Summary

Finally, the end of the semester! We made it!

Topic 1: Introduction and Overview



- 1. Why linear regression, why not t-test?
- 2. Basic concepts: Population, Sample, parameter, statistic
- 3. Statistical Activities: Parameter Estimation, Inference

Topic 2: Linear Algebra Review

- 1. Matrix operation, matrix addition, matrix multiplication ...
- 2. An orthogonal matrix is a square matrix with $A' = A^{-1}$.
- 3. Rules of Matrix Operation.
- 4. Linear Dependence and Rank, matrix determinant
- 5. Positive Definite and Semi-positive Definite Matrices
- 6. <u>Inverse</u> and Generalized Inverse
- 7. Eigenvalues, Eigenvectors. Suppose $\bf A$ is an symmetric matrix. Then there exists an orthogonal (column orthonormal) matrix $\bf V$ such that $\bf A = \bf V \Lambda \bf V'$.

8. Random Vectors and Matrices

$$E(\mathbf{AY} + \mathbf{b}) = \mathbf{A}E(\mathbf{Y}) + \mathbf{b} = \mathbf{A}\boldsymbol{\mu} + \mathbf{b}$$
$$Cov(\mathbf{AY} + \mathbf{b}) = \mathbf{A}Cov(\mathbf{Y})\mathbf{A}' = \mathbf{A}\boldsymbol{\Sigma}\mathbf{A}'.$$

- 9. Important Distributions for Linear Models. If $Z \sim N(0,1)$, $X_1 \sim \chi^2(n_1)$ and $X_2 \sim \chi^2(n_2)$, and X_1 and X_2 are independent. Construct random variables following t-distribution and F distribution.
- 10. Maximum Likelihood Estimates (MLE)

Topics 3 and 4: Simple Linear Regression and the General Linear Model: Estimation and Testing

1.
$$\mathbf{y}_{n\times 1} = \mathbf{X}_{n\times p} \, \boldsymbol{\beta}_{p\times 1} + \boldsymbol{\varepsilon}_{n\times 1}$$

- 2. Least Squares Estimation: $\widehat{\boldsymbol{\beta}} = \operatorname{argmin}_{\boldsymbol{\beta}} (\mathbf{y} \mathbf{X}\boldsymbol{\beta})'(\mathbf{y} \mathbf{X}\boldsymbol{\beta}).$ $E(\widehat{\boldsymbol{\beta}}) = \boldsymbol{\beta} \text{ and } \operatorname{Cov}(\widehat{\boldsymbol{\beta}}) = \sigma^2(\mathbf{X}'\mathbf{X})^{-1}.$
- 3. HILE Gauss
 - Existence Assumption
 - Linearity Assumption
 - Independence Assumption
 - Homogeneity Assumption
 - Gaussian Errors Assumption
- 4. $m{\beta}$ is the vector of primary parameters, and $m{ heta}_{a imes 1} = \mathbf{C}_{a imes p} \; m{eta}_{p imes 1}$ is

a vector of secondary parameters, defined by \mathbf{C} , the *contrast* matrix. Each row of \mathbf{C} defines a new scalar parameter in terms of the $\boldsymbol{\beta}$'s, e.g., $\beta_1 - \beta_2$. The general linear hypothesis is

$$H_0: \boldsymbol{\theta}_{a \times 1} = \boldsymbol{\theta}_0$$

$$H_A: \boldsymbol{\theta}_{a \times 1} \neq \boldsymbol{\theta}_0.$$

- 5. Estimability and Testability of a Parameter. If X is full rank, then $\widehat{\beta}$ exists (uniquely), β is estimable, and any (nonzero) C gives estimable θ . If C is full rank, β is testable.
- 6. Computation of Test Statistic and p-value. Let $\mathbf{M}_{a\times a}=\mathbf{C}(\mathbf{X}'\mathbf{X})^{-1}\mathbf{C}' \text{ and } SSH_{1\times 1}=(\widehat{\boldsymbol{\theta}}-\boldsymbol{\theta}_0)'\boldsymbol{M}^{-1}(\widehat{\boldsymbol{\theta}}-\boldsymbol{\theta}_0).$ The test-statistic is

$$F_{obs} = \frac{SSH/a}{SSE/(n-p)} = \frac{(\widehat{\boldsymbol{\theta}} - \boldsymbol{\theta}_0)' \boldsymbol{M}^{-1} (\widehat{\boldsymbol{\theta}} - \boldsymbol{\theta}_0)/a}{\widehat{\sigma}^2} = \frac{MSH}{MSE}$$

