

# Lecture on Penalized Regression

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# Issues Due to High Dimensionality

- The least squares estimates have relatively low bias and low variability especially when the relationship between  $Y$  and  $X$  is linear and the number of observations  $n$  is way bigger than the number of predictors  $p$ .
- But, when  $n \approx p$ , the least squares fit can have high variance and may result in over fitting and poor estimates on unseen observations.
- And, when  $n < p$ , the variability of the least squares fit increases dramatically, and the variance of these estimates is infinite



# Example in Bioinformatics

- Genome-Wide Association Studies
  - There are millions of SNPs predictors  $n \ll p$ .
- RNA Sequencing Data
  - $\approx 20,000$  protein coding genes  $n \approx p$ .

# Information Sparsity

- When we have a large number of variables  $X$  in the model there will generally be many that have little or no effect on  $Y$ .
- Leaving these variables in the model makes it harder to see the “big picture”, i.e., the effect of the “important variables”.
- The model would be easier to interpret by removing (i.e. setting the coefficients to zero) the unimportant variables.

# Solutions

- Subset Selection
  - Identifying a subset of all  $p$  predictors  $X$  that we believe to be related to the response  $Y$ , and then fitting the model using this subset
  - E.g. best subset selection and stepwise selection

# Solutions

- Dimension Reduction
  - Involves projecting all  $p$  predictors into an  $M$ -dimensional space where  $M < p$ , and then fitting linear regression model
  - E.g. Principle Components Regression

# Solutions

- Shrinkage
  - Involves shrinking the estimates coefficients towards zero
  - This shrinkage reduces the variance
  - Some of the coefficients may shrink to exactly zero, and hence shrinkage methods can also perform variable selection
  - E.g. Ridge regression, Lasso, and Elastic Net Model



# Penalized Regression

- Ordinary Least Squares (OLS) estimates by minimizing

$$RSS = \sum_{i=1}^n \left( y_i - \beta_0 - \sum_{j=1}^p \beta_j x_{ij} \right)^2$$

- Ridge Regression

$$\sum_{i=1}^n \left( y_i - \beta_0 - \sum_{j=1}^p \beta_j x_{ij} \right)^2 + \lambda \sum_{j=1}^p \beta_j^2 = RSS + \lambda \sum_{j=1}^p \beta_j^2$$

- Lasso Regression

$$\sum_{i=1}^n \left( y_i - \beta_0 - \sum_{j=1}^p \beta_j x_{ij} \right)^2 + \lambda \sum_{j=1}^p |\beta_j| = RSS + \lambda \sum_{j=1}^p |\beta_j|$$

# Another Formulation

Lasso (6.8) and Ridge (6.9) can be formulated as below with a constraint associated with parameter  $s$ .

$$\underset{\beta}{\text{minimize}} \left\{ \sum_{i=1}^n \left( y_i - \beta_0 - \sum_{j=1}^p \beta_j x_{ij} \right)^2 \right\} \quad \text{subject to} \quad \sum_{j=1}^p |\beta_j| \leq s$$

