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## 3 Leveraging Stacked Classifiers for Multi-task Executive Function 4 in Schizophrenia Yields Diagnostic and Prognostic Insights

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## *Abstract*

41        Cognitive impairment is a central characteristic of schizophrenia. Executive functioning  
42 (EF) impairments are often seen in mental disorders, particularly schizophrenia, where they  
43 relate to adverse outcomes. As a heterogeneous construct, how specifically each dimension of  
44 EF to characterize the diagnostic and prognostic aspects of schizophrenia remains opaque.  
45 We used classification models with a stacking approach on systematically measured EFs to  
46 discriminate 195 patients with schizophrenia from healthy individuals. Baseline EF  
47 measurements were moreover employed to predict symptomatically remitted or non-remitting  
48 prognostic subgroups. EF feature importance was determined at the group-level and the  
49 ensuing individual importance scores were associated with four symptom dimensions. EF  
50 assessments of inhibitory control (interference and response inhibitions), followed by  
51 working memory, evidently predicted schizophrenia diagnosis (area under the curve  
52 [AUC]=0.87) and remission status (AUC=0.81). The models highlighted the importance of  
53 interference inhibition or working memory updating in accurately identifying individuals  
54 with schizophrenia or those in remission. These identified patients had high-level negative  
55 symptoms at baseline and those who remitted showed milder cognitive symptoms at  
56 follow-up, without differences in baseline EF or symptom severity compared to non-remitting  
57 patients. Our work indicates that impairments in specific EF dimensions in schizophrenia are  
58 differentially linked to individual symptom-load and prognostic outcomes. Thus, assessments  
59 and models based on EF may be a promising tool that can aid in the clinical evaluation of this  
60 disorder.

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## 76 Introduction

77 Schizophrenia is a serious mental health condition that can severely impair an  
78 individual's functioning and quality of life. Individuals with schizophrenia present a wide  
79 range of symptoms of varying severity. These symptoms feature hallucinatory and delusional  
80 experiences termed “positive symptoms” as well as negative symptoms which manifest as  
81 atypical emotional and social behaviors. Distinct from negative symptoms, though related,  
82 cognitive symptoms involve the impairment of mental functions related to memory, attention,  
83 and executive tasks. This greatly affects the ability to live independently given the difficult to  
84 treat these symptoms using the currently available antipsychotic medications<sup>1</sup>. The cognitive  
85 symptoms in schizophrenia are a core aspect of psychopathology; they are considered a trait  
86 marker that emerges in the prodromal phase and persists throughout the illness<sup>2</sup>, unlike those  
87 manifested in affective psychotic disorders or drug-induced psychosis where cognitive  
88 deficits are epiphenomenal. However, cognitive performance in patients with schizophrenia is  
89 heterogeneous, and can vary from virtually unaffected to severely impaired<sup>3,4</sup>. While there is  
90 mixed evidence regarding the impairments in some cognitive domains, executive dysfunction  
91 is pervasively abnormal in schizophrenia. Previous studies have consistently indicated that  
92 mild-to-severe deficits in processes are related with executive functions (EFs)<sup>5</sup>.

93 EF represents a series of higher-order cognitive processes that involve impulse control  
94 and behavior orchestration<sup>6</sup>. Deficits in this domain can hinder goal-directed activity and  
95 contribute to aggression, violence, and poor compliance to medication in patients with  
96 schizophrenia, leading to worse clinical outcomes<sup>7</sup>. In the field of psychiatry, prognostic  
97 prediction remains a significant challenge in research and clinical practice, though it is crucial  
98 for early assessment and intervention. Previous studies employing baseline neuroimaging,  
99 genetic, or clinical data only approached chance-level accuracy in most cases<sup>8,9</sup>. This  
100 emphasized the lack of reliable markers for tracking the disease trajectory<sup>10,11</sup>. EF deficits  
101 emerge in the early stages (e.g., ultra-high risk, first episode) of schizophrenia<sup>12,13</sup> and have  
102 been related to disease progression, symptom severity, and recovery of social and  
103 occupational skills<sup>14</sup>. Hence, they might be potential markers for tracking the clinical courses  
104 and prognostic statuses of schizophrenia.

105 A broad range of EF impairments has been associated with this disorder, including  
106 increased difficulties in inhibiting automatic responses and switching to new ones, reduced  
107 cognitive flexibility, and disturbances in the maintenance and updating of goal-related or  
108 rule-based information in working memory. These actually align well with the three-factor  
109 model of EF proposed by Miyake et al.<sup>15</sup>, which characterizes 1) interference inhibition and  
110 response inhibition; 2) cognitive flexibility and switching; and (3) working memory updating  
111 and maintenance. This three-dimensional representation of EF functions robustly captures  
112 individual variation in EF subcomponents across a broad spectrum of age groups and clinical  
113 cohorts, including patients with psychiatric disorders<sup>16,17</sup>. These dimensions of EF differ in  
114 concepts and neurobiological substrates, highlighting the need to consider and assess these  
115 dimensions, while studying their correlations with diagnostic and prognostic aspects of  
116 schizophrenia. Such finer characterization of EF functions may assist in understanding  
117 different psychopathological processes in schizophrenia (e.g., disorganization symptoms),  
118 which manifest as difficulties in the goal-directed sequencing of thoughts and behaviors<sup>18,19</sup>.  
119 These symptoms are linked to the cognitive dimension in our recently introduced  
120 four-dimensional representation (positive, negative, cognitive, and affective) of schizophrenia  
121 psychopathology, as generalizable across populations and clinical settings<sup>20</sup>. Disturbances in  
122 EF have likewise been implicated in negative and positive symptoms. Firstly, failure in  
123 effectively monitoring volitional behaviors and inhibiting false inference in predictive  
124 processing would have consequences on positive symptoms (e.g., hallucinations and  
125 delusions)<sup>21,22</sup>. Secondly, cognitive rigidity hampers adjustment of thoughts and actions for  
126 environmental volatility<sup>23</sup>. Thirdly, impaired working memory updating and maintenance has  
127 been linked to poor abstract thinking<sup>24</sup>, which can increase the risk of developing negative  
128 symptoms (e.g., apathy and diminished expressive behavior) commonly observed in  
129 schizophrenia<sup>23,25</sup>. However, some dimensions of EF may be more severely affected than  
130 others in schizophrenia<sup>26</sup>. Currently, it remains unclearly the abnormalities in which EF  
131 dimensions characterize schizophrenia and would play a role in providing prognostic  
132 information.

133 Assessment of the underlying processes of the dimensions of EF is a challenge.

134 Currently available tools (e.g., Cambridge Neuropsychological Test Automated Battery<sup>27</sup> and  
135 the National Institutes of Health Toolbox<sup>28</sup>) do not feature a comprehensive assessment to  
136 cover the various dimensions of EF. Furthermore, assessments included in available  
137 neuropsychological batteries (e.g., Delis–Kaplan Executive Function System) are paper tests  
138 rather than experiments. Consequently, they do not offer a trial-by-trial based dynamic  
139 quantification, e.g., the reaction time in a sequential task. Trial-by-trial responses help detect  
140 subtle cognitive impairments in schizophrenia, including EF<sup>29,30</sup>. In addition, the cognitive  
141 symptom items routinely used in clinical practice, including EF rating such as in the Positive  
142 and Negative Syndrome Scale (PANSS), are retrospective based on information collected  
143 from patient interviews or contributions by relatives. Comparatively, task paradigms drawn  
144 from the cognitive psychology literature provide objective trial-by-trial tests that facilitate  
145 measuring particular cognitive functions with likely improved sensitivity and specificity<sup>16,31</sup>.  
146 Such a tailored assessment strategy would be ideal for investigating the diagnostic and  
147 prognostic value of EF dimensions in schizophrenia by establishing classification models.  
148 Previous psychiatric machine learning studies mainly considered single algorithms, such as  
149 support vector machine (SVM) or random forest (RF), comparing their respective  
150 accuracies<sup>32</sup>. Methods such as stacking, a mainstay multi-view learning approach, may be  
151 another strategy for improving model performance<sup>33</sup>.

152 In this study, we systematically applied six well-established behavioral paradigms to  
153 assess individual baseline functions along the three EF dimensions (i.e., inhibitory control,  
154 working memory maintenance and updating, and cognitive flexibility) to determine their  
155 consistency in characterizing patients with schizophrenia and their prognostic statuses at  
156 follow-up. This is tested via establishing diagnostic and prognostic classification models  
157 using machine learning methods: SVM, RF, Adaptive Boosting (AdaBoost), and their stacked  
158 model with a stringent nested cross-validation (CV) and independent testing. The importance  
159 of EF feature contributions to classification models was determined via the SHapley Additive  
160 exPlanations (SHAP) approach, which facilitates the identification of important features at  
161 both the group- and individual-levels. This approach enables a link between feature  
162 importance and individual psychopathology (Fig. 1).

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163 **Results**

164 **Participant identification, screening, and follow-up**

165 In this hospital-based study, we initially identified 270 individuals with schizophrenia  
166 (International Classification of Diseases, tenth edition [ICD-10], additional screening with the  
167 Structured Clinical Interview for DSM-IV axis I Disorders) aged 18–65 years, from a total of  
168 580 inpatients in the psychiatric department of the Third People's Hospital of Lanzhou  
169 (Lanzhou, China) (Fig. 2). These 270 patients with schizophrenia were clinically stable  
170 (change in total PANSS score<20% with a particular type of antipsychotics at a maintenance  
171 dosage within the last 6 weeks; details in Supplementary Table S1). Furthermore, 75 of the  
172 270 patients with schizophrenia were excluded due to illiteracy (N=35) or refusal to  
173 participate (N=40). Finally, 195 patients were included. The study was approved by the ethics  
174 committees of Northwest Normal University and the Third People's Hospital of Lanzhou  
175 (Lanzhou, China). They underwent evaluations that included the electronic medical records,  
176 the Positive and Negative Syndrome Scale (PANSS), six EF behavioral tasks, and the fluid  
177 intelligence (Raven's Progressive Matrices). To enable comparative analyses, 169  
178 demographically matched healthy control participants who were free of a history of mental  
179 illness or brain injury were recruited and underwent the same assessments, except the  
180 PANSS.

181 Among the 195 patients with schizophrenia, 86 participants completed follow-up  
182 assessments (PANSS and EF tests) in 4–6 weeks after the initial evaluation. The loss of 109  
183 participants at follow-up was primarily attributed to hospital discharge (N=98) or concerns  
184 regarding the potential impact of assessment results on their discharge timing (N=11)  
185 (Supplementary Table S2). Fifty-eight participants at follow-up identified as being in  
186 remission; who had scores  $\leq 3$  on key PANSS items (P1, P2, P3, N1, N4, N6, G5, G9; the  
187 Remission in Schizophrenia Working Group [RSWG] criteria<sup>34</sup>). The other 28 participants  
188 were not in remission (Supplementary Table S2). Alternatively, defining the remission status  
189 by a reduction in PANSS total score at follow-up assessment relatively to baseline showed  
190 that<sup>35–37</sup>: 1) a 25% reduction (68 remitted vs. 18 non-remitted); 2) a 35% reduction (48  
191 remitted vs. 38 non-remitted); 3) a 50% reduction (16 remitted vs. 70 non-remitted)

192 (Supplementary Table S2). In addition, prognostic statuses may also be reflective in a change  
193 of subtype membership from baseline to the end of follow-up, particularly surrounds the  
194 negative symptom subtype. This is because many patients tend to experience increased  
195 negative symptoms and diminished positive symptoms due to standard antipsychotic  
196 treatments and related factors at follow-up, while others feature primary (stable) negative  
197 symptoms<sup>38</sup> which relate to poorer clinical outcomes<sup>39</sup>. By using a subtyping system  
198 (<http://webtools.inm7.de/sczDCTS/>) from our previous work, patients with schizophrenia were  
199 assigned based on their symptom patterns as predominantly negative, positive, or ambiguous  
200 cases<sup>20</sup>. Of the 86 patients, 64 non-negative subtype patients (50 ambiguous and 14 positive)  
201 at baseline had a negative subtype assignment at follow-up assessment. Thirteen patients with  
202 schizophrenia maintained their negative subtype membership over time.

