Demographic Forecasting

Lecture 4: the Lee-Carter method

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Course overview

- Lecture 1: direct extrapolation by (generalized) linear models
- Lecture 2: direct extrapolation by time-series methods
- Lecture 3: parametric approaches
- Lecture 4: the Lee-Carter method



Mortality forecasting

- Crucial for sustainability of pensions, insurances, elderly care; predicting population ageing and projecting populations; ...
- Until the 1980s, the methods used to forecast mortality were deterministic, based on mathematical formulae or expert judgment
- Revived interest in recent years following the introduction of the Lee-Carter method in 1992
- One of the firstly introduced stochastic mortality models ⇒ the model "revolutionized probabilistic mortality and population forecasting" (Raftery 2023)



The Lee-Carter method (1992)

- Proposed in 1992 to model and forecast US mortality
- After 30+ years, Lee-Carter (LC) still widely employed by variety of users: governments, private companies, international organizations, ...
- The landmark model in mortality forecasting
- An extrapolation method:
 - model the mortality surface over age and time
 - extrapolate trends in the future, assuming that observed trends will continue
- Simplicity, robustness and objectivity have made the model so successful
- Nonetheless, some limitations of the model have stimulated several extensions over the years (see Basellini et al. (2023) for a recent comprehensive review)



The LC method

ullet A simple log-bilinear functional form for mortality rates $m_{x,t}$ at age x and time t

$$\ln(m_{x,t}) = \alpha_x + \beta_x \kappa_t + \epsilon_{x,t}$$

where:

- α_x is the general shape of log-mortality at age x
- lacksquare is the rate of mortality improvement at age x
- ullet κ_t is the general level of mortality at time t
- $\epsilon_{x,t}$ is the error term with mean 0 and variance σ^2_{ϵ} , reflecting residual agespecific influences not captured by the model
- Modelling log-rates ⇒ fitted and forecast rates constrained to be positive
- Log transformation partially counters heteroscedasticity of observed rates



The LC method

- The model is undetermined: if $m{ heta}_1 = [m{lpha}, m{eta}, m{\kappa}]$ is a solution, then for any scalar c:
 - $m{\theta}_2 = [m{lpha} m{eta} c, m{eta}, m{\kappa} + c]$ is also a solution
 - $m{m{\theta}}_3 = [m{lpha}, m{eta}c, \kappa/c]$ is also a solution
- Two constraints introduced to ensure model identification:

$$\sum_{x} \beta_{x} = 1$$
 and $\sum_{t} \kappa_{t} = 0$



The LC method: a schematic view

$$\ln(m_{x,t}) \simeq \alpha_x + \frac{\beta_x}{\kappa_t} \kappa_t$$

$$egin{pmatrix} \ln(m_{x_1,t_1}) & \ln(m_{x_1,t_2}) & \dots & \ln(m_{x_1,t_n}) \ \ln(m_{x_2,t_1}) & \ln(m_{x_2,t_2}) & \dots & \ln(m_{x_2,t_n}) \ \ln(m_{x_3,t_1}) & \ln(m_{x_3,t_2}) & \dots & \ln(m_{x_3,t_n}) \ & dots & dots & \ddots & dots \ \ln(m_{x_m,t_1}) & \ln(m_{x_m,t_2}) & \dots & \ln(m_{x_m,t_n}) \end{pmatrix} \simeq$$

$$egin{pmatrix} lpha_{x_1} \ lpha_{x_2} \ lpha_{x_3} \ dots \ lpha_{x_m} \end{pmatrix} + egin{pmatrix} eta_{x_2} \ eta_{x_3} \ dots \ eta_{x_m} \end{pmatrix} \left(\kappa_{t_1} & \kappa_{t_2} & \ldots & \kappa_{t_n}
ight)$$

$$\underbrace{n}_{ ext{years}} imes \underbrace{m}_{ ext{ages}} = \underbrace{mn}_{ ext{cells}} \simeq \underbrace{m}_{lpha_i} + \underbrace{m}_{eta_i} + \underbrace{n}_{\kappa_j} - \underbrace{2}_{ ext{constraints}} = \underbrace{2m+n-2}_{ ext{parameters}}$$





Model estimation

• The model is estimated by minimizing the residual sum of squares:

$$\sum_{x,t} \left(\ln(m_{x,t}) - \alpha_x - \beta_x \kappa_t \right)^2 \tag{1}$$

- A singular value decomposition (SVD) is employed to minimize Equation 1 & derive an ordinary least squares (OLS) solution:
 - \hat{lpha}_x is the average of the observed $\ln(m_{x,t})$



Exercise

Exercise

Open your R session. Load the MortSWE. Rdata dataset, and consider only data from 1950 onward for males aged $0 \le x \le 100$. Derive the LC estimate of the α_x parameter

Hint: Replace the only cell with 0 deaths with a tiny number of deaths (e.g., 1), and compute again the matrix of log-mortality rates. This will be particularly helpful later on.

