

## K-means Recap ...



Randomly initialize k centers

$$\square$$
  $\mu^{(0)} = \mu_1^{(0)}, \dots, \mu_k^{(0)}$ 

Classify: Assign each point j∈{1,...m} to nearest center:

$$\square$$
  $C^{(t)}(j) \leftarrow \arg\min_{i} ||\mu_i - x_j||^2$ 

• Recenter:  $\mu_i$  becomes centroid of its point:

$$\square \quad \mu_i^{(t+1)} \leftarrow \arg\min_{\mu} \sum_{j:C(j)=i} ||\mu - x_j||^2$$

 $\square$  Equivalent to  $\mu_i \leftarrow$  average of its points!

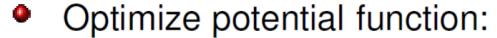
# What is K-means optimizing?

• Potential function  $F(\mu,C)$  of centers  $\mu$  and point allocations C:

$$F(\mu, C) = \sum_{j=1}^{m} ||\mu_{C(j)} - x_j||^2$$

- Optimal K-means:
  - $\square$  min<sub> $\mu$ </sub>min<sub>C</sub> F( $\mu$ ,C)

# K-means algorithm



$$\min_{\mu} \min_{C} F(\mu, C) = \min_{\mu} \min_{C} \sum_{i=1}^{k} \sum_{j:C(j)=i} ||\mu_i - x_j||^2$$

- K-means algorithm:
  - (1) Fix  $\mu$ , optimize C

$$\min_{C(1),C(2),...,C(m)} \sum_{j=1}^{m} \|\mu_{C(j)} - x_j\|^2$$

$$= \sum_{j=1}^{m} \min_{C(j)} \|\mu_{C(j)} - x_j\|^2$$

Exactly first step – assign each point to the nearest cluster center



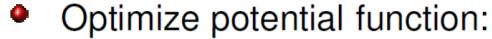
Optimize potential function:

$$\min_{\mu} \min_{C} F(\mu, C) = \min_{\mu} \min_{C} \sum_{i=1}^{k} \sum_{j:C(j)=i} ||\mu_i - x_j||^2$$

- K-means algorithm:
  - (2) Fix C, optimize  $\mu$

$$\begin{split} \min_{\mu_1,\mu_2,...\mu_K} \sum_{i=1}^K \sum_{j:C(j)=i} \|\mu_i - x_j\|^2 \\ &= \sum_{i=1}^K \min_{\mu_i} \sum_{j:C(j)=i} \|\mu_i - x_j\|^2 \\ &= \sum$$

# K-means algorithm



$$\min_{\mu} \min_{C} F(\mu, C) = \min_{\mu} \min_{C} \sum_{i=1}^{k} \sum_{j:C(j)=i} ||\mu_i - x_j||^2$$

K-means algorithm: (coordinate ascent on F)

(1) Fix  $\mu$ , optimize C

**Expectation step** 

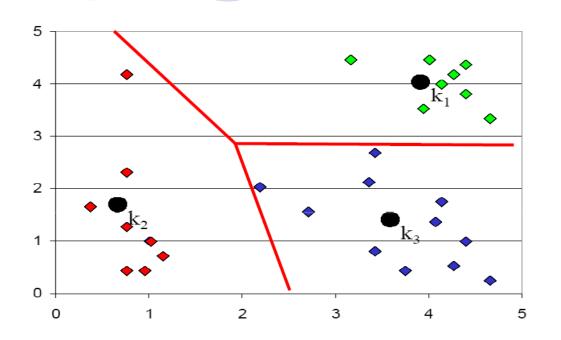
(2) Fix C, optimize  $\mu$ 

**Maximization step** 

Today, we will see a generalization of this approach:

