}) T_i^{Q_i}(\mathbf{x}_{-i}, x'_i \rightarrow \mathbf{x}_{-i}, x_i)}{P_{\Phi}(x_i, \mathbf{x}_{-i}) T_i^{Q_i}(\mathbf{x}_{-i}, x_i \rightarrow \mathbf{x}_{-i}, x'_i)} \right]. \end{aligned}$$

$$\begin{aligned} \frac{P_{\Phi}(x'_i, \mathbf{x}_{-i})}{P_{\Phi}(x_i, \mathbf{x}_{-i})} &= \frac{P_{\Phi}(x'_i | \mathbf{x}_{-i}) P_{\Phi}(\mathbf{x}_{-i})}{P_{\Phi}(x_i | \mathbf{x}_{-i}) P_{\Phi}(\mathbf{x}_{-i})} \\ &= \frac{P_{\Phi}(x'_i | \mathbf{x}_{-i})}{P_{\Phi}(x_i | \mathbf{x}_{-i})}. \end{aligned}$$

As for Gibbs sampling, we can use the observation that each variable X_i is conditionally independent of the remaining variables in the network given its Markov blanket. Letting U_i denote $\text{MB}_{\mathcal{K}}(X_i)$, and $\mathbf{u}_i = (\mathbf{x}_{-i}) \langle U_i \rangle$, we have that:

$$\frac{P_{\Phi}(x'_i | \mathbf{x}_{-i})}{P_{\Phi}(x_i | \mathbf{x}_{-i})} = \frac{P_{\Phi}(x'_i | \mathbf{u}_i)}{P_{\Phi}(x_i | \mathbf{u}_i)}.$$

Mixing Time

- How long does it take for a Markov chain to “mix” or “burn in”, *i.e.* distribution is within ϵ of π
- Analytical derivations are skipped
- Intuitively, highly skewed distributions will mix slowly
 - Hard to transition thru low probability valleys
- In general, mixing times can be rather long
- Data Driven MCMC
 - Can help drive the chain to high probability areas
 - We can use observations (data) to define the Q function

A Computer Vision Example

- Example follows
- For illustration of DDMMCMC only; material not included for assignments or exams

Model-based segmentation: A Bayesian Approach

- Problem Statement
 - Given a *foreground blob* (moving pixels) consisting of moving humans, estimate the position, size and the pose of the humans as a secondary objective
- Issues
 - Given a configuration (number, size, position and pose of hypotheses), we can evaluate its goodness (likelihood)
 - However, search space is too large for exhaustive search
 - Gradient ascent and similar methods can easily be locked into local maxima



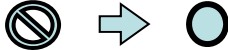




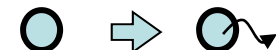
Prior probabilities and Likelihood Function

- Define $P(X)$ as a product of several terms: penalize large number of objects, assume some distribution of heights and other parameters
 - Details unimportant for today's discussion
- Likelihood function
 - Given a sample for $X = x$ (*i.e.* given number of humans, their positions and other parameters), we can compute overlap between predicted blob and observed blob
 - Likelihood is a function of this overlap and some other blob properties

Computing the MAP by MCMC

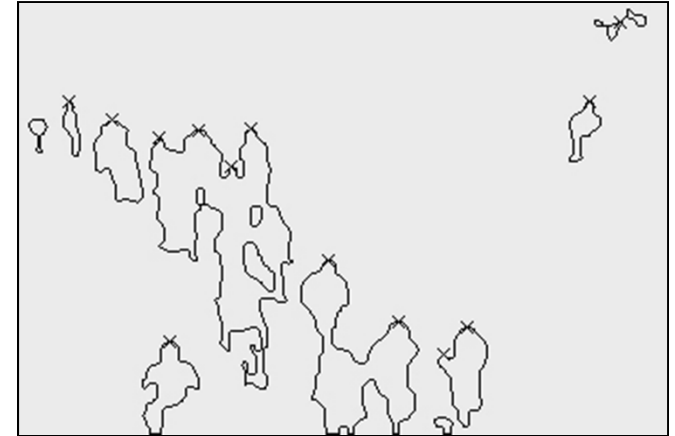
- Define a transition function that creates a regular chain.
- Use an *informed* proposal distribution (probability of this distribution is likely to be higher where the probability of the actual distribution is also high)
- Data driven MCMC (DDMCMC) uses the data to define function Q
 - For this problem, we use various heuristic cues, primarily an estimate of head like shapes present in the image

Reversible Markov chain dynamics

- Dynamics to explore the solution space
 - Adding an object 
 - Removing an object 
 - Split an object into two 
 - Merge two objects into one 
 - Switching between different models 
 - Stochastic diffusion 
- *Jumps* between subspaces of different object number and *diffuses* within each subspace
- In each iteration, one action is chosen randomly
- Results in a Markov chain which is *reversible*, *irreducible* and *aperiodic*.

