

Lecture 7: Hierarchical Modeling

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Announcements

- Fill out ESCI evaluations online!



- What did you like? Would you like to see more courses like this?
- What can be improved?

- Reading: Chapters 15 and 16

Hierarchical

- HW 5 out, due 19th.
- Quiz out, due 8th at 2pm. (MCMC)
(1 quiz next).

Comparing Multiple Related Groups

- Hierarchy of nested populations
- Models which account for this are called *hiearchical* or *multi-level* models

Some examples:

- Patient outcomes within several different hospitals
- People within counties in the United States (e.g. Asthma mortality example)
- Athlete performance in sports
- Genes within a group of animals

Eight schools example

- A study was performed for the Educational Testing Service (ETS) to evaluate the effects of coaching programs on SAT preparation
- Each of eight different schools used a short-term SAT prep coaching program
- Compute the average SAT score in those who did take the program minus those that did not participate in the program
- We observe the average difference varies by school. What accounts for these differences?

SO students
per school,
randomly assign
25

y_i , $i = 1, \dots, 8$

difference in SAT from
training vs control.

- Access varies by school
 - Income
 - Quality of the school
 - populations are different
 - Motivation
 - Language.
-

- Chance!

+ By Chance strong students
in 1 group & weaker in
the other.

Eight schools example

- Interested in "real" differences due to training
- Want to reduce ~~effect~~ ^{influence} of chance variability
- How do we estimate the effect of the program in each of the schools?
- Two extremes:
 - Estimate the effect of the program in every school independently
 - A separate prior distribution for each school effect
 - Or assume the effect is the same in every school
 - Combine all the data
 - A compromise between the above 2 options?

Eight Schools Example

$$\underline{y_j \sim N(\theta_j, \sigma_j^2)}, \quad \theta_j \sim N(\mu, \frac{\sigma^2}{K})$$

- y_j is the observed effects of the program in school j
 - Based on a sample of test scores from those in the program and those not in the program
- θ_j are the true *unknown* effects of the program in school j
- Variances, σ_j^2 , are known
 - Determined by the number of students in the sample

$$\sigma_j^2 = \frac{\sigma^2}{n_j}$$

$$P(\theta_j | y_j) \sim N(w y_j + (1-w)\mu, w \sigma_j^2)$$

$$w = \frac{n_j}{n_j + K}$$

Eight Schools Example

```
J <- 8
y = c(28, 8, -3, 7, -1, 1, 18, 12)
sigma <- c(15, 10, 16, 11, 9, 11, 10, 18)  $\sigma_j^2$ 
```

- Assuming the effect of the program on each school is identical.
- What are the chances of seeing a value as large as 28?
- As small as -3?

$$y_j \sim N(\theta, \sigma_j^2)$$

$$L(\theta) \propto \prod_{i=1}^8 \frac{1}{\sqrt{2\pi\sigma_j^2}} e^{-\frac{(y_i - \theta)^2}{2\sigma_j^2}}$$

$$\hat{\theta}_{MLE} = \frac{\sum_{i=1}^8 1/\sigma_j^2 y_j}{\sum 1/\sigma_j^2}$$

Why is MLE
not \bar{y} ?

Eight Schools Example

If effects are actual equal, what is it?

```
## Compute the precision from each school
prec <- 1/sigma^2

## global estimate is a weighted average
mu_global <- sum(prec * y / sum(prec))
mu_global
```

```
## [1] 7.685617
```

$\hat{\theta}_{MLE} \approx 7.7$ if no
variation between schools.

Eight Schools Example

- Assume the effect of the program on each school is identical, i.e.

$$\theta_j = \theta = 7.7$$

- What are the chances of seeing a value as large as 28?
- As small as -3?

```
# 1000 datasets with mean mu_global but different sigmas  
  
## Chance of seeing a value greater than 28  
mean(sapply(1:1000, function(x)  
  max(rnorm(J, mean=mu_global, sd=sigma))) >= 28)
```

```
## [1] 0.358
```

