Lecture 1: Review and Background

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2023-01-12

Class Resources

Required Textbook

Bayes Rules: https://www.bayesrulesbook.com/

Course Pages

- Class website on Canvas: (https://https://www.canvas.ucsb.edu/)
- Nectir for course related questions and discussion: https://ucsb.nectir.io/group/pstat115-w23
 - On Canvas site
- Gradescope: https://www.gradescope.com/courses/344698 (Enrollment code: BBRN6B)
 - On Canvas site

Grades

- 35% expect approximately 5 homeworks
- 20% Midterm (February 9)
- 10% Quizzes
- 5% Participation (Section attendance)
- 30% Final exam (March 21)

Homework

- There will be approximately 6 homeworks (35% of your grade total)
- You will typically have 1-2 weeks to complete the homeworks
- You are allowed to work with a partner
 - Add partners name to your assignment
- Every student *must* submit their own assignment on gradescope
- Homework turned in within 24 hrs after the deadline without prior approval will receive a 10 pt deduction (out of 100)
- Homework will not be accepted more than 24 hrs late.

Homework submission format

- All code must be written to be reproducible in Rmarkdown
- All derivations can be done in any format of your choosing (latex, written by hand) but must be legible and *must be integrated into your Rmarkdown pdf*.
- All files must be zipped together and submitted to Gradescope
- Ask a TA *early* if you have problems regarding submissions.

Labs and Quizzes

- There will be a handful of "pop" quizzes throughout the quarter.
- There are no makeups, but the lowest quiz grade will be dropped from your final score.
- Quizzes (10%) will be multiple choice and will test your comprehension of the basic concept.
- Participation (5%). Includes lecture attendance, section attendance, and nectir posts.

Class Policies

• All questions should be posted on nectir, *not by email* (unless they are personal or grade-related)

RStudio Cloud Service

- Log on to pstat115.lsit.ucsb.edu
 - Cloud based rstudio service
 - Log in with your UCSB NetID
- Use [https://tinyurl.com/32ra4at4] https://tinyurl.com/32ra4at4) to sync new material (BOOKMARK THIS)
- Make sure you can write and compile an R markdown (Rmd) document online
- Text formatting is minimal but syntax is simple

Logistics

- First homework out by end of week
 - Due January 22 (Sunday)
- Try pstat115.lsit.ucsb.edu
 - Cloud based rstudio service
 - Log in with your UCSB NetID
- Canvas website

Resources

Look at the resources folder in cloud for

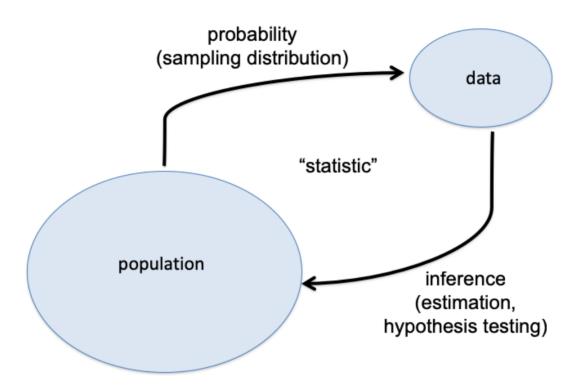
- A fantastic probability review sheet
- Probability density information
- Bayes Rules: Chapters 1 and 2

Logistics

- Homework 1 will be out by the end of the week in winter23/homework/homework1.Rmd
- Do not change the name of the file or the directory
- Autograding

```
o Leave code cells that look like . =
ottr::check("tests/qla.R")
```

Population and Sample



Population and Sample

- The *population* is the group or set of items relevant to your question
 - Usually very large (often conceptualize a population as infinite)
- Sample: a finite subset of the population
 - How is the sampling collected (representative?)
 - \circ Denote the sample size with n

Population and Sample

- Our goal is (usually) to learn about the population from the sample
 - Population parameters encode relevant quantities
 - The **estimand** is the thing we what to infer and is usually a function of the population parameters

