

1 best known equations for HRmax = (220 - age) bpm

Experiment to ascertain validity of the (220-age) equation

Healthy volunteers aged 18-80 were reconsted and had HR max determined by exertise phys.

N = sample site

A: = age) of ith subject

Hi = HR max

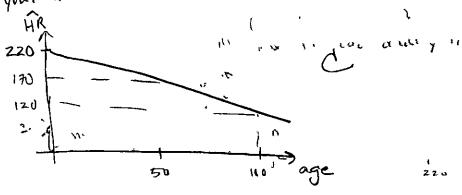
a) Write down a linear model relating HPR to age,

E(HR max) = witnest age & age in years

? predicted HR based on age

when age = 6 thop in HR max wheath addlyear

b) Assuming (220 - age) is correct, state values of the parameters in your livear model.



fit model to data => assuming model from A

$$\hat{H}R = \beta_1 + \beta_1 \text{ age}$$

$$X = \begin{pmatrix} 1 & \alpha_1 \\ \alpha_1 \end{pmatrix} \qquad (X \times X) = \begin{pmatrix} 1 & -1 \\ \alpha_1 & -1 \end{pmatrix} \begin{pmatrix} 1 & \alpha_1 \\ \alpha_1 & -1 \end{pmatrix}$$

$$\hat{B} = \begin{pmatrix} 207.0444 \\ -0.6605 \end{pmatrix} = \hat{\beta},$$

$$A) Test whether data are consistent
$$M / \beta_1 = -1 \qquad (\Delta \text{ implied } W / \text{ formula})$$

$$\hat{B} = \begin{pmatrix} -1 & 0 & 0 \\ 0 & 0 \end{pmatrix} = \begin{pmatrix} -1 \end{pmatrix}$$

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 $\frac{\beta_1 - 1}{\sqrt{6r\beta_1}} =$

est 50 yr old -> HR max = 160 different from expectation? Ho: Bo+Bilto)=160 expected = 207.044 - 0.6605(\$00) C= (1 20) = MANY = 167.414 Ho: HRS0 = 160 A= 167,44 $t = \frac{160-167414}{167414} \leq don't like this chief <math>c = (1 60) = 0$ C= (1 60) 0=160 F= 11.69 1,50 west, it is diff f) add interation g) new & Ai = 1880 +47 EA; = 101374 +472 ZA·H = 322 284 + (47.196) EH = 7040 + 196

then recall as in ()

for the others, as it and men explain

Mustina) age, educ. MS-a 2012 y= β0 + β, age + β, educ + ξ, N(0, 82) 1=20 $\int d_{0} = \frac{\sum X_{1} \cdot Y_{1} - N \cdot \overline{X} \cdot \overline{Y}}{\sum X_{1} \cdot \overline{Z} - N \cdot \overline{X}^{2}} = \frac{88328.1 - 20(\frac{127}{20})^{20}}{29199 - 20(\frac{127}{20})^{2}}$ $d_{0} = \overline{Y} - \beta_{1} \cdot \overline{X} = \frac{1}{29199 - 20(\frac{127}{20})^{2}}$ yi - yi $\vec{\xi} \sim N_{20}(0, I_{20}(\sigma^2))$ y~ N2(xβ, I, 62) b) $Ey = d_0 + d_1$ ask Zy; = 2301.6 $\hat{\chi} = (\chi'\chi)^{-1} (\chi'\chi)$ $= \begin{cases} (\chi'\chi)^{-1} (\chi'\chi) \\ \chi_{1} & \chi_{20} \end{cases} \begin{cases} (\chi_{1} & \chi_{20}) \\ \chi_{20} & \chi_{20} \end{cases}$ Exiy: = 883281 £x;2 = 29199 $= \begin{pmatrix} 20 & 2x_i \\ 2x_i & 2x_i^2 \end{pmatrix}^{-1} \begin{pmatrix} 2y_i \\ 2x_i y_i \end{pmatrix}$ $= \begin{pmatrix} 31, 108474 \\ 2, 218534 \end{pmatrix} =$ $= \begin{pmatrix} 20 & 757 \\ 757 & 29199 \end{pmatrix}^{-1} \begin{pmatrix} 2301.6 \\ 88328.1 \end{pmatrix}$

(2×1)7 1×2 = 2×2

c) consider two simple regs $\dot{m} = age$, $\dot{m} = educ$ 0.67 $0.71 = K^2$ $\dot{z}' = \dot{r}' \dot{h} \dot{h}$

NO=> individually, they explain that much. however, age and education may are not independent. Thus we cannot assume their effect together would be additive.

$$(x^{2}x)^{-1} = \begin{pmatrix} 3 \times 3 \end{pmatrix}$$

$$C = \begin{pmatrix} 1 & -1 & 0 \\ 1 & 0 & -1 \end{pmatrix} \quad \theta_0 = \begin{pmatrix} 0 \\ 0 \end{pmatrix} \qquad \qquad \beta_0 - \beta_1 = 0$$

$$\hat{\beta} = C\hat{\beta} =$$

$$don't have \hat{\beta}$$

(e)
$$\omega(r(\hat{\beta}, \hat{\beta}_2)) = \frac{\omega(\hat{\beta}, \hat{\beta}_2) - cell(1, 2)}{\sqrt{var\hat{\beta}, var\hat{\beta}_2}} ell(x, x)^{-1} e^{-2})$$

€ same approach diff. #'s

f) the x's dor't need to be normal.