1/
7.7

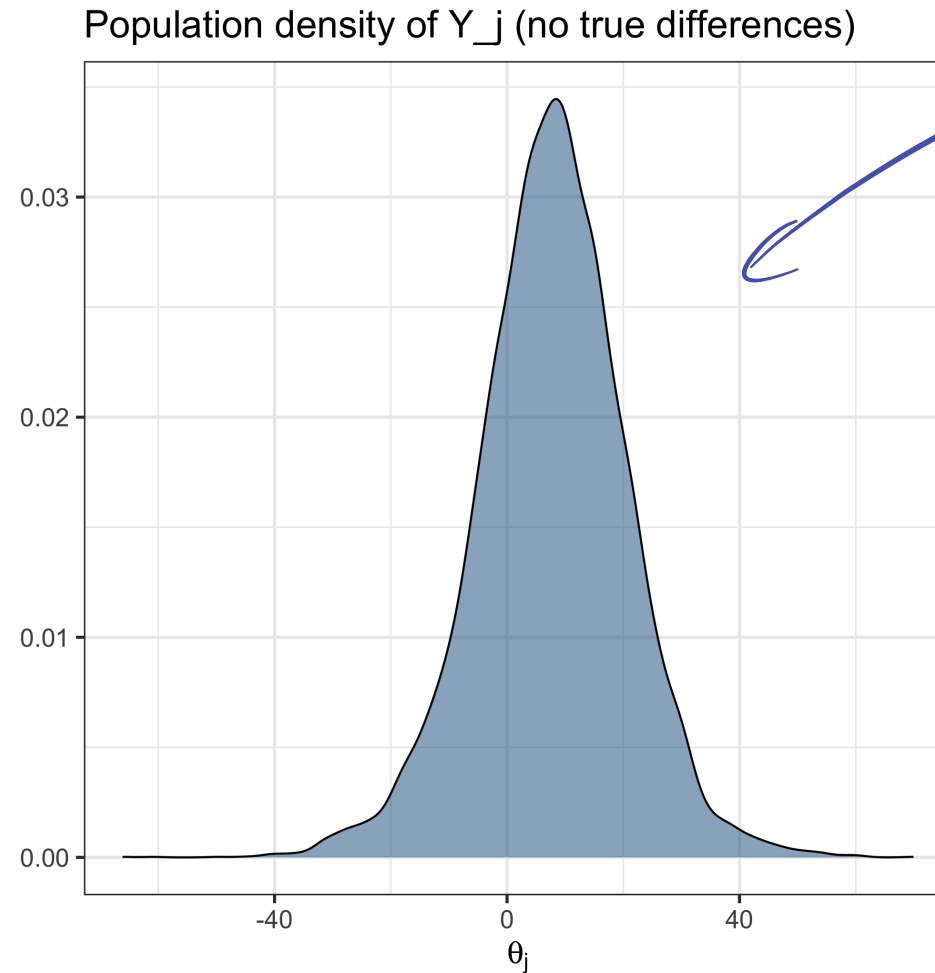
σ_j

```
## Chance of seeing a value less than -3  
mean(sapply(1:1000, function(x)  
  min(rnorm(J, mean=mu_global, sd=sigma))) <= -3)
```

```
## [1] 0.799
```


Eight Schools Example

11.7.7
Density of $Y_j \sim N(\theta_{\text{pooled}}, \sigma_j^2)$



*How much
y can
vary due
to chance.*

Eight Schools Example

$$y_j \sim N(\theta_j, \sigma_j^2)$$


- θ_j are the true unknown effects of the program in school j
- y_j is the observed effects of the program in school j
 - Based on a sample of test scores from those in the program and those not in the program
 - Number of people in the sample determine the magnitude of σ_j^2

$$\hat{\theta}_{j,MLE} = y_j$$

Eight Schools Example

- How do we estimate θ_j ?

- Independent: $\hat{\theta}_j^{(MLE)} = y_j$ is the MLE

- Identical effects: $\hat{\theta}_j^{(pool)} = \frac{\sum_i \frac{1}{\sigma_i^2} y_i}{\sum_i \frac{1}{\sigma_i^2}} = \theta^{(pool)}$ for all j ≈ 7.7.

