$$(Y_1, ..., Y_n \sim N(M, \sigma^2), \sigma^2 \text{ is known.}$$
 $M \sim N(M_0, \frac{\sigma^2}{K_0})$  (Prior)

 $(Prior \text{ prameters})$ 

# Bayes Estimators

$$P(ul_{1},...,n_{n},\sigma^{2}) \sim N(ul_{n},\sigma^{2})$$

$$u_{1} = wy + (1-w)u_{0}, \quad w = \frac{n}{n+k},$$

$$\sigma^{2} = w\sigma^{2}$$

$$n+k_{0} = w\sigma^{2}$$

E[ $M|y_1, y_1, \sigma^2$ ] =  $W\overline{y} + (1-w)M_0$  (estimate)

What are the frequentist properties

of  $E[M|Y_1, Y_1, \sigma^2]$  estimator...

Capital =

FIRE By Say 15

#### **Estimators: Bayes / Frequentist Unification**

- Bayesian inference provides a straightforward procedure for producing estimators given your prior beliefs.
  - 1. Compute posterior distribution
  - 2. Summarize the posterior distribution with a point estimator (e.g. posterior mean or posterior mode) and a probability interval
- Frequentists provide tools for evaluating the sampling properties of an estimator.
  - Bias, variance and MSE of an estimator
  - Well-calibrated probability intervals
- Both are useful!

#### The Bias-Variance Tradeoff

Reminder: an estimator is a random variable, an estimate is a constant

- *Bias*: systematic sampling error of the estimator
- *Variance*: variance of the estimator (from sampling & measurement error)
- Often we evaluate an estimator in terms of mean square error:  $MSE(\hat{\theta}) = E_Y(\hat{\theta} \theta)^2$
- The Bias-Variance tradeoff:  $MSE(\hat{\theta}) = Var(\hat{\theta}) + Bias(\hat{\theta})^2$

#### The Bias-Variance Tradeoff

- Variance of an estimator comes sampling from a population
  - If you were to repeatedly draw new samples of the same size how much would your estimates vary?
  - $\circ \,\,$  e.g. if  $y_i \sim N(\mu, \sigma^2)$  then  $\mathrm{Var}(ar{Y}) = \sigma^2/n$

$$\mathcal{J}_{MLE} = \vec{Y}$$

$$\mathcal{E}[\vec{Y}] = \mathcal{U} \quad (unbiased)$$

$$Var(\vec{Y}) = Var(\vec{X}) = \frac{1}{N} \mathcal{E} Var(\vec{Y})$$

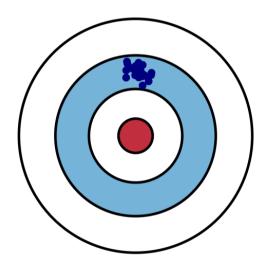
$$\mathcal{E}(\vec{Y}) = \frac{1}{N} \mathcal{E} Var(\vec{Y})$$

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#### Bias

The expected difference between the estimate and the response

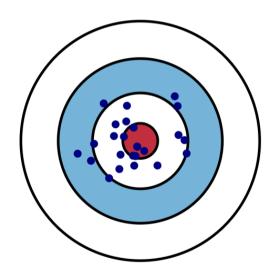


Statistical definition of bias:

$$E_Y[\hat{ heta}- heta]$$

#### Variance

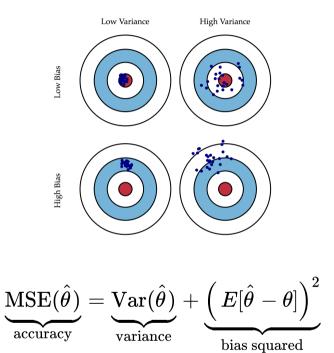
How variable is the prediction about its mean?



Statistical definition of variance:

$$E_Y[\hat{ heta}-E_Y[\hat{ heta}]]^2$$

#### Bias and Variance



#### The Bias-Variance Tradeoff

- The prior distribution (usually) makes your estimator biased...
- But the prior distribution also (usually) reduces the variance!
- Example: compute the frequentist mean and variance of the posterior mean.