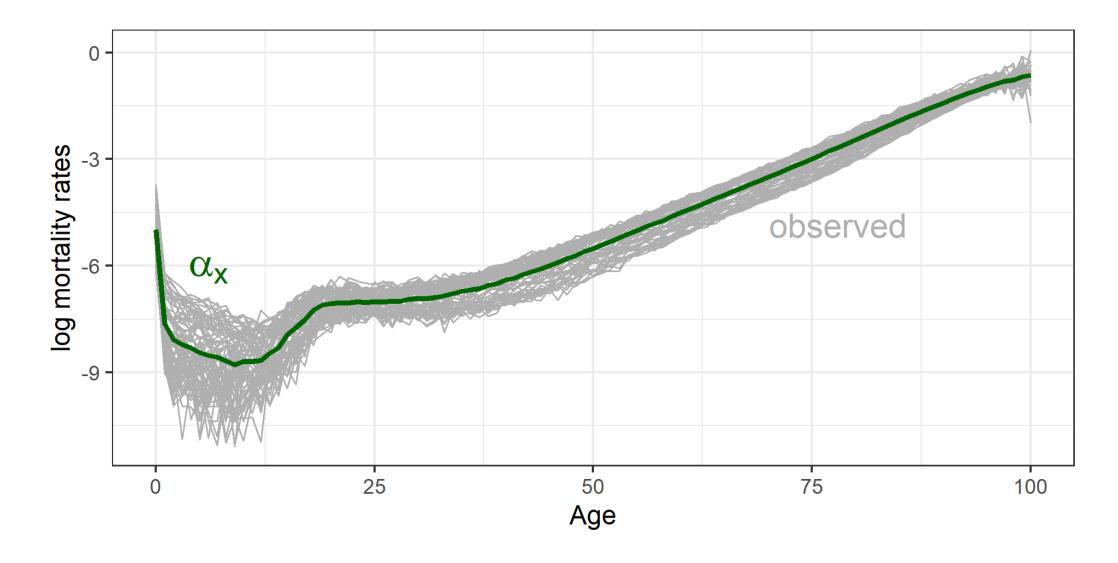


One possible solution

```
## cleaning the workspace
 2 rm(list=ls(all=TRUE))
 3 ## load useful packages
 4 library(tidyverse)
 5 ## loading data
 6 setwd(dirname(rstudioapi::getActiveDocumentContext()$path))
 7 load("data/MORTSWE.Rdata")
 8 ## subset
   my.df \leftarrow MORT.SWE \%>\% filter(Year>=1950, Sex=="Male", Age \leftarrow 100)
10 ## extract data
11 x <- unique (my.df$Age)
12 t <- unique (my.df$Year)
13 n \leftarrow length(t)
14 \text{ m} < - \text{length}(x)
15 ## matrices
16 DEATHS <- matrix (my.df$Deaths, m, n)
17 EXPOS <- matrix (my.df$Exposures, m, n)
18 RATES <- matrix(my.df$Rates,m,n)</pre>
```



One possible solution





Model estimation

The model is estimated by minimizing the residual sum of squares:

$$\sum_{x,t} \left(\ln(m_{x,t}) - \alpha_x - \beta_x \kappa_t \right)^2 \tag{2}$$

- A singular value decomposition (SVD) is employed to minimize Equation 2 & derive an ordinary least squares (OLS) solution:
 - ullet \hat{lpha}_x is the average of the observed $\ln(m_{x,t})$
 - $\hat{m{\beta}}_x$ and $\hat{\kappa}_t$ are the first left- and right-singular vectors of the SVD of the matrix $\ln(m_{x,t}) \hat{lpha}_x$, adjusted to comply with the two constraints

$$\sum_{x} oldsymbol{eta}_{x} = 1$$
 and $\sum_{t} \kappa_{t} = 0$



Exercise

Exercise

Now perform a singular value decomposition of the matrix of "centered" mortality rates $\ln(m_{x,t}) - \hat{\alpha}_x$, and plot the first left- and right-singular vectors

Hint: Use the command svd(x=.,nu=1,nv=1) to perform the SVD of your centred mortality rates matrix, and then extract the vectors u and v from the SVD object. Remember to use the modified log-rates matrix (which excludes infinite values).

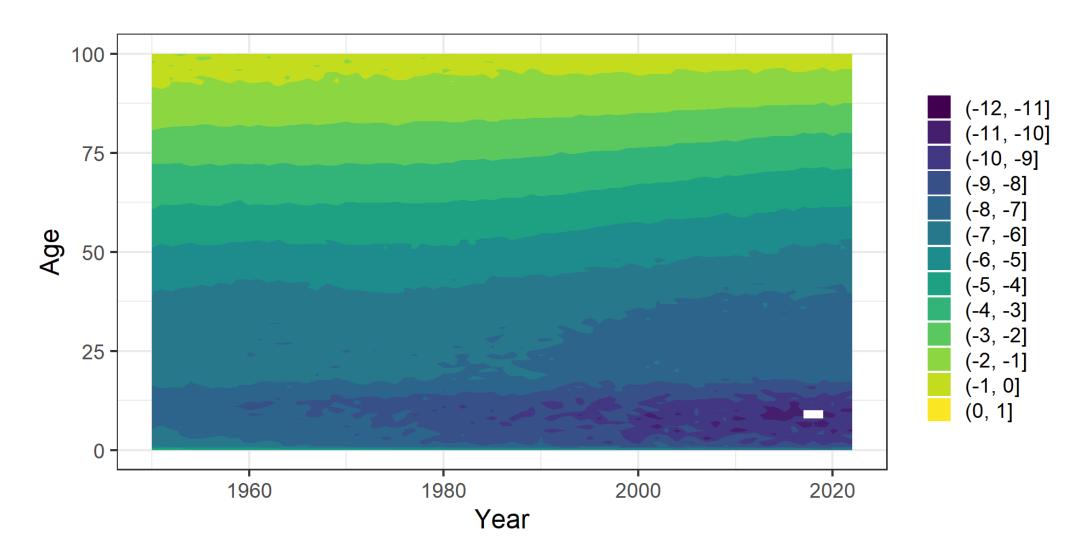


One possible solution

```
1 ##--- step 2: derive BETA and KAPPA ----
 2 ## derive matrix of residuals ("centred" mortality rates)
 3 LMXres <- LRATES1 - Alpha
 4 ## plotting the matrix
 5 library(fields)
 6 image.plot(t,x,t(LMXres),axes="F",xlab="",ylab="")
7 axis(1);axis(2,las=2);box()
 8 mtext("ages", side=2, line=2); mtext("year", side=1, line=2)
 9 dev.off()
10 ## performing the SVD
11 LCsvd <- svd(LMXres, nu=1, nv=1)
12 ## extract first left- and right-singular vectors of svd
13 Beta <- c(LCsvd$u)
14 Kappa1 \leftarrow c(LCsvd$v)
15 ## plotting singular vectors
16 par (mfrow=c(1,2))
17 plot (x, Beta)
18 plot(t, Kappa1)
    ## : a a 1 -- al : a a - + b a a a a a + - - a - : a +
```