Topic 5: Some Distributional Results for the GLM

- If ${\bf X}$ is full rank, $\widehat{{m eta}} \sim \mathcal{N}_p({m eta}, \sigma^2({\bf X}'{\bf X})^{-1})$
- $\theta = \mathbf{C}_{a \times p} \boldsymbol{\beta}$, then $\widehat{\boldsymbol{\theta}} \sim N_a(\boldsymbol{\theta}, \sigma^2 \mathbf{C}(\mathbf{X}'\mathbf{X})^{-1} \mathbf{C}')$.
- Predicted Values: Conditional Means and Future Observations

$$-\widehat{\mathbf{y}} = \mathbf{X}\widehat{\boldsymbol{\beta}} = \left[\mathbf{X}(\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}'\right]\mathbf{y} = \mathbf{H}\mathbf{y},$$

$$-E(\widehat{\mathbf{y}}) = \mathbf{X}\boldsymbol{\beta},$$

$$-\operatorname{cov}(\widehat{\mathbf{y}}) = \sigma^2 \mathbf{X} (\mathbf{X}' \mathbf{X})^{-1} \mathbf{X}'.$$

- Definitions and Properties of Residuals
- Residual Variance $\widehat{\sigma}^2 = \frac{SSE}{n-p} = \frac{\widehat{\boldsymbol{\varepsilon}}'\widehat{\boldsymbol{\varepsilon}}}{n-p} = \frac{\mathbf{y}'(\mathbf{I} \mathbf{H})\mathbf{y}}{n-p}$

Topic 6: Multiple Regression: General Consideration

Basic Sum Squares:

$$USS(total) = USS(model) + SSE, \quad y'y = y'Hy + y'(I - H)y.$$

$$CSS(\text{total}) = CSS(\text{model}) + SSE$$

$$\mathbf{y}' \left[\mathbf{I} - \frac{1}{n} \mathbf{J}_n \mathbf{J}'_n \right] \mathbf{y} = \mathbf{y}' \left[\mathbf{H} - \frac{1}{n} \mathbf{J}_n \mathbf{J}'_n \right] \mathbf{y} + \mathbf{y}' (\mathbf{I} - \mathbf{H}) \mathbf{y}.$$

$$\begin{split} \bullet \ \ F_{obs} &= \frac{MS(\mathsf{hypothesis})}{MSE} = \frac{SSH/dfH}{SSE/dfE} \\ &= \frac{[SSE(\mathsf{reduced}) - SSE(\mathsf{full})]/[dfE(\mathsf{reduced}) - dfE(\mathsf{full})]}{SSE(\mathsf{full})/dfE(\mathsf{full})} \\ &= \frac{CSS(\mathsf{Regression})/(p-1)}{SSE(full)/(n-p)} \,. \end{split}$$

Reject the hypothesis if $F_{obs} \geq F_F^{-1}(1-\alpha, p-1, n-p) = f_{crit}$. The usual test of overall regression assumes model spans an intercept and excludes the intercept from the test.

- ANOVA table.
- Usual "Corrected" R^2 : $R_{\rm c}^2 = \frac{CSS({\rm Regression})}{CSS({\rm Regression}) + SSE({\rm full})} = \frac{CSS({\rm Regression})}{CSS({\rm total})}. \ R_c^2 \ {\rm estimates} \ \rho_{\rm c}^2,$ the population ratio of model to total variance, with $0 \le \rho_{\rm c}^2 \le 1$ and $0 \le R_{\rm c}^2 \le 1$.
- The corrected test for overall regression,

$$H_0: \beta_1 = \beta_2 = \ldots = \beta_{p-1} = 0$$

holds if and only if $H_0: \rho_{\mathsf{c}}^2 = 0$

Topic 7: Testing Hypotheses in Multiple Regression

- All tests compare two models: the full model and the reduced model (this is the basic idea of likelihood ratio tests, called the *likelihood ratio principle*).
- Overalltest: $F_{obs} = \frac{CSS(\beta_1,...,\beta_{p-1})/(p-1)}{SSE(\beta_0,...,\beta_{p-1})/(n-p)}$.
- Added-Last Test: the added-last test seeks to assess the usefulness of one predictor, above and beyond all others.
 Coefficient Estimates/t-test table, Type III table. The F statistic is

$$F_{obs} = \frac{\frac{SSE(\text{reduced}) - SSE(\text{full})}{dfE(\text{reduced}) - dfE(\text{full})}}{SSE(\text{full})/dfE(\text{full})} = \frac{(\widehat{\boldsymbol{\theta}} - \boldsymbol{\theta}_0)'\mathbf{M}^{-1}(\widehat{\boldsymbol{\theta}} - \boldsymbol{\theta}_0)/dfH}{\mathbf{y}'(\mathbf{I} - \mathbf{H})\mathbf{y}/dfE},$$

where
$$\mathbf{C} = \begin{bmatrix} 0 & \dots & 0 & 1 & 0 & \dots & 0 \end{bmatrix}_{1 \times p}$$

• Added-in-Order Test: the *added-in-order test* seeks to assess the

contribution of predictor j above and beyond all of the preceding j-1 predictors (without the j+1, j+2, etc.predictors in the model).