Random variables

- A random variable, Y has variability, can take on several different values (possibly infinitely may), and is associated with a distribution.
 - The distribution determines the probability that the r.v. will take a specific value.
- Notation:
 - Y (uppercase) denotes a random variable
 - *y* (lowercase) is a *realization* of that random variable and is not random

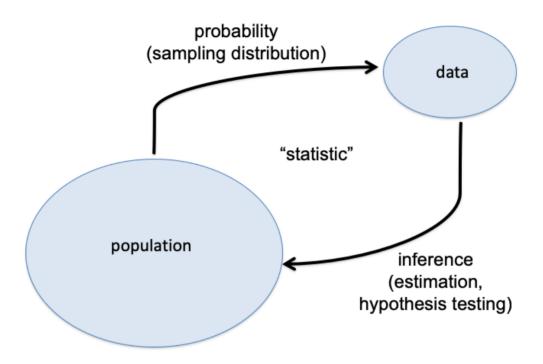
Constants

- Constants: quantities with 0 variance.
 - Constants can be *known* (e.g. observed data)
 - Constants can be *unknown* (not observed)

Setup

- The sample space \mathcal{Y} is the set of all possible datasets we could observe. We observe one dataset, y, from which we hope to learn about the world.
- The parameter space Θ is the set of all possible parameter values θ
- θ encodes the population characteristics that we want to learn about
- Our sampling model $p(y \mid \theta)$ describes our belief about what data we are likely to observe for a given value of θ .

Notation



The Likelihood Function

- The likelihood is the "probability of the observed data" expressed as a function of the unknown parameter:
- A function of the unknown constant θ .
- ullet Depends on the observed data $y=(y_1,y_2,\ldots,y_n)$

Independent Random Variables

- Y_1, \ldots, Y_n are random variables
- We say that Y_1, \ldots, Y_n are *conditionally* independent given θ if
- Conditional independence means that Y_i gives no additional information about Y_j beyond that in knowing θ

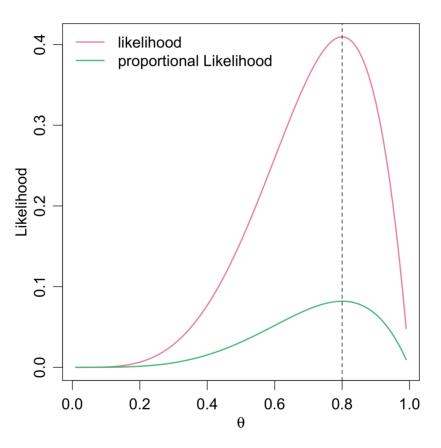
Example: A binomial model

- Assume I go to the basketball court and takes 5 free throw shots
- Model the number of made shots I make using a $Bin(5, \theta)$
 - What are the key assumptions that make these a reasonable emodel?
- θ represents my true skill (the fraction of shots I make)
- How can we estimate my true skill?

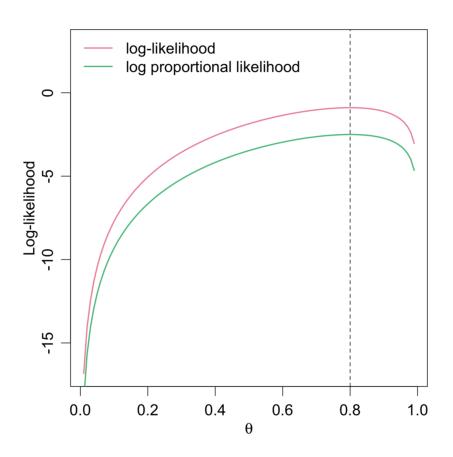
Likelihood:

