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Financial Econometrics Slides-09: Volatility Modeling

<https://tutorcs.com>

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Lecture Plan

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- Motivation for modeling return volatility
- Measures of return volatility
- Conditional volatility via smoothing
- ARCH
 - Conditional variance is a function of info set;
 - It captures "clustering" in return series;
 - It explains non-normality of return, to some extent;
 - It can be used to improve interval forecasts and VaR (Value at Risk);
 - Estimation and testing.

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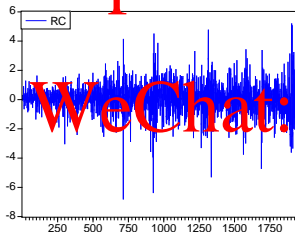
Introduction and Motivation

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eg. Volatility in NYSE Composite Index return

- Clustering.
- Squared returns are strongly autocorrelated.

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Sample: 1 1931
Included observations: 1930

Autocorrelation	Partial Correlation	AC	PAC	Q-Stat	Prob
1	0.179	0.179	62.044	0.000	
2	0.168	0.140	116.47	0.000	
3	0.175	0.131	176.90	0.000	
4	0.109	0.044	198.89	0.000	
5	0.184	0.130	264.80	0.000	
6	0.113	0.035	289.14	0.000	
7	0.136	0.067	325.10	0.000	
8	0.12	0.06	359.60	0.000	
9	0.105	0.039	372.11	0.000	
10	0.095	0.015	389.80	0.000	
11	0.107	0.042	411.90	0.000	
12	0.102	0.028	432.21	0.000	
13	0.066	-0.006	440.72	0.000	
14	0.053	-0.016	446.26	0.000	
15	0.041	-0.016	449.68	0.000	
16	0.097	0.052	467.86	0.000	
17	0.040	-0.018	470.96	0.000	
18	0.069	0.024	479.94	0.000	
19	0.063	0.014	487.77	0.000	
20	0.045	0.005	491.70	0.000	

Motivation

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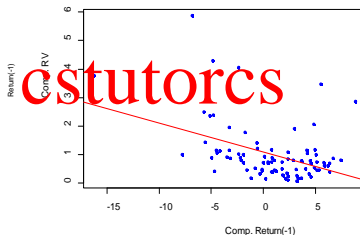
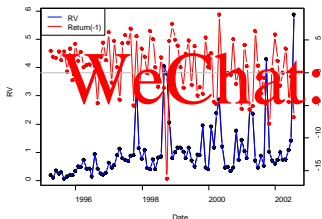
eg. Volatility in NYSE Composite index return

- Monthly realised variance:

RV = sample mean of squared daily returns in a month

- RV is negatively correlated to lagged monthly return.

$\text{Cor}(RV, \text{Return}(-1)) = -0.429$



Motivation

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- Importance of return volatility

Asset pricing, risk management and portfolio selection

Substantial dependence structure in volatility

- Clustering:

- strong autocorrelations in squared returns,
- large variations tend to be followed by large variations

- Asymmetry:

negative returns tend to cause more volatility than positives

- ARMA are unable to capture these features

Conditional variance is constant in ARMA.

Amend ARMA with a suitable conditional variance: ARCH and GARCH models.

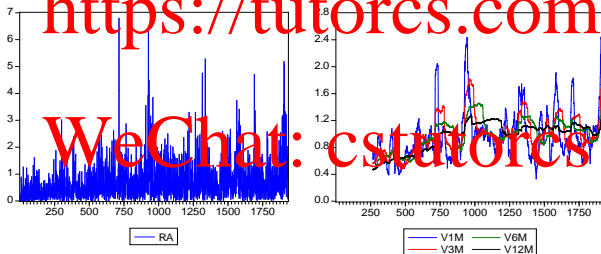
Volatility

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Measures of return volatility (tendency of variation)

- Historical volatility: Sample variance or Stddev

eg. NYSE composite return: Sample Stddev



Realized Volatility

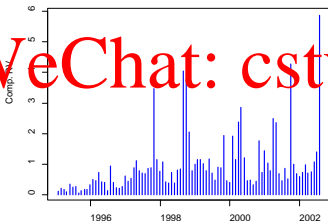
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- **Realised volatility:** Realised variance = Sample mean of squared higher frequency returns