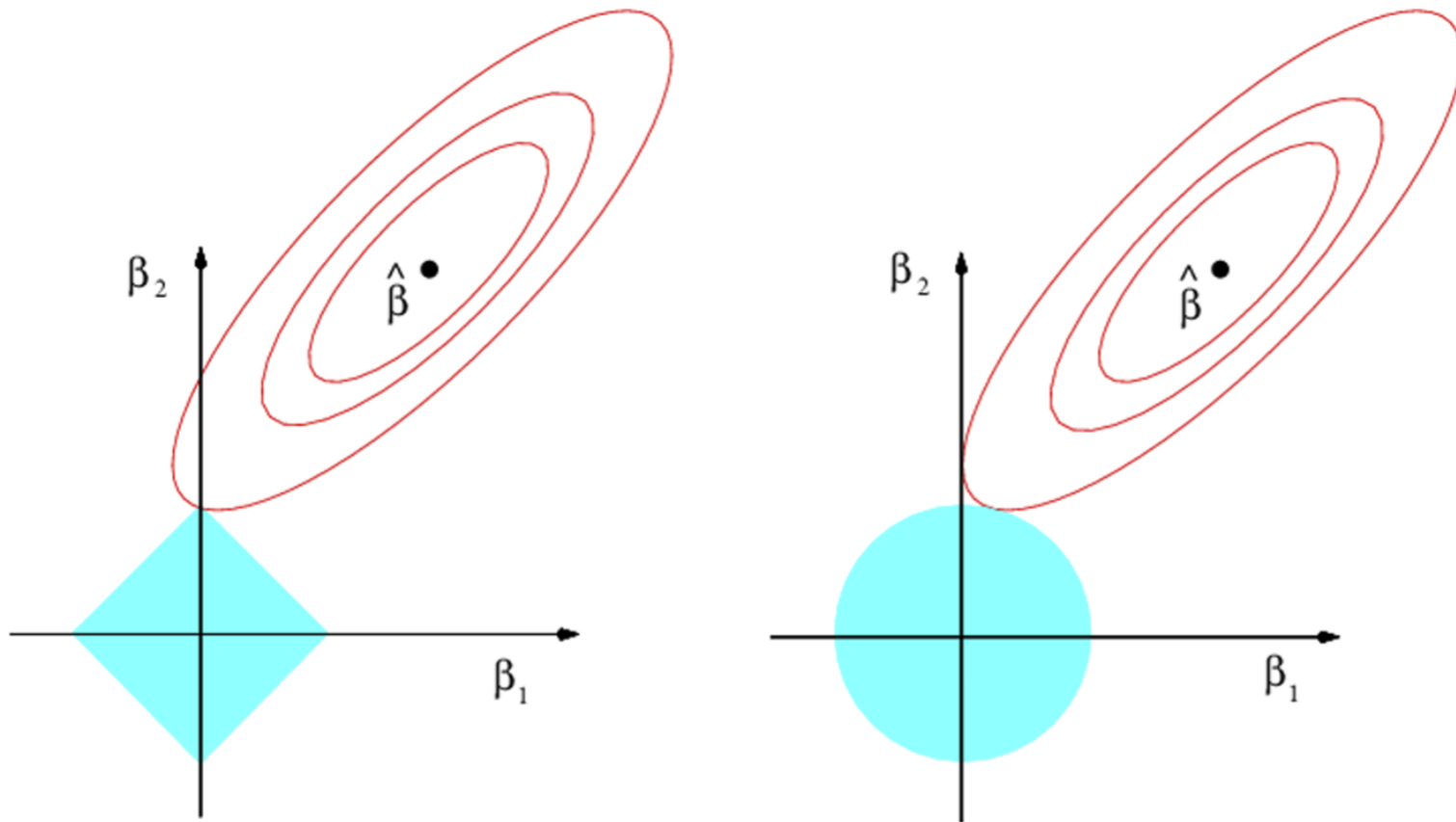
(6.8)

and

$$\underset{\beta}{\text{minimize}} \left\{ \sum_{i=1}^n \left( y_i - \beta_0 - \sum_{j=1}^p \beta_j x_{ij} \right)^2 \right\} \quad \text{subject to} \quad \sum_{j=1}^p \beta_j^2 \leq s,$$

(6.9)

# Geometry Illustration



**FIGURE 6.7.** Contours of the error and constraint functions for the lasso (left) and ridge regression (right). The solid blue areas are the constraint regions,  $|\beta_1| + |\beta_2| \leq s$  and  $\beta_1^2 + \beta_2^2 \leq s$ , while the red ellipses are the contours of the RSS.

# R Package

glmnet **4.1-7**

Get started

Reference

Articles

Changelog

## An Introduction to glmnet

Trevor Hastie

Junyang Qian

Kenneth Tay

March 27, 2023

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# R Package

This vignette describes basic usage of glmnet in R. There are additional vignettes that should be useful:

- “Regularized Cox Regression” describes how to fit regularized Cox models for survival data with `glmnet`.
- “GLM family functions in glmnet” describes how to fit custom generalized linear models (GLMs) with the elastic net penalty via the `family` argument.
- “The Relaxed Lasso” describes how to fit relaxed lasso regression models using the `relax` argument.

`glmnet` solves the problem

$$\min_{\beta_0, \beta} \frac{1}{N} \sum_{i=1}^N w_i l(y_i, \beta_0 + \beta^T x_i) + \lambda \left[ (1 - \alpha) \|\beta\|_2^2 / 2 + \alpha \|\beta\|_1 \right],$$