The binomial likelihood

I make 4 out of 5



The log-likelihood



Maximum Likelihood Estimation

- The maximum likelihood estimate (MLE) is the value of θ that makes the data the most "likely", that is, the value that maximizes $L(\theta)$
- To compute the maximum likelihood estimate:
 - 1. Write down the likelihood and take its log:

$$\log(L(\theta)) = \ell(\theta)$$

2. Take the derivative of $\ell(\theta)$ with respect to θ :

$$\ell'(heta) = rac{d\ell(heta)}{d heta}$$

3. Solve for $\hat{\theta}$ such that $\ell'(\theta) = 0$

Maximum Likelihood Estimation

Example: Binomial

- Assume we are polling the presidential race in the upcoming election
- We poll 25 random students in the class $Y_1, \ldots Y_n$ from n=25
- Y_i is either 0 (Trump) or 1 (Biden)
- $Y_i \sim \text{Bern}(\theta)$, where $\text{Bern}(\theta)$ is equivalent to $\text{Bin}(1, \theta)$
 - Bernoulli random variables is a binomial with one trial
 - Assume our class is a simple random sample of the population
- How do we estimate θ for multiple observations?

Example: the likelihood for independent Bernoulli's

$$egin{aligned} p(y_1,y_2,\ldots,y_n|1, heta)&=p(y_1,y_2,\ldots,y_n| heta)\ &=p(y_1| heta)p(y_2| heta)\ldots p(y_n| heta)\ &=\prod_{i=1}^n p(y_i| heta)\ &=\prod_{i=1}^n inom{1}{y_i} heta^{y_i}(1- heta)^{(1-y_i)}\ &=igg[\prod_{i=1}^n inom{1}{y_i}igg]\, heta^{\sum_{i=1}^n y_i}(1- heta)^{n-\sum_{i=1}^n y_i}\ &=L(heta) \end{aligned}$$

Sufficient Statistics

- Let $L(\theta) = p(y_1, \ldots y_n \mid \theta)$ be the likelihood and $s(y_1, \ldots y_n)$ be a statistic
- s(y) is a sufficient statistic if we can write:

$$L(heta) = h(y_1, \dots y_n) g(s(y), heta)$$

- \circ g is only a function of s(y) and θ only
- \circ h is *not* a function of θ
- This is known as the factorization theorem
- $L(\theta) \propto g(s(y), \theta)$

Sufficient Statistics

- Intuition: a sufficient statistic contains all of the information about θ
 - Many possible sufficient statistics
 - Often seek a statistic of the lowest possible dimension (minimal sufficient statistic)
 - What are some sufficient statistics in the previous binomial example?

Estimators and Estimates

- In classical (frequentist) statistics, θ is an unknown constant
- An **estimator** of a parameter θ is a function of the random variables, Y
 - $\circ ext{ E.g. for Binomial}(1, heta): \hat{ heta}(Y) = rac{\sum_i Y_i}{n}$
 - An estimator is a random variable
 - Interested in properties of estimators (e.g. mean and variance)

Estimators and Estimates

- $\hat{\theta}(y)$ as a function of realized data is called an **estimate**
 - \circ Plug in observed data $y=(y_1,\ldots y_n)$ to estimate heta
 - An estimate is a non-random constant (it is has 0 variability)
 - E.g. in the binomial $(1, \theta)$, $\hat{\theta} = \bar{y} = \frac{\sum_{i} y_{i}}{n}$ is the maximum likelihood estimate for the binomial proportion.

Bias and Variance

• Estimators are random variables. What are some r.v. properties that are desirable?

Bias and Variance

- Estimators are random variables. What are some r.v. properties that are desirable?
- Bias: $E[\hat{\theta}] \theta = 0$
 - $\circ \ E[\hat{ heta}] heta = 0$ means the estimator is unbiased
 - E.g. expectation of the binomial MLE: $E[\hat{\theta}] = E[\frac{\sum Y_i}{n}] = \theta$
- $\operatorname{Var}(\hat{ heta}) = E[(\hat{ heta} E[\hat{ heta}])^2]$