EM algorithm

### K-means Decision boundaries



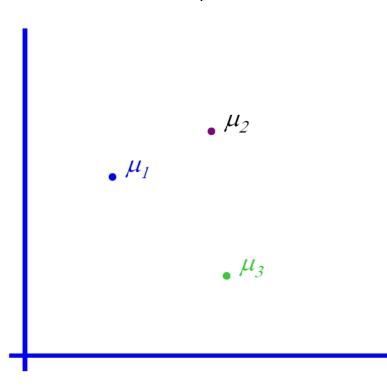
"Linear"
Decision
Boundaries

#### Generative Model:

Assume data comes from a mixture of K Gaussians distributions with same variance

#### Mixture of K Gaussians distributions: (Multi-modal distribution)

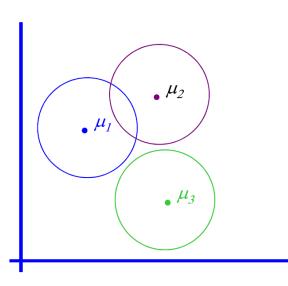
- There are k components
- Component i has an associated mean vector μ<sub>i</sub>



#### Mixture of K Gaussians distributions: (Multi-modal distribution)

- There are k components
- Component i has an associated mean vector μ<sub>i</sub>
- Each component generates data from a Gaussian with mean μ<sub>i</sub> and covariance matrix σ<sup>2</sup>I

Each data point is generated according to the following recipe:

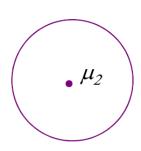


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 Pick a component at random: Choose component i with probability P(y=i)

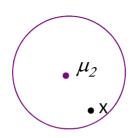


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- 2) Datapoint  $x \sim N(\mu_i, \sigma^2 I)$

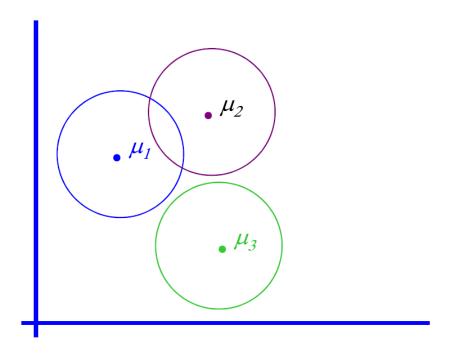


Mixture of K Gaussians distributions: (Multi-modal distribution)

$$p(x|y=i) \sim N(\mu_i, \sigma^2 I)$$

$$p(x) = \sum_i p(x|y=i) P(y=i)$$

$$\downarrow \qquad \qquad \downarrow$$
Mixture Mixture component proportion



Mixture of K Gaussians distributions: (Multi-modal distribution)

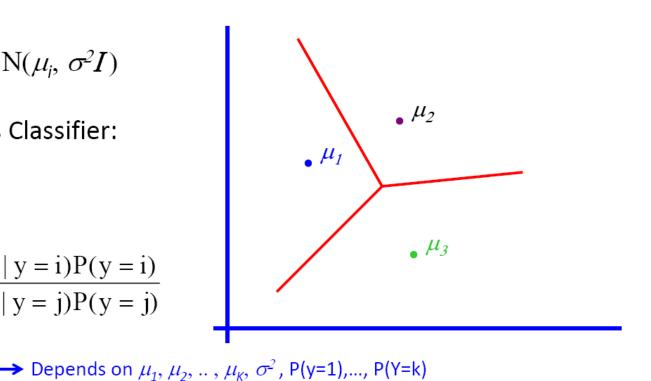
$$p(x|y=i) \sim N(\mu_i, \sigma^2 I)$$

Gaussian Bayes Classifier:

$$\log \frac{P(y=i \mid x)}{P(y=j \mid x)}$$

$$= \log \frac{p(x \mid y=i)P(y=i)}{p(x \mid y=j)P(y=j)}$$

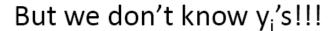
$$= \mathbf{W}^{T} \mathbf{X}$$



"Linear Decision boundary" – Recall that second-order terms cancel out

Maximum Likelihood Estimate (MLE)