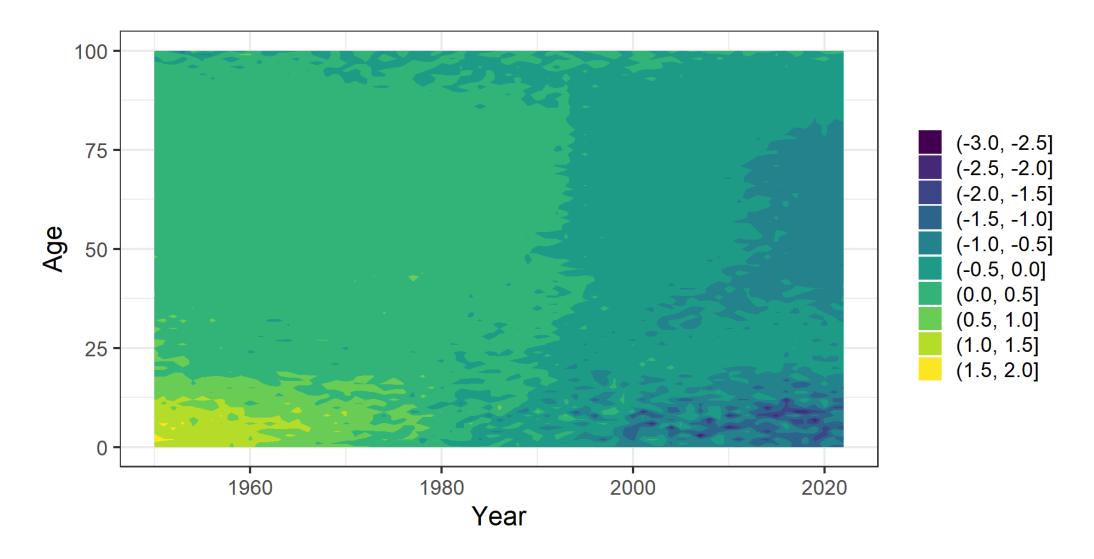


Log-mortality rates



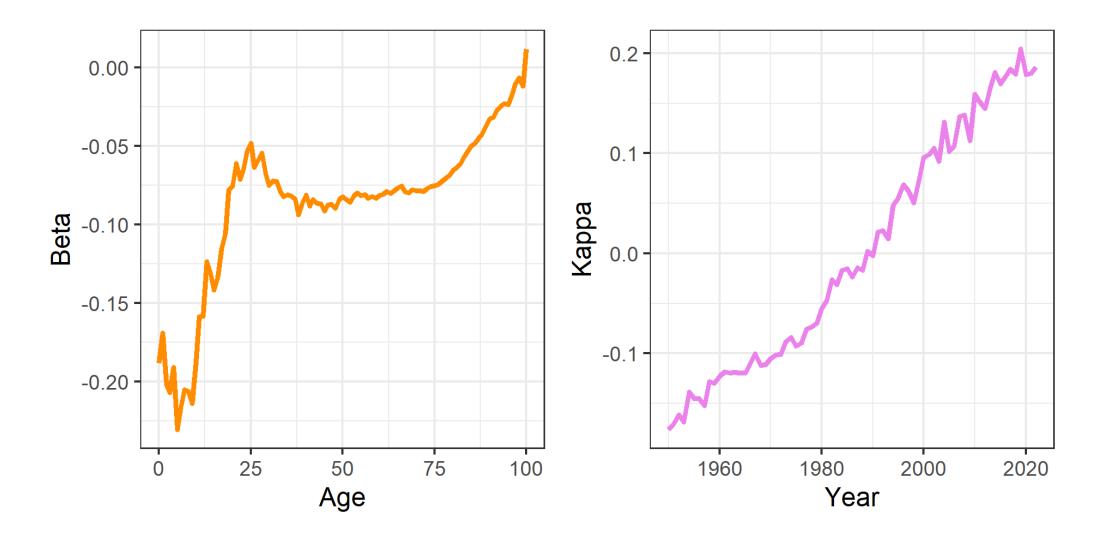


Centred log-mortality rates



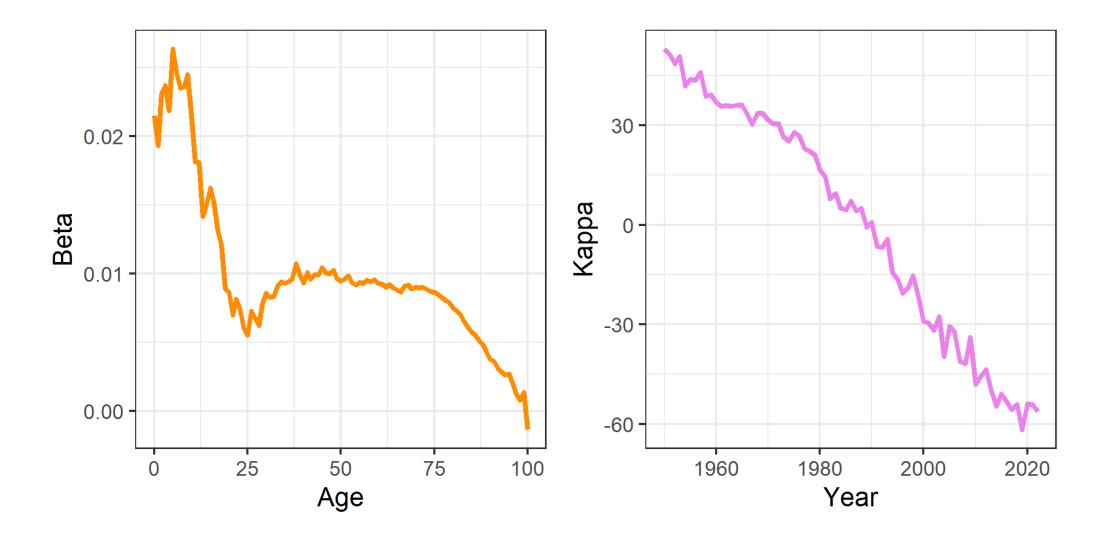


\hat{eta}_x and $\hat{\kappa}_t$ from SVD





\hat{eta}_x and $\hat{\kappa}_t$ with constraints





Second step estimation of $\hat{\kappa}_t$

• In a second-step estimation, $\hat{\kappa}_t$ is adjusted so that the fitted deaths match the observed deaths in all years, i.e.