$$E[\mathcal{M}|Y_{i,1}, Y_{i,n}, \sigma^{2}] = \mathcal{M}_{PM} = \mathcal{W}\overline{Y} + (1-\mathcal{W})\mathcal{M}_{0}$$

$$Bias: E[\mathcal{W}\overline{Y} + (1-\mathcal{W})\mathcal{M}_{0} - \mathcal{M}] = \mathcal{W}E[\overline{Y}] + (1-\mathcal{W})E[\mathcal{M}_{0}] - \mathcal{M}$$

$$= \mathcal{W}\mathcal{M} + (1-\mathcal{W})\mathcal{M}_{0} - \mathcal{M} = (1-\mathcal{W})(\mathcal{M}_{0} - \mathcal{M})$$

$$\frac{V\omega}{v}: V\omega(u)/(v,v_0) = V\omega(w) + (1-w)M_0 = \frac{v^2 V\omega(v)}{v^2 + (1-w)^2 V\omega(v)} = 0 \quad (const)$$

$$= W^2 Var(\overline{Y}) \leq Var(\overline{X}) = Var(\widehat{\mathcal{U}}_{MLE})$$

#### **Example: IQ scores**

- Scoring on IQ tests is designed to yield a N(100, 15) distribution for the general population
- We observe IQ scores for a sample of n individuals from a particular town and estimate  $\mu$ , the town-specific IQ score
- If we lacked knowledge about the town, a natural choice would be  $\mu_0=100$
- Suppose the true parameters for this town are  $\mu=112$  and  $\sigma=13$ 
  - The town is smarter on average than the general population

#### **Example: IQ scores**

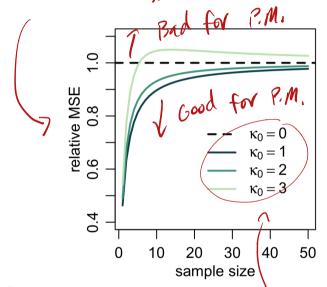
- What is the mean squared error of the MLE? MSE of the posterior mean?
- ullet  $ext{MSE}[\hat{\mu}_{MLE}] = ext{Var}[\hat{\mu}_{MLE}] = rac{\sigma^2}{n} = rac{169}{n}$
- ullet MSE $[\hat{\mu}_{PM}| heta_0]=w^2rac{169}{n}+(1-w)^2144$
- Reminder:  $w = \frac{n}{\kappa_0 + n}$ . For what values of n and  $\kappa_0$  is the MSE smaller for the posterior mean estimator than the maximum likelihood?

$$\frac{7}{169} > \frac{169}{194} + (1-w)^{2}144 = 78\pi s^{2} = (1-w)^{2}144$$

MSE ( $\frac{1}{2}$  MSE ( $\frac{1}{2}$  MSE ( $\frac{1}{2}$  MSE)

MSE ( $\frac{1}{2}$  MSE)

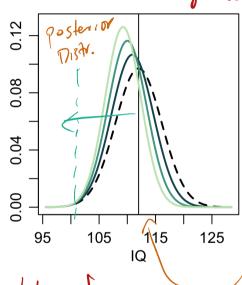
#### **Example: IQ scores**



0 10 20 30 40 sample size

MSE PM

N-90, W-71 eg. Dem-).



Strength of
prior knowledge
(# of pseudo - obs,
tetermines w)

## **Decision Theory**

#### Why the posterior mean?

- Often times we need to make a "decision" by providing a single estimate
- The posterior provides a full distribution over  $\theta$ , which can be summarized in infinitely many ways
- Specify a *loss function* which describes the cost of estimating  $\hat{\theta}$  when the truth is  $\theta$

### **Bayes Estimators**

penalize the difference Letween estimate of truth

- The loss function:  $L(\hat{\theta}, \theta)$ 
  - Squared error:  $L(\hat{\theta}, \theta) = (\hat{\theta} \theta)^2$ Absolute error:  $L(\hat{\theta}, \theta) = |\hat{\theta} \theta|$
- The **Bayes risk** is the posterior expected loss:

Squared error: 
$$L(\hat{\theta}, \theta) = (\hat{\theta} - \theta)^2$$
• Absolute error:  $L(\hat{\theta}, \theta) = |\hat{\theta} - \theta|$ 
• The **Bayes risk** is the posterior expected loss:
$$E_{\theta|y}[L(\hat{\theta}, \theta)] = \int L(\hat{\theta}, \theta)p(\theta \mid y)d\theta$$

$$E_{\theta|y}[L(\hat{\theta}, \theta)] = \int L(\hat{\theta}, \theta)p(\theta \mid y)d\theta$$
Choose an estimator of  $\theta$  based on minimizing the Bayes risk.

