# Lab Report: Lab5 - Computational Statistics

Dhanush Kumar Reddy Narayana Reddy (dhana004), Udaya Shanker Mohanan Nair (udamo524)

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## Introduction

Implementation of 2 Assignment questions of Computational Statistics Lab 5.

### Contributions

Member: Dhanush Kumar Reddy Narayana Reddy, Liu Id: dhana004, Contribution: Report writing and coding of question 1.

Member: Udaya Shanker Mohanan Nair, Liu Id: udamo524, Contribution: Report writing and coding of question 2.

# Question 1

### Part A

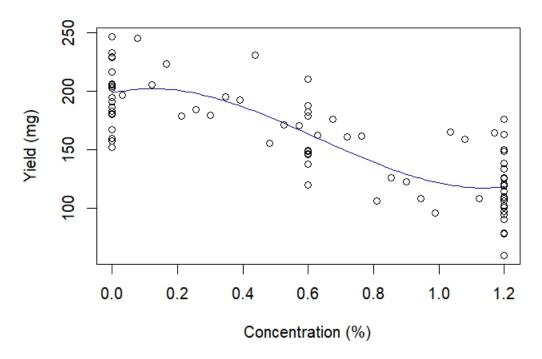
$$y = \beta_0 + \beta_1 x + \beta_2 x^2 + \beta_3 x^3 + \epsilon$$

The summary of Cubic regression model:

```
Call:
lm(formula = yield ~ poly(concentration, 3, raw = TRUE), data = data)
Residuals:
    Min
             1Q Median
                             3Q
                                   Max
-58.909 -17.239 -3.878 18.138 58.091
Coefficients:
                                    Estimate Std. Error t value Pr(>|t|)
                                                 5.583 35.564
(Intercept)
                                     198.568
                                                                 <2e-16 ***
poly(concentration, 3, raw = TRUE)1
                                     66.714
                                                72.748
                                                         0.917
                                                                 0.3620
poly(concentration, 3, raw = TRUE)2 -305.221
                                               166.937
                                                        -1.828
                                                                 0.0714 .
poly(concentration, 3, raw = TRUE)3 161.634
                                                92.076
                                                         1.755
                                                                 0.0832 .
Signif. codes: 0 '*** 0.001 '** 0.01 '* 0.05 '.' 0.1 ' 1
Residual standard error: 26.58 on 76 degrees of freedom
Multiple R-squared: 0.6406,
                              Adjusted R-squared: 0.6264
F-statistic: 45.15 on 3 and 76 DF, p-value: < 2.2e-16
```

Part B

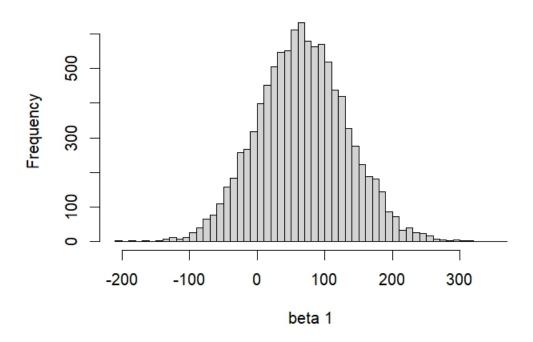
# Yield vs. Concentration



 $95\%\text{-}\mathrm{Confidence}$  Interval Analytical (lm): -78.17667 211.605

### Part C

# **Bootstrap Distribution for beta 1**



95%-Bootstrap Confidence Interval Bootstrap Percentile (manual): -61.1966 198.2414

### Part D

```
# BOOTSTRAP CONFIDENCE INTERVAL CALCULATIONS
# Based on 10000 bootstrap replicates
#
# CALL :
# boot.ci(boot.out = bs_result, type = c("perc", "bca"))
#
# Intervals :
# Level Percentile BCa
# 95% (-63.34, 199.09 ) (-59.78, 200.63 )
# Calculations and Intervals on Original Scale
```

The 95% confidence interval for beta1 using the percentile method is (-63.34, 199.09), while the BCa method provides a slightly adjusted interval of (-59.78, 200.63).

#### Part E

### Confidence Interval Comparison

Method	95% Confidence Interval for beat1
Analytical (lm)	[-78.17667, 211.60502]
Bootstrap Percentile (manual)	[-61.1966, 198.2414]
Bootstrap Percentile (boot package)	[-63.34, 199.09]
Bootstrap BCa (boot package)	[-59.78, 200.63]

#### Comparision:

**Analytical vs. Bootstrap Approaches:** (1) The analytical confidence interval (from lm) is wider than all the bootstrap-based intervals.