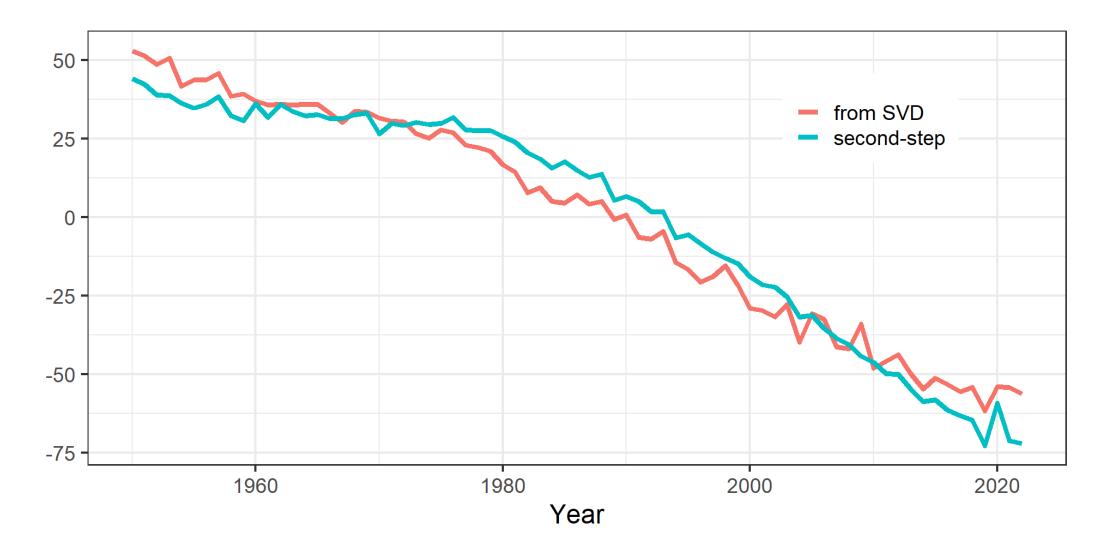
$$\sum_x \hat{y}_{x,t} = \sum_x y_{x,t} \quad ext{for all } t$$



One possible solution

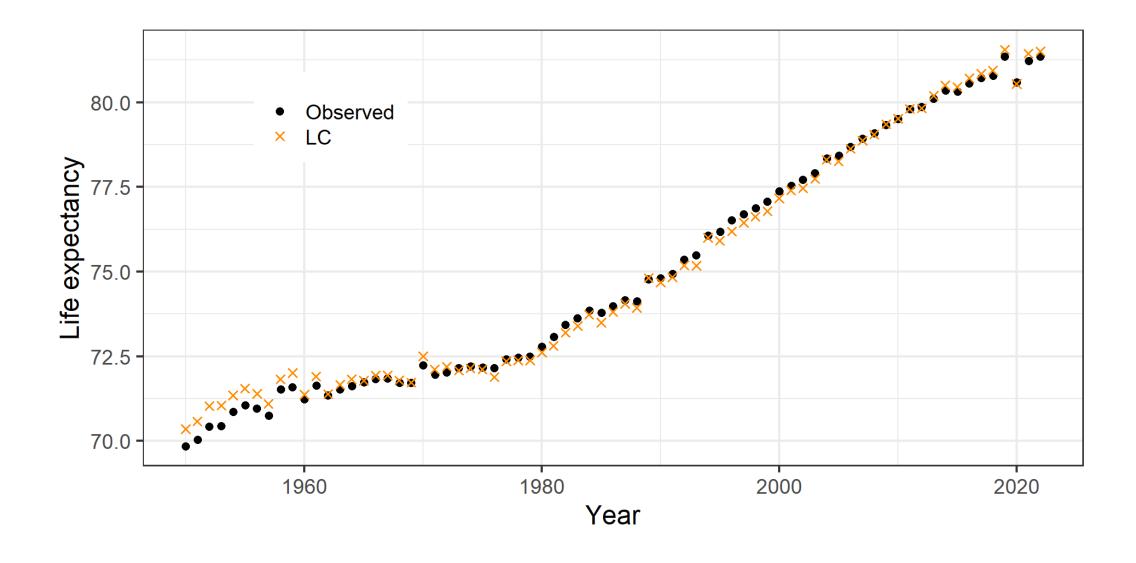
```
##--- step 3: adjust KAPPA ----
  ## function to compute difference between observed and fitted LC deaths
   koptim <- function(par,alpha,beta,sum.dx,Exp) {</pre>
     kappa <- par[1]
     lmx.lc <- alpha+beta*kappa</pre>
     z.lc <- exp(lmx.lc) *Exp
    sum.z.lc <- sum(z.lc)
    diff.lc <- abs(sum.dx-sum.z.lc)
     return (diff.lc)
10
   ## adjust Kappa every year
   Kappa <- numeric(n)</pre>
   for (i in 1:n) {
14
     KappaSecStep \leftarrow optimize(f=koptim,interval=c(-100,100),alpha=Alpha,
    beta=Beta, sum.dx=sum(DEATHS[,i]), Exp=EXPOS[,i])
16
    Kappa[i] <- KappaSecStep$minimum</pre>
17 }
18 ## plotting
```

Second step estimation of $\hat{\kappa}_t$





Goodness-of-fit





Intermezzo: animated visualizations & shiny-app

In some instances, it can be useful to introduce dynamic plots, animations and flexible outputs in your presentations or documents:

