

Question 1

$$Y_1 \dots Y_n \sim N(\mu, \sigma^2)$$

$$\Sigma = \sigma^2 I \text{ is } n \times n$$

$$X \text{ is } 1_n \text{ is } n \times 1$$

$$R_1 = Y_i - Y_1$$

$$R_{n \times 1}$$

$$R_2 = Y_i - \bar{Y} \quad \bar{Y} = \frac{1}{n} \sum_{i=1}^n Y_i$$

$$\ell_{\text{REML}}(\theta, R) = -\frac{1}{2} \log |\Sigma| - \frac{1}{2} \log |X^T \Sigma^{-1} X| - \frac{1}{2} Q(R, \Sigma, X)$$

$$X^T \Sigma^{-1} X = [1 \dots 1] \frac{1}{\sigma^2} I \begin{bmatrix} 1 \\ \vdots \\ 1 \end{bmatrix} = \frac{n}{\sigma^2}$$

$$\Rightarrow \ell_{\text{REML}}(\theta, R) = -\frac{1}{2} \log(\sigma^2 I) - \frac{1}{2} \log\left(\frac{n}{\sigma^2}\right) - \frac{1}{2} Q(R, \Sigma, X)$$

$$Q(R, \Sigma, X) = R^T \{ \Sigma^{-1} - \Sigma^{-1} X (X^T \Sigma^{-1} X)^{-1} X^T \Sigma^{-1} \} R$$

$$= \{ R^T \Sigma^{-1} - R^T \Sigma^{-1} X (X^T \Sigma^{-1} X)^{-1} X^T \Sigma^{-1} \} R$$

$$= R^T \Sigma^{-1} R - R^T \Sigma^{-1} X (X^T \Sigma^{-1} X)^{-1} X^T \Sigma^{-1} R$$

$$= R^T \frac{1}{\sigma^2} I R - R^T \frac{1}{\sigma^2} I X \cdot \left(\frac{\sigma^2}{n}\right) X^T \left(\frac{1}{\sigma^2}\right) R$$

$$= \frac{1}{\sigma^2} R^T R - \frac{1}{\sigma^2} \frac{\sigma^2}{n} \frac{1}{\sigma^2} R^T X X^T R$$

$$= \frac{1}{\sigma^2} R^T R - \frac{1}{\sigma^2 n} R^T J R = \frac{1}{\sigma^2} \sum_{i=1}^n R_i^2 - \frac{1}{\sigma^2 n} \left(\sum_{i=1}^n R_i \right)^2$$

$$\ell_{\text{REML}}(\theta, R) = -\frac{1}{2} \log(\sigma^2 I) - \frac{1}{2} \log\left(\frac{n}{\sigma^2}\right) - \frac{1}{2} \left[\frac{1}{\sigma^2} \sum_{i=1}^n R_i^2 - \frac{1}{\sigma^2 n} \left(\sum_{i=1}^n R_i \right)^2 \right]$$

\Rightarrow if R 's vary the $Q(R, \Sigma, X)$ changes

For $R_1 = Y_i - Y_1$

$$\ell_{\text{REML}}(\theta, R_1) = -\frac{1}{2} \log(\sigma^2 I) - \frac{1}{2} \log\left(\frac{n}{\sigma^2}\right) - \frac{1}{2} \left[\frac{1}{\sigma^2} \sum_{i=1}^n (Y_i - Y_1)^2 - \frac{1}{\sigma^2 n} \left(\sum_{i=1}^n (Y_i - Y_1) \right)^2 \right]$$

$$= \sum_{i=1}^n (Y_i - Y_1)(Y_i - Y_1) = \sum_{i=1}^n Y_i^2 - 2Y_1 Y_i + Y_1^2$$

$$= \sum_{i=1}^n Y_i^2 - 2 \sum_{i=1}^n Y_i Y_1 + \sum_{i=1}^n Y_1^2$$

$$\sum_{i=1}^n Y_i^2 - 2Y_1 \sum_{i=1}^n Y_i + nY_1^2$$

$$\left(\sum_{i=1}^n Y_i - \sum_{i=1}^n Y_1 \right)^2$$

$$\left(\sum_{i=1}^n Y_i \right)^2 - 2 \sum_{i=1}^n Y_i \sum_{i=1}^n Y_1 + \left(\sum_{i=1}^n Y_1 \right)^2$$

$$\left(\sum_{i=1}^n Y_i \right)^2 - 2Y_1 n \sum_{i=1}^n Y_i + (nY_1)^2$$

$$\begin{aligned}
 Q(R_1, \Sigma, x) &= \frac{1}{\sigma^2} \left(\sum_{i=1}^n Y_i^2 - 2Y_1 \sum_{i=1}^n Y_i + nY_1^2 \right) - \frac{1}{\sigma^2 n} \left[\left(\sum_{i=1}^n Y_i \right)^2 - 2Y_1 n \sum_{i=1}^n Y_i + (nY_1)^2 \right] \\
 &= \frac{1}{\sigma^2} \sum_{i=1}^n Y_i^2 - \frac{2}{\sigma^2} Y_1 \sum_{i=1}^n Y_i + \frac{n}{\sigma^2} Y_1^2 - \frac{1}{\sigma^2 n} \left(\sum_{i=1}^n Y_i \right)^2 + \frac{2}{\sigma^2} Y_1 \sum_{i=1}^n Y_i - \frac{n}{\sigma^2} Y_1^2 \\
 &= \frac{1}{\sigma^2} \sum_{i=1}^n Y_i^2 - \frac{1}{\sigma^2 n} \left(\sum_{i=1}^n Y_i \right)^2
 \end{aligned}$$