203 **EF dimensions**

204 ***EF dimensions are consistently and differentially affected in schizophrenia***

205 Six behavioral paradigms (e.g., Zhao et al., 2023)<sup>40</sup>, were administered to assess the  
206 three EF dimensions (inhibitory control, working memory maintenance and updating, and  
207 cognitive flexibility; Fig. 3A), based on 14 measurements: 1) The inhibitory control  
208 dimension: four measurements for the interference control function based on the *Stroop* task,  
209 three measurements for the response inhibition function based on the *Go/No-Go* task; 2) The  
210 working memory dimension: two measurements for working memory updating function  
211 based on the *running memory* task, three measurements for numeric working memory  
212 maintenance capacity based on the *Corsi block* test and the *digit span backward* task; 3) The  
213 cognitive flexibility dimension: two measurements (switch cost and mixing cost) in the  
214 *number-letter switching* task. Besides a conceptual formulation of the 14 task measurements  
215 according to the three-dimensional representation of EF, we supplemented five composite  
216 scores. This included an *Inhibition* composite score, an abbreviated version for representing  
217 general EF functions, and three cross-dimensional composite scores (*Inhibition/Switching*,  
218 *Inhibition/Working memory updating* and *Switching/Working memory updating*). Among  
219 these measurements, all reaction times, switching cost, and three composite scores (abbreviated

220 general EFs, *Inhibition/Switching*, *Inhibition/Updating*) are the higher the worse EF  
221 performance, while the remaining the higher the better.

222 To evaluate whether schizophrenia differentially affected the three-dimensional EF  
223 measurements and the five EF composite scores, we performed a mixed-model analysis of  
224 covariance (ANCOVA). This has revealed a significant two-way interaction between group  
225 (schizophrenia vs. HC) and EF measurements ( $p<0.001$ ). Follow-up one-way ANCOVAs  
226 revealed that, except for 2 measurements which assess working memory maintenance based  
227 on the *Corsi block* test, the accuracy in No-Go trails, the difference of reaction time between  
228 the congruent and the incongruent condition trials (i.e., the interference effect) in the Stroop  
229 task, the mixing cost in the switching task and the two composite scores of  
230 *Inhibition/Working memory updating* and *Inhibition*, other measurements and composite  
231 scores were all significantly different between patients with schizophrenia and healthy controls  
232 (all  $p<0.05$ , false discovery rate [FDR] corrected) (Table 2; Fig. 3B). These differed EF  
233 performance cover all of the three EF dimensions with the reaction times for Go trails  
234 (response inhibition) in the *Go/No-Go* task presenting the largest effect size ( $\eta^2= 0.160$ ) (Fig.  
235 3C, Table 2).

236 ***Remitted patients show improved interference inhibition at follow-up, without difference in***  
237 ***any EF dimension at baseline, compared with non-remitted patients***

238 Further analyses were conducted to examine differences in baseline EF dimensions  
239 between the remission and non-remission groups (RSWG criteria); there were no significant  
240 between-group differences observed on any EF dimension (Supplementary Table S3). When  
241 comparing respective changes in these dimensions from baseline to follow-up  
242 (Supplementary Table S5), significant differences were only observed in the remission group.  
243 Specifically, on the Stroop task, the remission group demonstrated shorter reaction times  
244 under neutral, congruent, and incongruent conditions compared with baseline (i.e., better  
245 interference inhibition ability; all  $p<0.05$ ).

246 Regarding clinical outcomes defined by subtype transmission, there were no significant  
247 differences in baseline EF measurements among the three subgroups (i.e., positive, negative,  
248 and ambiguous) of patients with schizophrenia (all  $p>0.05$ , FDR corrected). In patients with

249 more prominent secondary negative symptoms (i.e., baseline non-negative subtype with  
250 transition to a negative subtype), the Stroop task revealed significantly reduced reaction times  
251 at follow-up compared with baseline for incongruent stimuli ( $p=0.022$ ), congruent stimuli  
252 ( $p=0.007$ ), and neutral stimuli ( $p=0.037$ ). Additionally, the mixing cost in the switching task  
253 was significantly lower at follow-up ( $p=0.048$ ). However, patients of the stable negative  
254 subtype during follow-up did not show significant differences between baseline and  
255 follow-up across all EF measurements. There were no significant differences observed in  
256 baseline EF measurements between the secondary negative group and the primary negative  
257 group (all  $p> 0.05$ ).

258 **Psychopathology**

259 ***Psychopathology dimensions specifically correlated with different EF measurements***

260 Previous studies showed that EF, and specifically their dimensional measurements,  
261 could be associated with different aspects of symptomatology<sup>41</sup>. We probed the potential  
262 association of multifaceted aspects of psychopathology with different EF functions at the  
263 baseline assessment using Pearson correlation analysis (Supplementary Fig. S4). Using the  
264 four-dimensional structure of the PANSS<sup>20</sup>, we observed significant reductions in severity  
265 across negative, positive, cognitive, and affective symptoms in patients with schizophrenia at  
266 follow-up ( $p<0.001$ , Supplementary Table S6). Among the 86 follow-up patients, the ability  
267 to inhibit conflict, as reflected by reaction times to incongruent ( $r=0.212$ ,  $p=0.050$ ) and  
268 congruent stimuli ( $r=0.246$ ,  $p=0.022$ ) in the *Stroop* task, were significant correlated with  
269 baseline positive symptom scores. Patient capacity to maintain and shift mental sets,  
270 quantified by switch cost ( $r=0.298$ ;  $p=0.005$ ) and mixing cost ( $r=0.265$ ;  $p=0.014$ ) in the  
271 *number-letter switching* task, was significantly correlated with baseline cognitive and  
272 positive symptom scores, respectively, within the follow-up subset but not in the overall  
273 patient sample. For comparison, correlation analyses were repeated using PANSS  
274 three-original subscale scores; these yielded similar results, except for correlations with  
275 *Stroop* metrics (Supplementary Fig. S5).

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276 ***EF dimensions show consistent and distinct impairments in schizophrenia and offer***  
277 ***prognostic insights via machine learning classification models***

278       Using a multivariate approach, classification models can be employed to identify feature  
279 variables (e.g., EF measurements) which can reliably differentiate target cases (i.e., patients  
280 with schizophrenia) from reference cases (i.e., HC). Additionally, they can provide insights  
281 into future categories (e.g., prognostic status) based on baseline assessments. Three methods  
282 (i.e., RF, SVM, and AdaBoost), and their stacked assembler, were used to construct  
283 classification models. The original data was repeatedly split into discovery and test sets, with  
284 each discovery set nested for hyperparameters tuning and model validation as a CV design.  
285 Next, the ensuing best model was applied to the test set from each repeat to obtain  
286 out-of-sample performance. The whole procedure was repeated for 100 tests, resulting in 100  
287 hold-out, test sets. This approach has been demonstrated to effectively gauge generalization  
288 while balancing practical acquisitions of clinical sample data<sup>42-44</sup>.

289 **1) Diagnostic classification**

290       For our classification experiments, two feature sets were used: 1) 19 EF assessments,  
291 which reflect three EF dimensions measured by six behavioral paradigms (Table 2); and 2) 32  
292 features, which added 13 sociodemographic variables to the 19 EF variables in feature set 1  
293 (Table 2). Performance metrics were assessed on the 100 test sets (Supplementary Table S7).  
294 For the feature set relying on only EF assessments (i.e., feature set 1), we aimed for a model  
295 to classify new patients, regardless of sociodemographics. The highest out-of-sample  
296 classification was achieved by the stacking model (area under the curve [AUC]=0.87). The  
297 feature set 2, which also included sociodemographic variables, was aimed at incorporating  
298 information from routine clinical interviews that indicates disease susceptibility<sup>45</sup>. With the  
299 addition of sociodemographic variables, improved model performance was observed  
300 (stacking model AUC=0.91). In addition, repeating the entire CV process on the EF  
301 assessments after controlling for the effects of sociodemographic variables largely maintained  
302 the performance (stacking model AUC=0.80) (Fig. 4A).

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303     **2) Prognostic classification**

304       For prognostic classification, we applied the same models (RF, SVM, AdaBoost, and  
305       their stacking, as described in the diagnostic classification above) on two feature sets, to  
306       assess their performance in discriminating remission status: (1) 19 baseline EF variables; (2)  
307       32 features (the 19 baseline EF variables plus 13 sociodemographic variables). The stacking  
308       model achieved the highest classification accuracy based on the EF assessments-only feature  
309       set ( $AUC=0.81$ ), and the performance was identical to that observed when the EF plus  
310       sociodemographic features set was used.

311       Additionally, to test the influence of demographic variables and medication on our  
312       models, we conducted several control analyses (Fig. 5A). By regressing out the effects of  
313       sociodemographic variables on EF assessments, we found that the model performance was  
314       decreased to  $AUC=0.65$ , indicating a poorer setup (Fig. 5A). However, there was no  
315       significant difference in any of the sociodemographic variables adjusted in our classification  
316       models between the remitted and non-remitted patient subgroups (all  $p>0.05$ ) (Supplementary  
317       Table S3). Additionally, controlling for medication effects using an olanzapine (OZP)  
318       equivalent dosage (which did not differ significantly between remitted and non-remitted  
319       patients;  $p=0.15$ ) diminished the prognostic classification accuracy to an  $AUC$  of 0.72. By  
320       investigating only those patients ( $N=46$ ) treated with a commonly effective OZP-equivalent  
321       dosage of 10–20 mg/day in clinical practice produced a similar prognostic classification  
322       performance ( $AUC=0.82$ ) to that recorded in the group of followed up patients ( $N=86$ ).

323       We supplemented classifications to discriminate prognostic subgroups defined by the  
324       reduction in re-assessed total PANSS score. The results showed a poorer discriminative  
325       power, with AUCs of 0.74, 0.58, and 0.73 (Fig. 5A; Supplementary Table S8) for a 25%,  
326       35%, and 50% reduction in the PANSS score, respectively. Alternatively, we established a  
327       classification model for the clinical outcomes of patients based on subtype-membership  
328       transition to distinguish between 1) baseline non-negative subtype with a transition to  
329       negative subtype (secondary) and 2) stable negative subtype during follow-up (primary). A  
330       promising classification performance was revealed ( $AUC=0.81$ ).

331       Moreover, two sensitive analyses were performed to take the attribution condition in our

332 study into consideration. The demographic variables, symptoms, OZP-equivalent dosage, and  
333 EF measurements did not differ significantly between patients who were followed up and  
334 those who were not (all  $p>0.05$ , Supplementary Table S3). Furthermore, by treating the  
335 attribution patients as best cases (i.e., all remitted), the classification accuracy decreased  
336 slightly to an AUC of 0.75, while treating the attribution patients as worst cases (none  
337 remitted) yielded a further decrease (AUC=0.68).

338 ***Feature importance and association with individual psychopathology***

339 SHAP analysis was performed on the best-performing classifiers trained on the 19 EF  
340 dimension assessments (for the feature sets including sociodemographic variables please refer  
341 to Supplementary Table S9)<sup>46</sup>. The goal was to determine the directional contribution of EF  
342 dimensions for classification informed by decision path (i.e., better or worse performance in  
343 an EF dimension drives the model to assign a schizophrenia or HC label). Besides  
344 group-level determination of important dimensions, the Shapley value for each EF feature in  
345 the best-performing classifiers was calculated for each individual<sup>47,48</sup>, facilitating a link to the  
346 expression level of patients along several psychopathological dimensions.

347 **1) Group-level feature importance and decision path**

348 Absolute Shapley values derived from the best-performing classifier (highest AUC)  
349 identified in the 100 test sets were used to rank each EF dimension to indicate its importance  
350 in discriminating patients with schizophrenia, and those remitted at follow-up (Figs. 4B, 5B).  
351 Including only the 19 EF dimensions scores as the feature set in both diagnostic and  
352 prognostic classifications, the inhibition control dimension—comprised of response and  
353 interference inhibitions—ranked highest in both classifying schizophrenia group participants  
354 and their follow-up remission status. Within the inhibition control dimension, important  
355 features were from the *Go/No-Go task*, which assesses response inhibition (Go trial accuracy  
356 and reaction time), and the Stroop task, which assesses interference inhibition (reaction time  
357 for neutral stimuli in the diagnostic model, and reaction time for incongruent stimuli in the  
358 prognostic model). Adding sociodemographic variables to the diagnostic and prognostic  
359 classifiers generally replicated inhibition control as the strongest contributing dimension

360 (Supplementary Figs S6A, S7A).