E.g. variance of the binomial MLE is

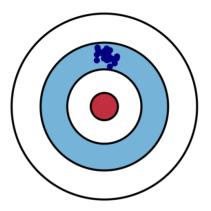
$$ext{Var}[\hat{ heta}] = ext{Var}(rac{\sum Y_i}{n}) = rac{ heta(1- heta)}{n}$$

Bias and Variance

- Want estimators that have low bias and variance because this implies low overall error
- Mean squared error equals bias² + variance

Bias

The average difference between the prediction and the response

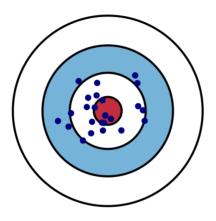


Statistical definition of bias:

$$E[\hat{ heta} - heta]$$

Variance

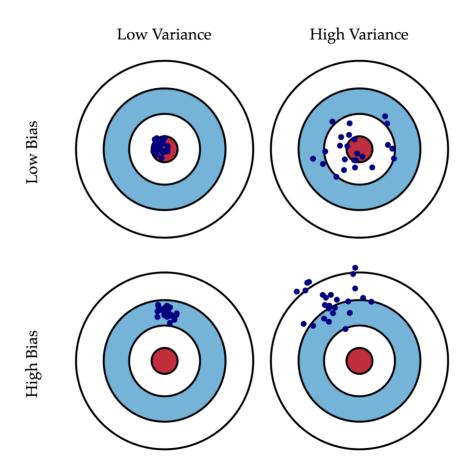
How variable is the prediction about its mean?



Statistical definition of variance:

$$E[\hat{ heta}-E[\hat{ heta}]]^2$$

Bias and Variance



Maximum Likelihood Estimators

Under relatively weak conditions:

- The MLE is *consistent*. It converges to the true value as the sample size goes to infinity.
 - Need bias and variance to go to 0 as sample size increases
- The MLE is *asymptotically optimal*. For "large" sample sizes is has the lowest variance.
- Equivariance: if $\hat{\theta}$ is the MLE for θ then $g(\hat{\theta})$ is the MLE for $g(\theta)$

Confidence Interval Simulations

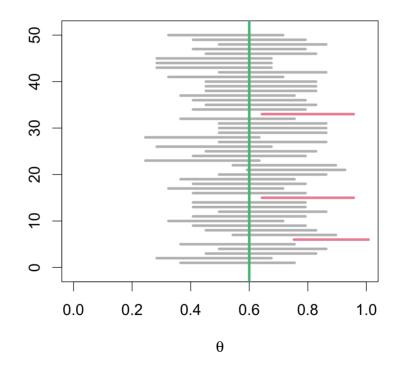
Let's do 50 hypothetical replications to illustrate confidence intervals

```
for i in 1 to 50:
    - Draw Y_i from Bin(25, 0.6)
    - Compute and plot the 95% confidence interval
```

- Will have 50 confidence intervals based on 50 simulated datasets.
- A 95% interval means that on average 95% of these 50 intervals will cover the true value

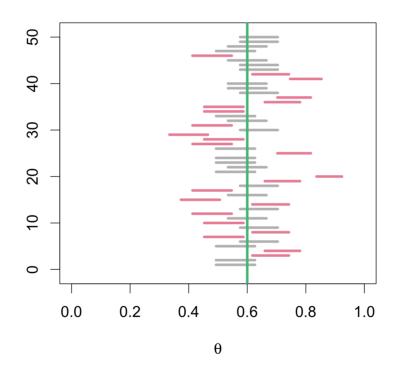
95% Confidence Intervals

In truth, 60% of the population will vote for "candidate 1"



We expect $0.05 \times 50 = 2.5$ of the intervals to *not* cover the true parameter, p = 0.6, on average