(eg. daily RV = Sample mean of squared 5-min returns)

eg. NYSE composite return: Monthly realised variance

RV = Sample mean of squared daily returns in a month



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Realized Volatility

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Measures of return volatility

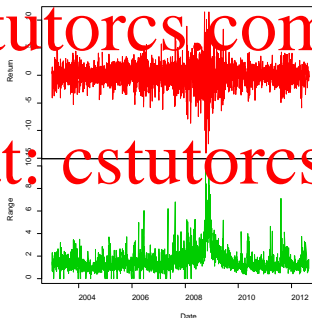
- **Range (high/low):**

$100 \times \ln(\text{high/low})$ in a time interval (eg, a day)

eg. BRP

daily return and range

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Implied Volatility

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Implied volatility:

standard deviation derived from options prices

- Option of an asset, the right to buy/sell the asset at a future time (maturity) at a fixed price (strike).
- Given the price of an option, maturity, strike and risk-free interest rate, the std deviation can be recovered from Black-Scholes formula, known as IV.
- IV represents market's opinions on the return's std deviation.

Black-Scholes formula:

price of an option = $f(\text{stdev}, \text{maturity}, \text{strike}, r_f - \text{rate})$

Implied Volatility

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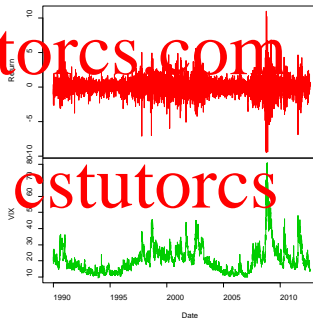
• Implied volatility:

eg. **VIX**: index of IVs of a set of options on the SP500 index

SP500 daily return & VIX

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Conditional Volatility

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Conditional variance of return

- $\sigma_{t+1|t}^2 = \text{Var}(r_{t+1}|\Omega_t)$,

where $r_{t+1} = 100\ln(P_{t+1}/P_t)$ is the return and

Ω_t is the information set at the end of period t .

- It should capture “clustering” or autocorrelations in squared returns, and facilitate predicting the return volatility

Knowing it helps to

- assess the risk of an asset via value-at-risk;
- price options;
- form mean-variance efficient portfolios.

Conditional Volatility

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Exponentially weighted moving average (EWMA)

- The squared returns $\{r_t^2, r_{t-1}^2, \dots, r_1^2\}$ carry info about the volatility as $E(r_t^2) \equiv \text{variance}$.
- A weighted average of squared returns is an approximation to the conditional variance. Recent observations should weigh more.
- EWMA: for $0 < \lambda < 1$,

$$\sigma_{t+1|t}^2 = (1 - \lambda)(r_t^2 + \lambda r_{t-1}^2 + \lambda^2 r_{t-2}^2 + \dots)$$

- weights decay exponentially;
- weights sum up to 1.
- RiskMetrics recommend $\lambda = 0.94$

EWMA

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- EWMA: alternative formulation

- $\sigma_{1|0}^2 = r_1^2$

- $\sigma_{t+1|t}^2 = (1 - \lambda)r_1^2 + \lambda\sigma_{t|t-1}^2$, for $t = 1, 2, 3, \dots$

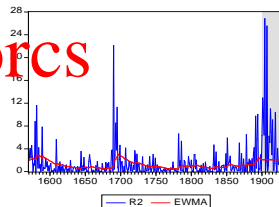
– Quick and easy,

– Can be used as 1-step ahead prediction.

eg. NYSE Composite return:

$\lambda = 0.94$

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ARCH (autoregressive conditional heteroskedasticity) Engle (1982) – Nobel price winner 1993

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Autoregressive conditional heteroscedasticity (ARCH) models are a class of models where the conditional variance evolves according to an autoregressive process.

First define the conditional variance of the error term u_t to be

$$\sigma_t^2 = \text{var}(\mu_t | \mu_{t-1}, \mu_{t-2}, \dots) = E((\mu_t - E(\mu_t))^2 | \mu_{t-1}, \mu_{t-2}, \dots)$$

As it is usually assumed that $E(\mu_t) = 0$

$$\sigma_t^2 = \text{var}(\mu_t | \mu_{t-1}, \mu_{t-2}, \dots) = E(\mu_t^2 | \mu_{t-1}, \mu_{t-2}, \dots) = E_{t-1}(\mu_t^2)$$

The ARCH(1) model assumes

$$\sigma_t^2 = E_{t-1}(\mu_t^2) = \alpha_0 + \alpha_1 \mu_{t-1}^2$$

The conditional variance captures 'clustering': large past shock leads to large conditional variance.

