## Lab 4

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Based on Applied Statistics with R (appliedstats) by David Dalpiaz (https://github.com/daviddalpiaz/appliedstats)

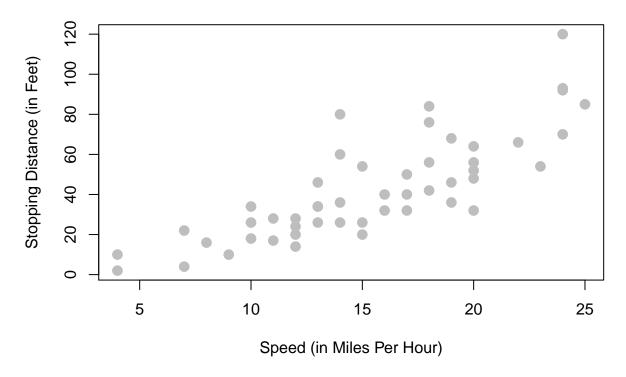
#### cars Example

We use a simple example of how the speed of a car affects its stopping distance, that is, how far it travels before it comes to a stop. To examine this relationship, we will use the cars dataset which, is a default R dataset. Thus, we don't need to load a package first; it is immediately available.

```
View(cars)
```

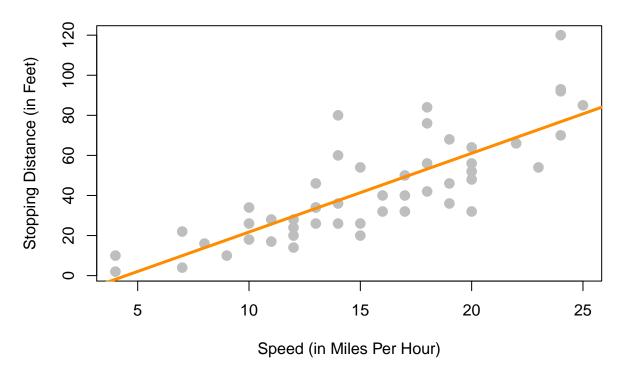
Reading the documentation we learn that this is data gathered during the 1920s about the speed of cars and the resulting distance it takes for the car to come to a stop. The interesting task here is to determine how far a car travels before stopping, when traveling at a certain speed. So, we will first plot the stopping distance against the speed.

### **Stopping Distance vs Speed**



The line on the plot below seems to summarize the relationship between stopping distance and speed quite well. As speed increases, the distance required to come to a stop increases. There is still some variation about this line, but it seems to capture the overall trend.

## **Stopping Distance vs Speed**



#### Fit the model

Fit the model using lm() then use summary() to view the results in greater detail.

```
stop_dist_model = lm(dist ~ speed, data = cars)
summary(stop_dist_model)
```

```
##
## Call:
## lm(formula = dist ~ speed, data = cars)
##
##
  Residuals:
##
       Min
                1Q
                   Median
                                3Q
                                       Max
  -29.069 -9.525
                    -2.272
                             9.215
                                    43.201
##
##
## Coefficients:
##
               Estimate Std. Error t value Pr(>|t|)
                                    -2.601
                                             0.0123 *
## (Intercept) -17.5791
                            6.7584
## speed
                 3.9324
                            0.4155
                                     9.464 1.49e-12 ***
##
                  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
## Signif. codes:
##
## Residual standard error: 15.38 on 48 degrees of freedom
## Multiple R-squared: 0.6511, Adjusted R-squared: 0.6438
## F-statistic: 89.57 on 1 and 48 DF, p-value: 1.49e-12
```

#### Tests in R

We will now discuss the results displayed called Coefficients. First recall that we can extract this information directly.

#### names(summary(stop\_dist\_model))

```
## [1] "call" "terms" "residuals" "coefficients"
## [5] "aliased" "sigma" "df" "r.squared"
## [9] "adj.r.squared" "fstatistic" "cov.unscaled"
```

summary(stop\_dist\_model)\$coefficients

```
## Estimate Std. Error t value Pr(>|t|)
## (Intercept) -17.579095 6.7584402 -2.601058 1.231882e-02
## speed 3.932409 0.4155128 9.463990 1.489836e-12
```

The names() function tells us what information is available, and then we use the \$ operator and coefficients to extract the information we are interested in. Two values here should be immediately familiar.

$$b_0 = -17.5790949$$

and

$$b_1 = 3.9324088$$

which are our estimates for the model parameters  $\beta_0$  and  $\beta_1$ .

