

Lecture 6: The Normal Model

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2023-02-07

Announcements

- Reading: Section 5.3.3 and 5.3.4 (Normal Model, Start after exam).
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- + Quiz 2, out, due 2pm Wed.
 - + No Thurs Section.
 - + Lauren: Wed 2-4, Thurs 8:30-9:30.
(St 542)
 - Doris: Zoom Wed 5-6.

- Likelihood, simplify, drop normalizing constant.
- Conjugate Priors, meaning of the parameters (matching units)
- Prior & Posterior Mean for Beta / Gamma.

$$\frac{\alpha}{\alpha + \beta} \quad \alpha / \beta$$
- Write posterior mean as $w \hat{\theta}_{MLE} + (1-w) \hat{\theta}_{prior}$
 + How we derived that.
 + what w is
- Basic idea behind M.C.
 + Inversion / Rejection sampling.
- Posterior Predictive Distrn.

$$P(\tilde{y} / y) = \int P(\tilde{y} / \theta) P(\theta / y) d\theta$$

- MC: sampler $\theta^s \sim P(\theta|y)$
 sample $\tilde{y} \sim P(\tilde{y}|\theta^s)$

- Uncertainty Quant.

- Quantile and HPD strategies

$$P(\tilde{y}|\cancel{y}) = \int P(\tilde{y}|\theta) P(\theta|\cancel{y}) d\theta$$

Prior Predictive.

$$\frac{y \sim \text{Bin}(n, \theta)}{\text{Beta}(a, b) \rightarrow \binom{n}{y} \theta^y (1-\theta)^{n-y}} \quad \binom{a+b-1}{a-1} \theta^{a-1} (1-\theta)^{b-1}$$

$$\theta^y (1-\theta)^{n-y} \times \theta^{a-1} (1-\theta)^{b-1}$$

$$\binom{n+a-1}{y+a-1} \theta^{y+a-1} (1-\theta)^{b-1}$$

The Normal Distribution

- One of the most utilized probability models in data analysis
- Central Limit Theorem
- Separate parameters for the mean and the variance (intuitive)

$$Y \sim N(\mu, \sigma^2)$$

$$Z_1, \dots, Z_n \sim \text{iid}$$

- 2 parameter model

$\bar{Z} \approx \text{normal}$.

Need: $P(\mu, \sigma^2)$

Get: $P(\mu, \sigma^2 | y_1, \dots, y_n)$

Normal Distribution

- Symmetric with mode = median = mean = μ
- Approximately 95% of the population lies within two standard deviations of the mean

$$\bar{y} \pm 1.96 \hat{\sigma} \quad (95\% \text{ CI})$$

- Density:

$$p(y|\mu, \sigma^2) = \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{1}{2}\left(\frac{y-\mu}{\sigma}\right)^2}, \quad -\infty < y < \infty$$

- $X \sim N(\mu_x, \sigma_x^2)$ and $Y \sim N(\mu_y, \sigma_y^2)$ with X and Y independent then

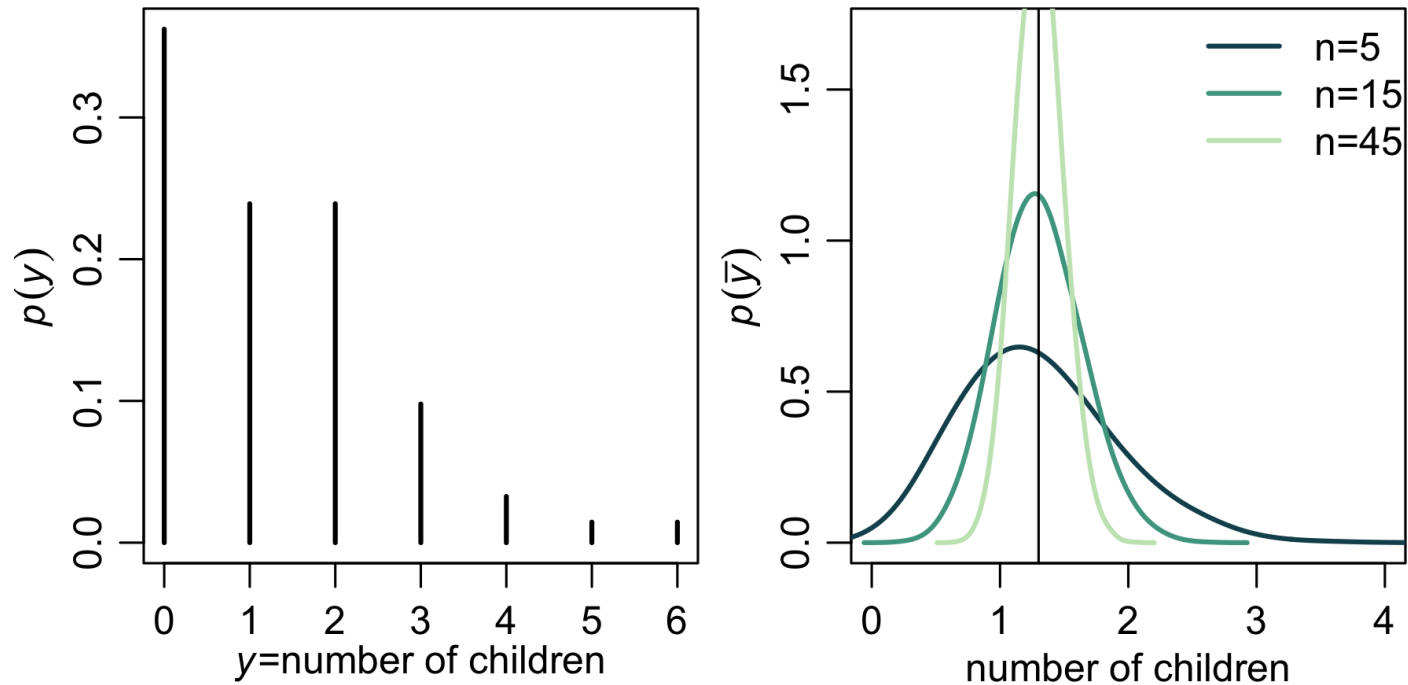
$$aX + bY \sim N(a\mu_x + b\mu_y, a^2\sigma_x^2 + b^2\sigma_y^2)$$

