

## Multivariate Regression (chap. 6)

1. **(6.3)** A student stated: “Adding predictor variables to a regression model can never reduce  $R^2$ , so we should include all available predictor variables in the model.” Comment.

**SOLUTION.**  $R^2$  is a fair measure of comparison for models with the same number of variables. For models of different ranks  $p$ , there are other measures of comparison such as adjusted  $R^2$  because  $R^2$  can only increase when variables are added to the model, even if they are completely irrelevant.

2. **(6.4)** Why is it not meaningful to attach a sign to the coefficient of multiple correlation  $R$ , although we do so for the coefficient of simple correlation  $r_{12}$ ?

**SOLUTION.** Coefficient of multiple correlation measures the strength of linear relationship among several variables. In a space of dimension more than 1, there are many directions, and not just negative or positive. Thus,  $R$  shows how strong the mutual relationship is, but does not indicate any direction.

3. **(6.27)** In a small-scale regression study, the following data were obtained,

$Y$	$X_1$	$X_2$
42.0	7.0	33.0
33.0	4.0	41.0
75.0	16.0	7.0
28.0	3.0	49.0
91.0	21.0	5.0
55.0	8.0	31.0

Assume the standard multiple regression model with independent normal error terms. Compute  $\mathbf{b}$ ,  $\mathbf{e}$ ,  $\mathbf{H}$ ,  $SSE_{err}$ ,  $R^2$ ,  $s^2_{\mathbf{b}}$ ,  $\hat{Y}$  for  $X_1 = 10, X_2 = 30$ . You can do the computations using software or by hand, although it would be lengthy to do them by hand.

**SOLUTION.** These answers are based on the R code and output in the end of these solutions.

$$\mathbf{b} = \begin{pmatrix} 33.93 \\ 2.78 \\ -0.25 \end{pmatrix}, \quad \mathbf{e} = \begin{pmatrix} -2.70 \\ -1.23 \\ -1.64 \\ -1.33 \\ -0.90 \\ 6.99 \end{pmatrix}, \quad \mathbf{H} = \begin{pmatrix} 0.23 & 0.25 & 0.21 & 0.15 & -0.05 & 0.21 \\ 0.25 & 0.31 & 0.09 & 0.27 & -0.15 & 0.22 \\ 0.21 & 0.09 & 0.70 & -0.32 & 0.10 & 0.20 \\ 0.15 & 0.27 & -0.32 & 0.61 & 0.14 & 0.15 \\ -0.05 & -0.15 & 0.10 & 0.14 & 0.94 & 0.02 \\ 0.21 & 0.22 & 0.20 & 0.15 & 0.02 & 0.20 \end{pmatrix},$$

$$SSE_{err} = 62.07, \quad R^2 = 0.98, \quad s^2_{\mathbf{b}} = \begin{pmatrix} 715.47 & -34.16 & -13.59 \\ -34.16 & 1.66 & 0.64 \\ -13.59 & 0.64 & 0.26 \end{pmatrix}, \quad \hat{Y} = 53.85.$$

4. (Computer project, **#6.5—#6.8**) Dataset “Brand preference” is available on our Blackboard, on <http://statweb.lsu.edu/EXSTWeb/StatLab/DataSets/NKNWData/CH06PR05.txt>, and here:

