## **Deletion Diagnostics**

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Ordinary least-quares: The setup is the usual one with design matrix X of dimensions  $n \times p$ , parameters  $\beta$ , OLS estimates  $\hat{\beta}$  and residual variance estimate  $s^2$ . Estimates obtained from the data with the ith observation deleted are denoted (i) as in  $\hat{\beta}_{(i)}$ . The diagonal elements of the hat matrix are  $h_i = x_i^{\mathsf{T}} (X^{\mathsf{T}} X)^{-1} x_i$ . Since  $X^{\mathsf{T}} X = \sum_{i=1}^n x_i x_i^{\mathsf{T}}$  and  $X^{\mathsf{T}} y = \sum_{i=1}^n y_i x_i$ , after deleting the ith observation,  $X^{\mathsf{T}} X$  becomes  $X^{\mathsf{T}} X - x_i x_i^{\mathsf{T}}$  and  $X^{\mathsf{T}} y$  becomes  $X^{\mathsf{T}} y - y_i x_i$ . The fitted values are  $\hat{\mu}_i = x_i^{\mathsf{T}} \hat{\beta}$ , and the residuals are  $r_i = y_i - \hat{\mu}_i$ . The standardized residuals are

$$\frac{r_i}{\sqrt{(1-h_i)}}$$

and the Studentized standardized residuals are the scaled version

$$\frac{r_i}{s\sqrt{(1-h_i)}}.$$

A matrix formula: For matrix A and (column) vectors b and c,

$$(A + bc^{\mathsf{T}})^{-1} = A^{-1} - \frac{A^{-1}bc^{\mathsf{T}}A^{-1}}{1 + b^{\mathsf{T}}A^{-1}c},$$

assuming the dimensions are conformable and the inverses all exist.

In the OLS setup, take  $A = X^{\top}X$ ,  $b = x_i$  and  $c = -x_i$ . This gives an expression for  $(X^{\top}X - x_ix_i^{\top})^{-1}$ , which is then used to obtain,

$$\hat{\beta} - \hat{\beta}_{(i)} = (X^{\top} X)^{-1} x_i \frac{r_i}{1 - h_i}.$$

Cook's distance is a measure of influence of an observation on the parameter estimates  $\hat{\beta}$ . It is usually defined as (the quadratic form)

$$D_{i} = \frac{1}{ps^{2}} (\hat{\beta} - \hat{\beta}_{(i)})^{\top} (X^{\top} X) (\hat{\beta} - \hat{\beta}_{(i)}).$$

Note: There are various related but slightly different forms of Cook's distance in the literature. The above reduces to

$$D_i = \frac{r_i^2}{ps^2} \frac{h_i}{(1 - h_i)^2},$$

which shows that influence is essentially the product of residual  $(r_i^2 \text{ or } (r_i/s)^2)$  times leverage  $(h_i/(1-h_i)^2 = h_{(i)}/(1-h_i))$ .

Note that the above formulae give exact results, not approximations. There is no need to refit the model n times. Once the fitted values and the hat matrix are available, all the above diagnostics are easily computed.

Weighted least-quares (WLS): The problem is transformed to OLS by defining  $t_i = y_i \sqrt{w_i}$  and  $z_i = x_i \sqrt{w_i}$ . Now, all the OLS formulae given above apply to (t, Z).

Generalized linear models for independent outcomes: These models require iteration to find estimates. Approximations (usually accurate) to the quantities that would be computed after full iteration are possible: The iterative weights are used as if they were the weights in WLS.

**GEE**: Here two deletions are possible; either one observation or one cluster. So there are two versions of each of the quantities defined above; one for observation deletion and one for cluster deletion.

References: The (minimum) recommended reading for OLS is chapter 13 in the BIOS 663 textbook. The classic books on the topic are Cook & Weisberg (1982) and Belsley, Kuh & Welsch (1980) - both highly recommended. For generalized linear models for independent outcomes, section 12.7 of the book by McCullagh & Nelder (1989, 2nd ed.) gives the basics. For dependent outcomes and GEE, see Preisser & Qaqish (1996, Biometrika).