- can grab the audience's attention
- show time-series evolutions
- to describe your model
- sensitivity analysis of your results



Animated images

- \animategraphics in Beamer (see LaTeX workshop)
- gganimate package in Quarto

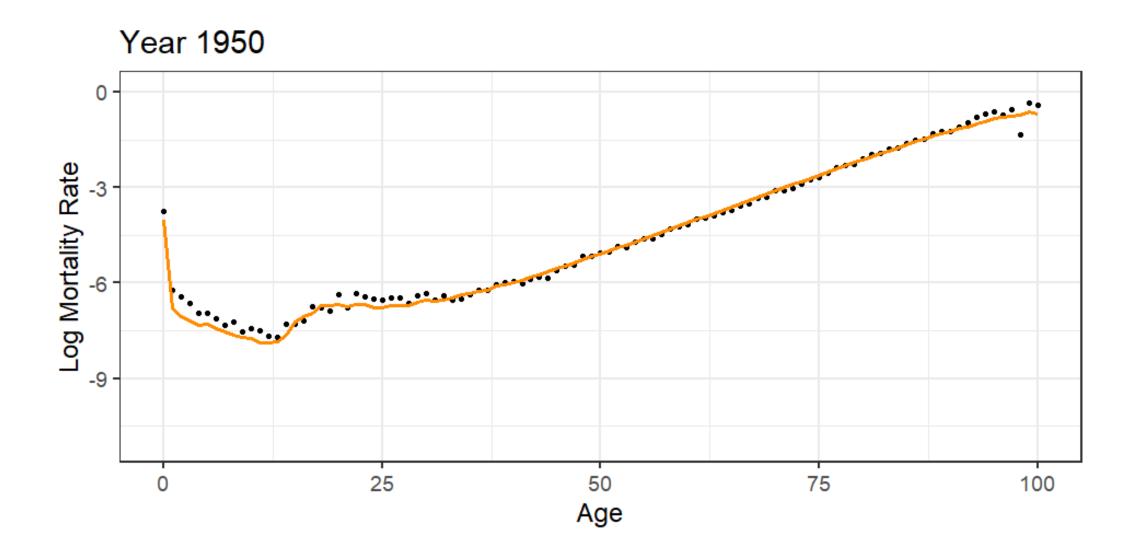


Animation plots with gganimate

```
1 ## fitted log-mortality
 2 Ones \leftarrow matrix(1,n)
 3 ETAlc <- Alpha%*%t(Ones) + Beta%*%t(Kappa)</pre>
 4 ## basic plot
  q <- my.df %>%
     mutate(logRates=case when(
       is.infinite(logRates)~NA,
    TRUE~logRates),
    Fitted=c(ETAlc)) %>%
10
    ggplot(aes(x=Age, group=Year))+
11
    geom point(aes(y=logRates))+
12
    geom line(aes(y=Fitted), color="darkorange", linewidth=1.2) +
13
    scale color viridis c() +
    theme_bw(base size = 18) +
14
     labs (y= "Log Mortality Rate")
   ## animating with gganimate
  library(gganimate)
18 gg <- g + transition time(Year) +
```



Goodness-of-fit





Shiny apps

- shiny is an R package that makes it easy to build interactive web apps straight from R
- keep them on webpages or embed them in R Markdown documents or Quarto presentations
- a user-friendly interface to interact with your R analysis and show your results (my own example)
- it is composed by
 - an UI (user interface), where you can create the inputs for your and decide the outputs to display
 - a server, where you assemble the outputs from your given inputs
 - the shinyApp, putting the two together
- to learn more, visit https://shiny.posit.co/ to get started, plenty of videos and written tutorials

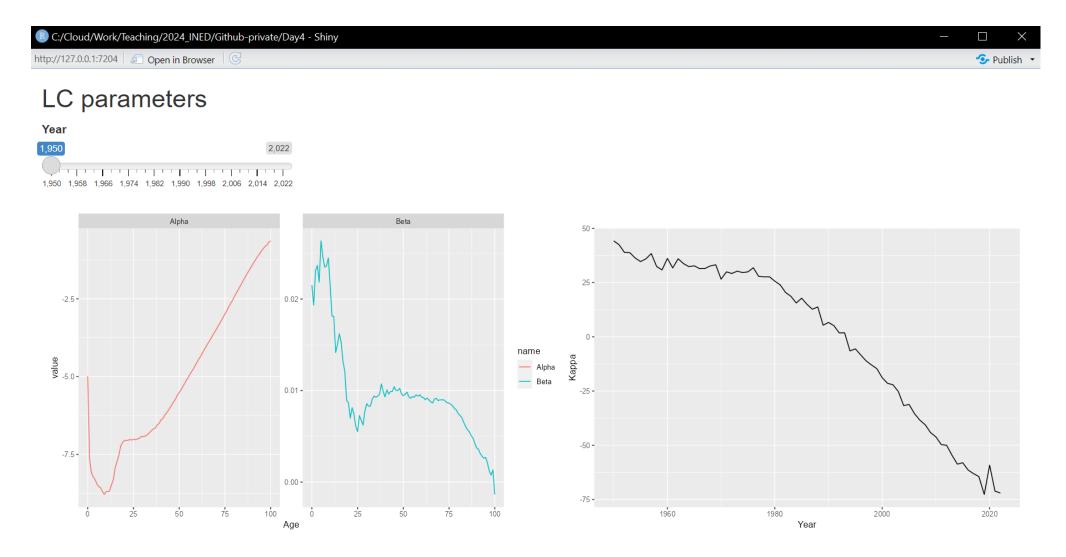


Shiny app: an example

```
##--- SHINY APP EXAMPLE ----
 3 ## cleaning the workspace
 4 rm(list=ls(all=TRUE))
 5 ## load useful packages
 6 library(tidyverse)
7 library (patchwork)
 8 library(shiny)
  ## loading data
10 setwd(dirname(rstudioapi::getActiveDocumentContext()$path))
11 load ("data/MORTSWE.Rdata")
12 ## subset
13 my.df <- MORT.SWE %>% filter(Sex=="Male", Age <= 100)
14
   ## build your user interface
16 ui <- fluidPage(
17 ## title of your shiny
18 titlePanel ("LC parameters"),
```



