

Problem 1

$$X_1, X_2, X_3 \sim \exp(\lambda)$$

$$Y \sim \max(X_1, X_2, X_3)$$

$$P(Y \leq y) = P(\max(x_1, x_2, x_3) \leq y)$$

$$= P(X_1 \leq y \text{ and } X_2 \leq y \text{ and } X_3 \leq y)$$

Since the components are independent:

$$= P(X_1 \leq y)P(X_2 \leq y)P(X_3 \leq y)$$

Each components probability is:

$$P(X \leq y) = \int_0^y \lambda e^{-\lambda x} dx = 1 - e^{-\lambda y} \quad 0 < y < \infty \quad \lambda > 0$$

$$P(Y \leq y) = (1 - e^{-\lambda y})^3 \quad 0 < y < \infty$$

$$f_y(y) = \begin{cases} 3(1 - e^{-\lambda y})^2 e^{-\lambda y} & y > 0 \\ 0 & \text{otherwise} \end{cases}$$

Problem 2

Want to find $g(S^2)$ such that $E[g(S^2)] = \sigma$

Let $g(S^2) = c\sqrt{S^2}$ then:

$$E[(c\sqrt{S^2})] = E\left(c\sqrt{S^2} \sqrt{\frac{\sigma^2}{n-1}} \sqrt{\frac{n-1}{\sigma^2}}\right)$$

$$= c\sqrt{\frac{\sigma^2}{n-1}} E\left(\sqrt{\frac{S^2(n-1)}{\sigma^2}}\right)$$

$$\text{Let } Z = \frac{S^2(n-1)}{\sigma^2}$$

$$Z \sim \chi_{n-1}^2$$

Use transformation $Y = \sqrt{Z}$

$$Y = \sqrt{\frac{S^2(n-1)}{\sigma^2}}$$

$$Z = Y^2 \quad \frac{dy}{dz} = \frac{1}{2y}$$

$$\begin{aligned}
f_Y(y) &= f_Z(g^{-1}(z)) \left| \frac{dx}{dy} \right| \\
&= \frac{1}{2^{(n-1)/2} \Gamma((n-1)/2)} (y^2)^{((n-1)/2-1)} e^{-y^2/2} |2y| \\
f_Y(y) &= \frac{2}{2^{(n-1)/2} \Gamma((n-1)/2)} y^{n-2} e^{-y^2/2} \\
E[(c\sqrt{S^2})] &= c \sqrt{\frac{\sigma^2}{n-1}} \int_0^\infty 2y \frac{1}{2^{(n-1)/2} \Gamma((n-1)/2)} y^{n-2} e^{-y^2/2} dy \\
&= c \sqrt{\frac{\sigma^2}{n-1}} \int_0^\infty \frac{1}{2^{(n-1)/2} \Gamma((n-1)/2)} y^{n-2} e^{-y^2/2} (2y) dy \\
&\quad \text{Let } w = y^2 \quad dw = 2y dy \quad y = w^{1/2} \\
&= c \sqrt{\frac{\sigma^2}{n-1}} \int_0^\infty \frac{1}{2^{(n-1)/2} \Gamma((n-1)/2)} w^{n/2-1} e^{-w/2} dw \\
&= c \frac{\sqrt{2}\Gamma(n/2)}{\Gamma((n-1)/2)} \sqrt{\frac{\sigma^2}{n-1}} \int_0^\infty \frac{1}{2^{n/2} \Gamma(n/2)} w^{n/2-1} e^{-w/2} dw \\
E[(c\sqrt{S^2})] &= \sigma = c \frac{\sqrt{2}\Gamma(n/2)}{\Gamma((n-1)/2)} \sqrt{\frac{\sigma^2}{n-1}} \\
c &= \sigma * 1 / \left(\frac{\sqrt{2}\Gamma(n/2)}{\Gamma((n-1)/2)} \sqrt{\frac{\sigma^2}{n-1}} \right) \\
c &= \frac{\sqrt{n-1} \Gamma((n-1)/2)}{\sqrt{2} \Gamma(n/2)}
\end{aligned}$$

Problem 3

$$\begin{aligned}
P(Z > z) &= \sum_{x=1}^{\infty} P(Z > z|x) P(X = x) \\
&= \sum_{x=1}^{\infty} P(U_1 > z, \dots, U_x > z|x) P(X = x) \\
&\quad \text{Since } U_i \text{ are independent:} \\
&= \sum_{x=1}^{\infty} \prod_{i=1}^x P(U_i > z) P(X = x)
\end{aligned}$$

$$= \sum_{x=1}^{\infty} P(U_i > z)^x P(X = x)$$

$$= \sum_{x=1}^{\infty} (1-z)^x \frac{1}{x!(e-1)}$$

$$= \frac{1}{e-1} \sum_{x=1}^{\infty} \frac{(1-z)^x}{x!}$$

$$\text{Since } \sum_{x=1}^{\infty} z^x/x! = e^z - 1$$

$$\sum_{x=1}^{\infty} (1-z)^x/x! = e^{1-z} - 1$$

Thus we have:

$$= \frac{e^{1-z} - 1}{e - 1} \quad 0 < z < 1$$

Problem 4

(a)

$$T|V = v \sim \frac{U}{\sqrt{v/p}} \sim N(0, p/v)$$

$$f_{T|V=v}(t) = \frac{\sqrt{v}}{\sqrt{2p\pi}} e^{\left(\frac{-t^2 v}{2p}\right)}$$

$$f_{T,V}(t, v) = f_{T|V}(t) f_V(v)$$

$$f_{T,V}(t, v) = \frac{\sqrt{v}}{\sqrt{2p\pi}} e^{\left(\frac{-t^2 v}{2p}\right)} \frac{1}{\Gamma(p/2) 2^{p/2}} v^{p/2-1} e^{-v/2}$$

$$f_{T,V}(t, v) = \frac{1}{\Gamma(p/2) \sqrt{p\pi} 2^{(p-1)/2}} e^{(-v/2(1+t^2/p))} v^{(p-1)/2}$$

$$f_T(t) = \int_0^{\infty} f_{T,V}(t, v) dv$$

$$\frac{1}{\Gamma(p/2) \sqrt{p\pi} 2^{(p-1)/2}} \int_0^{\infty} e^{(-v/2(1+t^2/p))} v^{(p-1)/2} dv$$

$$\text{Let } z = v/2(1+t^2/p) \quad dz = 1/2(1+t^2/p) dv \text{ then } v = \frac{2z}{1+t^2/p}$$

$$\begin{aligned}
f_T(t) &= \frac{2^{(p-1)/2+1}}{\sqrt{p\pi}} \frac{(1+t^2/p)^{-(p-1)/2+1}}{\Gamma(p/2)2^{(p+1)/2}} \int_0^\infty e^{-z} z^{(p-1)/2} dz \\
&= \frac{(1+t^2/p)^{-(p+1)/2}}{\Gamma(p/2)\sqrt{p\pi}} \int_0^\infty e^{-z} z^{(p+1)/2-1} dz \\
f_T(t) &= \frac{\Gamma((p+1)/2)}{\Gamma(p/2)\sqrt{p\pi}} (1+t^2/p)^{-(p+1)/2}
\end{aligned}$$

Which is the t-distribution with p degrees of freedom

(b)

$$\begin{aligned}
E(T) &= E\left(\frac{U}{\sqrt{V/p}}\right) \\
&= E\left(E\left(\frac{U}{\sqrt{V/p}} \middle| V = v\right)\right) \\
&= E\left(\frac{1}{\sqrt{v/p}} E(U)\right) \\
&= E\left(\sqrt{\frac{p}{v}} * 0\right) = 0 \\
E(T) &= 0 \\
Var(T) &= E\left(Var\left(\frac{U}{\sqrt{V/p}} \middle| V = v\right)\right) + Var\left(E\left(\frac{U}{\sqrt{V/p}} \middle| V = v\right)\right) \\
&= E\left(\frac{p}{V} Var(U)\right) + 0 \\
&= E\left(\frac{p}{V} * 1\right) \\
&= pE(1/V) \\
Var(T) &= pE(v^{-1}) = p \int_0^\infty v^{-1} \frac{1}{\Gamma(p/2)2^{p/2}} v^{p/2-1} e^{-v/2} dv \\
&= p \int_0^\infty \frac{1}{\Gamma(p/2)2^{p/2}} v^{(p/2-1)-1} e^{-v/2} dv \\
&= p \frac{\Gamma(p/2-1)2^{p/2-1}}{\Gamma(p/2)2^{p/2}} \int_0^\infty \frac{1}{\Gamma(p/2-1)2^{p/2-1}} v^{(p/2-1)-1} e^{-v/2} dv \\
Var(T) &= p \frac{\Gamma(p/2-1)2^{p/2-1}}{\Gamma(p/2)2^{p/2}} \\
&= \frac{p}{2} \frac{\Gamma(p/2-1)}{\Gamma(p/2)}
\end{aligned}$$

Since $p/2 - 1$ and $p/2$ are whole numbers we have:

$$\begin{aligned}
 &= \frac{p}{2} \frac{(p/2 - 1)!}{(p/2 - 1)!} \\
 &= \frac{p}{2(p/2 - 1)} = \frac{p}{p - 2} \\
 \text{Var}(T) &= \frac{p}{p - 2}
 \end{aligned}$$