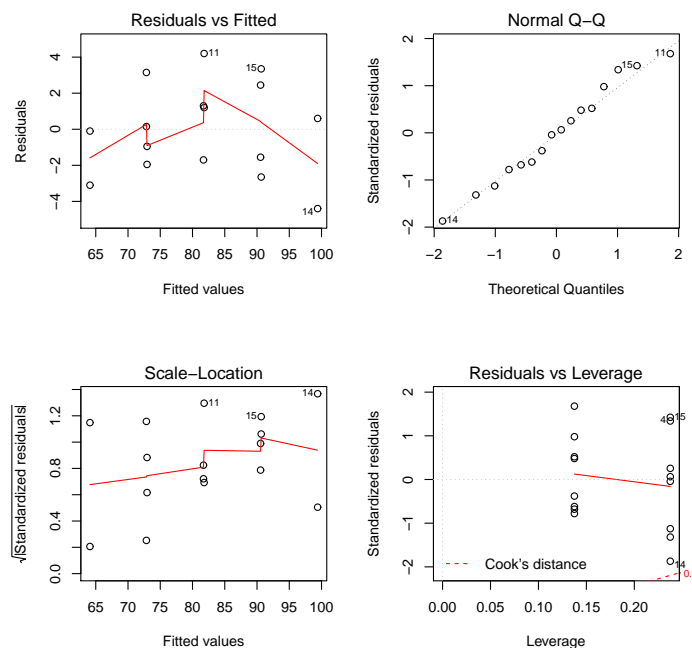
$Y_i$	64	73	61	76	72	80	71	83	83	89	86	93	88	95	94	100
$X_{i1}$	4	4	4	4	6	6	6	6	8	8	8	8	10	10	10	10
$X_{i2}$	2	4	2	4	2	4	2	4	2	4	2	4	2	4	2	4

It was collected to study the relation between degree of brand liking ( $Y$ ) and moisture content ( $X_1$ ) and sweetness ( $X_2$ ) of the product.

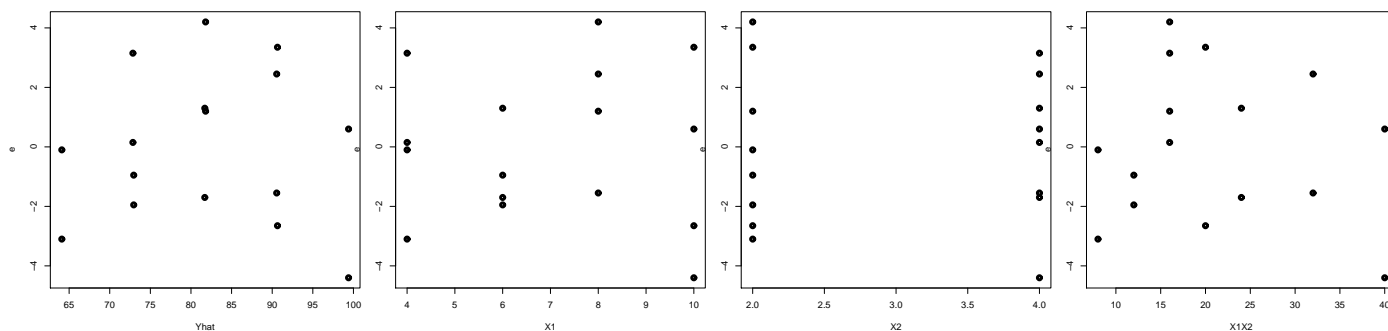
- Fit a regression model to these data and state the estimated regression function. Interpret  $b_1$ .
- Obtain residual plots. What information do they provide? Plot residuals against fitted values, against each predictor, and against the product of predictors.
- Test homoscedasticity.
- Conduct a formal lack of fit test.
- Test whether the proposed linear regression model is significant. What do the results of the ANOVA F-test imply about the slopes?
- Estimate both slopes simultaneously using the Bonferroni procedure with at least a 99 percent confidence level.
- Report  $R^2$  and adjusted  $R^2$ . How are they interpreted here?
- Calculate the squared correlation coefficient between  $Y_i$  and  $\hat{Y}_i$ . Compare with part (g).
- Obtain a 99% confidence interval for  $\mathbf{E}(Y)$  when  $X_1 = 5$  and  $X_2 = 4$ . Interpret it.
- Obtain a 99% prediction interval for a new observation  $Y$  when  $X_1 = 5$  and  $X_2 = 4$ . Interpret it.

