

Chapter 11: Fundamentals of Estimation

Ram Gopal, Dan Philps, and Tillman Weyde

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Load packages

```
library(bbmle)
```

Introduction

Suppose we have a sample of data as follows.

```
x = c(4,3,5,13,7,10,9,9,3,6,4,3,7,10,7,6,7,8,7,7)
```

We believe that this sample comes from a population with the data generation process that can be described by a Poisson distribution. This is depicted in the figure below.

Population

Data Generation Process = $Pois(\lambda)$

4,3,5,13,7,10,9,9,3,6,4,3,7,10,7,6,7,8,7,7

In **statistical estimation**, we want to ask the question: what is the estimate of λ that best explains the data that we observe? The most popular approach for estimation is called **maximum likelihood estimation**.

Likelihood Principle

Our objective is to find the value of λ that best explains the data we have. Let us start this quest with an initial value of 6 for λ .

Notice that the first observation in our sample has a value of 4. If λ is 6, what is the probability of finding a value of 4?

```
x = c(4,3,5,13,7,10,9,9,3,6,4,3,7,10,7,6,7,8,7,7)
lam = 6
dpois(x = 4, lambda = lam)
```

```
## [1] 0.1339
```

The second observation has a value of 3. What is the probability of observing this?

```
dpois(x = 3, lambda = lam)
```

```
## [1] 0.08924
```

Similarly, we can compute the probabilities of each of the values in the sample.

```
lam = 6
p = dpois(x = x, lambda = lam)
print(round(p,3))
```

```
## [1] 0.134 0.089 0.161 0.005 0.138 0.041 0.069 0.069 0.089 0.161 0.134 0.089
## [13] 0.138 0.041 0.138 0.161 0.138 0.103 0.138 0.138
```

Let us introduce some notation first. Let x_1, \dots, x_n denote the sample observations and $P(x_1|\lambda), \dots, P(x_n|\lambda)$ denote the corresponding probabilities for a given λ . If the observations are **independent**, then the probability of observing the entire sample is

This is termed the likelihood function.

The likelihood function as defined above invokes the **iid** assumption, which stands for independent and identically distributed observations.

```
lam = 6
p = dpois(x = x, lambda = lam)
L = prod(p)
L
```

You will notice that likelihood value is very small. Such low values are difficult to work with from a computational standpoint. It is more convenient to use the log of likelihood.

$$LL = \log(P(x_i|\lambda), \dots P(x_n|\lambda))$$

```
lam = 6
p = dpois(x = x, lambda = lam)
LL = sum(log(p))
LL
```

Let us package this into a function.

```
LLpois = function(lam){
  p = dpois(x = x, lambda = lam)
  LL = sum(log(p))
  return(LL)
}
LLpois(5)
```

```
## [1] -52.52
```

The **maximum likelihood principle** states that the best estimate of λ is one that maximizes the likelihood of the observed sample. This makes sense because L in essence is the joint probability of observing the entire sample of data. The best estimate of λ is the one that maximizes L . Note that maximizing L is equivalent to maximizing the loglikelihood LL .

In this simple example, it can be shown mathematically that the maximum likelihood estimate of λ is simply the sample average.