- Choose an estimator of  $\theta$  based on minimizing the Bayes risk.
- An estimator  $\theta$  is said to be a **Bayes estimator** if it minimizes the Bayes risk among all estimators.

$$\min_{\hat{ heta}} E_{ heta \mid y} (\hat{ heta} - heta)^2 = \min_{\hat{ heta}} \ \int (\hat{ heta} - heta)^2 p( heta \mid y) d heta$$

Differențiate with respect to  $\hat{\theta}$  and set equal to zero:

Tentrate with respect to 
$$\theta$$
 and set equal to zero.

$$\frac{d}{d\theta} \int_{0}^{\infty} \frac{(\partial - \theta)^{2} P(\theta | y) d\theta}{\int_{0}^{\infty} \frac{d\theta}{d\theta}} = \frac{(\partial - \theta)^{2} P(\theta | y) d\theta}{\int_{0}^{\infty} \frac{d\theta}{d\theta}} = \frac{(\partial - \theta) P(\theta | y) d\theta}{\int_{0}^{\infty} \frac{d\theta}{d\theta}} = \frac{(\partial - \theta) P(\theta | y) d\theta}{\int_{0}^{\infty} \frac{d\theta}{d\theta}} = \frac{(\partial - \theta) P(\theta | y) d\theta}{\int_{0}^{\infty} \frac{d\theta}{d\theta}} = \frac{(\partial - \theta) P(\theta | y) d\theta}{\int_{0}^{\infty} \frac{d\theta}{d\theta}} = \frac{(\partial - \theta) P(\theta | y) d\theta}{\int_{0}^{\infty} \frac{d\theta}{d\theta}} = \frac{(\partial - \theta) P(\theta | y) d\theta}{\int_{0}^{\infty} \frac{d\theta}{d\theta}} = \frac{(\partial - \theta) P(\theta | y) d\theta}{\int_{0}^{\infty} \frac{d\theta}{d\theta}} = \frac{(\partial - \theta) P(\theta | y) d\theta}{\int_{0}^{\infty} \frac{d\theta}{d\theta}} = \frac{(\partial - \theta) P(\theta | y) d\theta}{\int_{0}^{\infty} \frac{d\theta}{d\theta}} = \frac{(\partial - \theta) P(\theta | y) d\theta}{\int_{0}^{\infty} \frac{d\theta}{d\theta}} = \frac{(\partial - \theta) P(\theta | y) d\theta}{\int_{0}^{\infty} \frac{d\theta}{d\theta}} = \frac{(\partial - \theta) P(\theta | y) d\theta}{\int_{0}^{\infty} \frac{d\theta}{d\theta}} = \frac{(\partial - \theta) P(\theta | y) d\theta}{\int_{0}^{\infty} \frac{d\theta}{d\theta}} = \frac{(\partial - \theta) P(\theta | y) d\theta}{\int_{0}^{\infty} \frac{d\theta}{d\theta}} = \frac{(\partial - \theta) P(\theta | y) d\theta}{\int_{0}^{\infty} \frac{d\theta}{d\theta}} = \frac{(\partial - \theta) P(\theta | y) d\theta}{\int_{0}^{\infty} \frac{d\theta}{d\theta}} = \frac{(\partial - \theta) P(\theta | y) d\theta}{\int_{0}^{\infty} \frac{d\theta}{d\theta}} = \frac{(\partial - \theta) P(\theta | y) d\theta}{\int_{0}^{\infty} \frac{d\theta}{d\theta}} = \frac{(\partial - \theta) P(\theta | y) d\theta}{\int_{0}^{\infty} \frac{d\theta}{d\theta}} = \frac{(\partial - \theta) P(\theta | y) d\theta}{\int_{0}^{\infty} \frac{d\theta}{d\theta}} = \frac{(\partial - \theta) P(\theta | y) d\theta}{\int_{0}^{\infty} \frac{d\theta}{d\theta}} = \frac{(\partial - \theta) P(\theta | y) d\theta}{\int_{0}^{\infty} \frac{d\theta}{d\theta}} = \frac{(\partial - \theta) P(\theta | y) d\theta}{\int_{0}^{\infty} \frac{d\theta}{d\theta}} = \frac{(\partial - \theta) P(\theta | y) d\theta}{\int_{0}^{\infty} \frac{d\theta}{d\theta}} = \frac{(\partial - \theta) P(\theta | y) d\theta}{\int_{0}^{\infty} \frac{d\theta}{d\theta}} = \frac{(\partial - \theta) P(\theta | y) d\theta}{\int_{0}^{\infty} \frac{d\theta}{d\theta}} = \frac{(\partial - \theta) P(\theta | y) d\theta}{\int_{0}^{\infty} \frac{d\theta}{d\theta}} = \frac{(\partial - \theta) P(\theta | y) d\theta}{\int_{0}^{\infty} \frac{d\theta}{d\theta}} = \frac{(\partial - \theta) P(\theta | y) d\theta}{\int_{0}^{\infty} \frac{d\theta}{d\theta}} = \frac{(\partial - \theta) P(\theta | y) d\theta}{\int_{0}^{\infty} \frac{d\theta}{d\theta}} = \frac{(\partial - \theta) P(\theta | y) d\theta}{\int_{0}^{\infty} \frac{d\theta}{d\theta}} = \frac{(\partial - \theta) P(\theta | y) d\theta}{\int_{0}^{\infty} \frac{d\theta}{d\theta}} = \frac{(\partial - \theta) P(\theta | y) d\theta}{\int_{0}^{\infty} \frac{d\theta}{d\theta}} = \frac{(\partial - \theta) P(\theta | y) d\theta}{\int_{0}^{\infty} \frac{d\theta}{d\theta}} = \frac{(\partial - \theta) P(\theta | y) d\theta}{\int_{0}^{\infty} \frac{d\theta}{d\theta}} = \frac{(\partial - \theta) P(\theta | y) d\theta}{\int_{0}^{\infty} \frac{d\theta}{d\theta}} = \frac{(\partial - \theta) P(\theta | y) d\theta}{\int_{0}^{\infty} \frac{d\theta}{d\theta}} = \frac{(\partial - \theta) P(\theta | y) d\theta}{\int_{0}^{\infty} \frac{d\theta}{d\theta}} = \frac{(\partial - \theta) P(\theta | y) d\theta}{\int_{0}^{\infty} \frac{d\theta}{d\theta}} = \frac{(\partial - \theta) P(\theta | y) d\theta}{\int_{0}^{\infty} \frac{d\theta}{d\theta}} = \frac{(\partial - \theta) P(\theta | y) d\theta}{\int_{0}^{\infty$$