- Same effect for all schools: estimate using a weighted average of the observed effects

Eight Schools

Bias - Variance Tradeoff.

```
theta_j_mle <- y  
theta_j_mle
```

```
## [1] 28  8 -3  7 -1  1 18 12
```

```
theta_j_pooled <- rep(sum(1/sigma^2 * y) / sum(1/sigma^2), J)  
theta_j_pooled
```

```
## [1] 7.685617 7.685617 7.685617 7.685617 7.685617 7.685617 7.685617 7.685
```

Is there a middle ground between two extremes?

$$\hat{\theta}_{i, \text{partial}} = \tilde{w} y_i + (1 - \tilde{w}) \hat{\theta}_{\text{pooled}}$$

Eight schools example

A hierarchical model:

How variable
are the true SAT
effects across schools,

- Add a shared normal prior distribution to θ_j
- Assume the global mean, μ is also unknown parameters.
- How do we choose prior for μ ?

$\mu \sim N(\mu_0, \tau_0^2)$? or $p(\mu) \propto 1$?

(hyperprior)

- Need to estimate all of $(\mu, \theta_1, \dots, \theta_8)$ with MCMC
- τ^2 determines how much weight we put on the independent estimate vs the pooled estimate

θ_{pooled}

$$\theta_i \sim N(\mu, \tau^2)$$

$$y_j \sim N(\theta_j, \sigma_j^2)$$

$$\theta_{j, \text{PM}} = w y_j + (1-w) \mu$$

$$w = \frac{1/\sigma_j^2}{1/\sigma_j^2 + 1/\tau^2}$$

random
(not known)

9 - parameters
posterior.

$$P(\mu, \theta_1, \dots, \theta_8 / y_1, \dots, y_8, \sigma_j^2)$$

$$P(\mu, \theta_1, \dots, \theta_g / y_1, \dots, y_g, \sigma_j) \propto$$

$$\prod_{i=1}^g \left[\frac{1}{\sqrt{2\pi}\sigma_j} e^{-\frac{(y_i - \theta_j)^2}{2\sigma_j^2}} \frac{1}{\sqrt{2\pi}\tau^2} e^{-\frac{(\theta_j - \mu)^2}{2\tau^2}} \right] P(\mu)$$

- Compromise: $\hat{\theta}_j^{(shrink)} = w\theta_j^{(MLE)} + (1 - w)\theta_j^{(pool)}$
- How is this different from the standard normal-normal model we saw before?

Intuition behind shrinkage

- $Y_j = \theta_j + \epsilon_j$ and for simplicity assume that the variance of ϵ , σ^2 for all j
 - θ_j represents true differences between schools (signal)
 - ϵ_j is sampling variability (noise, chance variation)
- $Var(Y_j) = Var(\theta_j) + Var(\epsilon_j) = \tau^2 + \sigma_j^2$
 - The variance of the observed outcomes is the sum of signal variance, τ^2 , and the sampling variance σ_j^2
- Consequence: the observed outcomes always have higher variance across groups than the signal
 - $Var(Y_j) > Var(\theta_j)$
- Intuition: reduce the variance by shrinking Y_j 's closer together!
 - Want the variance of the shrunken estimates to be close to τ^2

Eight Schools examples

$$\theta_j \sim N(\mu, \tau^2)$$

Comments:

- The global average, μ , is a parameter so also has uncertainty
- How do we determine how much to shrink, e.g. how do we determine τ^2 ? \rightarrow signal variability.
- Is the training program effective in school j ?
 - What is $P(\theta_j > 0 \mid y)$?
- On average (over all schools) is the training program effective?
 - What is $P(\mu > 0 \mid y)$?

Eight schools example

- If τ^2 is large, the prior for θ_j is not very strong
 - If $\tau^2 \rightarrow \infty$ equivalent to the no pooling model
- If τ^2 is small, we assume a priori that θ_j are very close
 - if $\tau^2 \rightarrow 0$ equivalent to the complete pooling model, $\theta_j = \mu$

$$\theta_{j,pm} = w y_j + (1-w) \mu$$

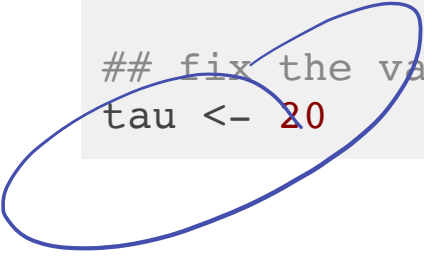
$$w = \frac{1/\sigma_j^2}{1/\sigma_j^2 + \frac{1}{\tau^2}}$$