$$R_2 = Y_i - \bar{Y}$$

$$\begin{aligned}
 Q(R_2, \Sigma, x) &= \frac{1}{\sigma^2} \sum_{i=1}^n R_i^2 - \frac{1}{\sigma^2 n} \left(\sum_{i=1}^n R_i \right)^2 \\
 &= \frac{1}{\sigma^2} \sum_{i=1}^n (Y_i - \bar{Y})^2 - \frac{1}{\sigma^2 n} \left(\sum_{i=1}^n Y_i - n\bar{Y} \right)^2 \\
 &= \frac{1}{\sigma^2} \sum_{i=1}^n (Y_i^2 - 2\bar{Y}Y_i + \bar{Y}^2) - \frac{1}{\sigma^2 n} \left(\sum_{i=1}^n Y_i - n\bar{Y} \right)^2 \\
 &= \frac{1}{\sigma^2} \left[\sum_{i=1}^n Y_i^2 - 2\bar{Y} \sum_{i=1}^n Y_i + n\bar{Y}^2 \right] - \frac{1}{\sigma^2 n} \left(\sum_{i=1}^n Y_i - n\bar{Y} \right)^2 \\
 &= \frac{1}{\sigma^2} \sum_{i=1}^n Y_i^2 - \frac{2}{\sigma^2} \bar{Y} \sum_{i=1}^n Y_i + \frac{n}{\sigma^2} \bar{Y}^2 - \frac{1}{\sigma^2 n} \left[\left(\sum_{i=1}^n Y_i \right)^2 - 2n\bar{Y} \sum_{i=1}^n Y_i + n^2 \bar{Y}^2 \right] \\
 &= \frac{1}{\sigma^2} \sum_{i=1}^n Y_i^2 - \frac{2}{\sigma^2} \bar{Y} \sum_{i=1}^n Y_i + \frac{n}{\sigma^2} \bar{Y}^2 - \frac{1}{\sigma^2 n} \left(\sum_{i=1}^n Y_i \right)^2 + \frac{2}{\sigma^2} \bar{Y} \sum_{i=1}^n Y_i - \frac{n}{\sigma^2} \bar{Y}^2 \\
 &= \frac{1}{\sigma^2} \sum_{i=1}^n Y_i^2 - \frac{1}{\sigma^2 n} \left(\sum_{i=1}^n Y_i \right)^2
 \end{aligned}$$

Conclusion: The REML for R_1 is the same as the REML for R_2 .

Question 2

$$\begin{bmatrix} \mu_1 \\ \mu_2 \\ \mu_3 \\ \mu_4 \end{bmatrix} = \begin{bmatrix} 1 \\ 1 \\ 1 \\ 1 \end{bmatrix} \beta_0 + \begin{bmatrix} t_1 \\ t_2 \\ t_3 \\ t_4 \end{bmatrix} \beta_1 + \begin{bmatrix} t_1^2 \\ t_2^2 \\ t_3^2 \\ t_4^2 \end{bmatrix} \beta_2 + \begin{bmatrix} t_1^3 \\ t_2^3 \\ t_3^3 \\ t_4^3 \end{bmatrix} \beta_3$$

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$\bar{y} \cdot P$ = orthogonal polynomial estimate

$$\begin{bmatrix} 0.5 & 0.5 & 0.5 & 0.5 \\ -0.67 & -0.22 & 0.22 & 0.67 \\ 0.5 & -0.5 & -0.5 & 0.5 \\ -0.22 & 0.67 & -0.67 & 0.22 \end{bmatrix} \begin{bmatrix} 1.14 \\ 2.54 \\ 2.99 \\ 3.47 \end{bmatrix} = \begin{bmatrix} 5.67 \\ 1.66 \\ -0.46 \\ 0.22 \end{bmatrix}$$

P

$$\begin{bmatrix} \bar{y}_1 \\ \bar{y}_2 \\ \bar{y}_3 \\ \bar{y}_4 \end{bmatrix} \begin{bmatrix} \beta_0 \\ \beta_1 \\ \beta_2 \\ \beta_3 \end{bmatrix}$$

→ the second row of P is the contrast of linear component (β_1) $[-0.67 \ -0.22 \ 0.22 \ 0.67] = 0$

→ the third row of P is the contrast of quadratic component (β_2)

→ the fourth row of P is the contrast of the cubic component (β_3)