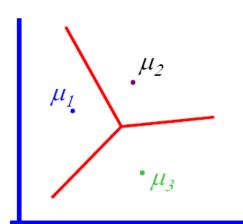
$$\underset{\substack{\mu_1,\ \mu_2,\ \dots,\ \mu_{\mathit{K}},\sigma^2,\\ \mathsf{P}(\mathsf{y}=1),\dots,\ \mathsf{P}(\mathsf{Y}=\mathsf{k})}}{\operatorname{argmax} \prod_i \mathsf{P}(\mathsf{y}_i,\mathsf{x}_i)}$$





$$\underset{\text{argmax } \prod_{j} \sum_{i=1}^{K} P(y_{j}=i,x_{j})}{\text{argmax } \prod_{j} \sum_{i=1}^{K} P(y_{j}=i,x_{j})}$$

$$= \underset{\text{argmax } \prod_{j} \sum_{i}^{K} P(y_{j}=i) p(x_{j} \mid y_{j}=i)}{\text{proposed for } P(y_{j}=i) p(x_{j} \mid y_{j}=i)}$$



Maximize marginal likelihood:

$$argmax \prod_{j} P(x_{j}) = argmax \prod_{j} \sum_{i=1}^{K} P(y_{j}=i,x_{j})$$
$$= argmax \prod_{j} \sum_{i=1}^{K} P(y_{j}=i)p(x_{j}|y_{j}=i)$$

$$P(y_j = i, x_j) \propto P(y_j = i) \exp \left[ -\frac{1}{2\sigma^2} ||x_j - \mu_i||^2 \right]$$

If each  $x_j$  belongs to one class C(j) (hard assignment), marginal likelihood:

$$P(y_i=i) = 1 \text{ or } 0$$
 1 if  $i = C(j)$ 

$$\prod_{j=1}^{m} \sum_{i=1}^{k} P(y_{j} = i, X_{j}) \propto \prod_{j=1}^{m} exp \left[ -\frac{1}{2\sigma^{2}} \left\| X_{j} - \mu_{C(j)} \right\|^{2} \right] = \sum_{j=1}^{m} -\frac{1}{2\sigma^{2}} \left\| X_{j} - \mu_{C(j)} \right\|^{2}$$

Same as K-means!!!

# (One) bad case for K-means

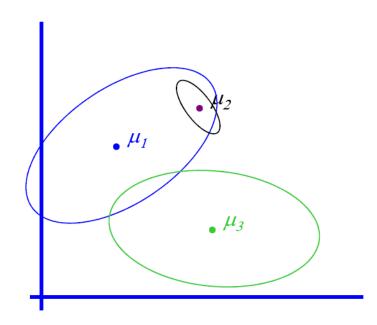
- Clusters may not be linearly separable
- Clusters may overlap
- Some clusters may be "wider" than others

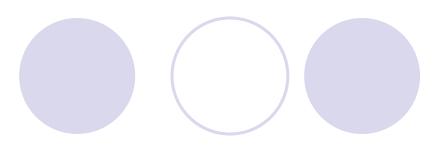
GMM – Gaussian Mixture Model (Multimodal distribution)

- There are k components
- Component i has an associated mean vector µ<sub>i</sub>
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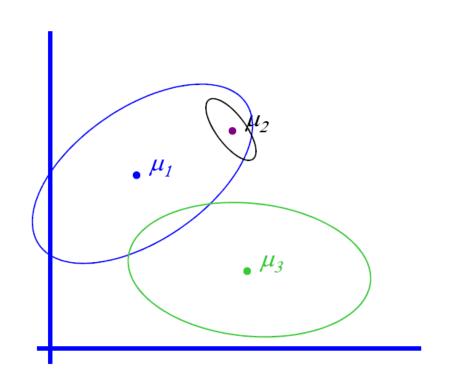