Shiny app: an example





Forecasting with LC

- Forecasting "made simple": choose an appropriate time-series model for $\hat{\kappa}_t$ and extrapolate it
- The forecast $\hat{\kappa}_{T+h}$ allows one to derive the entire age-pattern of mortality at time T+h:

$$\ln(\hat{m}_{x,T+h}) = \hat{lpha}_x + \hat{eta}_x \hat{\kappa}_{T+h}$$

- LC suggest a random walk model (i.e. ARIMA(0,1,0)) with drift: $\kappa_t = \kappa_{t-1} + c + e_t$ where c is a constant (drift) and e_t the error term (purely random process)
- ullet For this time-series model: $\hat{\kappa}_{T+h|T} = \hat{\kappa}_T + c\,h$
- Simulated future trajectories of \hat{k}_{T+h} to construct prediction intervals for $\hat{m}_{x,T+h}$ and other summary measures (e.g., $\hat{e}_{0,T+h}$)
- (Coale and Guo (1989) adjustment for forecast rates at ages 85+)

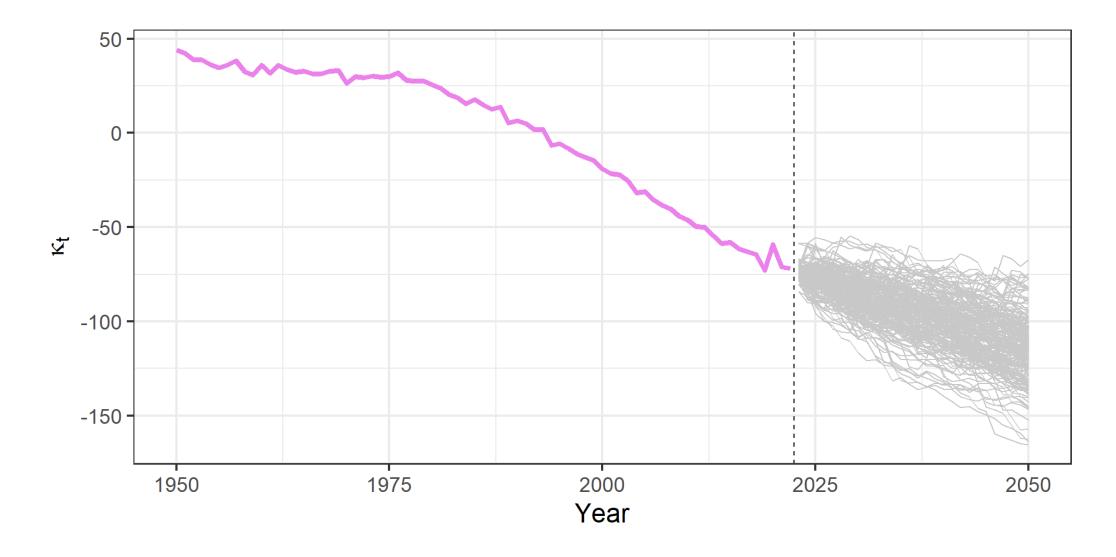


Forecasting with LC: a schematic view



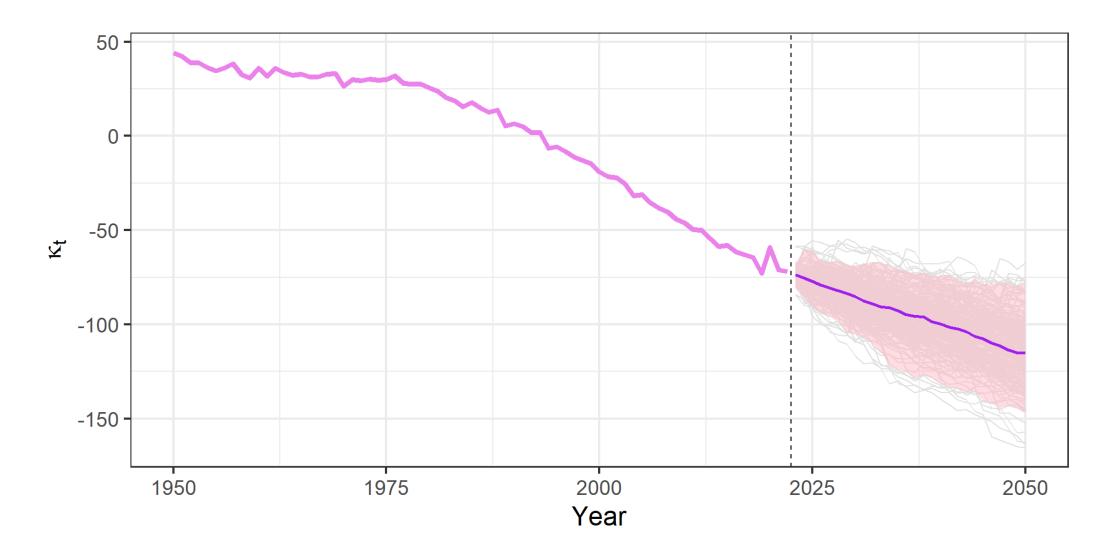


$\hat{\kappa}_t$ simulations



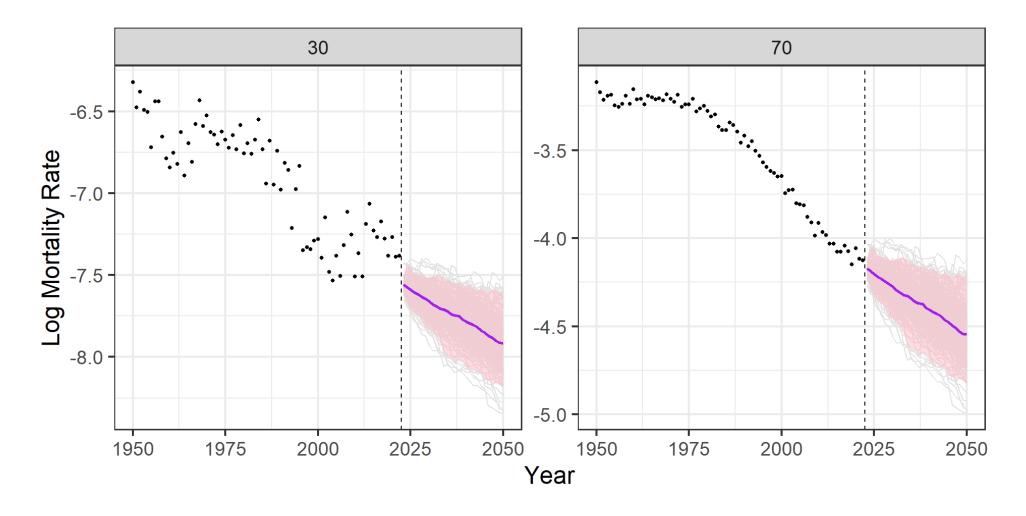


$\hat{\kappa}_t$ 95% PIs



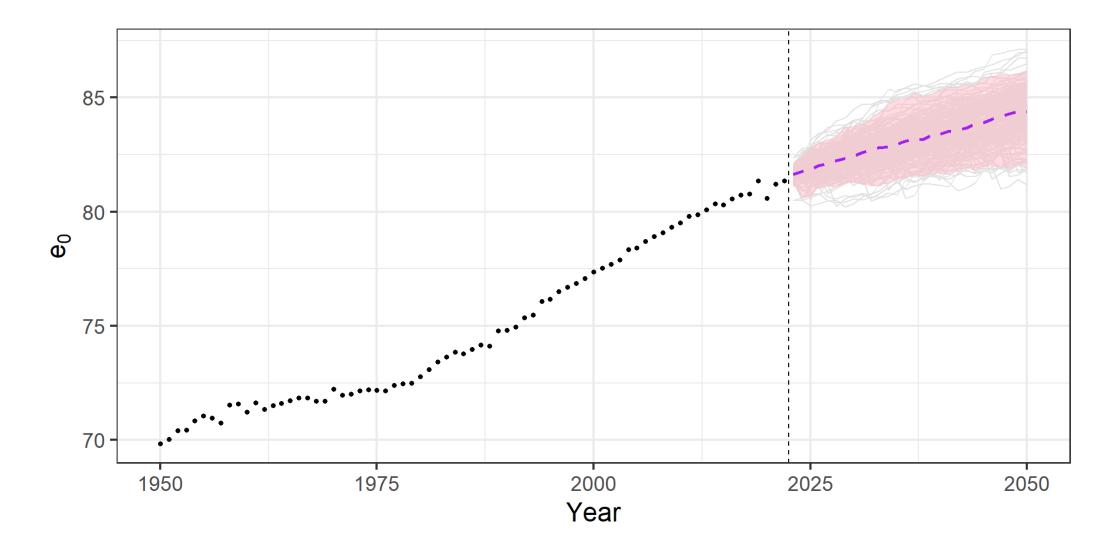


From $\hat{\kappa}_t$ to $\ln(\hat{m}_{xt})$





From $\hat{\kappa}_t$ to \hat{e}_0





The LC model: a summary

Advantages:

- simple & powerful method: forecast rates derived by modeling single time index by standard time-series model
- linear time index very often captures well historical decline in mortality
- stochastic model ⇒ probabilistic intervals
- extrapolative approach, no expert opinions more accurate than previous methodologies

• Disadvantages:

- jump-off bias
- Normality assumption (from SVD)
- jagged fitted and forecast age profile, lacking smoothness
- fixed age-pattern of mortality decline
- rigid structure
- Several extensions proposed to overcome some of these issues



