

Converge in Probability

$X_n \xrightarrow{p} X$ as $n \rightarrow \infty$
 $\lim_{n \rightarrow \infty} P(|X_n - X| < \epsilon) = 1$

WLLN

a) X_i iid **b)** $E(X_i) = \mu$ **c)** $Var(X_i) = \sigma^2 < \infty$

Let $\bar{X}_n = \sum_{i=1}^n X_i$ Then for any $\epsilon > 0$:

$\lim_{n \rightarrow \infty} P(|\bar{X}_n - \mu| < \epsilon) = 1$

$\bar{X}_n \xrightarrow{p} \mu$ (consistency of \bar{X}_n)

The condition $E(X_i)$ exists and is finite is sufficient in WLLN

Converge in Distribution

Let F_{X_n} be cdf of X_n

Then $X_n \xrightarrow{d} X$ as $n \rightarrow \infty$

If $\lim_{n \rightarrow \infty} F_{X_n}(x) = F_X(x)$ at all points where $F_X(x)$ is continuous

For large n , the cdf of X_n becomes close to the cdf of X

Convergence in distribution does not imply X_n and X approximate each other

CLT

a) X_i iid **b)** $E(X_i) = \mu$ **c)** $Var(X_i) = \sigma^2 < \infty$

Let $\bar{X}_n = \sum_{i=1}^n X_i$ (so far same setup as WLLN)

Let $Z_n = \sqrt{n}(\bar{X}_n - \mu)/\sigma$ and let G_n denote cdf of Z_n

Then for any $-\infty < z < \infty$: $\lim_{n \rightarrow \infty} G_n(z) = \Phi(z)$

Z_n has a limiting standard normal distribution: $Z_n \xrightarrow{d} N(0, 1)$ as $n \rightarrow \infty$

Relationships between Modes of Convergence

$X_n \xrightarrow{p} X \Rightarrow X_n \xrightarrow{d} X$

Slutsky's Theorem

If **a)** $X_n \xrightarrow{d} X$ and **b)** $Y_n \xrightarrow{p} a$ (a is finite constant)

Then: **1)** $Y_n X_n \xrightarrow{d} aX$ **2)** $Y_n + X_n \xrightarrow{d} a + X$ **3)** $X_n/Y_n \xrightarrow{d} X/a$ (if $a \neq 0$)

X_n and Y_n don't have to be independent

Slutsky's allows substituting consistent estimators when proving convergence in distribution

Convergence of Transformed Sequences

If $X_n \xrightarrow{p} X$ then $h(X_n) \xrightarrow{p} h(X)$

If $X_n \xrightarrow{d} X$ then $h(X_n) \xrightarrow{d} h(X)$

h only has to be continuous on range of X

$S_n^2 \xrightarrow{a.s.} \sigma^2$ as $n \rightarrow \infty$ $\bar{X}_n^2 \xrightarrow{a.s.} \mu^2$ as $n \rightarrow \infty$

Delta Method - Univariate

Suppose $\{T_n\}$ is a random sequence with

$\sqrt{n}(T_n - \theta) \xrightarrow{d} N(0, \sigma^2)$ and

g is a function with $g'(\theta)$ exists and $\neq 0$ Then:

$\sqrt{n}\{g(T_n) - g(\theta)\} \xrightarrow{d} N(0, \{g'(\theta)\}^2 \sigma^2)$

θ is the asymptotic mean of T_n

θ may or may not be mean of T_n or mean may not even exist