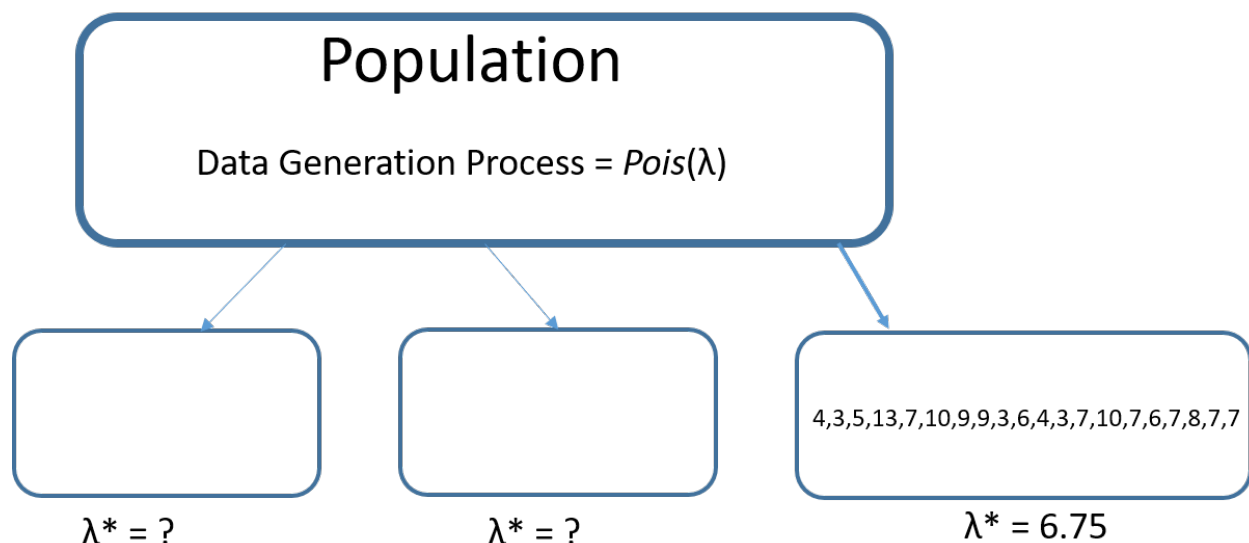
```
mean(x)
```

```
## [1] 6.75
```

In more complicated settings, we may not be able to solve the optimization problem analytically. In such cases, we use numerical computational methods to get the maximum likelihood estimates.

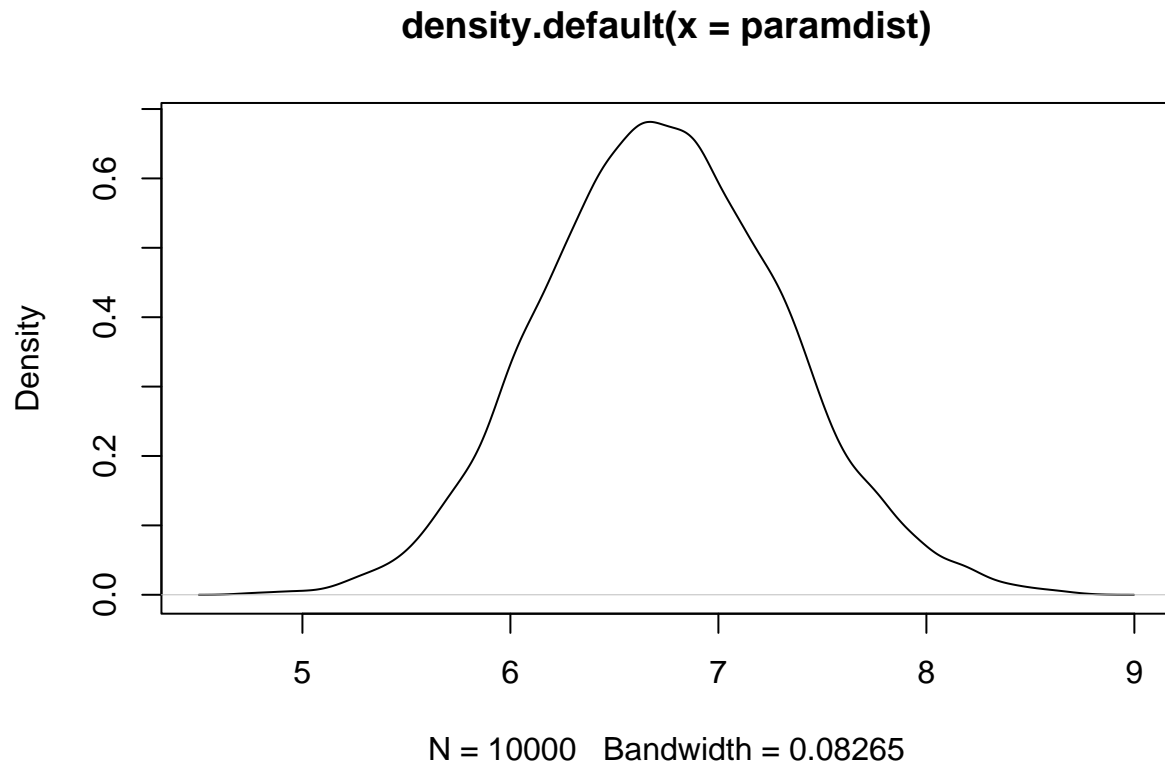
Yet Another Distribution!