(c)

$$\begin{aligned}
 U &\sim N(0, 1) \quad V \sim \chi_p^2 \\
 f_{UV}(u, v) &= f_U(u)f_V(v) = \frac{1}{\sqrt{2\pi}} e^{-u^2/2} \frac{1}{\Gamma(p/2)2^{p/2}} v^{p/2-1} e^{-v/2} \\
 \text{Let } T &= \frac{U}{\sqrt{V/p}} \quad W = V \\
 V &= W \quad U = T\sqrt{W/p} \\
 J &= \begin{bmatrix} \sqrt{w/p} & -t \\ 0 & 1 \end{bmatrix} = |\sqrt{w/p}| \\
 f_{TW}(t, w) &= f_{UV}(t\sqrt{w/p}, w)(\sqrt{w/p}) \\
 &= \frac{1}{\sqrt{2\pi}} e^{-(t\sqrt{w/p})^2/2} \frac{1}{\Gamma(p/2)2^{p/2}} w^{p/2-1} e^{-w/2} (\sqrt{w/p}) \\
 &= \frac{1}{\sqrt{2\pi}} e^{-\left(\frac{t^2 w}{2p}\right)} \frac{1}{\Gamma(p/2)2^{p/2}} w^{(p-1)/2} e^{-w/2} (1/\sqrt{p}) \\
 f_T(t) &= \int_0^\infty f_{TW} dw \\
 &= \frac{1}{\sqrt{2p\pi}} \frac{1}{\Gamma(p/2)2^{p/2}} \int_0^\infty e^{(-w/2)(t^2/p+1)} w^{(p-1)/2} dw \\
 \text{Let } z &= (1 + t^2/p)(w/2) \quad dz = 1/2(1 + t^2/p) dw \quad \text{Then } w = \frac{2z}{1 + t^2/p} \\
 f_T(t) &= \frac{2^{(p-1)/2+1}}{\sqrt{p\pi}} \frac{(1 + t^2/p)^{-[(p-1)/2+1]}}{\Gamma(p/2)2^{(p+1)/2}} \int_0^\infty e^{-z} z^{(p-1)/2} dz \\
 &= \frac{(1 + t^2/p)^{-(p+1)/2}}{\Gamma(p/2)\sqrt{p\pi}} \int_0^\infty e^{-z} z^{(p+1)/2-1} dz \\
 f_T(t) &= \frac{\Gamma((p+1)/2)}{\Gamma(p/2)\sqrt{p\pi}} (1 + t^2/p)^{-(p+1)/2}
 \end{aligned}$$

Which is the t-distribution with p degrees of freedom

Problem 5

(a)

$$U \sim \min(X_1, X_2)$$

$$P(U \geq t) = P(\min(x_1, x_2) \geq t)$$

$$\pi_0(t) = P(X_1 \geq t, X_2 \geq t)$$

Since X_i are independent:

$$= P(X_1 \geq t)P(X_2 \geq t)$$

Each probability is:

$$P(X \geq t) = 1 - \int_0^t \alpha e^{-\alpha x} dx$$

$$= 1 + \left| e^{-\alpha x} \right|_0^t = e^{-\alpha t}$$

$$\pi_0(t) = 1 - F_U(t) = e^{(-\alpha t)^2} = e^{-2\alpha t} \quad t \geq 0$$

$\pi_0(t)$ is the probability that both organs are still functioning at time t

(b)

$$1 - f_U(u) = \frac{d}{du} 1 - e^{-2\alpha u} = 2\alpha e^{-2\alpha u}$$

$$f_{UV}(u, v) = f_V(V|U = u)f_U(u)$$

$$f_{UV}(u, v) = \beta e^{-\beta(v-u)}(2\alpha)e^{-2\alpha u}$$

$$= 2\alpha\beta e^{-\beta v} e^{-(2\alpha-\beta)u}$$

$$\pi_1(t) = P(U \leq t, V \geq t) = \int_0^t \int_t^\infty f_{UV}(u, v) dv du$$

$$= \int_0^t \int_t^\infty 2\alpha\beta e^{-\beta v} e^{-(2\alpha-\beta)u} dv du$$

$$= 2\alpha\beta \int_0^t \left| \frac{-1}{\beta} e^{-\beta v} e^{-(2\alpha-\beta)u} \right|_t^\infty du$$

$$= 2\alpha \int_0^t e^{-\beta t} e^{-(2\alpha-\beta)u} du$$

$$= 2\alpha \left| \frac{-1}{2\alpha-\beta} e^{-\beta t} e^{-(2\alpha-\beta)u} \right|_0^t$$

$$\begin{aligned}
&= \frac{2\alpha}{\beta - 2\alpha} e^{-\beta t} (e^{-(2\alpha - \beta)t} - 1) \\
\pi_1(t) &= \frac{2\alpha}{\beta - 2\alpha} (e^{-2\alpha t} - e^{-\beta t}) \quad t \geq 0
\end{aligned}$$

Which is the probability that exactly one kidney is functioning at time t

(c)

$$\begin{aligned}
f_V(t) &= f_T(t) \\
f_V(t) &= \int_0^t f_{UV}(u, v) \, du \\
&= \int_0^t 2\alpha\beta e^{-\beta t} e^{-(2\alpha - \beta)u} \, du \\
&= \left[\frac{2\alpha\beta}{\beta - 2\alpha} e^{-\beta t} e^{-(2\alpha - \beta)u} \right]_0^t \\
&= \frac{2\alpha\beta}{\beta - 2\alpha} e^{-\beta t} (e^{-(2\alpha - \beta)t} - 1) \\
f_T(t) = f_V(t) &= \frac{2\alpha\beta}{\beta - 2\alpha} (e^{-2\alpha t} - e^{-\beta t}) \quad t \geq 0
\end{aligned}$$