**SOLUTION.** These answers are based on the R code and output in the end of these solutions.

- $\hat{Y} = 37.65 + 4.425X_1 + 4.375X_2$ . The slope  $b_1 = 4.425$  means that the brand liking is expected to increase by 4.425 when the product moisture content increases by 1 while sweetness is unchanged.
- Looking at the standard residual plots, there is some indication of a nonlinear trend; the Q-Q plot looks fairly straight, so probably, no problem with Normality; the variance of responses does not seem to change with the increase of their mean.



Looking at residual  $e_i$  plotted against fitted values  $\hat{Y}$ , predictors  $X_1$  and  $X_2$ , and against the product of predictors  $X_1X_2$  (the 4 plots below), there may be a concave nonlinear trend as a function of  $X_1$  and no visible nonlinear relation with  $X_2$  or  $X_1X_2$ .



- (c) There is no significant evidence against the hypothesis of a constant variance  $H_0 : \sigma^2 = \text{const}$ , with the test statistic  $\chi^2 = 0.626$  and p-value  $p = 0.4288$ .
- (d) There is no significant evidence of a nonlinear trend, with the test statistic  $F = 1.045$  and p-value  $p = 0.384$ .
- (e) Significance of the whole model is tested by  $H_0 : \beta_1 = \beta_2 = 0$  vs  $H_1 : \beta_1 \neq 0$  or  $\beta_2 \neq 0$ . The ANOVA F-test shows that the model is significant, with the test statistic  $F = 129.1$  and p-value  $p = 2.66 \cdot 10^{-9}$ . This means a significant evidence that at least one of the slopes is not 0.
- (f) Using the Bonferroni adjustment for two simultaneous confidence intervals, the alpha level 0.01 is divided by 2. We obtain confidence intervals

$$[3.41, 5.44] \text{ for } \beta_1 \quad \text{and} \quad [2.10, 6.65] \text{ for } \beta_2.$$

- (g)  $R^2 = 0.9521$  is the proportion of the total variation  $SST_{\text{tot}}$  explained by the two variables  $X_1$  and  $X_2$  combined. It measures goodness of fit, but it can only be used to compare models of the same rank  $p$ .

$R^2_{\text{adj}} = 0.9447$  is the measure of a goodness of fit that can be used to compare models of different ranks, that is, different numbers of  $X$ -variables.

- (h)  $r_{Y_i \hat{Y}_i} = 0.9521 = R^2$ . Apparently, this is a general result, see exercise #5.
- (i) A 99% confidence interval for  $\mathbf{E}\{Y \mid X_1 = 6, X_2 = 4\}$  is  $[73.88, 80.67]$ . There is a 99% confidence that this interval covers the mean of responses with these values of  $X_1$  and  $X_2$ . That is, in a long run of intervals computed from different samples, 99% of these intervals contain  $\mathbf{E}\{Y \mid X_1 = 6, X_2 = 4\}$ .
- (j) A 99% prediction interval for  $Y$  when  $X_1 = 6, X_2 = 4$  is  $[68.48, 86.07]$ . There is a 99% confidence that this interval covers the actual responses  $X_1 = 6$  and  $X_2 = 4$ . That is, in a long run of intervals computed from different samples and random responses  $Y$ , 99% of these intervals will cover this response.

5. (# 6.26, Stat-615 only) Show that the squared sample correlation coefficient between  $Y$  and  $\hat{Y}$  equals  $R^2$ .

Remark. Now you can check if you did #3h correctly.

Hints. First, show that the sample averages of  $Y_i$  and  $\hat{Y}_i$  are the same. Then, write the sample correlation coefficient between  $Y$  and  $\hat{Y}$  as

$$r_{Y\hat{Y}} = \frac{\sum(Y_i - \bar{Y})(\hat{Y}_i - \bar{Y})}{\sqrt{\sum(Y_i - \bar{Y})^2 \sum(\hat{Y}_i - \bar{Y})^2}} = \frac{\sum(\hat{Y}_i - \bar{Y} + Y_i - \hat{Y}_i)(\hat{Y}_i - \bar{Y})}{\sqrt{\sum(Y_i - \bar{Y})^2 \sum(\hat{Y}_i - \bar{Y})^2}}$$

and use known properties of residuals  $\sum e_i = 0$ ,  $\sum X_{ij}e_i = 0$ ,  $\sum \hat{Y}_i e_i = 0$ .