Let's now focus on the second row of output, which is relevant to  $\beta_1$ .

```
summary(stop_dist_model)$coefficients[2,]
```

```
## Estimate Std. Error t value Pr(>|t|)
## 3.932409e+00 4.155128e-01 9.463990e+00 1.489836e-12
```

Again, the first value, Estimate is

$$b_1 = 3.9324088.$$

The second value, Std. Error, is the square root of the estimated variance (standard error) of  $b_1$ ,

$$s\{b_1\} = \frac{\sqrt{\text{MSE}}}{\sqrt{\sum (X_i - \bar{X})}} = 0.4155128.$$

The third value, t value, is the value of the test statistic for testing  $H_0: \beta_1 = 0$  vs  $H_1: \beta_1 \neq 0$ ,

$$t = \frac{b_1 - 0}{s\{b_1\}} = 9.46399.$$

Lastly, Pr(>|t|), gives us the p-value of that test.

p-value = 
$$1.4898365 \times 10^{-12}$$

Note here, we are specifically testing whether or not  $\beta_1 = 0$ .

The first row of output reports the same values, but for  $\beta_0$ .

```
summary(stop_dist_model)$coefficients[1,]
```

```
## Estimate Std. Error t value Pr(>|t|)
## -17.57909489 6.75844017 -2.60105800 0.01231882
```

In summary, the following code stores the information of summary(stop\_dist\_model)\$coefficients in a new variable stop\_dist\_model\_test\_info, then extracts each element into a new variable which describes the information it contains.

```
stop_dist_model_test_info = summary(stop_dist_model)$coefficients

b_0 = stop_dist_model_test_info[1, 1] # Estimate

b_0_se = stop_dist_model_test_info[1, 2] # Std. Error

b_0_t = stop_dist_model_test_info[1, 3] # t value

b_0_pval = stop_dist_model_test_info[1, 4] # Pr(>/t/)

b_1 = stop_dist_model_test_info[2, 1] # Estimate

b_1_se = stop_dist_model_test_info[2, 2] # Std. Error

b_1_t = stop_dist_model_test_info[2, 3] # t value

b_1_pval = stop_dist_model_test_info[2, 4] # Pr(>/t/)
```

#### Task

Verify some equivalent expressions: the t test statistic for  $b_1$  and the two-sided p-value associated with that test statistic.

```
(b_1 - 0) / b_1_se

## [1] 9.46399

b_1_t

## [1] 9.46399

2 * pt(abs(b_1_t), df = length(resid(stop_dist_model)) - 2, lower.tail = FALSE)

## [1] 1.489836e-12

b_1_pval

## [1] 1.489836e-12
```

#### Significance of Regression, t-Test

We pause to discuss the **significance of regression** test. First, note that based on the above distributional results, we could test  $\beta_0$  and  $\beta_1$  against any particular value, and perform both one and two-sided tests.

However, one very specific test,

$$H_0: \beta_1 = 0 \text{ vs } H_1: \beta_1 \neq 0$$

is used most often. Let's think about this test in terms of the simple linear regression model,

$$Y_i = \beta_0 + \beta_1 X_i + \epsilon_i.$$

If we assume the null hypothesis is true, then  $\beta_1 = 0$  and we have the model,

$$Y_i = \beta_0 + \epsilon_i$$
.

In this model, the response does **not** depend on the predictor. So then we could think of this test in the following way,

- Under  $H_0$  there is not a significant linear relationship between X and Y.
- Under  $H_1$  there is a significance linear relationship between Y and Y.