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ARCH (autoregressive conditional heteroskedasticity)

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Extension:

- An ARCH(q) model is given by

$$\sigma_t^2 = \alpha_0 + \alpha_1 \mu_{t-1}^2 + \alpha_2 \mu_{t-2}^2 + \dots + \alpha_q \mu_{t-q}^2$$

• Under ARCH, the conditional mean equation can take any form. An example of a full model would be

$$y_t = \beta_1 + \beta_2 x_{2t} + \beta_3 x_{3t} + \beta_4 x_{4t} + \mu_t \quad \mu_t \sim N(0, \sigma_t^2)$$

$$\sigma_t^2 = \alpha_0 + \alpha_1 \mu_{t-1}^2$$

Alternative notation

$$y_t = \beta_1 + \beta_2 x_{2t} + \beta_3 x_{3t} + \beta_4 x_{4t} + \mu_t$$

$$\mu_t = \nu_t \sigma_t \quad \nu_t \sim N(0, 1)$$

$$\sigma_t^2 = \alpha_0 + \alpha_1 \mu_{t-1}^2$$

Properties of ARCH(1)

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- ARCH(1): $\mu_t | \Omega_{t-1} \sim N(0, \sigma_t^2)$, $\Omega_{t-1} = \{y_{t-1}, \mu_{t-1}, y_{t-2}, \mu_{t-2}, \dots\}$ is the info set at the end of period $t-1$.
 $\sigma_t^2 = \alpha_0 + \alpha_1 \mu_{t-1}^2$, $\alpha_0 > 0$, $0 \leq \alpha_1 < 1$
 - Its conditional variance is time varying: $\text{Var}(\mu_t | \Omega_{t-1}) = \sigma_t^2$, $\text{CI}(95\%) = ?$
 - It is WN: (Us: LIE) $E(\mu_t) = 0$, $\text{Var}(\mu_t) = \frac{\alpha_0}{1 - \alpha_1}$, $\text{Cov}(\mu_t, \mu_{t-j}) = 0$
 But it is NOT independent WN or iid WN. Why?

Proof of properties

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Definition (Law of Iterated Expectations)

For a random variable Y and information sets Ω_1 and Ω_2 , the the LIE states that

$$E(Y|\Omega_1) = E(E(Y|\Omega_2)|\Omega_1),$$

where information set Ω_1 is included in information set Ω_2 .

Example: $E(Y_t|\Omega_{t-2}) = E(E(Y_t|\Omega_{t-1})|\Omega_{t-2})$

Special Case: If Ω_1 is empty set, then $E(Y) = E(E(Y|\Omega_2))$.

$$\mu_t = \nu_t \sigma_t = \nu_t \sqrt{\alpha_0 + \alpha_1 \mu_{t-1}^2}, \text{ where } \nu_t \text{ is } N(0, 1)$$

- ① Unconditional Expectation of μ_t . We have that $\mu_t|\Omega_{t-1} \sim N(0, \sigma_t^2)$:

$$E(\mu_t) = E[E(\mu_t|\Omega_{t-1})] \quad (1)$$

$$E(\mu_t|\Omega_{t-1}) = 0 \quad (2)$$

$$E(\mu_t) = 0. \quad (3)$$

Proof of properties

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$\mu_t = \alpha_0 + \alpha_1 \mu_{t-1}^2 = \nu \sqrt{\alpha_0 + \alpha_1 \mu_{t-1}^2}$, where ν_t is $N(0, 1)$