Estimating parameters

- MH: Need to generate a proposal from a 9-dimensional posterior distribution
 - Eight parameters for θ_j and one for μ
- Gibbs: sample each of the 9 parameters from the complete conditionals
 - Sample $p(\theta_j \mid \theta_{-j}, \mu)$
 - Sample $p(\mu \mid \theta_1, \dots, \theta_8, \mu)$
- Stan

Eight Schools Estimation

```
J <- 8  
y = c(28, 8, -3, 7, -1, 1, 18, 12)  
sigma <- c(15, 10, 16, 11, 9, 11, 10, 18)  
  
## fix the variance of the prior to a number  
tau <- 20
```



Eight Schools in Stan

```
// saved as 8schools.stan
data {
  int<lower=0> J;           // number of schools
  real y[J];               // estimated treatment effects
  real<lower=0> sigma[J];  // standard error of effect estimates
  real<lower=0> tau;       // shrinkage standard deviation
}
parameters {
  real mu;                // population treatment effect (global average)
  vector[J] eta;          // unscaled deviation from mu by school  $\theta_j$ 
}
transformed parameters {
  vector[J] theta = mu + tau * eta; // school treatment effects
}
model {
  target += normal_lpdf(eta | 0, 1); // prior log-density
  target += normal_lpdf(y | theta, sigma); // log-likelihood
}
```

A number I choose.

(global average)

theta

θ_j

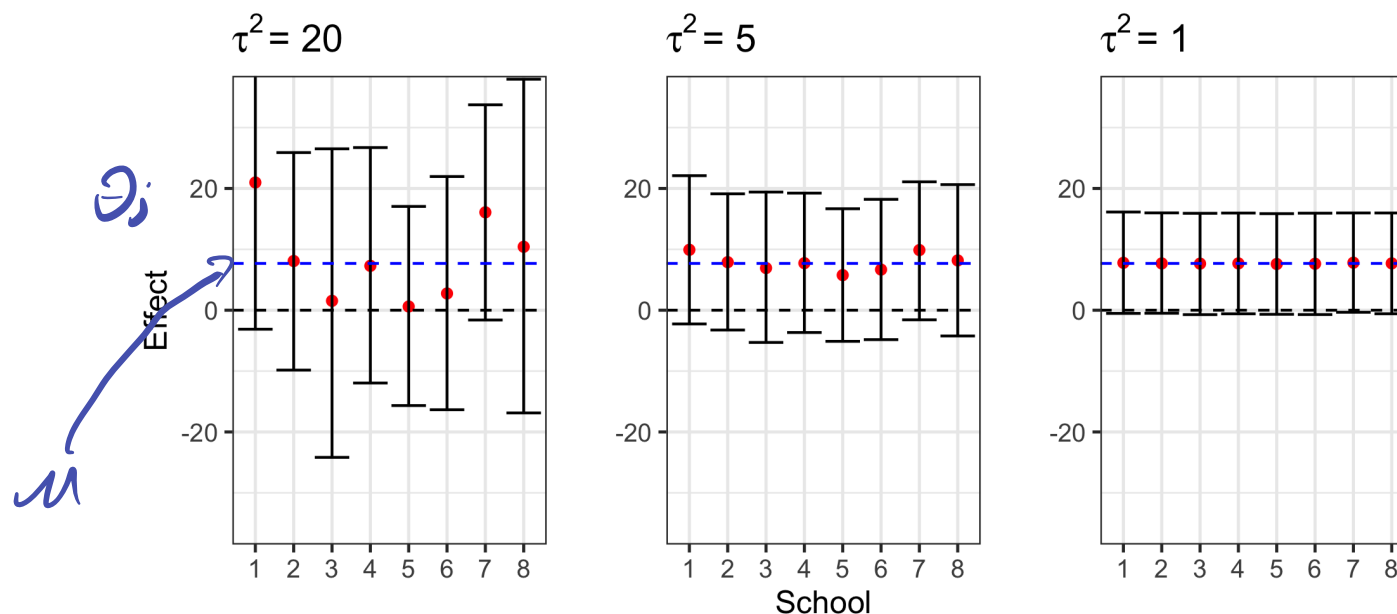


$y \sim \text{normal}(\theta, \sigma)$

$\theta \sim \text{normal}(\mu, \tau)$

($\mu \propto \text{const}$, flat a prior)