For the cars example,

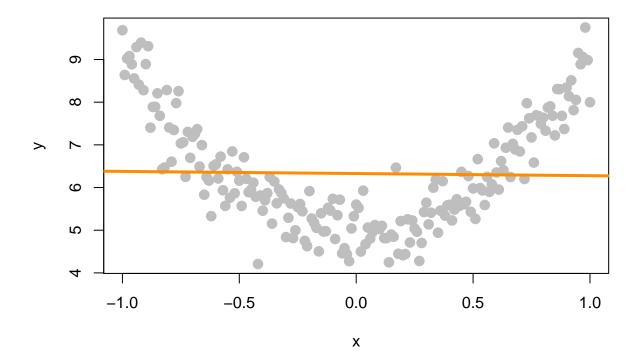
- Under  $H_0$  there is not a significant linear relationship between speed and stopping distance.
- Under  $H_1$  there is a significant linear relationship between speed and stopping distance.

Again, that test is seen in the output from summary(),

p-value = 
$$1.4898365 \times 10^{-12}$$
.

With this extremely low p-value, we would reject the null hypothesis at any reasonable  $\alpha$  level, say for example  $\alpha = 0.01$ . So we say there is a significant **linear** relationship between speed and stopping distance. Notice that we emphasize **linear**.

```
set.seed(42)
x = seq(-1, 1, 0.01)
y = 5 + 4 * x ^ 2 + rnorm(length(x), 0, 0.5)
plot(x, y, pch = 20, cex = 2, col = "grey")
abline(lm(y ~ x), lwd = 3, col = "darkorange")
```



In this plot of simulated data, we see a clear relationship between x and y, however it is not a linear relationship. If we fit a line to this data, it is very flat. The resulting test for  $H_0: \beta_1 = 0$  vs  $H_1: \beta_1 \neq 0$  gives a large p-value, in this case 0.7564548, so we would fail to reject and say that there is no significant linear relationship between x and y. We will see later how to fit a curve to this data using a "linear" model, but for now, realize that testing  $H_0: \beta_1 = 0$  vs  $H_1: \beta_1 \neq 0$  can only detect straight line relationships.

#### Confidence Intervals in R

Using R we can very easily obtain the confidence intervals for  $\beta_0$  and  $\beta_1$ .

This automatically calculates 99% confidence intervals for both  $\beta_0$  and  $\beta_1$ , the first row for  $\beta_0$ , the second row for  $\beta_1$ .

For the cars example when interpreting these intervals, we say, we are 99% confident that for an increase in speed of 1 mile per hour, the average increase in stopping distance is between 2.8179187 and 5.0468988 feet, which is the interval for  $\beta_1$ .

Note that this 99% confidence interval does **not** contain the hypothesized value of 0. Since it does not contain 0, it is equivalent to rejecting the test of  $H_0: \beta_1 = 0$  vs  $H_1: \beta_1 \neq 0$  at  $\alpha = 0.01$ , which we had seen previously.

You should be somewhat suspicious of the confidence interval for  $\beta_0$ , as it covers negative values, which correspond to negative stopping distances. Technically the interpretation would be that we are 99% confident that the average stopping distance of a car traveling 0 miles per hour is between -35.7066103 and 0.5484205 feet, but we don't really believe that, since we are actually certain that it would be non-negative.

Note, we can extract specific values from this output a number of ways.

```
confint(stop_dist_model, level = 0.99)[1,]
##
         0.5 %
                    99.5 %
## -35.7066103
                 0.5484205
confint(stop_dist_model, level = 0.99)[1, 1]
## [1] -35.70661
confint(stop_dist_model, level = 0.99)[1, 2]
## [1] 0.5484205
confint(stop_dist_model, parm = "(Intercept)", level = 0.99)
                            99.5 %
                   0.5 %
## (Intercept) -35.70661 0.5484205
confint(stop dist model, level = 0.99)[2,]
##
      0.5 %
              99.5 %
## 2.817919 5.046899
confint(stop_dist_model, level = 0.99)[2, 1]
## [1] 2.817919
```

#### Task

Verify that calculations that R is performing for the  $\beta_1$  interval.

```
# store estimate
b_1 = coef(stop_dist_model)[2]

# store standard error
b_1_se = summary(stop_dist_model)$coefficients[2, 2]

# calculate critical value for two-sided 99% CI
crit = qt(0.995, df = length(resid(stop_dist_model)) - 2)

# est - margin, est + margin
c(b_1 - crit * b_1_se, b_1 + crit * b_1_se)
```

## speed speed ## 2.817919 5.046899