50% Confidence Intervals



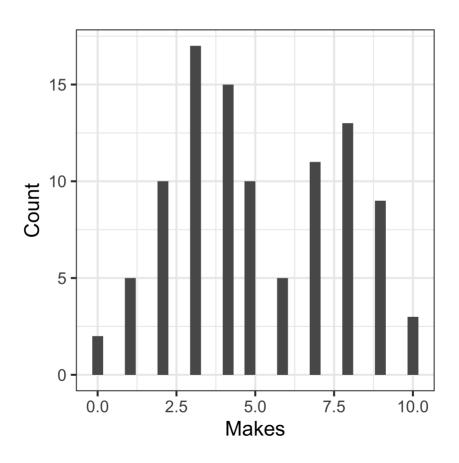
We expect $0.50 \times 50 = 25$ of the intervals to *not* cover the true parameter, 0.6

- DGP: a statistical model for how the observed data might have been generated
- Often write the DGP using pseudo-code:

```
for (i in 1:N)
  - Generate y_i from a Normal(0, 1)
return y = (y_1, ... y_N)
```

- The DGP should tell a story about how the data came to be
- Can translate the DGP into a statistical model

Assume everybody in this class goes to a basketball court and takes 10 free throw shots:





Tell a plausible story: some students play basketball and some don't. Before you take your shots we record whether or not you have played before.

```
assume theta_1 > theta_0
for (i in 1:100)
  - Generate z_i from Bin(1, phi)
  - p_i = theta_0 if z_i=0
  - p_i = theta_1 if z_i=1
  - Generate y_i from a Binom(10, p_i)
return y = (y_1, ... y_100) and z = (z_1, ..., z_100)
```

Is this a reasonable model?

Mixture Models

$$Z_i = egin{cases} 0 & ext{if the } i^{th} ext{ if student doesn't play basketball} \ 1 & ext{if the } i^{th} ext{ if student does play basketball} \ Z_i \sim Bin(1,\phi) \ Y_i \sim egin{cases} ext{Bin}(10, heta_0) & ext{if } Z_i = 0 \ ext{Bin}(10, heta_1) & ext{if } Z_i = 1 \end{cases}$$

Mixture Models

$$Z_i = egin{cases} 0 & ext{if the } i^{th} ext{ if student doesn't play basketball} \ 1 & ext{if the } i^{th} ext{ if student does play basketball} \end{cases}$$

$$Z_i \sim Bin(1,\phi)$$
 $Y_i \sim egin{cases} ext{Bin}(10, heta_0) & ext{if } Z_i = 0 \ ext{Bin}(10, heta_1) & ext{if } Z_i = 1 \end{cases}$

- \bullet ϕ is the fraction of students that have experience playing basketball
- θ_1 is the probability of making a shot for an experienced player
- θ_0 is the probability of making a shot for an inexperienced player

Table of relevant quantities

- Can be a fixed constant (no variability) or a random variable (has variability)
- Can be observed (known) or unobserved (unknown)
- Helpful for to keep track of all of the relevant quantities

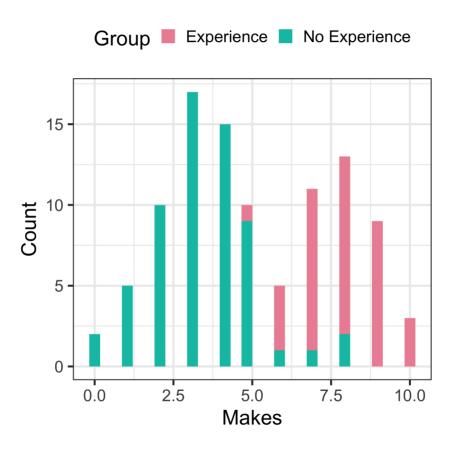
Mixture models

- A mixture model is a probabilistic model for representing the presence of subpopulations
- The subpopoluation to which each individual belongs is not necessarily known
 - e.g. do we ask: "have you played basketball before?"