② Unconditional variance of μ_t . We have that

$$E(\mu_t^2) = E[E(\mu_t^2 | \Omega_{t-1})] \quad (4)$$

$$= E[E(\nu^2(\alpha_0 + \alpha_1 \mu_{t-1}^2) | \Omega_{t-1})] \quad (5)$$

$$= E[(\alpha_0 + \alpha_1 \mu_{t-1}^2) E(\nu^2 | \Omega_{t-1})] \quad (6)$$

$$= E[\alpha_0 + \alpha_1 \mu_{t-1}^2] = E[\alpha_0 + \alpha_1 E(\mu_{t-1}^2 | \Omega_{t-2})] \quad (7)$$

$$= \alpha_0 + \alpha_1 E[\alpha_0 + \alpha_1 \mu_{t-2}^2] \quad (8)$$

$$= \dots = \alpha_0 (1 + \alpha_1 + \alpha_1^2 + \dots + \alpha_1^{t-1}) + \alpha_1^t E(\mu_0^2) \quad (9)$$

As $t \rightarrow \infty$, the unconditional variance converges if $\alpha_1 < 1$ to:

$E(\mu_t^2) = \frac{\alpha_0}{1-\alpha_1}$. \rightarrow Unconditionally, the process μ_t is **homoskedastic**.

Properties of ARCH(1)

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- It can be alternatively expressed as: $\mu_t = \sigma_t v_t$, $v_t \sim i.i.d N(0, 1)$, where $v_t = \mu_t / \sigma_t$ is the **standardised shock**.
- When model is correct, v_t^2 should have no autocorrelation
- The unconditional distribution of μ_t is **NOT normal**, with heavy tails (kurtosis > 3).

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MLE of ARCH(1)

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- An example: $AR(1) = ARCH(1)$

$$y_t = c + \phi_1 y_{t-1} + \mu_t, \mu_t | \Omega_{t-1} \sim N(0, \sigma_t^2), \quad (10)$$

$$\sigma_t^2 = \alpha_0 + \alpha_1 \mu_{t-1}^2, \quad (11)$$

$$\alpha_0 \geq 0, \alpha_1 \leq 1. \quad (12)$$

- Likelihood of $\{y_1, y_2, \dots, y_{T-1}, y_T\}$:

$$L(\Theta) = f(y_T | \Omega_{T-1}) f(y_{T-1} | \Omega_{T-2}) \cdots f(y_2 | \Omega_1) f(y_1)$$

$$f(y_t | \Omega_{t-1}) = (2\pi\sigma_t^2)^{-1/2} \exp\left\{-\frac{(y_t - c - \phi_1 y_{t-1})^2}{2\sigma_t^2}\right\}. \quad (13)$$

- ML Estimator maximises the Log likelihood function

$$\ln L(\Theta) = -\frac{T}{2} \ln(2\pi) - \frac{1}{2} \sum_{t=2}^T \left[\ln(\sigma_t^2) + \frac{(y_t - c - \phi_1 y_{t-1})^2}{\sigma_t^2} \right].$$

MLE of ARCH(1)

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- ML estimators are generally consistent with an asymptotic normal distribution.

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- The above holds even when the conditional normality $\mu_t | \omega_{t-1} \sim N(0, \sigma_t^2)$ is **incorrectly** assumed, as long as the conditional mean and conditional variance are correctly specified.

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- With robust quasi ML standard errors, inference is standard.

ML Estimation

Example

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eg. NYSE composite return: AR(1)-ARCH(s)

Dependent Variable: RC
Method: ML - ARCH (Marquardt) - Normal distribution

Sample (adjusted): 3 1931
Included observations: 1929 after adjustments
Convergence achieved after 17 iterations
Bollerslev-Wooldridge test for conditional heteroskedasticity
Variance based on 1929 observations
GARCH = $C + C(1) \text{RESID}(-1)^2 + C(5) \text{RESID}(-2)^2 + 0(6) \text{RESID}(-3)^2 + C(7) \text{RESID}(-4)^2 + C(8) \text{RESID}(-5)^2$

	Coefficient	Std. Error	z-Statistic	Prob.
C	0.073023	0.020198	3.615442	0.0003
AR(1)	0.109876	0.026312	4.175893	0.0000

Variance Equation

	Coefficient	Std. Error	z-Statistic	Prob.
C	0.313844	0.034075	9.2124	0.0000
RESID(-1) ²	0.13226	0.01011	3.38769	0.0007
RESID(-2) ²	0.5362	0.01578	4.48731	0.0000
RESID(-3) ²	0.0411	0.01243	2.87400	0.0041
RESID(-4) ²	1.6056	0.0364	4.40983	0.0000
RESID(-5) ²	0.059128	0.023724	2.492331	0.0127