- In R: dnorm, rnorm, pnorm, qnorm.

- Warning: the argument to the norm functions R is σ not σ^2 !

$$E[aX + bY] = aE[X] + bE[Y]$$

The Central Limit Theorem



$$\text{Var}\left(\frac{\sum Y_i}{n}\right) = \frac{1}{n^2} n \text{Var}(Y_i) = \frac{1}{n}$$

$$\text{CLT: } \bar{y} \approx N(E[Y], \text{Var}[Y]/n)$$

$$Y_i \sim \text{Pois}(\lambda) ; \quad \bar{y} = \frac{\sum Y_i}{n} \approx N\left(\lambda, \frac{1}{n}\right)$$

Bayesian inference in the normal model

- Assume $y_1, \dots, y_n \sim N(\mu, \sigma^2)$ with σ^2 a known constant
- Lets start with a non-informative, improper prior: $p(\mu) \propto \text{const}$
- What is the posterior distribution $p(\mu \mid y_1, \dots, y_n, \sigma^2)$?

"flat"

$$P(\mu \mid \sigma^2, y_1, \dots, y_n) \propto L(\mu) \times \text{const}$$

$$P(\mu | \dots) \propto \prod_{i=1}^n \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{(y_i - \mu)^2}{2\sigma^2}}$$

$$\propto \left(\frac{1}{\sqrt{2\pi\sigma^2}} \right)^n e^{-\frac{\sum (y_i - \mu)^2}{2\sigma^2}}$$

$$\propto e^{-\frac{(\sum y_i^2 - 2\mu \sum y_i + n\mu^2)}{2\sigma^2}} \quad (\text{expand the sq})$$

$$\propto \cancel{e^{-\frac{\sum y_i^2}{2\sigma^2}}} e^{\frac{2\mu \sum y_i - n\mu^2}{2\sigma^2}}$$

(complete the sq)

Aside

$$\begin{aligned} (ax^2 - bx) &= a\left(x^2 - \frac{b}{a}x\right) \\ &= a\left(x - \frac{b}{2a}\right)^2 - \frac{b^2}{4a} \end{aligned}$$

$$\propto e^{\frac{-n\mu^2 - 2\mu \sum y_i}{2\sigma^2}}$$

$$\begin{aligned} a &= n \\ b &= 2\sum y_i \\ x &= \mu \end{aligned}$$

$$\propto e^{\frac{-n\left(\mu - \frac{\sum y_i}{n}\right)^2}{2\sigma^2}} \cancel{e^{-\frac{(2\sum y_i)^2}{4n}}}$$

$$\propto e^{\frac{-(\mu - \bar{y})^2}{2\sigma^2/n}}$$

$$P(\mu | y_1, \dots, y_n) \sim N(\bar{y}, \frac{\sigma^2}{n})$$

$$E[\mu | y_1, \dots, y_n] = \bar{y} = \hat{\mu}_{MLE}$$

$$\text{Var}(\mu | y_1, \dots, y_n) = \sigma^2/n$$

$$\int_{-\infty}^{\infty} e^{-\frac{(\mu - \bar{y})^2}{2\sigma^2/n}} d\mu < \infty$$

Bayesian inference in the normal model

- Assume $y_1, \dots, y_n \sim N(\mu, \sigma^2)$ with σ^2 a known constant
- The normal prior distribution is conjugate for μ in the normal sampling model
- Sampling distribution, prior distribution and posterior distribution are all normal.
- Assume the prior is $p(\mu) \sim N(\mu_0, \tau^2)$
 - Prior guess about μ (mean)*
 - How certain*
- What are the parameters of the posterior $p(\mu \mid y_1, \dots, y_n, \sigma^2)$?