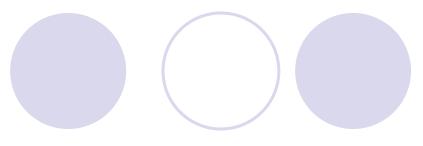
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$$p(x|y=i) \sim N(\mu_i, \Sigma_i)$$

$$p(x) = \sum_{i} p(x/y=i) P(y=i)$$

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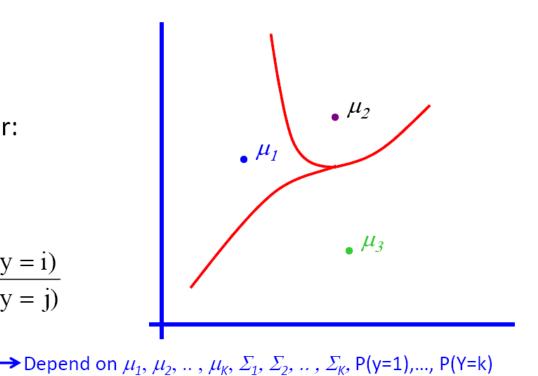
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Gaussian Bayes Classifier:

$$\log \frac{P(y=i \mid x)}{P(y=j \mid x)}$$

$$= \log \frac{p(x \mid y=i)P(y=i)}{p(x \mid y=j)P(y=j)}$$

$$= x^{T}Wx + W^{T}x$$



"Quadratic Decision boundary" - second-order terms don't cancel out



$$\underset{\text{argmax } \prod_{j} \sum_{i=1}^{K} P(y_{j}=i,x_{j})}{\text{argmax } \prod_{j} \sum_{i=1}^{K} P(y_{j}=i,x_{j})}$$

$$= \underset{\text{argmax } \prod_{j} \sum_{i=1}^{K} P(y_{j}=i)p(x_{j}|y_{j}=i)}{\text{P(y_{j}=i)}}$$

Uncertain about class of each  $x_j$  (soft assignment),  $P(y_j=i) = P(y=i)$ 

$$\prod_{j=1}^{m} \sum_{i=1}^{k} P(y_j = i, \boldsymbol{x}_j) \propto \prod_{j=1}^{m} \sum_{i=1}^{k} P(y = i) \frac{1}{\sqrt{det(\boldsymbol{\Sigma}_i)}} exp \Bigg[ -\frac{1}{2} (\boldsymbol{x}_j - \boldsymbol{\mu}_i)^T \, \boldsymbol{\Sigma}_i (\boldsymbol{x}_j - \boldsymbol{\mu}_i) \Bigg]$$

How do we find the  $\mu_i$ 's which give max. marginal likelihood?

\* Set  $\frac{\partial}{\partial \mu_i}$  log Prob (....) = 0 and solve for  $\mu_i$ 's. Non-linear non-analytically solvable

\* Use gradient descent: Often slow but doable

# Expectation-Maximization (EM)

#### A general algorithm to deal with hidden data, but we will study it in the context of unsupervised learning (hidden labels) first

- EM is an optimization strategy for objective functions that can be interpreted as likelihoods in the presence of missing data.
- It is much simpler than gradient methods:

No need to choose step size. Enforces constraints automatically. Calls inference and fully observed learning as subroutines.

• EM is an Iterative algorithm with two linked steps:

E-step: fill-in hidden values using inference M-step: apply standard MLE/MAP method to completed data

 We will prove that this procedure monotonically improves the likelihood (or leaves it unchanged). Thus it always converges to a local optimum of the likelihood.

# Expectation-Maximization (EM)

#### A simple case:

We have unlabeled data  $\mathbf{x}_1 \mathbf{x}_2 \dots \mathbf{x}_m$ 

We know there are k classes

We know P(y=1), P(y=2) P(y=3) ... P(y=K)

We don't know  $\mu_1 \mu_2 ... \mu_k$ 

We know common variance  $\sigma^2$ 

We can write P( data  $| \mu_1 .... \mu_k$ )

$$= \mathbf{p}(x_1...x_m | \mu_1...\mu_k)$$

$$= \prod_{j=1}^{m} p(x_j | \mu_1 ... \mu_k)$$

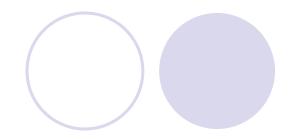
Independent data

$$= \prod_{j=1}^{m} \sum_{i=1}^{k} p(x_j | \mu_i) P(y = i)$$

Marginalize over class

$$\propto \prod_{j=1}^{m} \sum_{i=1}^{k} \exp \left(-\frac{1}{2\sigma^{2}} \|x_{j} - \mu_{i}\|^{2}\right) P(y = i)$$