The normality assumption of HILE Gains Conly applies to the unditural atstributions and error

Busting Insatment efficacy MS-2 2012 = Po + trt + age + gender + trtxage + trtxgender med or comsel [- n=100, 25 pur gender/trt combo (2 genders, 2 troatments) a) With the relation bother age and tot efficacy freach معهم Relation) TRT <u>Gender</u> Bo + Page med female male med Bo + Bage + Bgender female Po + Ptrt + Page + Prtxage counsel male Bo + Btr+ + Bage + Bgender + Btr+xage + Btr+xgender counsel b) Reg. Wells obtained. Interpret wells: (ME = A in extracted, for (non-neg or I unit meneral) in x 5 5016-1 WT = addl Expects of an niesse in age / being male for thou on total (addl great of BOTH) To test whether there are gender diff in tot em given they are the same ag, we should test whether $\hat{\Theta} = c\hat{\beta}$ H.: Part xgender=0 C= (0 0 0001) 0=0 F= (6-0) -2, -30-50 +4 or Be wta4 F~ 1, 100 - 6=94 c) added last. (Wold test of well=0) U non-sign terms are tr+xage, age (culfre), and trt. probably want to keep tot given to gender is sigh.

. .

...

d) Fill in ANCOVA table Type I/MSE Type 1 df MS Source - 45.72234 48.584 6,6181 age 0.5817 0,43 gender 113,8701 < 0.-- 107.163 trtxage 0,3580 = MS 6,55 0.3369) 1 trt xgender 12. 84505b 13,649 94 error from b) = 0.9410987 last added in order= total 99 mse*df added last F-stat, and df, for testing · E(y) = \beta_0 + \beta_1 trt + \beta_2 age + \beta_3 gender + \beta_4 trt x age + \beta_5 trt x g ' **Y**S` E(y) = po + p, trt + p2 age + p3 gender => Ho: By= Bs=0 $\hat{g} = c\hat{\beta} = \begin{pmatrix} -0.004731 \\ 1.529518 \end{pmatrix}$ or use SSH of By and Bs (ê-00) M-1 (ê-00)/af H = etechn w MSE etc. testing. MSE ~ 2,9 M-1= C(x'x)-1 C'

e)
$$m_m = mean med when $g=0$ age = 33
 $M_c = mean comset when $g=0$ age = 33$$$

$$H_0: \mu_m = \mu_c$$

$$\beta_0 + \beta_2(33) = \beta_0 + \beta_1 + \beta_2(33) + \beta_4(33)$$

$$\Rightarrow \beta_1 + \beta_4(33) = 0$$

C= (01 0 0 33 0) C/

$$\theta_0 = 0$$

2012 April Question 4) (125/m2) M5-2 -2012 current smolling status, VI.n cough that day a) Two separate logistic meg models are specified Provide the algebraic expressions for each as well as interpretations. is logit(congh) = mt + smokers as the many a cough (it non-smoker) Barole = 0.5415 (for every 1 unit increase in smoking freq, the logoddo of having a cough increases by 0.5415 (treating as linear logid (cough) = p. + paras. + pryclar Bo = log adds for coughins when x,=x2=0 (non-smoker) Bou : log OR for having a cough if an occas smoker compared W/a non-smoker regular amoher -b) Interpret the model w/ smoke as a covariate but without on interest term. (this model was n't run ... correct?)

hgit (x) = \beta_i (smoke)

increase in prob-of smoking assuming no curyling a mong non-smokers.

c) odv interval -> easin to understand (simple) is to disadv interval -> assumes the gimp from each level of smoking is =

adv cot -> allows the effect of each successive 'level' of conghing to differ ()

disadv cot -> might be harden to interpret /use.

ha, almost exactly the same as Craig

d) estimate P(conghing) for occassiving smokens i. $P(\text{congh}) = \frac{\exp(\beta_0 + \beta_1(1))}{1 + \exp(\beta_0 + \beta_1(0))} = 0.298615$

11. $f(wyh) = \exp(\beta_0 + \beta_1)$ $\frac{1}{1 + (\exp(\beta_0 + \beta_1))} = 0.287701$

e) Pille preforred model, with 2-31 sentences summarying