361 Next, group-level decision path analysis of each EF dimension identified that poor  
362 performance on any of these drove the model to correctly classify individuals as patients with  
363 schizophrenia (Fig. 4C). However, worse performance on the inhibition control dimension  
364 increased the likelihood of non-remission status classification (Fig. 5C). These results were  
365 replicated in additional models in which sociodemographic variables were included in  
366 diagnostic and prognostic classifiers (Supplementary Figs. S6B, S7B).

## 367 **2) Individual-level decision path and association with psychopathology**

368 Decision path analysis was likewise conducted at the individual level, to determine the  
369 relative performance of EF dimensions (as identified in group averages) for correctly  
370 assigning individuals. Among correctly identified participants, patients with schizophrenia  
371 scored below the averages of both HC and schizophrenia groups on at least one EF dimension  
372 (Fig. 4D). As expected for the prognostic classification model, remitted status for most  
373 participants in the schizophrenia group was correctly assigned based on higher baseline  
374 inhibitory control dimension (including both interference control and response inhibition)  
375 performance compared with the averages of both remitted and non-remitting participants.  
376 Specifically, remitted participants showed shorter reaction times to the incongruent condition  
377 in the *Stroop* task, and higher accuracy in the response to the Go trials during the *Go/No-Go*  
378 task. However, a few patients presented both worse performance in inhibitory control and  
379 higher baseline abilities in other dimensions such as working memory updating or shifting  
380 (Fig. 5D; Supplementary Fig. S9). These findings were replicated when sociodemographic  
381 variables were included in diagnostic classification models (Supplementary Materials).

382 Pearson correlation analysis was further performed on Shapley values for each EF  
383 feature and scores on the four symptom dimensions, across the overall schizophrenia group  
384 ( $N=195$ ) and follow-up subset ( $N=86$ ). Results showed that individual Shapley values of the  
385 interference inhibition function, as assessed by a difference in reaction time between the  
386 congruent and the incongruent trials in the *Stroop* task, that promoted a correct assignment of  
387 cases versus HC were significantly associated with the negative symptoms ( $r=0.439, p=0.042$ ,  
388 FDR corrected) for individuals with schizophrenia. The inference inhibition function, though

389 assessed by a different behavioral metric, was similarly identified as a factor of top  
390 importance at the group-level for the diagnostic model described above (Fig. 6A). After  
391 including the 13 sociodemographic variables, there was no significant correlation found.

392 For prognostic classification, the importance of working memory updating (assessed in  
393 the *running memory* task;  $r=0.618$ ,  $p=0.023$ , FDR corrected) and maintenance (assessed in  
394 the *Corsi block* test;  $r=0.597$ ,  $p= 0.031$ , FDR corrected) in the model that accurately assigned  
395 remitted patients was correlated with low-level cognitive symptoms at follow-up.  
396 Additionally, working memory updating function contributing to the accurate assignment of  
397 remitted patients was associated with more severe re-assessed negative symptoms ( $r=0.596$ ,  
398  $p=0.031$ , FDR correction). Following the inclusion of sociodemographic variables in the  
399 model, the significantly associated EF variables were changed (Fig. 6C).

400 Using PANSS three-subscale scores in correlation analyses did not reveal significant  
401 correlation with the importance scores of any EF features of the dimensions identified in the  
402 diagnostic models. Significant correlation patterns among the importance scores of EF  
403 dimension features in prognostic models were a subset of those reported (Supplementary Fig.  
404 S10) when using the four dimensions of psychopathology, as described above.

405

## 406 Discussion

407 This study is the first to investigate the classification power of comprehensively three  
408 EF dimensions (i.e., inhibitory control, working memory maintenance and updating,  
409 cognitive flexibility) for discriminating both patients with schizophrenia from HC and  
410 determine their remission status at follow-up. Importantly, our SHAP approach parsed the  
411 relative importance of each feature in these classification tasks, at the group and individual  
412 level. Collectively, we found that EF dimensions consistently and differentially characterize  
413 schizophrenia and are informative regarding the prognostic status, though certain dimensions  
414 are more closely linked to the disease trait and related psychopathology.

415 The four primary findings are as follows. Firstly, EF assessments could be used to both  
416 classify patients with schizophrenia ( $AUC=0.87$ ) and to identify those with remitted status

417 (AUC=0.81). Importantly, there is no significant difference in baseline EF and symptom  
418 severity between the two prognostic subgroups. Secondly, inhibition control was the most  
419 strongly contributing dimension to patient classification of both schizophrenia and remission  
420 outcome. Thirdly, at the individual level, correctly identified patients presented  
421 below-average performance on at least one EF dimension. However, except for a few patients  
422 with correctly assigned remission status who had worse performance in inhibitory control,  
423 others featured higher baseline performance in this dimension compared with the averages of  
424 both remitted and non-remitting patients. Finally, the EF dimension interference inhibition,  
425 which is important in promoting the correct classification of patients with schizophrenia by  
426 the model, was significantly associated with patient negative symptoms. The model  
427 importance of working memory for accurate remission assignment covaried with low-level  
428 follow-up cognitive symptoms.

429 The paradigms measuring EF and its dimensions are readily available through software  
430 platforms, such as *E-Prime* and Matlab-based Psychtoolbox. Thus, they can be implemented  
431 in clinical practice. The present findings may encourage the use of a feasible, low-cost, and  
432 effective approach to schizophrenia diagnosis and psychopathology evaluation.

### 433 **Schizophrenia, and its remission status, are classified by EF dimensions**

434 Previous machine learning studies employing neuropsychological test batteries (e.g., the  
435 Cambridge neuropsychological test automated battery, the Wechsler adult intelligence scale,  
436 and the brief assessment of cognition in schizophrenia) to differentiate patients with  
437 schizophrenia from HC have yielded accuracy rates <70%<sup>8,49</sup>. Neuroimaging-derived  
438 assessments are an alternative and broadly attempted approach<sup>50</sup>. While showing promise for  
439 improving classification performance<sup>51,52</sup>, this strategy is associated with other challenges<sup>53</sup>  
440 including heterogeneous data acquisition, high dimensionality due to large numbers of voxels  
441 or measures, and limited applicability in low-income countries and regions, as illustrated by  
442 our recent meta-analysis of global psychiatric neuroimaging data<sup>54</sup>. Another direction is  
443 systematic modeling based on readily attainable data with improved objectivity and reliability,  
444 namely behavioral EF tasks, for improved clinical translation<sup>10,55</sup>. By carefully assessing  
445 three EF dimensions via six tasks, and thus 19 variables, our stacking model achieved

446 reasonably high diagnostic accuracy. Furthermore, the inclusion of sociodemographic  
447 characteristics increased the AUC to 0.91. This is broadly consistent with previous findings  
448 of an association between sociodemographics and disease susceptibility, clinical course, and  
449 symptom expression in schizophrenia<sup>45</sup>. In clinical practice, early and accurate prediction of  
450 remission outcomes holds significant implications for effective treatments. Nevertheless, this  
451 remains difficult, with recent models based on baseline neuroimaging, genetics, and clinical  
452 factors often producing accuracy rates marginally better than the chance level (i.e., 50%).  
453 Comparatively, our prognosis classification model, incorporating only baseline EF  
454 assessments, demonstrated improved performance in denoting remission status among  
455 patients with schizophrenia at 4–6-week follow-up (AUC=0.81). We moreover tested the  
456 classification models by two methods for defining remission, i.e., based on a reduction in  
457 re-assessed PANSS scores calculated with specific items or all items, to assess the robustness  
458 of our findings to the definition of treatment response<sup>36</sup>. The discriminative power of the  
459 classification model was higher when specific items, compared with all items, of the PANSS  
460 were involved in defining prognostic statuses. This evidence implied specificities in mapping  
461 EF dimensions and symptom recovery in patients with schizophrenia.

462 **Classification-important EF dimensions and associations with individual  
463 psychopathology**

464 Leveraging the SHAP framework, we conducted feature importance analysis on our  
465 classification models to identify the relative contributions of each EF dimension variable.

466 **Diagnostic classification**

467 At the group level, we found that the inhibitory control dimension (i.e., interference  
468 inhibition and response inhibition) ranked highest in importance for classifying participants  
469 in the schizophrenia group. This is consistent with previous works showing abnormal  
470 alterations in the temporal and spatial characteristics of inhibition-related brain responses and  
471 behaviors in schizophrenia<sup>56,57</sup>. Moreover, our decision path plot pertaining to all correctly  
472 classified individuals indicated that those who performed poorly on any EF dimension tended  
473 to be classified in the schizophrenia group. This emphasizes the general EF impairments

474 among patients with schizophrenia, consistent with a previous meta-analysis showing that EF  
475 deficits within this patient population cover broad dimensions<sup>26</sup>. Considering individual  
476 variation in this context through decision path analysis for each participant, we noted that a  
477 few accurately identified patients with schizophrenia performed slightly above average on  
478 either of the three EF dimensions (e.g., inhibition, updating, or shifting). This aligns well with  
479 data showing that cognitive performance, including EF functions, in schizophrenia can vary  
480 from mild deficiency<sup>58,59</sup> to severely impaired<sup>60,61</sup>, and connects the neuropsychological and  
481 neurobiological heterogeneity systematically observed among these patients<sup>62,63</sup>.

482 By extending the SHAP framework to individual-level analyses, we further linked the  
483 importance of each EF feature in diagnostic classifier—quantified by Shapley values—to the  
484 landscape of individual psychopathology. Notably, we found Shapley values from the  
485 dimension identified as important at the group-level—interference inhibition—to be valuable  
486 within the model correctly classifying patients with schizophrenia, and significantly  
487 associated with a patient’s negative symptom expression. Consistent with our finding, earlier  
488 research showed that individuals with more severe negative symptoms tend to have  
489 diminished inhibitory control, as measured by the *Stroop* task<sup>64</sup>. This deficiency, particularly  
490 within interference inhibition, might play a role in the manifestation of negative symptoms,  
491 which may be understood through the lens of target-speech recognition deficit in  
492 schizophrenia<sup>64</sup>. Specifically, the interference inhibitory function assists individuals in  
493 accurately segregating target speech from noisy background environments in which multiple  
494 speakers are talking simultaneously<sup>65</sup>. An impairment in this function would thus be  
495 connected to the disorganized speech information processes in schizophrenia, including the  
496 inability to either inhibit unrelated speech signals or capture desired speech signals<sup>65</sup>. Such  
497 impairment has been correlated with the severity of negative symptoms, including poverty of  
498 speech and hypobulia<sup>66</sup>.

499 **Prognostic classification**

500 Interestingly, while inhibitory control contributed to diagnostic classification, it was also  
501 the top predictor of remission outcome at follow-up. Our decision path analysis showed that  
502 patients who were correctly classified as remitted generally (despite some exceptions)

503 showed good baseline performance on inhibitory control tasks. Previous work has  
504 consistently demonstrated an association between EF and long-term post-treatment remission  
505 outcomes in patients with schizophrenia<sup>67</sup>. Specifically, patients with higher EF performance  
506 are more likely to remit relative to those with lower EF performance<sup>68,69</sup>, especially on the  
507 inhibition control dimension<sup>70</sup>. A possible interpretation is that patients with better inhibitory  
508 control are more adherent to therapeutic plans, including pharmacological interventions and  
509 lifestyle modifications<sup>71,72</sup>. Alternatively, because inhibitory control assessments are closely  
510 related to specific clinical manifestations in patients with schizophrenia<sup>73–75</sup>, better  
511 performance on this EF dimension along with milder symptom expressions among such  
512 patients, implies higher chance of remission<sup>68</sup>. Our results showed significant post-treatment  
513 improvement in interference inhibition in the remission subgroup, but not in the  
514 non-remission subgroup (Supplementary Materials), corroborating a relationship between  
515 inhibitory performance and remission in schizophrenia. Interestingly, we did not find  
516 significant difference in any baseline EF assessment and symptom dimension score between  
517 remitted and non-reddited patients. This points to a dissociation between symptoms of  
518 schizophrenia and the construct of EF; nonetheless, it highlights the role of EF dimensions in  
519 predicting remission rather than merely acting as a marker of illness severity.

520 A few patients who were correctly classified as having symptom remission exhibited  
521 poor inhibitory control performance; nevertheless, they demonstrated better baseline abilities  
522 in other dimensions, such as working memory updating or shifting. Furthermore, we revealed  
523 that the importance of working memory updating and maintenance contributions to the  
524 prognostic model in classifying remission status covaried with poor patient cognition at  
525 follow-up. Supporting this observation, previous studies have shown that working memory  
526 (updating and maintenance) performance is superior among stably remitted patients with  
527 schizophrenia versus non-reddited patients<sup>68</sup>. Disrupted working memory in patients with  
528 schizophrenia has been linked to cognitive disorganization and poorer performance on tasks  
529 requiring abstract thinking<sup>24,76,77</sup>, which is similarly assessed within our cognitive symptom  
530 dimension based on PANSS.