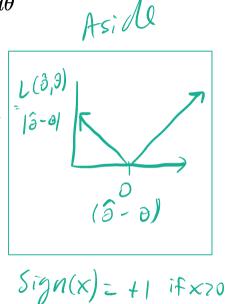
#### **Absolute loss**

$$\min_{\hat{ heta}} E_{ heta \mid y} |\hat{ heta} - heta | = \min_{\hat{ heta}} \ \int |\hat{ heta} - heta | p( heta \mid y) d heta$$

Differentiate with respect to  $\hat{\theta}$  and set equal to zero:

$$\frac{d}{d\hat{\theta}} \int |\hat{\partial} - \theta| P(\theta|y) d\theta =$$

$$\int \frac{d}{d\hat{\theta}} |\hat{\partial} - \theta| P(\theta|y) d\theta =$$



 $\begin{array}{l}
9 = \hat{\vartheta} \\
5 + 1 P(\vartheta | y) d\vartheta + 5 P(\vartheta | y) d\vartheta \\
9 = \hat{\vartheta} \\
\hline
P(\vartheta \angle \hat{\vartheta} | y) - (1 - P(\vartheta \angle \hat{\vartheta} | y)) = 0 \\
2 P(\vartheta \angle \hat{\vartheta} | y) = 1 \\
P(\vartheta \angle \hat{\vartheta} | y) = 1/2 \\
\vec{\vartheta} \text{ is the posterior median} \\
Buyes Estimator for absolute loss.}$ 

#### Loss functions in practice

- Squared error and absolute error are good default loss functions
  - Motivated largely by mathematical considerations
- In practice we should define a loss function specific to our problem
- Loss in dollars? Loss in "quality of life"?

#### Decision making: flu example

• The CDC produces estimates of the expected prevalance and severity of flu during flu season

• Assume  $\theta$  represents severity of the flu

•  $p(\theta \mid y)$  is CDC posterior distribution based on initial data about the upcoming flu season

•  $\hat{\theta}$  determines how much flu vaccine to make. How do we determine

 $L(\hat{\theta}, \theta)$ ?

+ Shelf life of vaccine

+ Too little = death

+ Too much is wasted money