(a) if the true population linear contrast = 0 it means the linear component has no effect on the overall mean

$$P_2 = [-0.67 \ -0.22 \ 0.22 \ 0.67] \div 0.22 = [-3 \ -1 \ 1 \ 3]$$

$$[-3 \ -1 \ 1 \ 3] \begin{bmatrix} \bar{y}_1 \\ \bar{y}_2 \\ \bar{y}_3 \\ \bar{y}_4 \end{bmatrix} = -3\bar{y}_1 - \bar{y}_2 + \bar{y}_3 + 3\bar{y}_4 = \bar{y}_3 + 3\bar{y}_4 - [3\bar{y}_1 + \bar{y}_2]$$

The difference between the average mean at time 3 and three times the mean at time 4, and three times the mean at time 1 and time 2.



(B) what does it mean to say the quadratic contrast is zero?
if contrast = 0 \Rightarrow quadratic component was no effect
on overall mean.

$$[0.5 \quad -0.5 \quad -0.5 \quad 0.5] \div 0.5 = [1 \quad -1 \quad -1 \quad 1]$$

$$[1 \quad -1 \quad -1 \quad 1] \begin{bmatrix} \bar{Y}_1 \\ \bar{Y}_2 \\ \bar{Y}_3 \\ \bar{Y}_4 \end{bmatrix} = \bar{Y}_1 - \bar{Y}_2 - \bar{Y}_3 + \bar{Y}_4$$
$$= (\bar{Y}_1 + \bar{Y}_4) - (\bar{Y}_3 + \bar{Y}_2)$$

\Rightarrow The difference between the average of the means of time 1 and time 4, and the average of the means at time 2 and time 3.



Question 3

(a)

- M1 - covariance is unstructured and different for each group ✓
- different means for each level of group and time, ✓
 - 8 different combinations of group and time that all have different intercepts,
- M2 - covariance is unstructured and not different for each group. ✓
- the mean structure is the same as model 1. ✓
- M3 - the covariance is unstructured and not different for each group. ✓
- the model assumes that the marginal mean follow a linear trend with time x.
 - each treatment group has its own line, one for placebo and one for active w/ two different slopes and two different intercepts. ✓
- M4 - covariance is unstructured and not different for each group. ✓
- the model assumes that the marginal mean follows a linear trend with time x.
 - two different slopes (placebo and active) and the same intercept. ✓
- M5 - covariance is unstructured and not different for each group. ✓
- 4 different combinations of group (placebo and active) and time (baseline and not baseline) that have different intercepts ✓

M6 - covariance is unstructured and the same for both groups. ✓

- model takes $t1, t4$ and active
- the intercept is the same for both active and placebo. ✓

M7 - covariance is unstructured and the same for both groups. ✓

- model uses different time points and active.
- the intercept (baseline) is the same for both treatments.
- means are different for each timepoint after baseline ($t1, t4, t6$) for each trt group. ✓

B) • $K=100$

• separated by group and time.

• $\mu_i = (\mu_{i1}, \mu_{i2}, \mu_{i3}, \mu_{i4})^T$ for the four different time points,

$$X_{ij} = (1, t1_{ij}, t4_{ij}, t6_{ij}, t1^* \text{active}_{ij}, t4^* \text{active}_{ij}, t6^* \text{active}_{ij})$$

$$i = 1, \dots, K \quad j = 1, 2, 3, 4$$

where

- $t1 = 1$ when $j=2$ $t1 = 0$ otw
 - $t4 = 1$ when $j=3$ $t2 = 0$ otw
- ect. ✓

$$\begin{bmatrix} \mu_1 \\ \mu_2 \\ \mu_3 \\ \mu_4 \end{bmatrix} = \begin{bmatrix} 1 & t_1 & t_4 & t_6 & t_1 * \text{active} & t_4 * \text{active} & t_6 * \text{active} \\ 1 & t_1 & t_4 & t_6 & t_1 * \text{active} & t_4 * \text{active} & t_6 * \text{active} \\ 1 & t_1 & t_4 & t_6 & t_1 * \text{active} & t_4 * \text{active} & t_6 * \text{active} \\ 1 & t_1 & t_4 & t_6 & t_1 * \text{active} & t_4 * \text{active} & t_6 * \text{active} \end{bmatrix} \begin{bmatrix} \beta_0 \\ \beta_1 \\ \beta_2 \\ \beta_3 \\ \beta_4 \\ \beta_5 \\ \beta_6 \end{bmatrix}$$

covariance structure? -2



(c)

t_6

" β_3 " is the mean incremental increase of lead blood levels at week 6 compare to baseline for placebo.

$t_6 * \text{active}$

" β_6 " is the mean incremental increase of interaction effect of week 6 and active.
i.e. difference btwn. placebo & active. -2

(d) To see the treatment effect you would look at the interaction parameters

($t_1 * \text{active}$, $t_4 * \text{active}$, $t_6 * \text{active}$). This would show the treatment effect because its adding an additional incremental increase solely because of the active treatment.