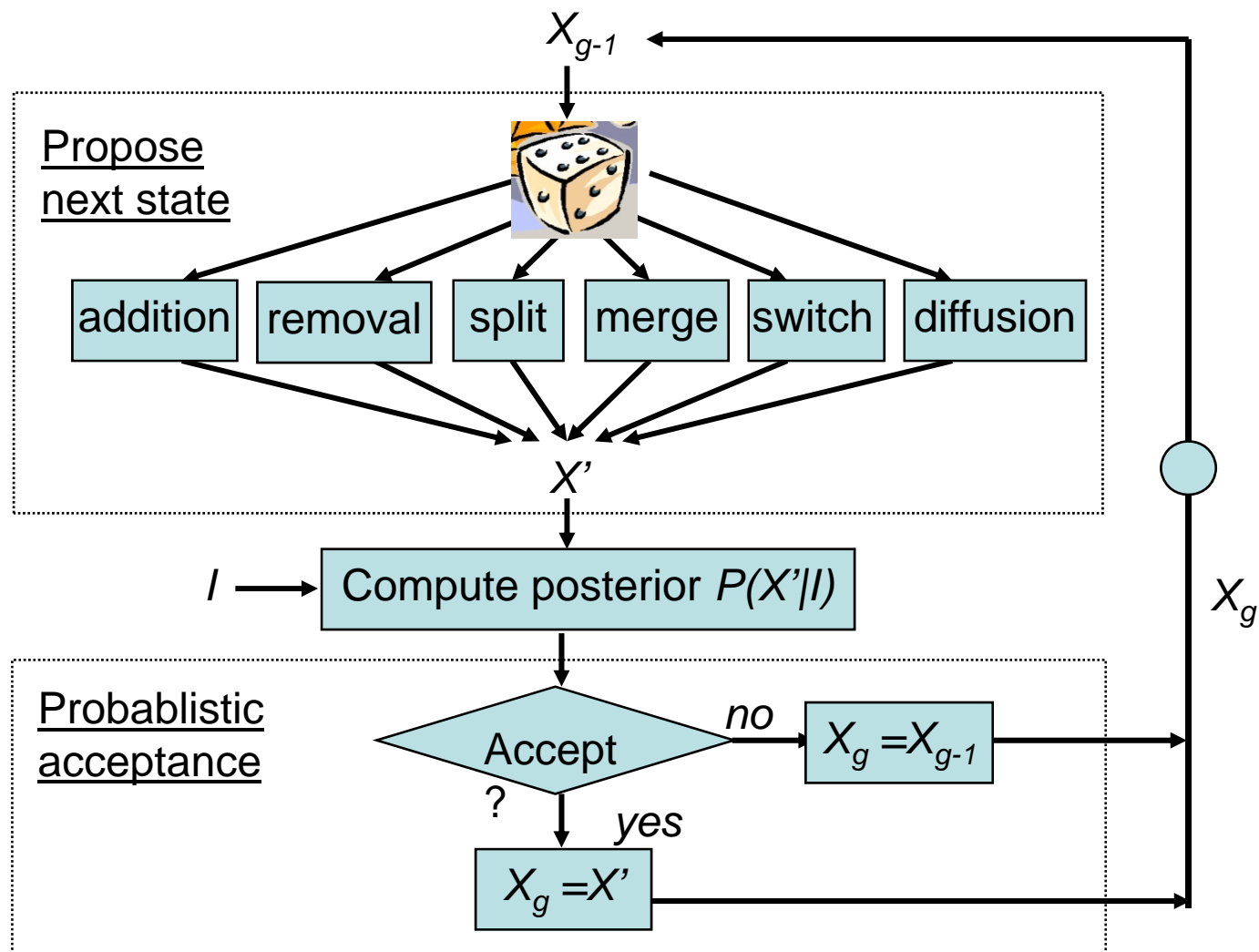
Informed proposals

- Addition 1: head candidates by foreground boundaries
 - Peaks of the foreground boundary
 - Does not work for interior heads
- Addition 2: head candidates by intensity edges



Ω

Summary of the algorithm



- The number of iterations needed depends on the complexity of the data

Result: sequence “Topping”



- 2000 iterations per frame

Next Class

- Read sections 6.2, 15.1 and 15.2 of the KF book