MATH60047: Stochastic Simulation Coursework 1

CID: 01938572

1 Q1: Sampling from Chi-Squared using Rejection Sampling

1.1 Compute $M_{\lambda} = \sup_{x} \frac{p_{\nu}(x)}{q_{\lambda}(x)}$

$$M_{\lambda} = \sup_{x} \frac{p_{\nu}(x)}{q_{\lambda}(x)}$$

$$\frac{p_{\nu}(x)}{q_{\lambda}(x)} = \frac{1}{2^{\frac{\nu}{2}}\Gamma(\frac{\nu}{2})} \cdot \frac{x^{\frac{\nu}{2}-1}e^{-\frac{x}{2}}}{\lambda e^{-\lambda x}} = \frac{1}{2^{\frac{\nu}{2}}\Gamma(\frac{\nu}{2})\lambda} x^{\frac{\nu}{2}-1}e^{-\frac{x}{2}+\lambda x}$$

$$\log(\frac{p_{\nu}(x)}{q_{\lambda}(x)}) = -\log(2^{\frac{\nu}{2}}\Gamma(\frac{\nu}{2})\lambda) + (\frac{\nu}{2} - 1)\log(x) - \frac{x}{2} + \lambda x$$

Taking the derivative of \log with respect to x and setting it to 0, we obtain

$$\frac{d}{dx}(\log(\frac{p_{\nu}(x)}{q_{\lambda}(x)})) = (\frac{\nu}{2} - 1)\frac{1}{x} - \frac{1}{2} + \lambda x = 0$$
$$(\frac{\nu - 2}{2})\frac{1}{x} = \frac{1 - 2\lambda}{2}$$
$$x^* = \frac{\nu - 2}{1 - 2\lambda}$$

To show that x^* is a maximum, we take the second derivative

$$\frac{d^2}{dx^2}\log(\frac{p_{\nu}(x)}{q_{\lambda}(x)}) = -(\frac{\nu}{2} - 1)\frac{1}{x^2}$$

$$x^* = \frac{\nu - 2}{1 - 2\lambda} \implies \frac{d^2}{dx^2}\log(\frac{p_{\nu}(x)}{q_{\lambda}(x)}) = -(\frac{\nu - 2}{2})(\frac{1 - 2\lambda}{\nu - 2})^2 = -\frac{(1 - 2\lambda)^2}{2(\nu - 2)} < 0$$

for $\nu > 2$ and $0 < \lambda < \frac{1}{2}$.

Placing $x = x^*$ in the ratio $\frac{p_{\nu}(x)}{q_{\lambda}(x)}$, we obtain

$$M_{\lambda} = \frac{1}{2^{\frac{\nu}{2}}\Gamma(\frac{\nu}{2})\lambda}(\frac{\nu-2}{1-2\lambda})^{\frac{\nu}{2}-1}e^{-(\frac{1-2\lambda}{2})(\frac{\nu-2}{1-2\lambda})} = \frac{1}{2^{\frac{\nu}{2}}\Gamma(\frac{\nu}{2})\lambda}(\frac{\nu-2}{1-2\lambda})^{\frac{\nu}{2}-1}e^{-\frac{\nu}{2}+1}$$

1.2 Find the optimal λ^* in terms of ν

$$\log(M_{\lambda}) = -\log(2^{\frac{\nu}{2}}\Gamma(\frac{\nu}{2})) - \log(\lambda) + (\frac{\nu}{2} - 1)(\log(\nu - 2) - \log(1 - 2\lambda)) - \frac{\nu}{2} + 1$$

Taking the derivative of log with respect to λ and setting it to 0, we obtain

$$\frac{d}{d\lambda}(\log(M_{\lambda})) = -\frac{1}{\lambda} + (\frac{\nu}{2} - 1)(\frac{2}{1 - 2\lambda}) = 0$$
$$-\frac{1}{\lambda} + (\frac{\nu - 2}{2})(\frac{2}{1 - 2\lambda}) = 0$$
$$-\frac{1}{\lambda} + \frac{\nu - 2}{1 - 2\lambda} = 0$$
$$\frac{-1 + 2\lambda + \nu\lambda - 2\lambda}{\lambda(1 - 2\lambda)} = 0$$

$$\nu\lambda = 1$$
$$\lambda^* = \frac{1}{\nu}$$

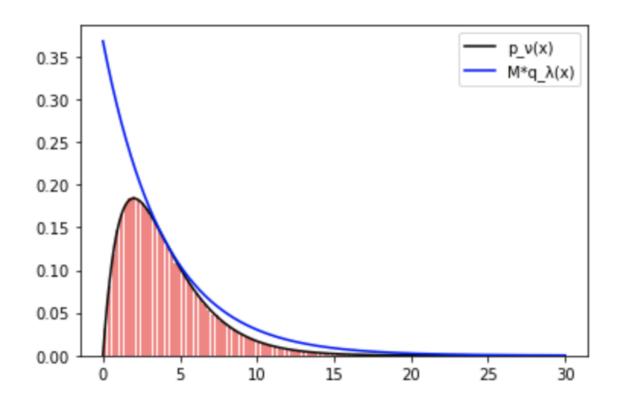
To show that λ^* is a minimum, we take the second derivative

$$\frac{d^2}{d\lambda^2}(\log(M_\lambda)) = \frac{1}{\lambda^2} + (\nu - 2)\frac{2}{(1 - 2\lambda)^2}$$