Based on our sample of data, the maximum likelihood estimate for λ is 6.75.



What if you take another sample from the same population? A different sample drawn from the population could potentially yield a different estimate for λ . Let us create other samples using the bootstrapping approach and compute the estimate for each.

```
set.seed(987654321)
paramdist = replicate(10000, mean(sample(x = x, replace = T)))
plot(density(paramdist))
```

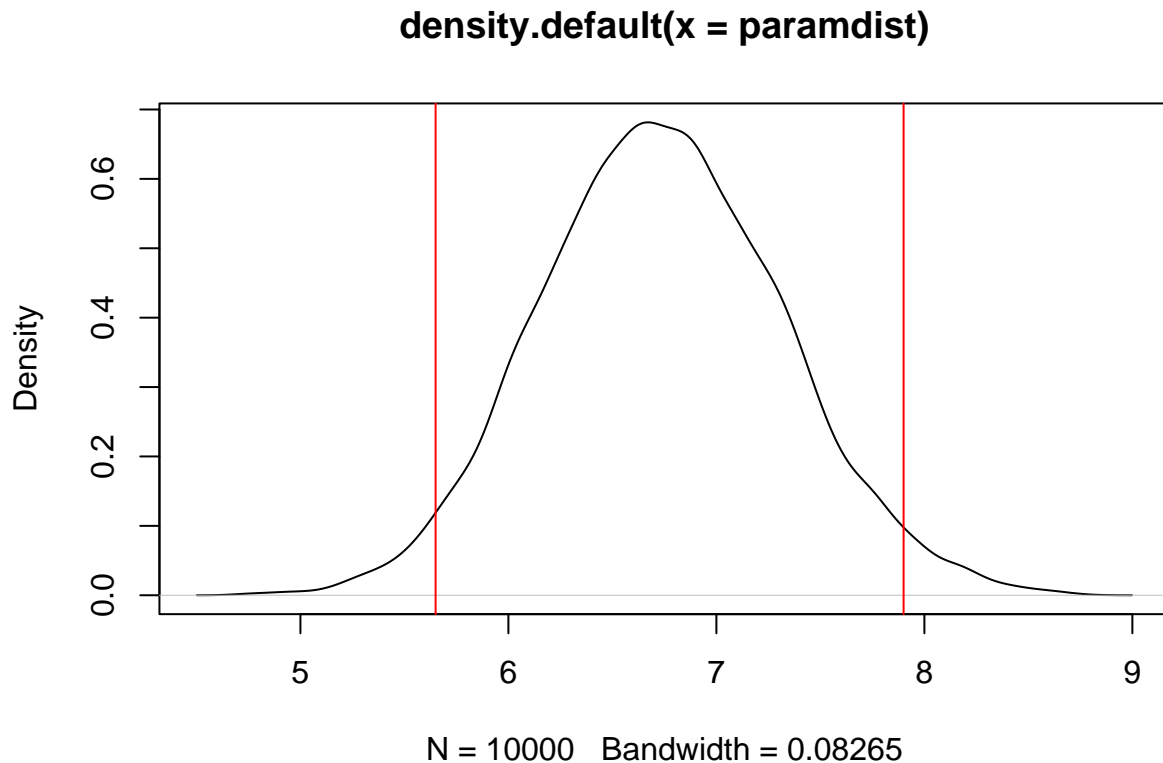


What we have is called a **parameter distribution**. Let us compute the 95% confidence interval and the standard deviation for this distribution.

```
sd(paramdist)
```

```
## [1] 0.5794
```

```
q2 = quantile(paramdist, c(.05/2, 1-(.05/2)))
plot(density(paramdist))
abline(v = q2, col = "red")
```



```
paste("95% Confidence interval = ", "[", q2[1], q2[2], "]")
```

```
## [1] "95% Confidence interval = [ 5.65 7.9 ]"
```

When you take a different sample, you will obviously get a different estimate. Based on the parameter distribution, we can say the following.

1. The maximum likelihood estimate is 6.75. This is also called the **point estimate**.
2. We are 95% confident that the population value of λ is between 5.6 and 7.9.
3. The standard deviation for the parameter distribution is 0.582. This is also called the **standard error** of the parameter estimate.

One final point to be made here is that in most cases, the parameter distribution is reasonably normal.

Optimizing Computationally

We will use the **bbmle** package to computationally optimize the likelihood function.