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531 **Limitations and considerations**

532 First, applied machine learning tends to rely on multiple datasets from independent  
533 medical centers for extensive tests. However, such concerns are moderated by our use of  
534 multiple random-splits to set aside a test ‘lock box’ in each repeat, while performing nested  
535 CVs on the remaining sample, as well-established previously<sup>50,53</sup>. This strategy effectively  
536 gauges the out-of-sample generalization performance, while balancing practical clinical data  
537 collection issues<sup>42,43,78,79</sup>. Nevertheless, future multisite and population-level EF studies may  
538 help expand the applicability. Second, patients with schizophrenia in our study had been  
539 treated with antipsychotics, reflecting typical clinical practice. In our sample, the OZP  
540 equivalent dosage did not show a significant correlation with most symptom scores (except  
541 for the affective symptom dimension) or EF measurements (except for one variable).  
542 Moreover, there was no significant difference in baseline OZP dosage between patients who  
543 were in remission and those who were not at follow-up. Nevertheless, the exact dosage of  
544 antipsychotic medication could have an impact on individual prognostic status<sup>80</sup>. As expected,  
545 the classification accuracy decreased when adjusting for individual variations in OZP  
546 equivalent dosage. However, a control analysis that only included patients who received the  
547 suggested starting dose for OZP (e.g., 10–20 mg/day)—a dose commonly effective in  
548 individuals with schizophrenia<sup>81</sup>—remained the classification performance as in our main  
549 experiments. Notwithstanding, future research involving drug-naïve patients at baseline and  
550 continuous assessments of EF function along with detailed records of medication usage over  
551 time would help establish the causal relationships between antipsychotic effects, prognostic  
552 statuses, and specific EF dimensions. Third, the attrition rate in our patient sample was  
553 similar to those reported previously.<sup>82–84</sup> This rate pertains to the representativeness of the  
554 findings derived from the patients who continued in the study, though we did not observe  
555 significant differences in all of the baseline characteristics between the followed up and  
556 drop-out patients (Supplementary Table S3). Furthermore, classification models established  
557 based on the two extreme conditions (i.e., treating the attribution patients as best-case [all  
558 remitted] or worst-case [non-remitted] scenarios) showed decreased prognostic  
559 discrimination accuracy. This was in line with the previous notion<sup>85,86</sup> that the worst case

560 analysis would lead to underestimated results<sup>87</sup>, though these might not reflect the true  
561 potential attrition bias. Future research may also incorporate outpatients to develop models  
562 representing the broader spectrum of patient populations, but managing the attrition rates  
563 remains a challenge.

564 To conclude, here we tested the classification power of six well-established behavioral  
565 paradigms, which assess three EF dimensions, for discriminating patients with schizophrenia  
566 from HC at baseline, as well as the remission status at follow-up. Results from robust  
567 validation and testing revealed promising performance and, thus, a consistent impairment in  
568 dimensions of EF to characterize individuals with schizophrenia and provide important  
569 prognostic information. Furthermore, different EF dimensions characterized diagnosis and  
570 prognosis to varying extents. Inhibitory control and working memory were identified as the  
571 most important factors for accurate classification of schizophrenia and remission status.  
572 Additionally, the classification strength of these EF dimension features was associated with  
573 specific psychopathologies. Thus, our research presents evidence that certain dimensions of  
574 EF are reliably compromised in individuals with schizophrenia. These deficits correlate with  
575 the severity of symptoms and can predict future outcomes. Hence, these measures may serve  
576 as valuable aids in the clinical assessment of this disorder.

577

## 578 **Methods**

### 579 **Ethics approval and consent**

580 This study was conducted in full compliance with the ethical guidelines and approved  
581 protocols of the Ethics Committee at the Third People's Hospital of Lanzhou City and  
582 Northwest Normal University (Lanzhou, China). Before participation, all participants  
583 involved in the study were provided with comprehensive information regarding the study  
584 aims, procedures, potential risks, and benefits. The study adhered to the principles outlined in  
585 the Declaration of Helsinki, and informed consent was provided by all participants.

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586 **Data collection**

587 ***Participant recruitment, clinical characterization, and definition of prognostic status***

588 The study was conducted from March to August, 2023. Participants were 195  
589 individuals who had been diagnosed with schizophrenia (of those, 86 were further assessed  
590 after a 4–6-week interval), and 169 healthy individuals (HC group) (Table 1). Participants in  
591 the schizophrenia group had received inpatient treatment at the Third People's Hospital in  
592 Lanzhou City within the past 2 years. Diagnoses were reached by two resident psychiatrists  
593 using the ICD-10 diagnostic criteria for schizophrenia (F20.9). The participants were further  
594 screened with the Structured Clinical Interview for DSM-IV axis I Disorders. The patients  
595 were in a stable condition and receiving consistent treatment, with no changes in medication  
596 expected during the study. Inclusion criteria were age 18–65 years and the ability to  
597 communicate effectively, complete experimental tasks, and voluntarily sign the informed  
598 consent form. Individuals with a severe physical disease, visual abnormality, or adverse drug  
599 reactions were excluded from the study (details for the inclusion and exclusion criteria are  
600 listed in Supplementary Table S1). Participants in the HC group, recruited through offline  
601 promotions and online advertisements, were matched with those in the schizophrenia group  
602 for age, sex, education level, and socioeconomic status. All HC group participants were  
603 physically healthy and did not have a history of mental illness or brain injury.

604 Symptom severity in each patient with schizophrenia was evaluated using the PANSS<sup>87</sup>.  
605 Scores for four symptom dimensions (i.e., positive, negative, affective, and cognitive factors)  
606 were derived for each patient via the Dimensions and Clustering Tool for Schizophrenia  
607 Symptomatology (DCTS; <http://webtools.inm7.de/sczDCTS/>). These dimensions have been  
608 previously identified as stable and generalizable across populations, regions, and clinical  
609 settings<sup>20</sup>. Higher scores denote more severe symptoms within each dimension. Patients were  
610 further categorized into the positive subtype, negative subtype, or ambiguous cases lying  
611 in-between these two subtypes based on their DCTS-derived symptom dimensional scores  
612 and membership values. We employed a heuristic membership degree of 0.6 as the cutoff  
613 value for the ambiguous cases subgroup given that these participants were not clearly  
614 assigned to any of the two more differentiated negative-positive subtypes.

615 For the 86 followed up patients with schizophrenia, remission versus non-remission  
616 prognostic statuses were determined based on the RSWG criteria<sup>34</sup>. Remission is defined by  
617 scores  $\leq 3$  on key PANSS items (P1, P2, P3, N1, N4, N6, G5, G9). Considering the current  
618 lack of consensus on a definition of clinical outcomes in schizophrenia, we utilized an  
619 alternative definition. This definition was based on the reduction of the total PANSS score,  
620 and three remitted or non-remitted conditions were specified by: 1) a 25% reduction; 2) a  
621 35% reduction; and 3) a 50% reduction<sup>36</sup>. Clinical outcomes of patients were moreover  
622 defined based on subtype-membership transition: 1) baseline non-negative subtype with  
623 transition to a negative subtype, and 2) stable negative subtype during follow-up. This  
624 definition helps in identifying patients that experience primary (and stable) negative  
625 symptoms or secondary symptoms due to antipsychotic treatments and related factors<sup>38,39</sup>.

626 **Assessments**

627 ***Sociodemographic and electronic records***

628 The standard 60-item Raven's Progressive Matrices test was used to assess fluid  
629 intelligence<sup>88</sup>. Family socioeconomic status was assessed using a family financial status  
630 questionnaire<sup>89</sup>. Thereafter, we collected detailed clinical information from the electronic  
631 medical records of patients (e.g., disorder onset, number of episodes, age at diagnosis,  
632 duration of illness, types and dosages of antipsychotic drugs, family medical history).

633 **EF**

634 This study was based on the influential model subdividing EF into three core dimensions:  
635 inhibitory control; working memory (updating and maintenance); and cognitive  
636 flexibility/shifting<sup>15,90</sup>. Working memory updating is the process of continuously replacing  
637 old information with new in working memory, according to current task requirements.  
638 Working memory span/maintenance is the ability to maintain and process information over a  
639 period of time, often directly linked to short-term memory capacity<sup>40</sup>. Inhibitory control  
640 involves the ability to suppress dominant responses and adapt to a changing environment,  
641 minimizing the impact of irrelevant information on ongoing information processing<sup>15</sup>.  
642 Therefore, inhibition is also divided into two dimensions: interference inhibition (or

643 interference control) and response inhibition (or behavioral inhibition)<sup>91</sup>. Cognitive flexibility  
644 is considered a single dimension, characterizing the ability to flexibly switch between  
645 different tasks and modes of thought<sup>92</sup>.

646 According to their complexity, we selected six behavioral tasks to measure these EF  
647 dimensions (Fig. 3)<sup>40</sup>: 1) *number running memory updating* task was used to examine  
648 working memory updating<sup>89</sup>; 2) *digit span backward* task was used to measure working  
649 memory maintenance (span)<sup>40</sup>; 3) *Corsi block* test was used to measure working memory  
650 maintenance (span), which more comprehensively assesses maintenance in the spatial  
651 dimension<sup>93</sup>; 4) *Stroop* task was used to measure interference inhibition<sup>94</sup>; 5) *Go/No-Go* task  
652 was used to measure response inhibition<sup>95</sup>; and 6) *number switching* task was used to measure  
653 shifting<sup>92</sup>. All behavioral tasks were performed using *E-Prime* 3.0 software (Psychology  
654 Software Tools, Inc., Pittsburgh, PA, USA). Accuracy and reaction time on each task can be  
655 weighted to derive 14 comprehensive assessment indicators from these six behavioral tests.

656 ***Inhibitory control***

657 The four measurements for assessing the interference control function included the  
658 reaction times for the incongruent, congruent, and neutral stimuli, and the difference of  
659 reaction time between the congruent and the incongruent condition trials (i.e., the interference  
660 effect) in the *Stroop* task. Three measurements for assessing the response inhibition function  
661 included the reaction time for the “Go” stimuli and the accuracy for the Go and No-Go  
662 stimuli in the *Go/No-Go* task.

663 ***Working memory***

664 Two measurements for assessing working memory updating function included the  
665 proportion of digits correctly recalled and placed in the correct sequence at two different  
666 speeds of presentation (1,750 ms and 750 ms per digit) in the *running memory* task. Three  
667 measurements for assessing the numeric working memory maintenance capacity included the  
668 length of the last correctly repeated sequence, the count of sequences correctly repeated until  
669 the conclusion of the test (i.e., the total number of successful trials) from the *Corsi block* test,  
670 and the maximal number of digits accurately recalled in the reverse order of the *digit span*

671 backward task.

672 **Cognitive flexibility**

673 Two measurements which included the difference in reaction time between the switch  
674 and the non-switch trials [switch cost], as well as the difference in the reaction time between  
675 the non-switch and the single-task trials [mixing cost]) measured in the *number-letter*  
676 *switching* task.

677 Besides a conceptual formulation of the 14 task measurements according to the  
678 three-dimensional representation of EF, we supplemented five composite scores calculated  
679 based on these measured variables<sup>96</sup>. Firstly, an inhibitory composite score was calculated by  
680 averaging: 1) the difference in reaction time between the congruent and the incongruent  
681 condition trials in the *Stroop* task; and 2) the accuracy for the No-Go stimuli in the *Go/No-Go*  
682 task as in previous studies<sup>97,98</sup>. The purpose of this approach was to denote the combined  
683 response and inference inhibitory functions. Furthermore, this inhibitory composite score was  
684 aggregated with the assessments in the *running memory* task (the proportion of digits  
685 correctly recalled and placed in the correct sequence at the speed of 1,750 ms per digit) and  
686 the *number-letter switching* task (switch cost) to form an abbreviated version for representing  
687 general EF functions. Such abbreviation is in compliance with the previous notion on a  
688 single-condition indicator that these trails require greater executive control demands<sup>99</sup>.  
689 Considering that EF functions interplay across conceptual constructs during cognitive  
690 engagement (e.g., problem-solving) for processing particular behaviors<sup>100</sup>, we additionally  
691 created three cross-dimensional EF composite scores by collapsing the cardinal three  
692 dimensions of EF, which the inhibitory composite score was similarly used: 1) Inhibition and  
693 Updating composite; 2) Inhibition and Switching composite; and 3) Switching and Updating  
694 composite. The measurements used to assess working memory updating and switching were  
695 employed to calculate the abbreviated version of the EF composite score. These composite  
696 measures would provide additional insights into the dimensional and cross-dimensional  
697 contributing features to diagnostic and prognostic classifications<sup>100</sup>. Consequently, 19  
698 EF-related indicators were used as input features for the machine learning models (Fig. 3).