# Expectation (E) step



If we know  $\mu_1,...,\mu_k \rightarrow \text{easily compute prob. point } x_j \text{ belongs to}$  class y=i

$$P(y = i | x_j, \mu_1...\mu_k) \propto exp(-\frac{1}{2\sigma^2} ||x_j - \mu_i||^2) P(y = i)$$

Simply evaluate gaussian and normalize

# Maximization (M) step



 $\rightarrow$  MLE for  $\mu_i$  is weighted average

imagine multiple copies of each  $x_j$ , each with weight  $P(y=i \mid x_j)$ :

$$\mu_{i} = \frac{\sum_{j=1}^{m} P(y=i|x_{j})x_{j}}{\sum_{j=1}^{m} P(y=i|x_{j})}$$

# EM for spherical, same variance **GMMs**

#### E-step

Compute "expected" classes of all datapoints for each class

$$P(y=i\big|x_{j},\mu_{1}...\mu_{k}) \propto exp\left(-\frac{1}{2\sigma^{2}}\big\|x_{j}-\mu_{i}\big\|^{2}\right) P(y=i) \qquad \text{In K-means "E-step"}$$
 we do hard assignment

In K-means "E-step"

EM does soft assignment

#### M-step

Compute Max. like  $\mu$  given our data's class membership distributions

$$\mu_{i} = \frac{\sum_{j=1}^{m} P(y=i|x_{j})x_{j}}{\sum_{j=1}^{m} P(y=i|x_{j})}$$

# 

Iterate. On iteration t let our estimates be

$$\lambda_{t} = \{ \mu_{1}^{(t)}, \mu_{2}^{(t)} \dots \mu_{k}^{(t)}, \sum_{1}^{(t)}, \sum_{2}^{(t)} \dots \sum_{k}^{(t)}, p_{1}^{(t)}, p_{2}^{(t)} \dots p_{k}^{(t)} \}$$

$$p_i^{(t)} = p^{(t)}(y=i)$$

#### E-step

Compute "expected" classes of all datapoints for each class

$$P(y = i | x_j, \lambda_t) \propto p_i^{(t)} p(x_j | \mu_i^{(t)}, \Sigma_i^{(t)})$$

Just evaluate a Gaussian at  $x_i$ 

#### M-step

Compute Max. like  $\mu$  given our data's class membership distributions

Compute Max. like 
$$\mu$$
 given our data's class membership distributions 
$$\mu_i^{(t+1)} = \frac{\displaystyle\sum_j P(y=i\big|x_j,\lambda_t)x_j}{\displaystyle\sum_j P(y=i\big|x_j,\lambda_t)}$$
 
$$p_i^{(t+1)} = \frac{\displaystyle\sum_j P(y=i\big|x_j,\lambda_t)}{m}$$
  $m = \#$  data points

### **EM** for general **GMMs**

Iterate. On iteration t let our estimates be

$$\lambda_t = \{ \, \mu_1^{(t)}, \, \mu_2^{(t)} \, ... \, \, \mu_k^{(t)}, \, \Sigma_1^{(t)}, \, \Sigma_2^{(t)} \, ... \, \, \Sigma_k^{(t)}, \, p_1^{(t)}, \, p_2^{(t)} \, ... \, p_k^{(t)} \, \}$$

 $p_i^{(t)}$  is shorthand for estimate of P(y=i) on t'th iteration

#### E-step

Compute "expected" classes of all datapoints for each class

$$P(y = i | x_j, \lambda_t) \propto p_i^{(t)} p(x_j | \mu_i^{(t)}, \Sigma_i^{(t)})$$