- Group Added-Last Tests
- Group Added-in-order Tests

Topic 8: Correlations

$$\rho = \operatorname{Corr}(X,Y) = \frac{\operatorname{Cov}(X,Y)}{\sqrt{\operatorname{Var}(X)\operatorname{Var}(Y)}}$$

$$R = \frac{\sum_{i=1}^{n} (X_i - \overline{X})(Y_i - \overline{Y})}{\sqrt{\left(\sum_{i=1}^{n} (X_i - \overline{X})^2\right)\left(\sum_{i=1}^{n} (Y_i - \overline{Y})^2\right)}}.$$

Partial correlations describe the strength of the linear relationship between two variables, Y and X, after controlling for the effects of other variables \mathbf{Z} .

Topic 9: GLM Assumption Diagnostics

- The First Step: Get to Know Your Data
- Homogeneity: violations seen in the pattern of residuals.
- Independence: assessed through logic of sampling scheme.
- Linearity: examine pattern of residuals.
- Existence: (finite sample...).
- Gaussian distribution: distributional assessment involves box plot of residuals, histogram of residuals, and test of Gaussian distribution of residuals. (The discrepancy between T and Gaussian random variables somewhat inflates the probability of rejecting the null...why?)
- Outliers: leverage, Influence: Cook's Distance

Topic 10: Computation Diagnostics

- Colinearity
- Eigenanalysis
- Condition Number and Condition Index: the *condition index* for the kth eigenvalue equals $\sqrt{\lambda_1/\lambda_k}$. The maximum condition index, called the *condition number*
- R_j^2 , Tolerance, and VIF $R_j^2 = R^2(X_j, \{X_1, \dots X_{j-1}, X_{j+1}, \dots X_{p-1}\})$

$$\mathsf{VIF}_j = \frac{1}{1 - R_j^2} = \frac{1}{\mathsf{tolerance}}.$$

- Leverage
- Cook's distance

Topic 11: Selecting the Best Model

- 1. Specify the maximum model under consideration.
- 2. Specify a criterion for model selection.
- 3. Specify a strategy for applying the criterion.
- 4. Conduct the analysis.

Topic 12: ANOVA

Coding schemes

$$\mathsf{Es}(\mathbf{X}_{ref}) = \begin{bmatrix} 1 & \mathbf{0} & \mathbf{0} \\ 1 & 1 & 0 \\ 1 & 0 & 1 \end{bmatrix}_{3\times3} \quad \mathsf{Es}(\mathbf{X}_{cell}) = \begin{bmatrix} 1 & 0 & 0 \\ 0 & 1 & 0 \\ 0 & 0 & 1 \end{bmatrix}_{3\times3}$$

$$\mathsf{Es}(\mathbf{X}_{anova}) = \begin{bmatrix} 1 & 1 & 0 & 0 \\ 1 & 0 & 1 & 0 \\ 1 & 0 & 0 & 1 \end{bmatrix}_{3\times4}$$

$$\mathsf{Es}(\mathbf{X}_{effect}) = \begin{bmatrix} 1 & -1 & -1 \\ 1 & 1 & 0 \\ 1 & 0 & 1 \end{bmatrix}$$

Step down test

Topic 13: Coding Schemes for Regression

$$oldsymbol{ullet} oldsymbol{eta} oldsymbol{(\mathsf{ANOVA})} oldsymbol{y} = egin{bmatrix} \mathbf{1} & \mathbf{1} & \mathbf{0} \ \mathbf{1} & \mathbf{0} & \mathbf{1} \ \mathbf{1} & \mathbf{0} & \mathbf{0} \end{bmatrix} egin{bmatrix} eta_0 \ eta_1 \ eta_2 \end{bmatrix} + oldsymbol{arepsilon}$$

ullet (Null) y=arepsilon

Topic 14: Logistic Regression

- Definition of odds, and odds ratio
- The general logistic regression model is given by

$$logit(p_i) = log\left(\frac{p_i}{1-p_i}\right)$$
$$= \beta_0 + \beta_1 x_{i1} + \beta_2 x_{i2} + \ldots + \beta_{p-1} x_{i,p-1}$$

with $y_i \sim \text{Bernoulli}(p_i)$, i = 1, ..., n, and the y's independent of each other.

- Interpretation of regression coefficients in terms of odds ratio.
- Model comparison by likelihood ratio test
- Logistic regression with categorical covariates and their interactions.
- Goodness of fit test

Topic 15: Mixed Effects Model

- When data are correlated and the independence assumption does not hold, mixed effects models are one way to adjust for the non-independence of observations
- Random effects may be introduced to account for the fact that observations within one subject (or more generally, within one cluster) may be more alike than observations from different clusters
- Forms of covariance matrices for clustered and repeated measurements
- Parameter interpretation of models for longitudinal data