**SOLUTION.** Following Hint 1, “First, show that the sample averages of  $Y_i$  and  $\hat{Y}_i$  are the same”.

We already know that  $\sum e_i = \sum Y_i - \sum \hat{Y}_i = 0$ . Therefore,  $\sum Y_i = \sum \hat{Y}_i$ , and dividing by  $n$ ,  $\bar{Y} = \bar{\hat{Y}}$ . Following Hint 2, “write the sample correlation coefficient between  $Y$  and  $\hat{Y}$  as ...”,

$$\begin{aligned} r_{Y\hat{Y}} &= \frac{\sum (Y_i - \bar{Y})(\hat{Y}_i - \bar{Y})}{\sqrt{\sum (Y_i - \bar{Y})^2 \sum (\hat{Y}_i - \bar{Y})^2}} = \frac{\sum (\hat{Y}_i - \bar{Y} + Y_i - \hat{Y}_i)(\hat{Y}_i - \bar{Y})}{\sqrt{\sum (Y_i - \bar{Y})^2 \sum (\hat{Y}_i - \bar{Y})^2}} \\ &= \frac{\sum (\hat{Y}_i - \bar{Y})^2 + \sum (Y_i - \hat{Y}_i)(\hat{Y}_i - \bar{Y})}{\sqrt{\sum (Y_i - \bar{Y})^2 \sum (\hat{Y}_i - \bar{Y})^2}} \\ &= \frac{SSReg + \sum e_i(\hat{Y}_i - \bar{Y})}{\sqrt{SSTot \cdot SSReg}} = \frac{SSReg + \sum \hat{Y}_i e_i - \bar{Y} \sum e_i}{\sqrt{SSTot \cdot SSReg}} = \frac{SSReg + 0 - 0}{\sqrt{SSTot \cdot SSReg}} = \sqrt{\frac{SSReg}{SSTot}} = \sqrt{R^2} \end{aligned}$$

So,  $r_{Y\hat{Y}}^2 = R^2$ .

### R Code and Output for Problem #3

```
# Enter the data
> Y = c(42,33,75,28,91,55)
> X1 = c(7,4,16,3,21,8)
> X2 = c(33,41,7,49,5,31)
> install.packages("matlib")
> library(matlib)

# Define the design matrix X
> X = matrix(c(1,1,1,1,1,1,X1,X2),6,3)
> X
      [,1] [,2] [,3]
[1,]    1    7   33
[2,]    1    4   41
[3,]    1   16    7
[4,]    1    3   49
[5,]    1   21    5
[6,]    1    8   31

# Compute the regression slope b
> b = inv(t(X) %*% X) %*% t(X) %*% Y
> b
      [,1]
[1,] 33.9321020
[2,]  2.7847707
[3,] -0.2643979

# Fitted values, residuals, and error sum of squares
> Yhat = X %*% b
> e = Y - Yhat
> e
      [,1]
[1,] -2.70036663
[2,] -1.23087135
```

```

[3,] -1.63764825
[4,] -1.33091751
[5,] -0.09029763
[6,] 6.98606687
> SSErr = sum(e^2)
> SSErr
      [,1]
[1,] 62.07354

# Hat matrix
> H = X%*%inv(t(X)%*%X)%*%t(X)
> H
      [,1]      [,2]      [,3]      [,4]      [,5]      [,6]
[1,] 0.23143639 0.25168006 0.21178834 0.1488734 -0.05475455 0.21099418
[2,] 0.25168006 0.31240977 0.09437951 0.2662835 -0.14787196 0.22314063
[3,] 0.21178834 0.09437951 0.70442097 -0.3191731 0.10446756 0.20412257
[4,] 0.14887339 0.26628346 -0.31917314 0.6142637 0.14143589 0.14834214
[5,] -0.05475455 -0.14787196 0.10446756 0.1414359 0.94040059 0.01632796
[6,] 0.21099418 0.22314063 0.20412257 0.1483421 0.01632796 0.19708945

# Compute $R^2$
> SSTot = sum((Y - mean(Y))^2)
> SSReg = SSTot - SSErr
> Rsq = SSReg/SSTot
> Rsq
      [,1]
[1,] 0.9797938

# Estimate VAR(b)
> s2 = SSErr/(6-3); # Estimated Var(Y)
> sb2 = s2*inv(t(X)%*%X); # Estimated VAR(b)
> sb2
      [,1]      [,2]      [,3]
[1,] 715.47117 -34.1589184 -13.5949378
[2,] -34.15892 1.6616665 0.6440674
[3,] -13.59494 0.6440674 0.2624678

# Prediction for the given $X_1$ and $X_2$
> X0 = c(1,10,30)
> Y0hat = X0%*%b
> Y0hat
      [,1]
[1,] 53.84787