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699 **Behavioral task and clinical scale analyses**

700 We performed a mixed-model ANCOVA with group (schizophrenia vs. healthy control)  
701 as a between-subjects factor and all EF measurements as well as composite scores as  
702 within-subjects factors, while controlling for sociodemographic variables, to examine  
703 whether schizophrenia has differentially affected EF performance. The mixed-model  
704 ANCOVA was followed by multiple one-way ANCOVAs to examine group differences in  
705 each EF measure. also controlling for Sociodemographic variables were likewise controlled  
706 and a Benjamini-Hochberg (BH) false discovery rate (FDR) correction was applied to adjust  
707 for multiple comparisons. Eta squared effect sizes were calculated to describe the magnitude  
708 of effect sizes, with certain values interpreted as small ( $\eta^2 \leq 0.01$ ), medium ( $0.01 < \eta^2 \leq 0.06$ ),  
709 or large ( $0.06 < \eta^2 \leq 0.14$ )<sup>101</sup>. Two-sample t tests were used to compare clinical symptoms,  
710 demographic characteristics, and EF between the remission and non-remission subgroups.  
711 Chi-squared tests were used to compare categorical variables between these groups. We used  
712 paired-samples *t*-tests to compare clinical symptoms between baseline and follow-up (after  
713 4–6 weeks of treatment) for the subset of followed participants. Finally, we used Pearson and  
714 Spearman correlation analyses to examine the relationships among clinical symptoms and EF,  
715 with the former applied to continuous variables and the latter applied to categorical variables.

716 **Classification modeling procedure**

717 **Features and models**

718 Participants were categorized into schizophrenia and HC groups. Two feature sets were  
719 tested for diagnostic classification accuracy: 1) 19 baseline EF assessments, subsuming the  
720 three EF dimensions measured by six behavioral paradigms (Table 2); and 2) 32  
721 features—the 19 baseline EF measures plus 13 routinely attainable sociodemographic  
722 variables. For prognostic classification, the same two feature sets were tested to distinguish  
723 patient remission status (remitted vs. non-remitted) after 4–6 weeks of antipsychotic  
724 treatment. SVM, RF, and AdaBoost, which are widely used in psychiatric machine-learning  
725 research<sup>102</sup>, were used for classification tasks, along with a synthesized stacking model of the  
726 three. Stacking models are a multi-view approach integrating classification weight estimates

727 from single classifiers to improve ultimate performance<sup>103</sup>.

728 **Complementary investigations**

729 ***Control analysis***

730 Several control analyses were performed, in which we controlled the effects of 1) 13  
731 demographic characteristics; 2) both the 13 demographic variables and the OZP equivalent  
732 dosage; 3) only the OZP equivalent dosage on the baseline EF assessments using regression  
733 approaches<sup>104</sup> when establishing the diagnostic and prognostic classification models. The  
734 continuous variables in the demographic dataset included age, years of education, body mass  
735 index, socioeconomic status, and fluid intelligence measured by Raven's Progressive  
736 Matrices test. The categorical variables consisted of sex, ethnicity, residence, employment  
737 status, only-child status, marital status, smoking history, and drinking history. One-hot  
738 encoding was applied to the categorical variables, converting them into binary vectors.  
739 Furthermore, we sought to mitigate the potential impact of antipsychotic drug dosage and  
740 evaluate the robustness of our classification results. Thus, we established the prognostic  
741 classification models with only those patients receiving a clinically standard, commonly  
742 effective OZP equivalent dosage of 10–20 mg/day.

743 ***Sensitivity analysis***

744 In our research, the majority of dropouts were due to hospital discharge. Therefore, we  
745 conducted a best-case sensitivity analysis, assuming that all participants who were lost to  
746 follow-up had positive treatment outcomes (remission). Additionally, we performed a  
747 worst-case analysis, assuming that all participants who were lost to follow-up had the least  
748 favorable treatment outcomes (no remission). These data help establish the potentially  
749 extreme scenarios due to attrition with respect to our main analyses with an observed  
750 prognostic status. Sensitivity analyses, such as best-worst (assuming all participants lost to  
751 follow-up in one group [referred as group 1] have had a beneficial outcome and all those with  
752 missing outcomes in the other group [group 2] have had a harmful outcome) and worst-best  
753 case (assuming that all participants lost to follow-up in group 1 have had a harmful outcome;  
754 and that all those lost to follow-up in group 2 have had a beneficial outcome), are utilized in

755 medical and psychiatric research to evaluate the reliability of results in relation to participants  
756 who do not complete the study<sup>105,106</sup>.

757 **Machine learning design**

758 Machine learning and CV were implemented using Python (version 3.10.11) and the  
759 scikit-learn package (version 1.3.0). Specifically, the original data were first preprocessed to  
760 accommodate missing values, outliers, and class imbalance issues (see Supplementary  
761 Materials for details). Thereafter, we randomly split the preprocessed data into a discovery  
762 dataset with 80% of the overall schizophrenia and HC groups ('training set') and a 'lock-box'  
763 test dataset with the remaining 20% of these samples ('test set') to determine out-of-sample  
764 classification performance(Fig. 1)<sup>4,42–44,107</sup>. The random split procedure was stratified for the  
765 outcome variable (diagnostic label or remission status), ensuring a balanced representation of  
766 labels in each dataset<sup>42,43</sup>. Using the discovery dataset, we performed a nested CV loop<sup>108</sup>  
767 (also termed double CV), which differentiates two CV roles to avoid 'circularity' introduced  
768 by overfitting when the same sample subset is used for both hyperparameter tuning and  
769 model validation<sup>53</sup>. Specifically, within a nested CV loop, the inner CV (k= 3), encompassing  
770 80% of the discovery sample, operates all data-dependent decisions while determining  
771 optimal hyperparameters. The outer CV (k= 5) is subsequently utilized for parameter  
772 assessment and model selection<sup>78</sup>. For optimal hyperparameter selection, we used the Optuna  
773 optimization technique (version 3.5.0) and selection based on the achievable AUC of  
774 candidate hyperparameters within the validation sets of the inner loop<sup>109</sup>. The AUC metric,  
775 representing the degree of separability, is widely used to evaluate model performance. In this  
776 investigation, it indicated the ability of the model to distinguish between the schizophrenia  
777 and HC groups, and between the patient remission and non-remission groups.  
778 Hyperparameters with the highest average performance over the 5 × 3 nested CV were used  
779 to train a model on the entire discovery sample without further modification; next, they were  
780 tested using the independent 'test set' sample<sup>50</sup>. In addition to AUC, sensitivity, specificity,  
781 and balanced accuracy performance metrics were assessed<sup>110</sup>. To avoid potential bias from  
782 random splitting, the aforementioned machine learning procedure was repeated 100  
783 times<sup>111,112</sup>.

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784 ***Feature importance analysis***

785 To evaluate the contributions of EF features to our classification models, we assigned an  
786 importance score (i.e., Shapley value) to each feature<sup>47</sup>. Specifically, we used the SHAP  
787 library (version 0.39.0) model-agnostic SHAP KernelExplainer approach, which is generally  
788 used to estimate Shapley values for prediction models<sup>113</sup>. SHAP KernelExplainer employs a  
789 Monte Carlo approach to randomly sample feature combinations based on input predictors.  
790 Initially, it estimates the importance of these combinations with varying features present in  
791 model predictions. Subsequently, individual Shapley values are calculated to denote the  
792 contribution of each feature to the target prediction based on a weighted linear regression  
793 model<sup>48</sup>.

794 After obtaining the individual-level Shapley values for each feature, we computed the  
795 mean absolute Shapley value (i.e., feature importance score) across all individuals (i.e., the  
796 group-level Shapley values) where larger Shapley values indicate stronger importance of this  
797 feature to the classification model. The group-level Shapley values were next used to depict a  
798 group-level decision path for each feature. Essentially, among those participants in the  
799 schizophrenia group who were correctly classified, performance by each EF dimension  
800 assessment was averaged across both the HC and schizophrenia groups for the diagnostic  
801 models, and across the remission and non-remission patient groups for the prognostic models.  
802 The averaged performance of each EF feature was then normalized and compared with a  
803 positive (or negative) value; higher (or lower) performance by an EF feature drove the model  
804 to more accurately classify true cases. Using individual Shapley values, we also plotted the  
805 decision path for each individual who was correctly classified to complement the group-level  
806 results; this was performed because individuals with schizophrenia have heterogeneous  
807 expressions across EF dimensions<sup>60</sup>. Finally, based on the individual Shapley values, we  
808 tested the extent to which the importance of the contribution of an EF feature to a  
809 classification was linked to individual psychopathology. This was conducted using Pearson  
810 correlation analysis based on individual scores (entire patient group  $N=195$ , diagnostic model;  
811 subset sample  $N= 86$ , prognostic model) along the four symptom dimensions which were  
812 assessed at baseline and follow-up, and the difference between the two. The FDR approach

813 was used for the correction of multiple comparisons to statistically rule out potential  
814 false-positive associations<sup>114</sup>.

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840 **Conflict of interest**

841 The authors declare no conflicts of interest.

842 **Data availability statement**

843 Information for the main sample used in the present study have been included in the  
844 Supplementary Materials. The raw data of our used sample are protected and are not publicly  
845 available due to data privacy. These data can be accessed upon reasonable request to the  
846 corresponding author (X. Z.). Derived data supporting the findings of this study are available  
847 from the corresponding authors (X. Z. or J. C.) upon request.

848 **Code availability statement**

849 Scripts to run the main analyses have been made publicly available and can be accessed  
850 at <https://doi.org/10.6084/m9.figshare.26086594.v1>.

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1148

1149 **Table 1 | Participant demographics and clinical characteristics**

Characteristic	Schizophrenia group (N= 195)	Healthy control group (N= 169)	p-value	Remission group at baseline (N= 58)	Non-remission group at baseline (N= 28)	p-value
<b>Demographic</b>						
Age	35.35 ± 9.35	37.69 ± 13.71	0.055	34.03 ± 8.69	34.54 ± 9.75	0.810
Sex			0.264			0.192
Male	114 (58.5%)	88 (52.1%)		39 (67.2%)	14 (50%)	
Female	81 (41.5%)	81 (47.9%)		19 (32.8%)	14 (50%)	
Ethnicity, Han	173 (88.7%)	152 (89.9%)	0.837	52 (89.7%)	24 (85.7%)	0.861
Education, years	11.12 ± 4.58	10.90 ± 3.94	0.615	10.52 ± 3.98	11.07 ± 4.24	0.555
BMI	23.21 ± 3.73	23.95 ± 4.32	0.494	23.60 ± 3.13	22.65 ± 3.78	0.217
Residence, urban	114 (58.5%)	98 (58.0%)	1.000	27 (24.6%)	16 (57.1%)	0.490
SES	23.21 ± 7.32	23.88 ± 5.80	0.338	22.72 ± 7.25	23.43 ± 8.02	0.684
RPM	32.56 ± 11.50	39.52 ± 9.63	< 0.001	34.84 ± 11.17	33.04 ± 12.49	0.500
Employed, yes	53 (27.2%)	111 (65.7%)	< 0.001	44 (75.9%)	21 (75.0%)	.
Only child, yes	63 (32.3%)	24 (14.2%)	< 0.001	42 (72.4%)	14 (50.0%)	0.072
Marital status			< 0.001			0.205
Unmarried	115 (59.0%)	54 (32.0%)		34 (58.6%)	19 (67.9%)	
Married	48 (24.6%)	108 (63.9%)		18 (31.0%)	4 (14.3%)	
Divorced	31 (15.9%)	7 (4.1%)		6 (10.3%)	5 (17.9%)	
Widowed	1 (0.5%)	0 (0.0%)		0 (0.0%)	0 (0.0%)	
Smoking history			0.079			0.731
Never	119 (61.0%)	122 (72.2%)		36 (62.1%)	19 (67.9%)	
1–3 years	19 (9.7%)	11 (6.5%)		15 (8.6%)	3 (10.7%)	
>3 years	57 (29.2%)	36 (21.3%)		17 (29.3%)	6 (21.4%)	