Just evaluate a Gaussian at  $x_i$ 

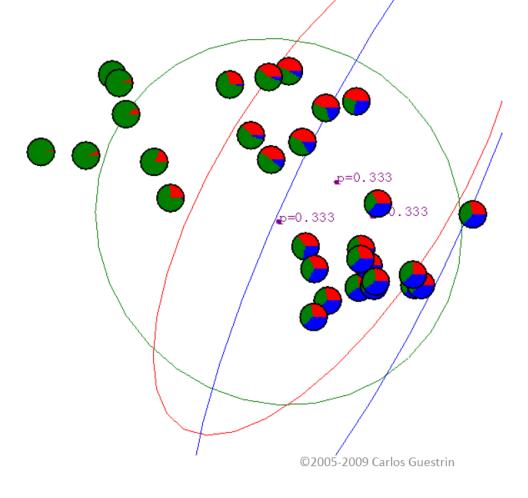
#### M-step

Compute MLEs given our data's class membership distributions (weights)

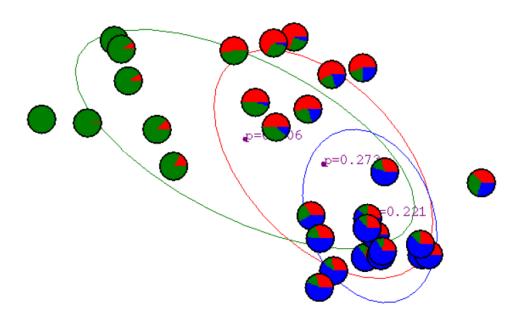
$$\mu_{i}^{(t+1)} = \frac{\sum_{j} P(y = i \big| x_{j}, \lambda_{t}) x_{j}}{\sum_{j} P(y = i \big| x_{j}, \lambda_{t})} \qquad \sum_{i} \frac{\sum_{j} P(y = i \big| x_{j}, \lambda_{t}) \left(x_{j} - \mu_{i}^{(t+1)}\right) \left(x_{j} - \mu_{i}^{(t+1)}\right)^{T}}{\sum_{j} P(y = i \big| x_{j}, \lambda_{t})}$$

$$p_{i}^{(t+1)} = \frac{\sum_{j} P(y = i \big| x_{j}, \lambda_{t})}{m} \qquad m = \# \text{data points}$$

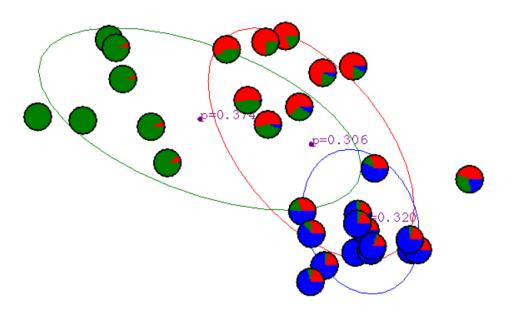
## EM for general GMMs: Example



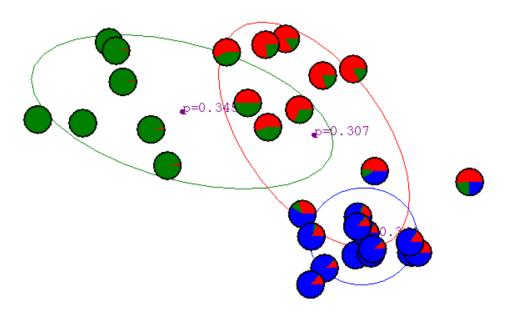
#### After 1<sup>st</sup> iteration



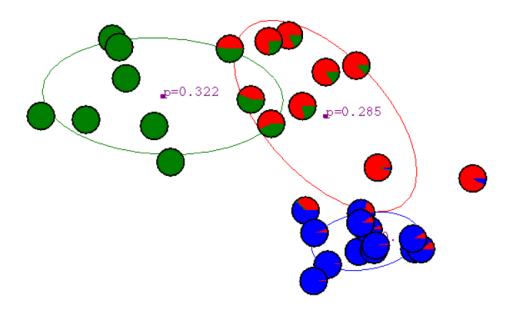
#### After 2<sup>nd</sup> iteration



### After 3<sup>rd</sup> iteration



### After 5<sup>th</sup> iteration



### After 20th iteration