Some LC extensions

- "jump-off" bias: use observed jump-off rates (Lee and Miller 2001)
- Normality assumption: Poisson LC (Brouhns et al. 2002)
- jagged fitted and forecast age profile, lacking smoothness: Smooth LC (Delwarde et al. 2007)
- fixed age-pattern of mortality decline: Li et al. (2013)
- rigid structure



Other LC extensions (single pop)

- Booth et al. (2002): adjusting κ_t to match the age-at-death distribution & determining optimal fitting period
- Renshaw and Haberman (2003): adding more than one principal components, i.e. $\ln(m_{x,t})=lpha_x+\sum_k eta_x^k \kappa_t^k$
- Koissi et al. (2006): residual bootstrap to include parameter uncertainty in forecasts
- Renshaw and Haberman (2006): including cohort effects, i.e.

$$\ln(m_{x,t}) = lpha_x + eta_x^{(1)} \kappa_t + eta_x^{(0)} \gamma_{t-x}$$

- Hyndman and Ullah (2007): smooth underlying data (functional data) & additional principal components
- Camarda and Basellini (2021): smoothing, decomposing and forecasting the three components of mortality (childhood, early-adulthood and senescence), i.e.

$$m_{x,t} = \sum_k \expig(lpha_x^k + eta_x^k \kappa_t^kig)$$

• for a comprehensive review, see Basellini et al. (2023)



Forecast evaluation

- How do we evaluate demographic forecasts?
- Option 1: wait until the future occurs (e.g., 2050) to find out
 - clearly not ideal :)
- Option 2: out-of-sample forecasting
 - from the observed time series $\boldsymbol{y} = [y_1, y_2, \dots, y_{n-h}, \dots y_T]$, suppose that you have observed only the first n-h data points, forecast h years in the future and compare the forecasts with the withheld data points $y_{n-h+1}, \dots y_T$
 - y can refer to any forecast measure (log-rates, rates, summary measures)