Alcohol consumption history						
Never	133 (68.2%)	99 (58.6%)	0.118	42 (72.4%)	18 (64.3%)	0.729
Occasionally	54 (27.7%)	64 (40.3%)		14 (24.1%)	9 (32.1%)	
Regularly	8 (4.1%)	6 (3.8%)		2 (3.4%)	1 (3.6%)	
<b>Clinical</b>						
Electronic medication records						
Age at onset	27.43 ± 8.63	.	.	27.02 ± 8.02	27.25 ± 9.39	0.906
Duration of disorder	9.74 ± 7.17	.	.	9.09 ± 5.98	10.48 ± 7.51	0.395
Frequency of episodes	5.98 ± 4.27			6.17 ± 3.84	6.32 ± 4.05	0.871
First episode, yes	10 (5.13%)	.	.	1 (1.72%)	1 (3.57%)	0.984
Family medical history, yes	33 (16.92%)	.	.	49 (84.5%)	23 (82.1%)	.
Dose equivalent to olanzapine (mg/day)	14.48 ± 6.41	.	.	14.30 ± 5.46	16.65 ± 7.66	0.154
Type of antipsychotic medication						0.829
First generation	9 (5%)	.	.	2 (3.4%)	2 (7.1%)	
Second generation	186 (95%)	.	.	56 (96.6%)	26 (92.9%)	
Clinical scale						
3 PANSS subscales						
PANNS-Negative	21.42 ± 6.48	.	.	21.59 ± 6.59	21.79 ± 5.99	0.893
PANNS-Positive	22.01 ± 4.46	.	.	21.64 ± 4.06	21.18 ± 5.99	0.676
PANNS-General	40.3 ± 6.81	.	.	39.67 ± 7.39	40.39 ± 6.20	0.657
PANNS-Total	83.79 ± 13.85	.	.	82.90 ± 13.97	83.36 ± 15.33	0.890
4 dimensions of PANSS						
Negative factor	8.22 ± 2.5	.	.	8.24 ± 2.70	8.71 ± 2.11	0.425
Positive factor	6.24 ± 1.74	.	.	5.99 ± 1.66	6.23 ± 2.07	0.560
Affective factor	5.65 ± 0.86	.	.	5.67 ± 0.93	5.57 ± 0.78	0.619
Cognitive factor	9.78 ± 1.72	.	.	9.66 ± 1.82	9.76 ± 1.61	0.823

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**Note:** Data are presented as the mean  $\pm$  standard deviation or n (%). The *p*-values in bold face indicate statistically significant differences ( $p < 0.05$ ). **Calculation:** BMI is calculated as weight (kg) divided by height squared (m<sup>2</sup>). Remission status (remission or non-remission) was determined based on the RSWG remission criteria. **Abbreviations:** BMI, body mass index; SES, socioeconomic status; RPM, Raven's Progressive Matrices. PANSS, Positive and Negative Syndrome Scale; RSWG, Remission in Schizophrenia Working Group.

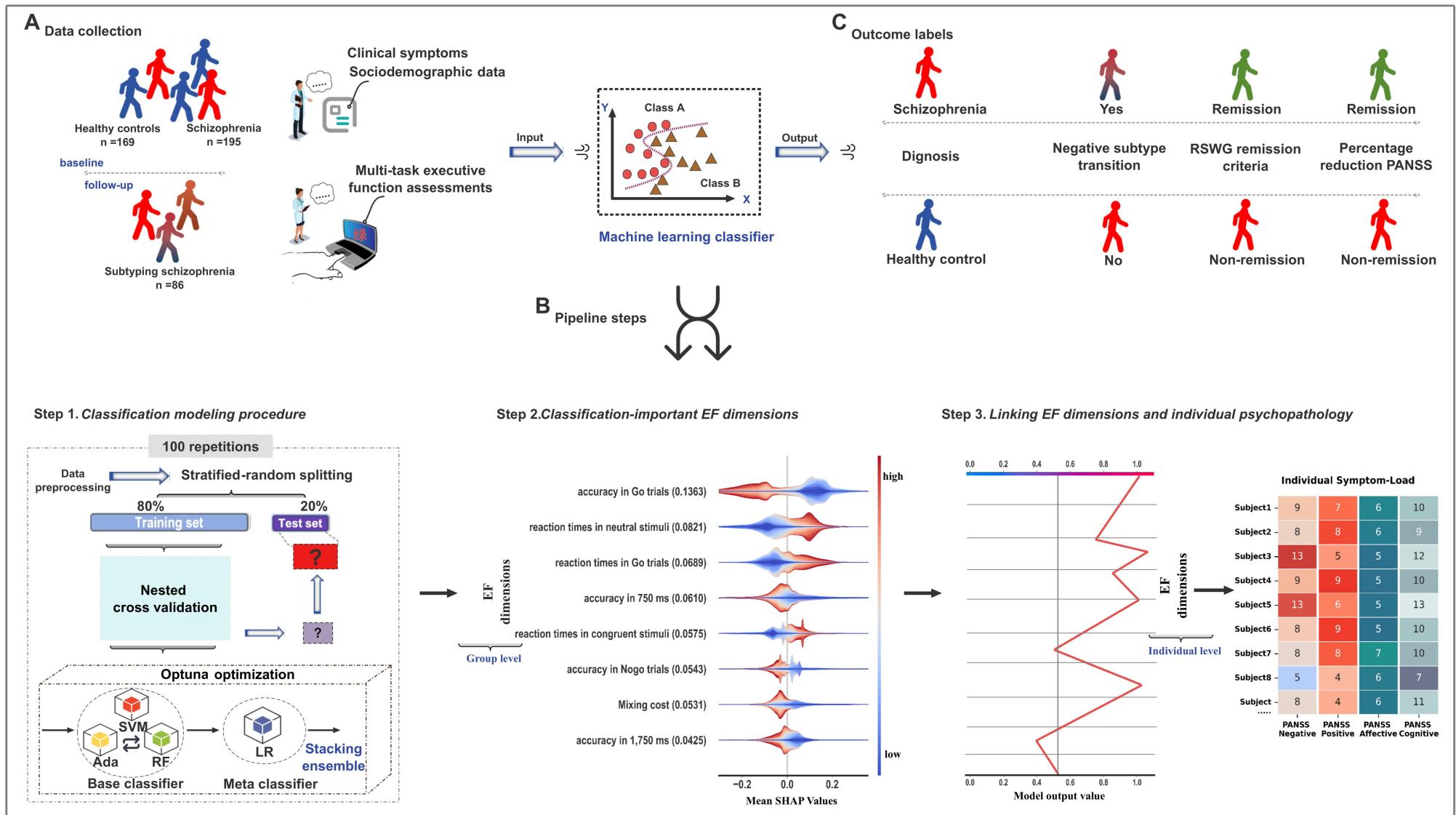
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**Table 2 | Comparisons of executive function dimensions between groups**

<b>Switching</b>	Switching cost	$304.80 \pm 303.40$	$198.95 \pm 238.37$	<b>0.001</b>	0.030	$292.36 \pm 300.00$	$277.50 \pm 286.88$	0.840
(Number switching task)	Mixing cost	$62.43 \pm 189.87$	$112.30 \pm 149.69$	0.278	0.010	$70.90 \pm 213.90$	$14.32 \pm 213.43$	0.258
<b>Composite scores</b>								
	Inhibition composite	$-11.09 \pm 36.83$	$-3.10 \pm 22.12$	0.217	0.010	$2.02 \pm 72.87$	$-3.85 \pm 31.75$	0.484
	Executive function composite	$98.10 \pm 102.30$	$65.53 \pm 80.06$	<b>0.001</b>	0.030	$98.31 \pm 103.95$	$91.39 \pm 94.78$	0.624
	Inhibition and updating composite	$-5.24 \pm 18.43$	$-1.18 \pm 11.05$	0.217	0.010	$1.29 \pm 36.43$	$-1.67 \pm 15.89$	0.481
	Inhibition and switching composite	$146.86 \pm 153.44$	$97.92 \pm 120.07$	<b>0.001</b>	0.030	$147.19 \pm 155.91$	$136.82 \pm 142.15$	0.625
	Updating and switching composite	$152.70 \pm 151.71$	$99.84 \pm 119.20$	<b>0.001</b>	0.030	$1146.46 \pm 150.01$	$139.01 \pm 143.46$	0.729

**Note:** The *p*-values in bold face indicate statistically significant differences ( $p < 0.05$ ). The *p*-values were adjusted using the false discovery rate (FDR) with Benjamini–Hochberg procedure.  $\eta^2$ , eta-squared effect size. The interference effect is quantified as the difference in reaction time between incongruent and congruent trials in the *Stroop* task.

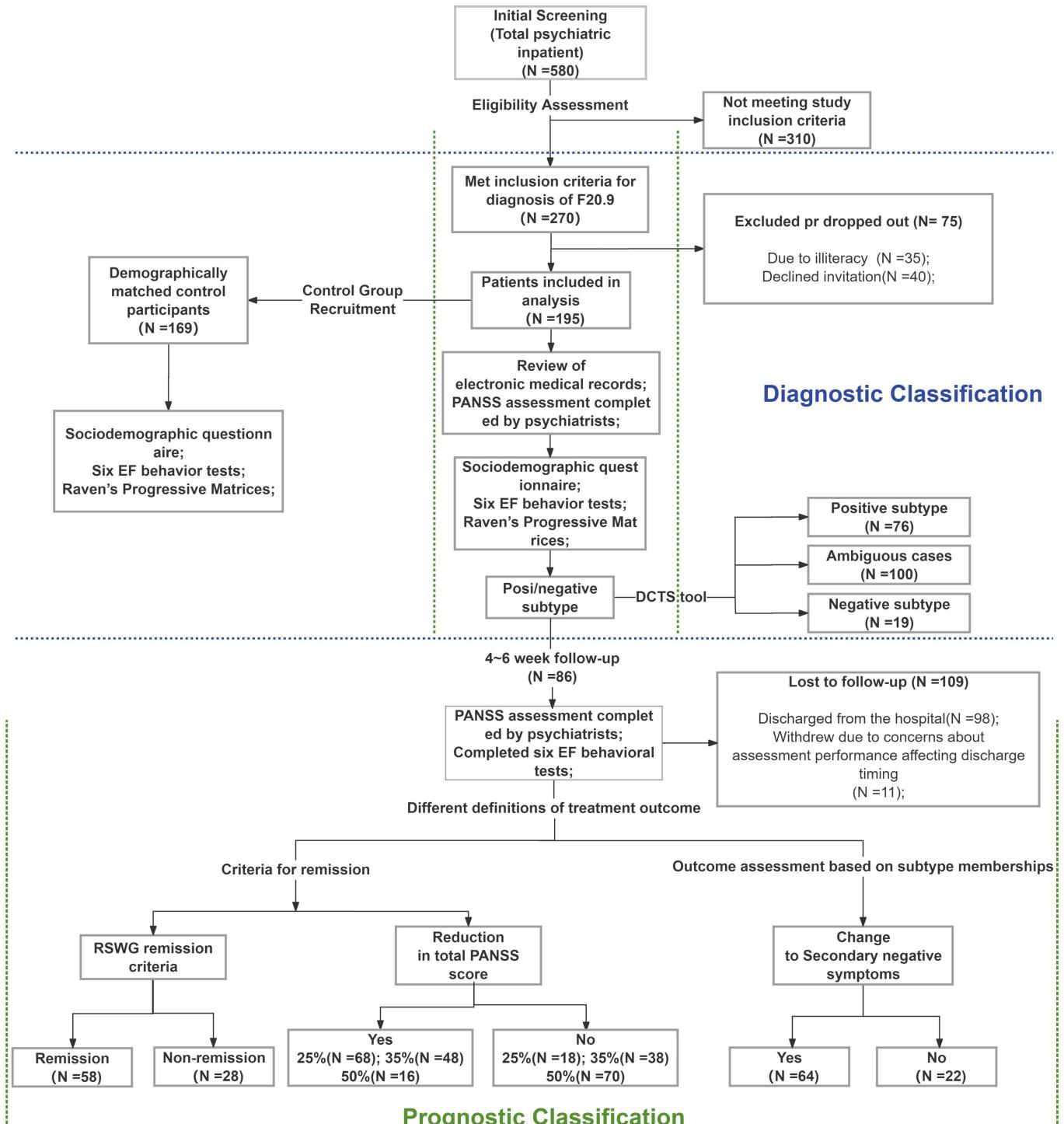
1153 Fig. 1 | Study overview.