(2) The analytical method assumes normality of residuals, which might not be valid if the error distribution is skewed or has outliers.

Bootstrap Methods (Percentile and BCa): (1) The manual percentile method and boot package percentile method give almost the same interval, which suggests that the bootstrap implementation is correct.

(2) The BCa (bias-corrected and accelerated) interval is slightly different from the percentile intervals, particularly at the lower bound (-59.78 vs -63.34). This suggests that the bootstrap distribution is slightly skewed, and the BCa method accounts for this bias.

Width of Confidence Intervals: (1) The analytical CI [-78.17667, 211.60502] is the widest.

- (2)The BCa method [-59.78, 200.63] is slightly narrower than the other bootstrap CIs, adjusting for bias and skewness.
- (3) The bootstrap CIs provide more precise estimates than the analytical CI, which suggests that the normality assumption in the analytical approach might not hold perfectly.

**Conclusion:** (1) The bootstrap confidence intervals are more reliable because they do not rely on normality assumptions and directly estimate variability from resampling.

- (2) The BCa interval is likely the best choice, as it adjusts for potential bias and skewness in the bootstrap distribution.
- (3) The analytical method overestimates uncertainty, possibly due to non-normal residuals or influential points.

# Question 2

Given a Gumbel distribution with scale parameter 1 and location parameter  $\mu + c$ , where  $c = \log(\log(2))$ . And this distribution has the following distribution function. For this distribution, median of a random variable is .

Now we are gone to generate random variables(Gumbel) using Inverse Transformation Method.

The CDF of the Gumbel distribution is given by

$$F(x) = \exp(-\exp(-(x-\mu-c)))$$

where

$$c = \log(\log(2))$$

Inverse Transformation of this function for a random variable is given by  $Y \sim Y(0,1)$ .

$$X = F^{-1}(Y)$$

$$F(x) = \exp(-\exp(-(x - \mu - c)))$$

Set Y = F(x), where  $Y \sim Y(0, 1)$ 

Take the natural logarithm on both sides:

$$\log(Y) = -\exp(-(x-\mu-c))$$

Take the logarithm again:

$$\log(-\log(Y)) = -(x - \mu - c)$$

Solve for x:

$$x = -\log(-\log(Y)) + \mu + c$$

## Loading required package: lattice

##

## Attaching package: 'BSDA'

## The following object is masked from 'package:datasets':

##

## Orange

Now we need to find out power of the test for median values ranging from 0 to 2, where for each value we will using number of observations, n as 13 and then used sign test(used SIGN.test in this case).

I have tested for 50 different values of median.

Jotting Power of few median values.

## Median Value: 0 Power: 0.03

## Median Value: 0.5306122 Power: 0.196

## Median Value: 1.020408 Power: 0.693

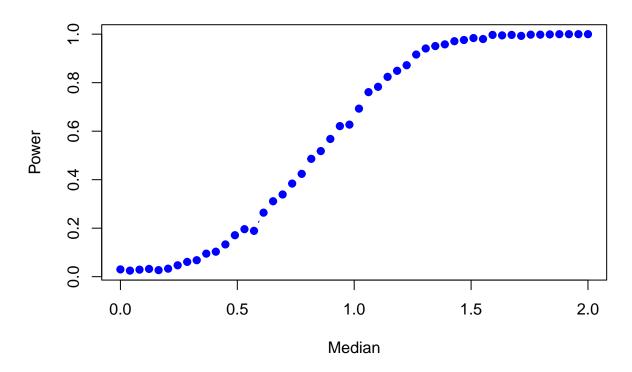
## Median Value: 1.510204 Power: 0.984

## Median Value: 2 Power: 1

from this we understand as the value of median increases the power also increases.