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- The subpopoluation to which each individual belongs is not necessarily known
 - e.g. do we ask: "have you played basketball before?"
- When z_i is not observed, we sometimes refer to it as a clustering model
 - unsupervised learning

A Mixture Model



Note: z is observed

Mixture Model Likelihood

Z is observed

Sufficient statistics When Z_i is observed

Together, the following quantities are sufficient for $(\theta_0, \theta_1, \phi)$

- $\sum y_i z_i$ (total number of shots made by experienced players)
- $\sum y_i(1-z_i)$ (total number of shots made by inexperienced players)
- $\sum z_i$ (total number experienced players)

```
for (i in 1:100)
  - Generate z_i from Bin(1, phi)
  - p_i = theta_1 if z_i=1
  - p_i = theta_0 if z_i=0
  - Generate y_i from a Binom(10, p_i)
return y = (y_1, ... y_100)
```

This time we don't record who has experience with basketball.

A Mixture Model

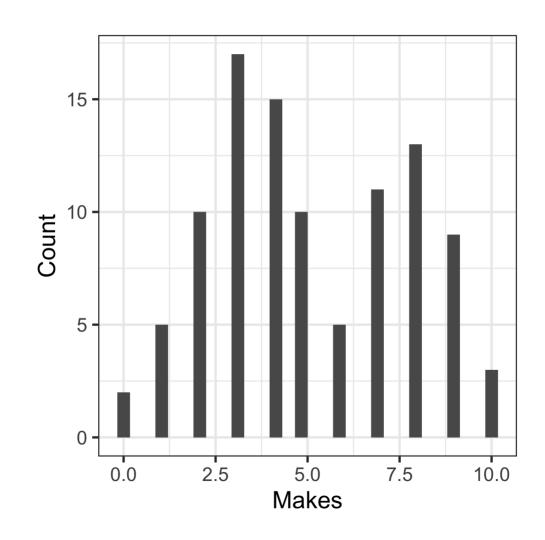


Table of Relevant Quantities

A finite mixture model

- Even if we don't observe Z, it's often useful to introduce it as a *latent* variable
- Write the *observed data likelihood* by integrating out the latent variables from the *complete data likelihood*

$$egin{aligned} p(Y \mid heta) &= \sum_{z} p(Y, Z = z \mid heta) \ &= \sum_{z} p(Y \mid Z = z, heta) p(Z = z \mid heta) \end{aligned}$$

In general we can write a K component mixture model as:

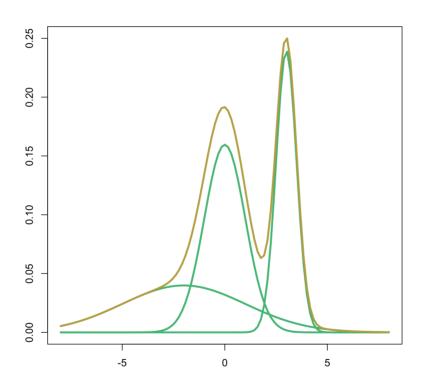
$$p(Y) = \sum_k^K \pi_k p_k(Y)$$

with
$$\sum \pi_k = 1$$

Mixture Model Likelihood

Z unobserved

Finite Mixture models



Infinite Mixture Models

- In the previous example the latent variable had finitely many outcomes
- Latent varibles can have infinitely many outcomes in which case we have any infinite mixture
- Example:

$$egin{aligned} \mu &\sim N(0, au^2) \ Y &\sim N(\mu,\sigma^2) \end{aligned} \ p(Y\mid\sigma^2, au^2) &= \int p(Y,\mu\mid\sigma^2, au^2) d\mu \end{aligned}$$

What is the *marginal* distribution of Y?

Bayesian Inference

- In frequentist inference, θ is treated as a fixed unknown constant
- In Bayesian inference, θ is treated as a random variable
- Need to specify a model for the joint distribution $p(y, \theta) = p(y \mid \theta)p(\theta)$

Bayesian Inference in a Nutshell

1. The *prior distribution* $p(\theta)$ describes our belief about the true population characteristics, for each value of $\theta \in \Theta$.