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1154 **A).** The diagnostic model included the recruitment of patients with schizophrenia, further stratified into positive and negative symptom subtypes, and healthy  
1155 controls. All participants underwent a comprehensive battery of tests and assessments tailored to their respective groups. The prognostic model was developed  
1156 from a cohort of 86 patients with schizophrenia who completed a standard treatment regimen within 4–6 weeks of hospitalization. At follow-up, patients were  
1157 evaluated using the RSWG criteria as the primary outcome measure. Additionally, three widely accepted definitions in the field were incorporated to  
1158 comprehensively classify treatment response (i.e., 25%, 35%, and 50% symptom reduction thresholds), as well as changes in positive and negative symptom  
1159 subtypes. **B).** Data underwent preprocessing, followed by a stratified random division into an 80% discovery dataset and a 20% test dataset, balanced for  
1160 diagnostic and remission outcomes. The discovery set was subjected to nested cross-validation, with model performance assessed on the test set using various  
1161 metrics. To reduce splitting variance, this procedure was repeated 100 times. **C).** SHAP values assigned to executive function features indicated their  
1162 importance in model predictions, with mean absolute values reflecting overall impact. Individual Shapley values highlighted feature influence on correct  
1163 classifications, which were also assessed for their relationships to psychopathology measures. FDR was used to control for false-positives in multiple  
1164 comparisons. Correlational analysis between individual-level Shapley values for each feature and individual psychopathology.  
1165 **Abbreviations:** Ada, AdaBoost; FDR, false discovery rate; RF, random forest; RSWG, Remission in Schizophrenia Working Group; SHAP, SHapley  
1166 Additive exPlanations; SVM, support vector machine. EF, executive function.

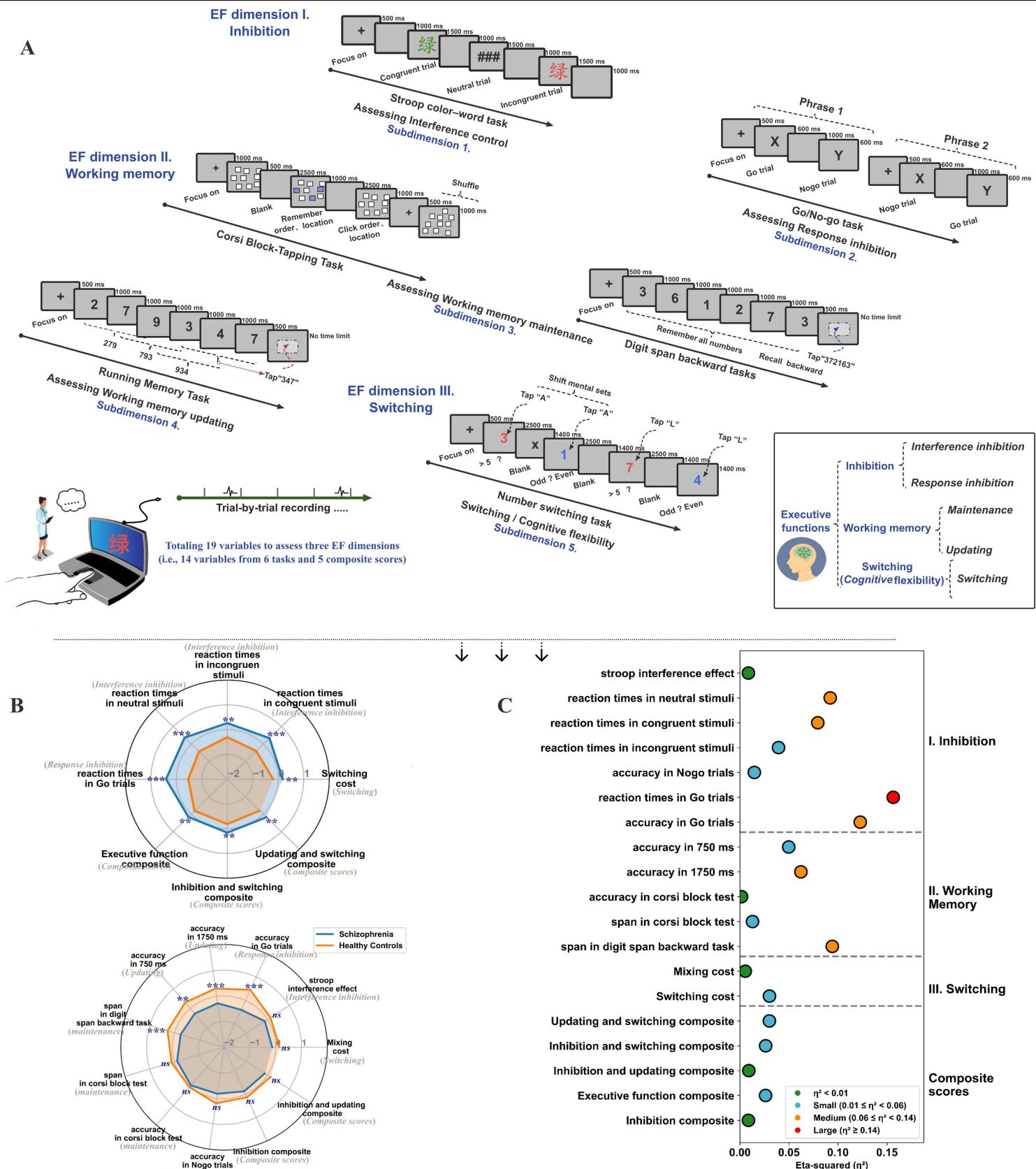
1167 Fig. 2 | Flowchart of patient identification



1168 Flow chart depicting the inclusion and exclusion criteria for participants.

1169 **Abbreviations:** DCTS, Dimensions and Clustering Tool for Schizophrenia Symptomatology;  
1170 EF, executive function; PANSS, Positive and Negative Syndrome Scale; RSWG, Remission  
1171 in Schizophrenia Working Group.

1172 Fig. 3 | Multi-task executive function dimension assessments



1173

1174 A). The six behavioral paradigms (e.g., Zhao et al., 2023)<sup>40</sup> used to assess the three EF

1175 dimensions (i.e., inhibitory control, working memory maintenance and updating, and  
1176 cognitive flexibility) based on 14 measurements. **B)** The radar plot was constructed based on  
1177 the Z-scores of the 14 EF measurements and the 5 composite scores, with annotated *p*-values  
1178 (FDR corrected) resulted from one-way ANCOVAs following a mixed model ANCOVA. **C)**.  
1179 Effect sizes are colored as small ( $0.01 \leq \eta^2 < 0.06$ ), medium ( $0.06 \leq \eta^2 < 0.14$ ), or large ( $\eta^2 \geq$   
1180 0.14) based on the guidelines proposed by Cohen (1988).

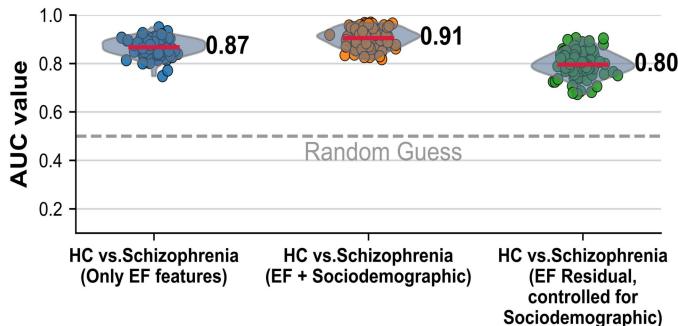
1181 **Abbreviations:** EF, executive function; FDR, false discovery rate; *ns*, not significant.

1182 **Note:** *ns*: Not significant difference. \*:  $p < 0.05$ . \*\*:  $p < 0.01$ . \*\*\*:  $p < 0.001$ .

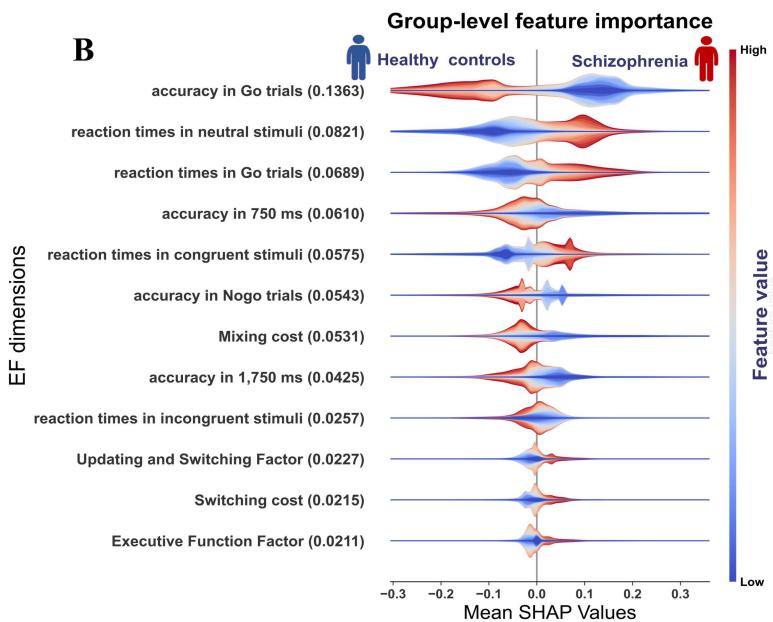
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1184 Fig. 4 | Model metrics and feature importance for diagnostic models

A



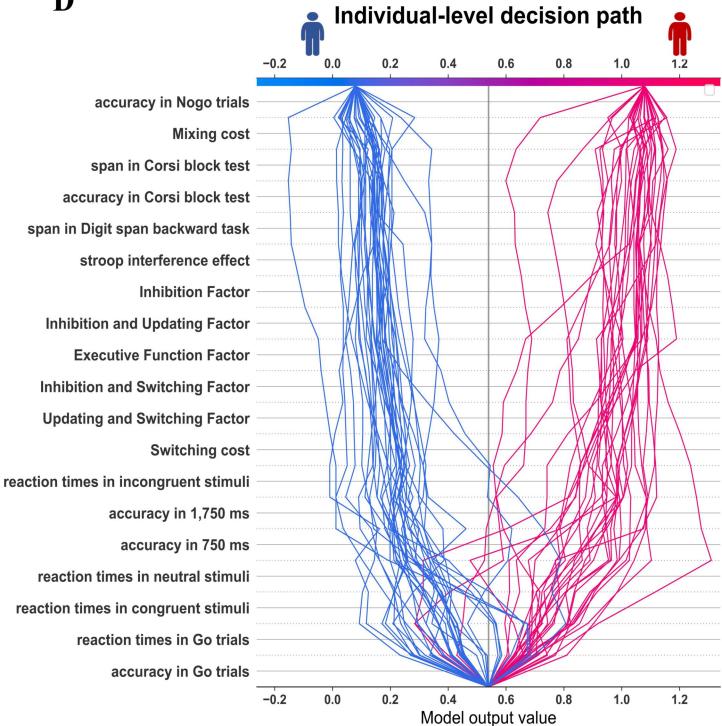
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C