Point forecast accuracy

- Let $e_t=\hat{y}_t-y_t$ denote the forecast error, with \hat{y}_t being the central forecast and y_t the withheld data point t
- ullet Let $p_t=100e_t/y_t$ be the percentage error
- Compute one of the following measures of point forecast accuracy:
 - lacksquare Mean Square Error (MSE) $= rac{1}{h} \sum_{t=1}^h e_t^2$
 - Root MSE (RMSE) = $\sqrt{\text{MSE}}$
 - lacksquare Mean Absolute Error (MAE) $=rac{1}{h}\sum_{t=1}^{h}|e_t|$
 - lacksquare Mean Absolute Percentage Error (MAPE) $=rac{1}{h}\sum_{t=1}^{h}|p_t|$
- The smaller the error, the better

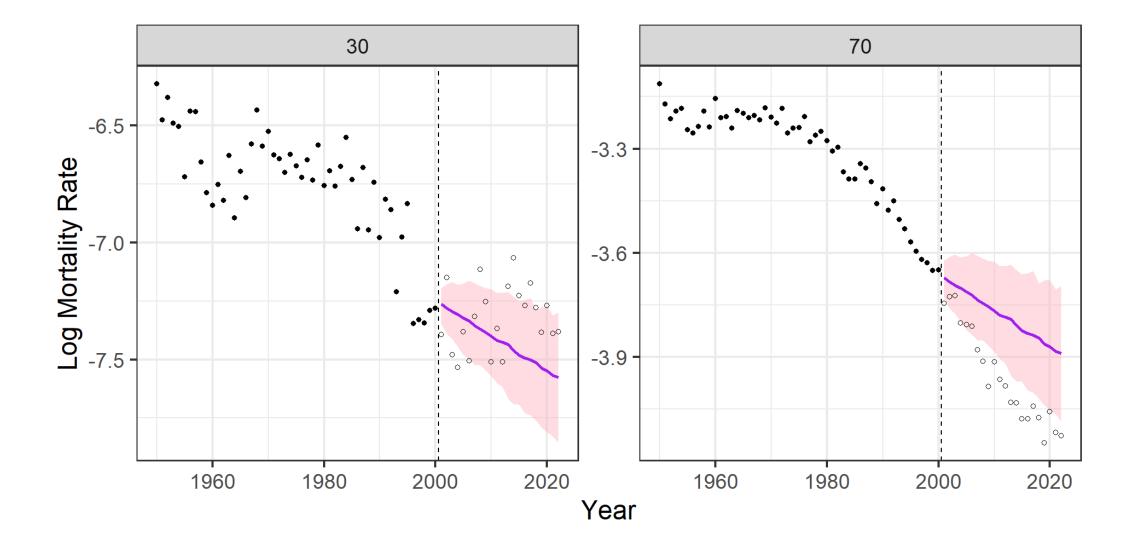


Interval forecast accuracy

- Compare withheld data points with the forecast α % prediction interval (PI):
 - Empirical Coverage Probability (ECP): proportion of observations falling into the PI
 - lacktriangledown Coverage Probability Deviance (CPD) $= |lpha \mathrm{ECP}|$, with $\mathrm{CPD} \in [0,lpha]$
- The smaller the CPD, the better

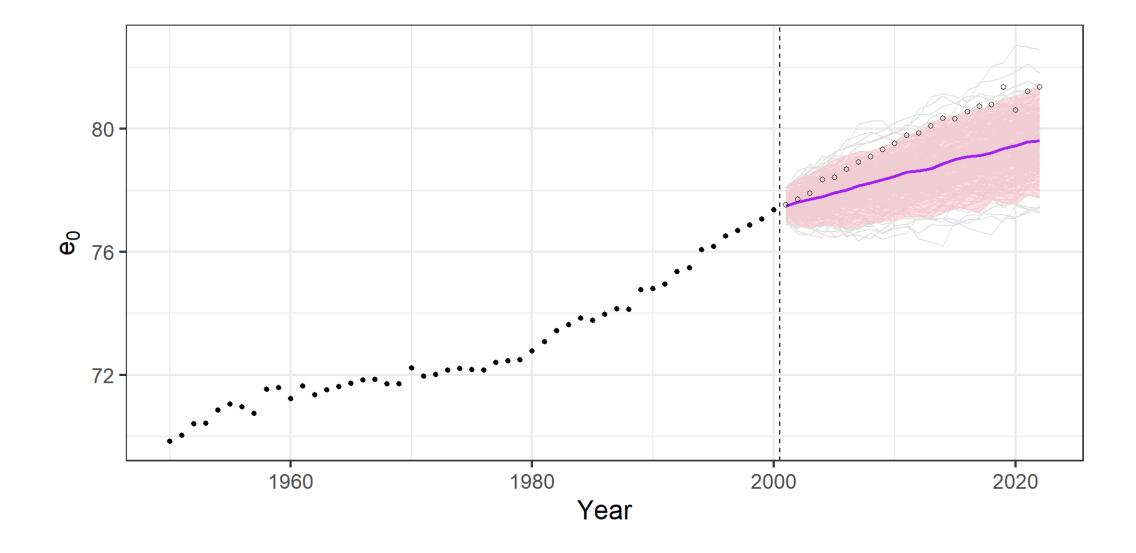


LC forecast evaluation





LC forecast evaluation





Day 4 assignment

Assignment

- 8. Load the mortality data MORTSWE . Rdata, and focus on male mortality from 1950 to 2000 for ages $x \leq 100$. Fit and forecast mortality with the LC method, and produce forecasts up to 2022. Evaluate the forecast accuracy of the method by computing a point forecast accuracy measure of your choice on a single measure of interest (e.g. log-rates or e_o).
- 9. Consider again the setting of Exercise 8 (i.e. male mortality from 1950 to 2000 for ages $x \leq 100$). Compare the forecast accuracy of two different methods that we have seen during this course. Which model is more accurate?