Now plotting the power curve

# **Power Curve**



# Appendix

# Question 1

```
# Question 1: Bootstrap for regression
# Part A
# Load the dataset
data <- read.table("C:/Users/dhanu/OneDrive/Documents/Computational_Statistics/Lab5/kresseertrag.dat", '</pre>
colnames(data) <- c("observation", "concentration", "yield")</pre>
# Fit a cubic regression model
model <- lm(yield ~ poly(concentration, 3, raw = TRUE), data = data)</pre>
# Display model summary
summary(model)
# Part B
\mbox{\it \#} The coefficients and 95%-confidence intervals
coefficients <- coef(model)</pre>
conf_intervals <- confint(model, level = 0.95)</pre>
cat("95%-Confidence Interval Analytical (lm): \n")
cat(conf_intervals[2, ])
# -78.17667 211.605
```

```
# Plot
plot(data$concentration, data$yield, main = "Yield vs. Concentration",
     xlab = "Concentration (%)", ylab = "Yield (mg)")
curve(coefficients[1] + coefficients[2]*x + coefficients[3]*x^2 +
        coefficients[4]*x^3, add = TRUE, col = "blue")
# Part C
set.seed(123)
b0 <- 10000
# bootstrap resampling function
bootstrap <- function(data, b0, parameter_index) {</pre>
  bs_beta1 <- numeric(b0)</pre>
  for (i in 1:b0) {
    bs_data <- data[sample(nrow(data), replace = TRUE), ]</pre>
    bs_model <- lm(yield ~ poly(concentration, 3, raw = TRUE), data = bs_data)
    bs_beta1[i] <- coef(bs_model)[2]</pre>
 return(bs_beta1)
bs_beta1 <- bootstrap(data, b0, 2)
# 95%-bootstrap confidence interval 95% percentile interval
bs_ci \leftarrow quantile(bs_beta1, probs = c(0.025, 0.975))
cat("95%-Bootstrap Confidence Interval Bootstrap Percentile (manual): \n")
cat(bs_ci)
# -61.1966 198.2414
# Plot histogram of bootstrap distribution
hist(bs_beta1, main = "Bootstrap Distribution for beta 1", xlab = "beta 1",
     breaks = 50)
# Part D
library(boot)
# Function to fit the model and extract the parameter of interest
bs fn <- function(data, indices) {</pre>
 bs data <- data[indices, ]
 bs_model <- lm(yield ~ poly(concentration, 3, raw = TRUE), data = bs_data)
 return(coef(bs_model)[2])
bs_result <- boot(data, statistic = bs_fn, R = b0)</pre>
# Percentile and BCa confidence intervals
ci_perc_bca <- boot.ci(bs_result, type = c("perc", "bca"))</pre>
cat("95%-Percentile and 95%-BCa Confidence Interval: \n")
ci_perc_bca
# BOOTSTRAP CONFIDENCE INTERVAL CALCULATIONS
# Based on 10000 bootstrap replicates
# CALL :
   boot.ci(boot.out = bs_result, type = c("perc", "bca"))
```

```
# Intervals :
             Percentile
 Level
                                    BCa
      (-63.34, 199.09)
                           (-59.78, 200.63)
# Calculations and Intervals on Original Scale
# Part E
# Comparision:
# Analytical vs. Bootstrap Approaches:
# (1)The analytical confidence interval (from lm) is wider than all the bootstrap-based intervals.
# (2) The analytical method assumes normality of residuals, which might not be valid if the error distri
# Bootstrap Methods (Percentile and BCa):
# (1)The manual percentile method and boot package percentile method give almost the same interval, whi
# (2) The BCa (bias-corrected and accelerated) interval is slightly different from the percentile interv
# Width of Confidence Intervals:
\# (1) The analytical CI [-78.17667, 211.60502] is the widest.
# (2)The BCa method [-59.78, 200.63] is slightly narrower than the other bootstrap CIs, adjusting for b
# (3) The bootstrap CIs provide more precise estimates than the analytical CI, which suggests that the n
# Conclusion:
# (1)The bootstrap confidence intervals are more reliable because they do not rely on normality assumpt
# (2) The BCa interval is likely the best choice, as it adjusts for potential bias and skewness in the b
# (3) The analytical method overestimates uncertainty, possibly due to non-normal residuals or influenti
```

### Question 2

```
library(BSDA)
c \leftarrow log(log(2))
gumbel_fn <- function(median,n = 13) {</pre>
  Y <- runif(n)
  X \leftarrow -\log(-\log(Y)) + \text{median} + c
  return(X)
}
median_values <- seq(0, 2, length.out = 50)</pre>
len <- length(median_values)</pre>
power <- numeric()</pre>
for (j in seq_along(median_values)) {
  median <- median_values[j]</pre>
  count <- 0
  for (i in 1:1000) {
    data <- gumbel_fn(median)</pre>
    test <- SIGN.test(data)</pre>
    if(test$p.value < 0.05){
       count <- count + 1
  power[j] <- count / 1000</pre>
cat("Median Value: ", median_values[1]," Power:",power[1])
```