

### 3.7 MSstatsTMT: Details of parameter estimation and testing for differential abundance

**Pairwise comparisons between conditions** Model-based testing for differentially abundant proteins between pairs of conditions is carried out through a contrast of the condition means. Denote  $\beta = (\beta_1, \beta_2, \dots, \beta_C)$  the parameters associated with the terms *Condition* in Supplementary Equation S1-5, and  $l = (l_1, l_2, \dots, l_C)$  a vector of coefficients where  $\sum l_c = 0$ . A contrast is defined as  $l^T \beta = \sum_c l_c \beta_c$ . For example, a pairwise comparison testing proteins for differential abundance between Condition 1 and Condition 2 can be expressed as a contrast with  $l = (1, -1, 0, \dots, 0)$ . We are interested in the null hypothesis  $l^T \beta = \sum l_c \beta_c = 0$ .

We estimate the parameters of the model using restricted maximum likelihood [7, 8, 9] to obtain  $\hat{\beta}$ , the contrast  $l^T \hat{\beta} = \sum_c l_c \hat{\beta}_c$ , and the corresponding t-statistic [10]

$$t = \frac{l^T \hat{\beta}}{\sqrt{l s^2 \hat{V} l^T}} \quad (6)$$

Here  $s^2$  is the estimate of  $\sigma^2$  in Supplementary Equation S1, and  $\hat{V}$  is the unscaled variance-covariance matrix of  $\hat{\beta}$ , such that  $s^2 \hat{V}$  is the variance-covariance matrix of  $\hat{\beta}$ , and  $\sqrt{l s^2 \hat{V} l^T}$  is the standard error of the contrast. The estimates  $s^2$  and  $\hat{V}$  are obtained by restricted maximum likelihood. The matrix  $\hat{V}$  is a function of estimates  $\hat{\sigma}_M^2$  for the random effect of mixtures,  $\hat{\sigma}_T^2$  for the random effect of technical replicates, and  $\hat{\sigma}_S^2$  for the random effect of subject. Therefore, the standard error of the contrast takes into account both technical variance  $\sigma_M^2$ ,  $\sigma_T^2$ ,  $\sigma^2$  and biological variance  $\sigma_S^2$ .

In general and unbalanced designs, the degrees of freedom of the t-statistic in Supplementary Equation S6 are derived by Satterthwaite approximation [11]

$$df = \frac{2(s^2 l^T \hat{V} l^T)^2}{[\text{VAR}(s^2 l^T \hat{V} l^T)]}. \quad (7)$$

The calculation of  $[\text{VAR}(s^2 l^T \hat{V} l^T)]$  is described in [12, 10]. In unbalanced design, different contrasts of a same protein may have different degrees of freedom.

**Empirical Bayes variance moderation** When the number of biological replicates in each condition is small, we adopt Empirical Bayes moderation from *limma* [13]. Briefly, we assume that the estimate of the error variance  $\sigma^2$  in Supplementary Equation S1  $s^2$  follows a scaled chi-square distribution with  $\nu$  degrees of freedom

$$s^2 | \sigma^2 \sim \frac{\sigma^2}{\nu} \chi_\nu^2, \quad (8)$$

The variance  $\sigma^2$  of each protein is in turn assumed to follow a scaled inverse chi-square prior distribution with prior degrees of freedom  $d_0$  and prior variance  $s_0^2$

$$\frac{1}{\sigma^2} \sim \frac{1}{d_0 s_0^2} \chi_{d_0}^2. \quad (9)$$

The degree of freedom  $\nu$  is estimated as implemented in the R package *lmerTest* [12]. The parameters  $d_0$  and  $s_0^2$  are estimated from the distribution of the observed  $s^2$  of all the proteins using an Empirical Bayes approach as implemented in the R package *limma* [14]. Finally, the posterior variance estimate is incorporated into the residual variance of each protein

$$\tilde{s}^2 = \frac{s_0^2 d_0 + s^2 \nu}{d_0 + \nu}. \quad (10)$$

Then, the moderated t-statistic for the contrast  $l$  becomes

$$\tilde{t} = \frac{l^T \hat{\beta}}{\sqrt{l \tilde{s}^2 \hat{V} l^T}}, \quad (11)$$

with  $df + d_0$  degrees of freedom.