$$\lambda^* = \frac{1}{\nu} \implies \frac{d^2}{d\lambda^2}(\log(M_\lambda)) = \nu^2 + (\nu - 2)\frac{2}{(1 - \frac{2}{\nu})^2} = \nu^2 + (\nu - 2)\frac{2}{(\frac{\nu - 2}{\nu})^2} = \nu^2 + \frac{2\nu^2}{\nu - 2} > 0$$
for $\nu > 2$.
Thus,
$$M_{\lambda^*} = \frac{\nu}{2^{\frac{\nu}{2}}\Gamma(\frac{\nu}{2})}(\frac{\nu - 2}{1 - \frac{2}{\nu}})^{\frac{\nu}{2} - 1}e^{-\frac{\nu}{2} + 1} = \frac{\nu}{2^{\frac{\nu}{2}}\Gamma(\frac{\nu}{2})}\nu^{\frac{\nu}{2} - 1}e^{-\frac{\nu}{2} + 1} = \frac{1}{2^{\frac{\nu}{2}}\Gamma(\frac{\nu}{2})}\nu^{\frac{\nu}{2}}e^{-\frac{\nu}{2} + 1}$$

1.3 Implement the rejection sampler for $\nu = 4$

1.3.1 Plot histogram, $p_{\nu}(x)$, and $M_{\lambda^*}q_{\lambda^*}(x)$



1.3.2 Compute the acceptance rate and compare it to the theoretical acceptance rate

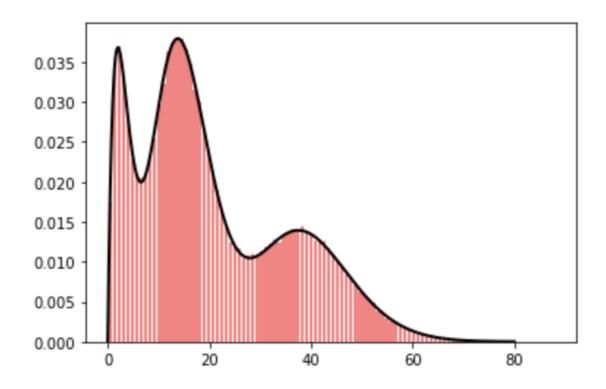
acceptance rate, $a=\frac{\text{number of accepted samples}}{\text{total number of samples}}=0.67917$ theoretical acceptance rate, $\hat{a}=\frac{1}{M_{\lambda^*}}=0.6795704571147613$

$$a - \hat{a} = -0.00040045711476122126$$

Thus, the computed acceptance rate and theoretical acceptance rate have very close values.

2 Q2: Sample from a Mixture of Chi-Squared

2.1 Plot the histogram and the density



3 Appendix

3.1 Code for Q1

```
import numpy as np
import matplotlib.pyplot as plt
def p(x, nu):
    return x ** (nu / 2 - 1) * np.exp(-x / 2) / (2 ** (nu / 2) * np.math.factorial(
                                                    int(nu / 2) - 1))
def q(x, lam):
    return lam * np.exp(-lam * x)
lam = 1 / nu # optimal lambda derived in 1.2
 M = nu ** (nu / 2) * np.exp(-nu / 2 + 1) / (2 ** (nu / 2) * np.math.factorial(int(nu / 2) - 1)) # optimal M derived in 1.2 
\# Rejection sampler for nu=4
n = 100000
x_accepted = np.array([])
count = 0
for i in range(n):
    # sample from exponential using inversion method
    u_1 = np.random.uniform(0, 1)
    x_proposed = -(1/lam) * np.log(1 - u_1)
a = p(x_proposed, nu) / (M * q(x_proposed, lam))
    u_2 = np.random.uniform(0, 1)
    if u_2 <= a:</pre>
        x_accepted = np.append(x_accepted, x_proposed)
         count += 1 \# count accepted samples
# plot histogram, p(x) and Mq(x)
```

3.2 Code for Q2

```
import numpy as np
import matplotlib.pyplot as plt
# function for rejection sampler
def chi_squared_sample(nu, n):
   lam = 1 / nu
   M = nu ** (nu / 2) * np.exp(-nu / 2 + 1) / (2 ** (nu / 2) * np.math.factorial(int))
                                               (nu / 2) - 1))
   x_accepted = np.array([])
    while len(x_accepted) < n:</pre>
        u_1 = np.random.uniform(0, 1)
        x_{proposed} = -(1/lam) * np.log(1 - u_1)
        a = p(x_proposed, nu) / (M * q(x_proposed, lam))
        u_2 = np.random.uniform(0, 1)
        if u_2 <= a:
           x_accepted = np.append(x_accepted, x_proposed)
    return x_accepted
# function for sampling from discrete distribution using inversion method
def discrete(w, s):
   cw = np.cumsum(w)
   u = np.random.uniform(0, 1)
   for k in range(len(cw)):
        if cw[k] > u:
           discrete_sample = s[k]
            break
    return discrete_sample
w = np.array([0.2, 0.5, 0.3])
s = np.array([0, 1, 2])
nu = np.array([4, 16, 40])
n = 100000
sample = np.array([])
# mixture sampling
for i in range(n):
   discrete_sample = discrete(w, s)
   x = chi_squared_sample(nu[discrete_sample], 1)
   sample = np.append(sample, x)
# plot histogram and density
def mixture_density(x, w, nu):
   return w[0] * p(x, nu[0]) + w[1] * p(x, nu[1]) + w[2] * p(x, nu[2])
xx = np.linspace(0, 80, 1000)
plt.plot(xx, mixture_density(xx, w, nu), color='k', linewidth=2)
plt.hist(sample, bins=100, density=True, rwidth=0.8, color='r', alpha=0.5)
plt.show()
```