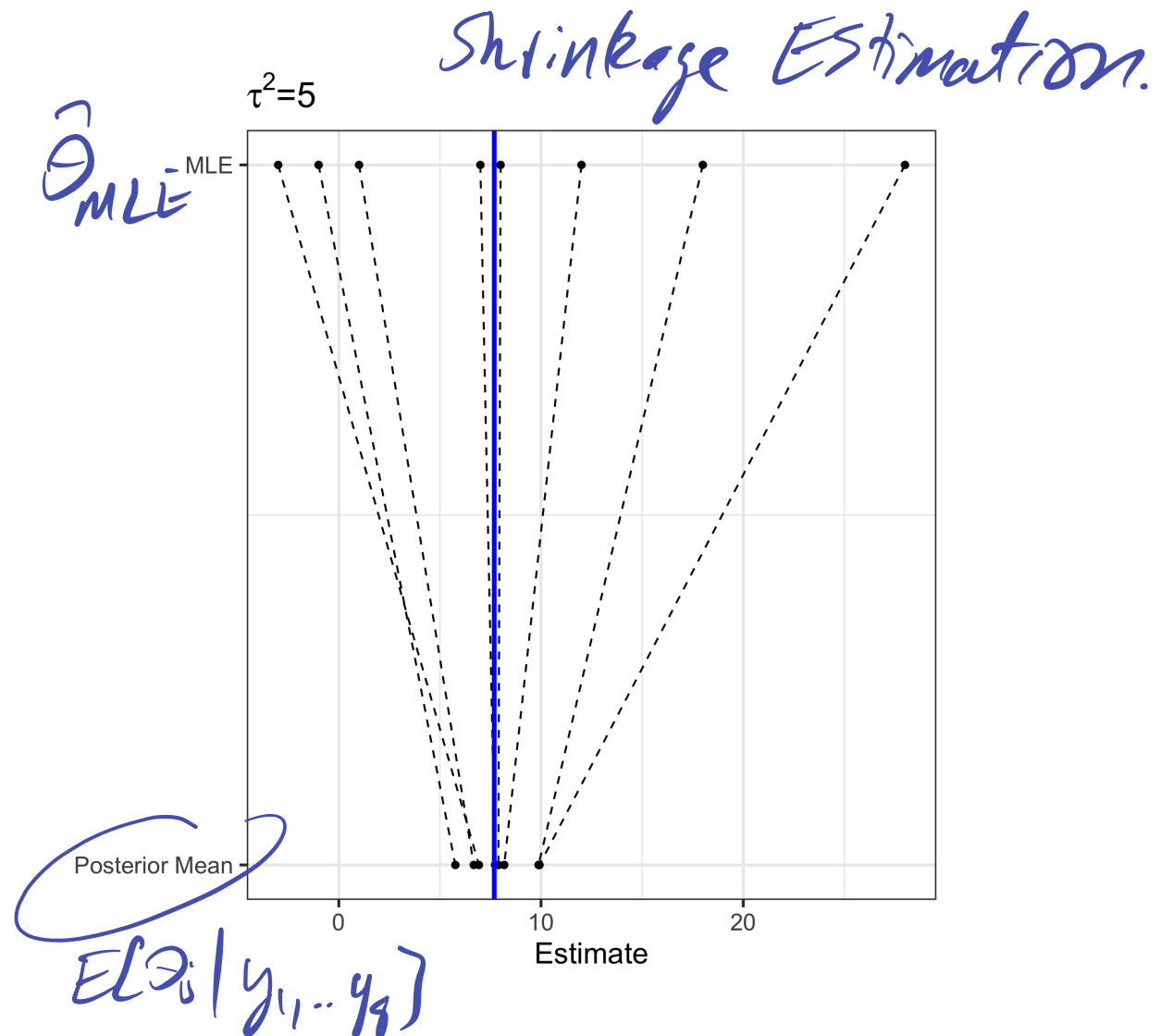
Eight Schools example



- Larger τ^2 means more variability in θ_j .
- Larger τ^2 means more posterior uncertainty.

$$\theta_j \sim N(w y_j + (1-w)\mu, w\sigma_j^2)$$

MLE vs Posterior Mean



Basketball Example

θ_i is
 FG%
 $y_i \sim \text{Fraction of makes}$
 μ is
 league average
 FG%