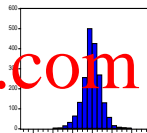
R-squared 0.001836 Mean dependent var 0.035168
Adjusted R-squared -0.001801 S.D. dependent var 1.006452
S.E. of regression 1.007357 Akaike info criterion 2.664282
Sum squared resid 1949.371 Schwarz criterion 2.687361
Log likelihood -2561.700 F-statistic 0.504880
Durbin-Watson stat 2.069097 Prob(F-statistic) 0.831451

Inverted AR Roots .11

Correlogram of E2

Sample: 1 1931
Included observations: 1929

Autocorrelation	Partial Correlation	AC	PAC	Q-Stat	Prob
1	0.216	0.216	90.395	0.000	
2	0.165	0.124	142.97	0.000	
3	0.187	0.137	210.51	0.000	
4	0.112	0.112	279.00	0.000	
5	0.038	0.038	347.50	0.000	
6	0.113	0.053	416.00	0.000	
7	0.100	0.033	484.50	0.000	
8	0.068	0.019	553.00	0.000	
9	0.104	0.041	621.50	0.000	
10	0.105	0.035	690.00	0.000	
11	0.051	0.012	758.50	0.000	
12	0.050	0.015	827.00	0.000	
13	0.046	0.007	895.50	0.000	
14	0.058	0.029	964.00	0.000	
15	0.086	0.059	1032.50	0.000	



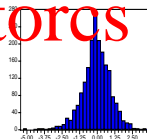
Series: E
Sample: 1 1931
Observations: 1929

Mean -0.033709
Median -0.037415
Maximum 5.362314
Minimum -6.772783
Std. Dev. 1.004962
Skewness -0.108864
Kurtosis 7.131158
Jarque-Bera 1384.431
Probability 0.000000

Correlogram of V2

Sample: 1 1931
Included observations: 1929

Autocorrelation	Partial Correlation	AC	PAC	Q-Stat	Prob
1	0.013	0.013	0.000	0.988	
2	-0.013	-0.013	0.361	0.832	
3	-0.012	-0.012	0.636	0.888	
4	-0.044	-0.044	1.3525	0.369	
5	0.023	0.023	2.069	0.154	
6	0.036	0.036	2.786	0.176	
7	0.025	0.025	3.503	0.181	
8	0.028	0.028	4.220	0.167	
9	0.020	0.019	4.937	0.189	
10	0.030	0.033	5.654	0.184	
11	0.025	0.020	6.371	0.165	
12	0.045	0.049	7.088	0.081	
13	0.019	0.014	7.805	0.109	
14	0.015	0.019	8.522	0.132	
15	0.091	0.094	9.239	0.174	
16	0.023	0.026	9.956	0.178	



Series: V
Sample: 1 1931
Observations: 1929

Mean -0.048030
Median -0.043219
Maximum 3.427025
Minimum -5.188789
Std. Dev. 0.999709
Skewness -0.415353
Kurtosis 4.558816
Jarque-Bera 250.3099
Probability 0.000000

Example

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eg. NYSE composite return: AR(1)-ARCH(5)

- Squared residuals (E^2) of AR(1) have strong autocorrelation.

Squared standardised residuals (V^2) are not autocorrelated

- Residuals (E) of AR(1) have larger kurtosis.

Standardised residuals (V) larger negative skewness.

- Normality is rejected for both E and V .

Two essential checks for the "adequacy" of a model

- ▶ Adequate mean equation: E (residuals) has no autocorrelation;
- ▶ Adequate variance equation: V^2 has no autocorrelation

Comments and limitations of ARCH

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Advantages of ARCH

- It is able to capture 'clustering' in return series or the autocorrelation in squared returns.
- It facilitates volatility forecasting.
- It explains partially, non-normality in return series.

Limitations of ARCH

- ▶ In ARCH(q), the q may be selected by AIC, SIC or LR test. The correct value of q might be very large. The model might not be parsimonious. (eg. ARCH(1) would not work for the composite return)
- ▶ The conditional variance σ_t^2 cannot be negative: Requires non-negativity constraints on the coefficients. Sufficient (but not necessary) condition is: $\alpha_i \geq 0$ for all $i = 0, 1, 2, \dots, q$. Especially for large values of q this might be violated