Bayesian Inference in a Nutshell

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- 2. Our sampling model $p(y \mid \theta)$ describes our belief about what data we are likely to observe if θ is true.

Bayesian Inference in a Nutshell

- 1. The *prior distribution* $p(\theta)$ describes our belief about the true population characteristics, for each value of $\theta \in \Theta$.
- 2. Our sampling model $p(y \mid \theta)$ describes our belief about what data we are likely to observe if θ is true.
- 3. Once we actually observe data, y, we update our beliefs about θ by computing the posterior distribution $p(\theta \mid y)$. We do this with Bayes' rule!

Bayes' Rule

$$P(A \mid B) = \frac{P(B \mid A)PAB)}{P(B)}$$

- $P(A \mid B)$ is the conditional probability of A given B
- $P(B \mid A)$ is the conditional probability of B given A
- P(A) and P(B) are called the marginal probability of A and B (unconditional)

Bayes' Rule for Bayesian Statistics

$$P(heta \mid y) = rac{P(y \mid heta)P(heta)}{P(y)}$$

- $P(\theta \mid y)$ is the posterior distribution
- $P(y \mid \theta)$ is the likelihood
- $P(\theta)$ is the prior distribution
- $P(y) = \int_{\Theta} p(y \mid \tilde{\theta}) p(\tilde{\theta}) d\tilde{\theta}$ is the model evidence

Bayes' Rule for Bayesian Statistics

$$P(\theta \mid y) = rac{P(y \mid \theta)P(\theta)}{P(y)} \ \propto P(y \mid \theta)P(\theta)$$

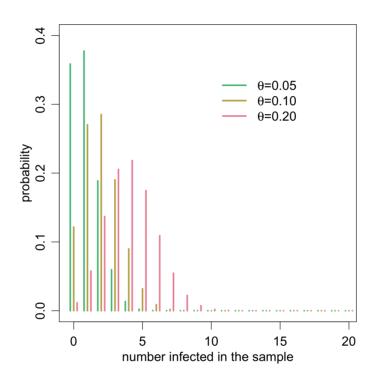
- Start with a subjective belief (prior)
- Update it with evidence from data (likelihood)
- Summarize what you learn (posterior)

The posterior is proportional to the likelihood times the prior!

- We need to estimate the prevalence of a COVID in Isla Vista
- Get a small random sample of 20 individuals to check for infection



- θ represents the population fraction of infected
- Y is a random variable reflecting the number of infected in the sample
- $\Theta = [0,1]$ $\mathcal{Y} = \{0,1,\ldots,20\}$
- Sampling model: $Y \sim \operatorname{Binom}(20, \theta)$



- Assume *a priori* that the population rate is low
 - The infection rate in comparable cities ranges from about 0.05 to 0.20
- Assume we observe Y = 0 infected in our sample
- What is our estimate of the true population fraction of infected individuals?

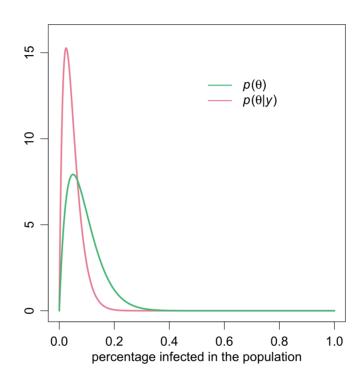


Table of Relevant Quantities

Summary

- Likelihood, log likehood in MLE
- Confidence intervals (how they are defined in frequentist inference)
- Sufficient statistics
- Mixture models

Summary

- In frequentist inference, unknown parameters treated as constants
 - Estimators are random (due to sampling variability)
 - Asks: "how would my results change if I repeated the experiment?"

Summary

- In Bayesian inference, unknown parameters are random variables.
 - Need to specify a prior distribution for θ (not easy)
 - Asks: "what do I *believe* are plausible values for the unknown parameters?"
 - Who cares what might have happened, focus on what did happen!

Assignments

- Read chapters 1 and 2 of BR
- Homework 1 due 1/22