```

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#### R Code and Output for Problem #4

```
# Enter the data and rename variables
```

```

> attach(A)
> Y=V1; X1=V2; X2=V3;

```

```

# Least squares estimation of regression slopes
> reg = lm( Y ~ X1 + X2 )
> reg
(Intercept)          X1          X2
      37.650       4.425       4.375

# Residual plots
> par(mfrow=c(2,2))
> plot(reg)
> e = residuals(reg); Yhat = predict(reg); X1X2 = X1*X2;
> par(mfrow=c(1,1))
> plot(Yhat,e,lwd=5)
> plot(X1,e,lwd=5)
> plot(X2,e,lwd=5)
> plot(X1X2,e,lwd=5)

# Testing for constant variance
> install.packages("car")
> library("car")
> ncvTest(reg)
Non-constant Variance Score Test
Variance formula: ~ fitted.values
Chisquare = 0.6261627, Df = 1, p = 0.42877

# Lack of fit test
> full.model = lm( Y ~ as.factor(X1) + as.factor(X2) )
> anova(reg, full.model)
Model 1: Y ~ X1 + X2
Model 2: Y ~ as.factor(X1) + as.factor(X2)
  Res.Df  RSS Df Sum of Sq    F Pr(>F)
1     13 94.30
2     11 79.25  2     15.05 1.0445 0.3843

# ANOVA F-test, R-square, and adjusted R-square
> summary(reg)
              Estimate Std. Error t value Pr(>|t|)
(Intercept)  37.6500      2.9961  12.566 1.20e-08 ***
X1           4.4250      0.3011  14.695 1.78e-09 ***
X2           4.3750      0.6733   6.498 2.01e-05 ***
---
Residual standard error: 2.693 on 13 degrees of freedom
Multiple R-squared:  0.9521,    Adjusted R-squared:  0.9447
F-statistic: 129.1 on 2 and 13 DF,  p-value: 2.658e-09

# Prediction. Confidence and prediction intervals
> confint(reg, level=0.995)
              0.25 %    99.75 %
(Intercept) 27.545738 47.754262
X1           3.409483  5.440517
X2           2.104236  6.645764

> (cor(Y,Yhat))^2

```

```
[1] 0.952059
```

```
> predict(reg, data.frame(X1=5,X2=4), interval="confidence", level=0.99)
```

```
      fit      lwr      upr  
1 77.275 73.88111 80.66889
```

```
> predict(reg, data.frame(X1=5,X2=4), interval="prediction", level=0.99)
```

```
      fit      lwr      upr  
1 77.275 68.48077 86.06923
```