D



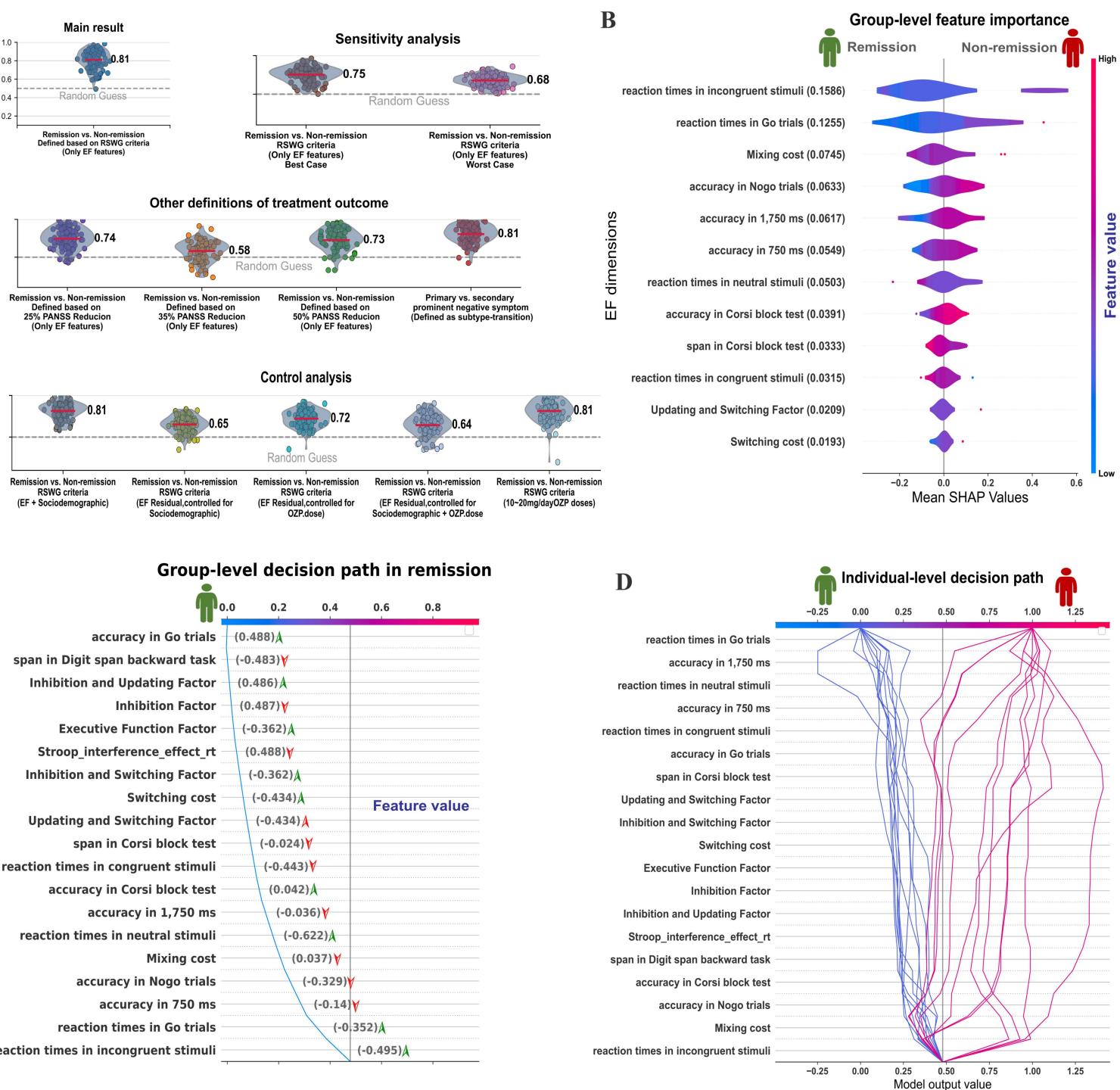
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1186 A). Violin plots show the values of area under the curve (AUC) for discriminating patients  
1187 with schizophrenia from healthy participants by diagnostic classification models. Each point  
1188 within the violin plots represents the AUC value derived from the hold-out test data of each  
1189 random split procedure (repeated 100 times) in our machine learning design. The red line  
1190 within each violin plot denotes the mean. B). The group-level feature importance plot ranks  
1191 EF features on the y-axis by their absolute average Shapley value across individuals,  
1192 representing their overall importance in the ability of the model to distinguish between  
1193 patients and healthy participants. The original feature weights for each EF variable were color

1194 coded, with blue color denoting a negative weight value and red color denoting a positive  
1195 weight value. Values along the x-axis indicate a positive or negative effect of an EF feature  
1196 on classifying an individual, with a negative and positive value promoting the model towards  
1197 a classification of "healthy" and "schizophrenia", respectively. Collectively, these findings  
1198 indicated that higher accuracy in the Go trials was associated with a higher likelihood of  
1199 classifying an individual as healthy. **C).** The group-level decision path, generated based on all  
1200 correctly classified schizophrenia samples. EF features are ranked from upper to bottom based  
1201 on their group-level importance present in b), the color bar denotes the impact of an EF  
1202 feature on model's classification towards "healthy" (blue) or "schizophrenia" (red), the red  
1203 curve shows the values for each EF feature coded in the color bar. The numbers in  
1204 parentheses represent the z-score standardized original measurement values of each EF  
1205 feature by the averages of this EF feature across the healthy and the patient groups in the  
1206 model test samples (a negative number to the right of the perpendicular line denotes the  
1207 measurement of an EF feature that is below the average in patients with schizophrenia). The  
1208 red or green arrows adjacent to the parentheses indicate the higher or lower values within the  
1209 parentheses that are associated with better or worse EF functions, respectively. This is  
1210 because some measurements from the EF tasks indicate better performance with higher values,  
1211 while others are more favorable with lower values. **D).** The decision path for each individual  
1212 in the test sets of the classification modeling iterations, that for each individual, how each of  
1213 these important EF features have promoted the diagnostic model to classify this individual as  
1214 a healthy participant or patient with schizophrenia, given the expression level (task  
1215 measurements) of this individual in each of the EF features. The blue and red lines indicate  
1216 accurate classifications as healthy participants and schizophrenia patients, respectively.

1217 **Abbreviations:** EF, executive function; HC, healthy control.

1218 Fig. 5 | Model metrics and feature importance for prognostic models



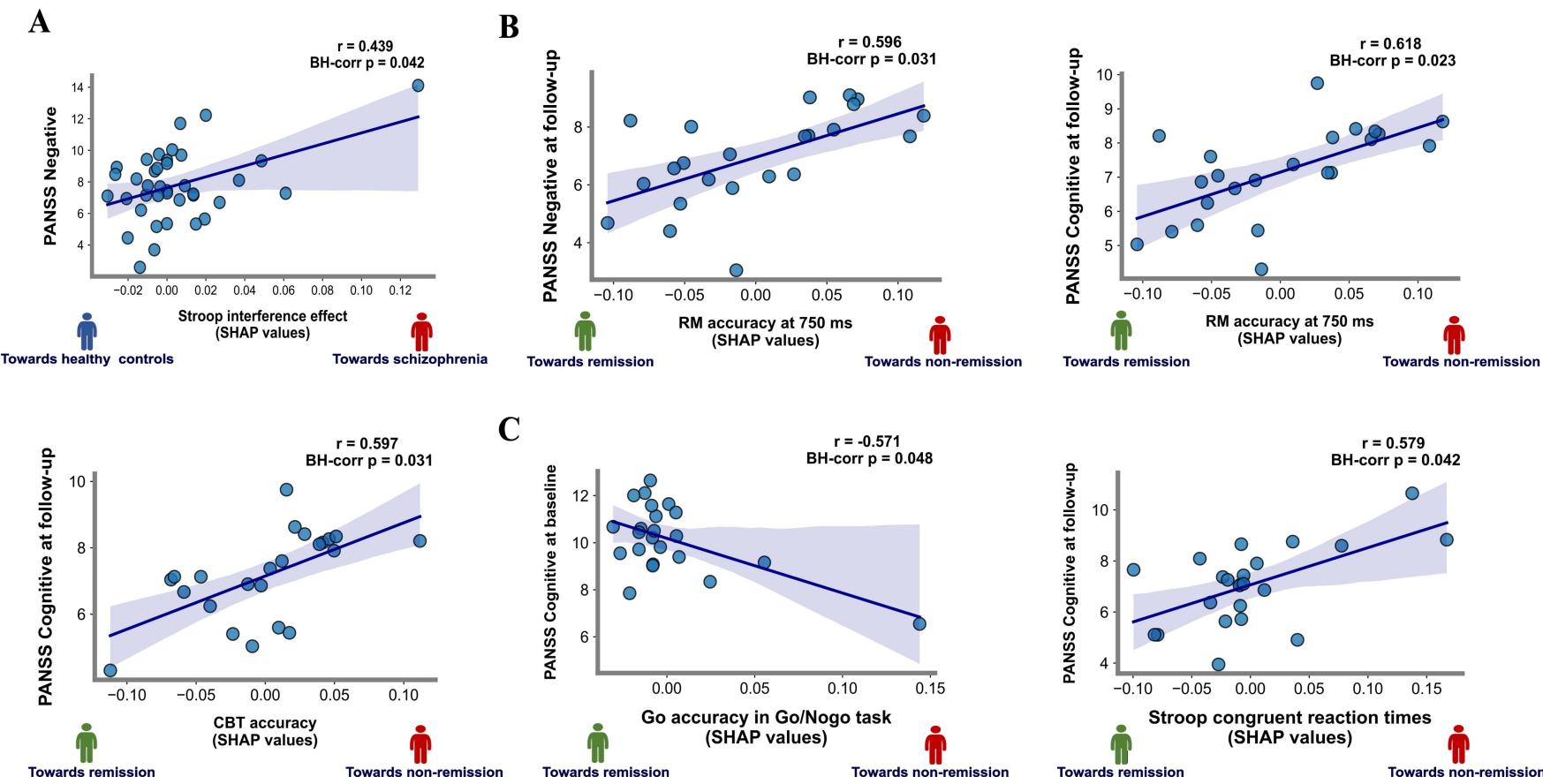
1219

1220 A). Violin plots show the values of area under the curve (AUC) for discriminating remitted  
1221 patients with schizophrenia from non-remitted patients by prognostic classification models.  
1222 Each point within the violin plots represents the AUC value derived from the hold-out test  
1223 data of each random split procedure (repeated 100 times) in our machine learning design. The  
1224 red line within each violin plot denotes the mean. B). The group-level feature importance plot  
1225 ranks EF features on the y-axis by their absolute average Shapley value across individuals,  
1226 representing their overall importance in the model's distinction between remitted and

1227 non-remitted patients (i.e., greater separation of the violin-like plots towards the extremes  
1228 denotes higher importance). The original feature weights for each EF variable are color-coded,  
1229 with blue indicating a negative weight value and red indicating a positive weight value.  
1230 Values along the x-axis indicate a positive or negative effect of an EF feature on classifying  
1231 an individual, with a negative and positive value promoting the model towards a classification  
1232 of "remission" and "non-remission", respectively. Collectively, these data suggest that longer  
1233 reaction times in incongruent stimuli are associated with a higher likelihood of classification  
1234 as non-remitted patient. **C)**. The group-level decision path, generated based on all correctly  
1235 classified remitted schizophrenia patients. EF features are ranked from upper to bottom based  
1236 on their group-level importance present in b), the color bar denotes the impact of an EF  
1237 feature on model's classification towards "remission" (blue) or "non-remission" (red), and the  
1238 blue curve shows the values for each EF feature coded in the color bar. The numbers in  
1239 parentheses represent the z-score standardized original measurement values of each EF  
1240 feature by the averages of this EF feature across the remitted and non-remitted patient  
1241 subgroups in the model test samples (a negative number to the left of the perpendicular line  
1242 denotes the measurement of an EF feature that is below the average in remitted patients). The  
1243 red or green arrows adjacent to the parentheses indicate the higher or lower values within the  
1244 parentheses that are associated with better or worse EF functions, respectively, as some  
1245 measurements from the EF tasks indicate better performance with higher values, while others  
1246 are more favorable with lower values. **D)**. The decision path for each individual in the test  
1247 sets of the classification modeling iterations, illustrating how each of these important EF  
1248 features influenced the prognostic model to classify each individual as remitted or  
1249 non-remitted, given their performance (task measurements) in each of the EF features. The  
1250 blue and red lines represent correct classifications as remitted and non-remitted patients,  
1251 respectively.

1252 **Abbreviations:** EF, executive function; OZP, olanzapine; PANSS, Positive and Negative  
1253 Syndrome Scale; RSWG, Remission in Schizophrenia Working Group; SHAP, SHapley  
1254 Additive exPlanations.

1255 **Fig. 6 | Correlation between the importance of an executive function feature and individual psychopathology along four symptom**  
 1256 **dimensions**



1257 **A)** Correlation in the diagnostic model using executive function features only. **B)** Correlation in the prognostic model using executive function features only.  
 1258 **C)** Correlation in the prognostic model using executive function features and sociodemographic features.

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1259 **Abbreviations:** BH, Benjamini–Hochberg correction; CBT, Corsi block test, assessing numeric working memory maintenance capacity; Go/No-Go task,  
1260 assessing response inhibition function; PANSS, Positive and Negative Syndrome Scale; RM, running memory task, assessing working memory updating  
1261 capability; SHAP, SHapley Additive exPlanations; Stroop, Stroop task, assessing interference control function.