Normal

$$P(\mu | y_1, \dots, y_n, \sigma^2) \propto \underline{L(\mu)} P(\mu)$$

$$\propto \underbrace{e^{-\frac{(\mu - \bar{y})^2}{2\sigma^2/n}}}_{L(\mu)} \underbrace{\frac{1}{\sqrt{2\pi\tau^2}} e^{-\frac{(\mu - \mu_0)^2}{2\tau^2}}}_{P(\mu)}$$

$$\propto \exp\left[-\left(\frac{\bar{y}^2 - 2\bar{y}\mu + \mu^2}{2\sigma^2/n} + \frac{\mu^2 - 2\mu\mu_0 + \mu_0^2}{2\tau^2}\right)\right]$$

$$\propto \exp\left[-\frac{1}{2} \left(\frac{n}{\sigma^2} \mu^2 - \frac{2n}{\sigma^2} \bar{y}\mu + \frac{1}{\tau^2} \mu^2 - \frac{2\mu\mu_0}{\tau^2} \right)\right]$$

$$\propto \exp\left[-\frac{1}{2} \left(\underbrace{\left(\frac{n}{\sigma^2} + \frac{1}{\tau^2}\right)}_a \mu^2 - 2 \underbrace{\left(\frac{n\bar{y}}{\sigma^2} + \frac{\mu_0}{\tau^2}\right)}_b \mu \right)\right]$$

$$\propto \exp\left[-\frac{1}{2} \underbrace{\left(\frac{n}{\sigma^2} + \frac{1}{\tau^2}\right)}_{\substack{\text{posterior} \\ \text{var.}}} \left(\mu - \frac{\frac{n\bar{y}}{\sigma^2} + \frac{\mu_0}{\tau^2}}{\underbrace{\frac{n}{\sigma^2} + \frac{1}{\tau^2}}_{\text{post. Mean}}} \right)^2\right]$$

$$E[\mu | y_1, \dots, y_n] = \underbrace{\frac{\frac{n}{\sigma^2} \bar{y}}{\frac{n}{\sigma^2} + \frac{1}{\tau^2}}}_{\substack{\text{(post mean)} \\ \text{MLE}}} + \underbrace{\frac{\frac{1}{\tau^2} \mu_0}{\frac{n}{\sigma^2} + \frac{1}{\tau^2}}}_{\substack{\text{prior mean.}}}$$

w

$1-w$

$$= w \bar{y} + (1-w) \mu_0$$

$$w = \frac{n/\sigma^2}{n/\sigma^2 + 1/\tau^2}$$

$$\text{Var}(\mu | y_1, \dots, y_n) = \frac{1}{\frac{n}{\sigma^2} + \frac{1}{\tau^2}}$$

$\frac{n}{\sigma^2}$ is precision of \bar{y}

$1/\tau^2$ is prior precision for μ .

$$p(\mu | \dots) \sim N(\mu_n, \sigma_n^2)$$

$$\mu_n = w \bar{y} + (1-w) \mu_0$$

$$w = \frac{n/\sigma^2}{n/\sigma^2 + 1/\tau^2}$$

$$\sigma_n^2 = w \frac{\sigma^2}{n}$$

A conjugate prior for the normal likelihood

- The normal distribution is conjugate for the normal likelihood
 - Often called the "normal-normal model"
- $Y_i \sim N(\mu, \sigma^2)$ and $\mu \sim N(\mu_0, \tau^2)$ implies that the posterior distribution $p(\mu \mid y)$ is also normally distributed:

$$\mu \mid Y \sim N(\mu_n, \tau_n^2)$$

$$\text{where } \mu_n = \frac{\frac{1}{\tau^2}\mu_0 + \frac{n}{\sigma^2}\bar{y}}{\frac{1}{\tau^2} + \frac{n}{\sigma^2}} \text{ and } \tau_n^2 = \frac{1}{\frac{1}{\tau^2} + \frac{n}{\sigma^2}}$$

The posterior mean and pseudo-counts

$$\begin{aligned}\mu_n &= \frac{\frac{1}{\tau^2}}{\frac{1}{\tau^2} + \frac{n}{\sigma^2}} \mu_0 + \frac{\frac{n}{\sigma^2}}{\frac{1}{\tau^2} + \frac{n}{\sigma^2}} \bar{y} \\ &= (1 - w) \mu_0 + w \bar{y}\end{aligned}$$

$$\text{where } w = \frac{\frac{n}{\sigma^2}}{\frac{1}{\tau^2} + \frac{n}{\sigma^2}}$$

Can we think about the normal prior parameters in terms of pseudo-counts?