The main function in the **bbmle** package that does the optimization is `mle2()`. The process is as follows.

1. Create a function that computes the log likelihood value for a given set of parameters. The parameters are the inputs to the function. The output of the function should be the negative of the loglikelihood as most optimization routines attempt to minimize a function.
2. Use the `mle2()` function to attempt to find the maximum likelihood parameter values. The basic syntax is

```
mle2(minuslogl=function to minimize, start=list("starting values of the parameters"))
```

Let us estimate the parameter for the example we just did.

```
LLpois = function(lam){
  p = dpois(x = x, lambda = lam)
  LL = sum(log(p))
  return(-1*LL)
}

res1 = mle2(minuslogl = LLpois, start = list(lam = 10))
summary(res1)

## Maximum likelihood estimation
##
## Call:
## mle2(minuslogl = LLpois, start = list(lam = 10))
##
## Coefficients:
##      Estimate Std. Error z value      Pr(z)
## lam      6.750      0.581    11.6 <0.0000000000000002 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## -2 log L: 94.02
```

You will notice that the results are the same as before (when we mathematically optimized likelihood).

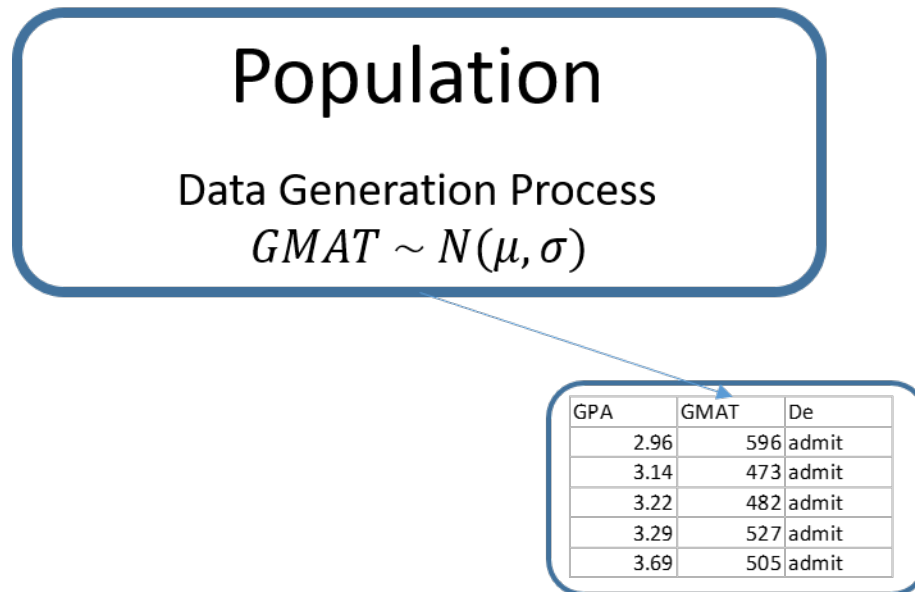
Numeric Outcome

Read the `admission.csv` file.

```
admission <- read.csv("../data/admission.csv")
head(admission)
```

```
##      GPA GMAT   De
## 1 2.96  596 admit
## 2 3.14  473 admit
## 3 3.22  482 admit
## 4 3.29  527 admit
## 5 3.69  505 admit
## 6 3.46  693 admit
```

Let us assume that the data generation process for the variable **GMAT** is normal. This is depicted in the figure. We want to obtain the maximum likelihood estimates of the population mean and standard deviation.



```
LLnorm = function(mean1, standdev){
  p = dnorm(x = admission$GMAT, mean = mean1, sd = standdev)
  LL = sum(log(p))
  return(-1*LL)
}

res1 = mle2(minuslogl = LLnorm, start = list(mean1 = 500, standdev = 100))
summary(res1)
```

```
## Maximum likelihood estimation
##
## Call:
## mle2(minuslogl = LLnorm, start = list(mean1 = 500, standdev = 100))
##
## Coefficients:
##           Estimate Std. Error z value      Pr(z)
## mean1      488.45      8.79    55.6 <0.000000000000002 ***
## standdev    81.04      6.22    13.0 <0.000000000000002 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## -2 log L: 988.4
```


Binary Outcome

We want to estimate the probability of **admit** based on the sample of data, as shown in the figure below.

Population

Data Generation Process

$$P(\text{admit}) = \pi$$



GPA	GMAT	De
2.96	596	admit
3.14	473	admit
3.22	482	admit
3.29	527	admit
3.69	505	admit

```
LLbinary = function(pi){
  p = ifelse(admission$De == "admit", pi, 1 - pi)
  LL = sum(log(p))
  return(-1*LL)
}

res1 = mle2(minuslogl = LLbinary, start = list(pi = .5))
summary(res1)
```

```
## Maximum likelihood estimation
##
## Call:
## mle2(minuslogl = LLbinary, start = list(pi = 0.5))
##
## Coefficients:
##      Estimate Std. Error z value      Pr(z)
## pi    0.3647     0.0522   6.99 0.0000000000028 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## -2 log L: 111.5
```

A Few Cautionary Notes

- This may be a bit challenging for some students to grasp. The instructor may want to skip this and return to it if/when issues in coding are encountered by students.

When using computational methods to find the maximum likelihood values, you should be aware of two potential issues.

1. Starting values.

The function `mle2()` requires you to specify starting values for the parameters to be optimized. For

complicated likelihood functions, it is important to specify good starting values so that the search for optimal values converges reasonably. As we will see in the next lesson, using starting values from a **baseline model** is a reasonable approach to take. We will explain the concept of a baseline model in the next lesson.

2. Errors.

In an attempt to get the optimal values of the parameters, optimization algorithms can potentially search in the entire domain of values from $-\infty$ to ∞ . Sometimes this can be a problem and lead to an error. Below is an example.

```
dpois(3,lambda = -10)
```

```
## Warning in dpois(3, lambda = -10): NaNs produced
```

```
## [1] NaN
```

Obviously, the mean of a Poisson distribution cannot be negative, but the optimization algorithms may try negative values in their search to find the optimal value of λ . This may throw an error and cause the routine to fail. If this happens, the best way to deal with it is to place bounds on the parameter values.

It also turns out that the **bbmle** package supports several different methods for optimization (see link). The optimization method **L-BFGS-B** allows box constraints, that is each variable can be given a lower and/or upper bound. We do so using the **lower** and **upper** commands in the function as shown in the example below.

```
res = mle2(minuslogl = LLmeansd,  
start = list(M=10,sigma=1),  
lower=c(M=-Inf,sigma=0),  
method = "L-BFGS-B")
```