# Hypothesis Testing, Power, Sample Size and Confidence Intervals (Part 1)

B.H. Robbins Scholars Series

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#### Outline

#### Introduction to hypothesis testing

Scientific and statistical hypotheses

Classical and Bayesian paradigms

Type 1 and type 2 errors

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#### Introduction

- Goal of hypothesis testing is to rule out chance as an explanation for an observed effect
- Example: Cholesterol lowering medications
  - 25 people treated with a statin and 25 with a placebo
  - ► Average cholesterol after treatment is 180 with statins and 200 with placebo.
- Do we have sufficient evidence to suggest that statins lower cholesterol?
- ► Can we be sure that statin use as opposed to a chance occurrence led to lower cholesterol levels?

Introduction to hypothesis testing

Scientific and statistical hypotheses

## Hypotheses

- Scientific Hypotheses
  - Often involve estimation of a quantity of interest
  - After amputation, to what extent does treatment with clonidine lead to lower rates of phantom limb pain than with standard therapy? (Difference or ratio in rates)
  - What is the average increase in alanine aminotransferase (ALT) one month after doubling the dose of medication X? (Difference in means)
- Statistical Hypothesis
  - ► A statement to be judged. Usually of the form: population parameter X is equal to a specified constant
  - ▶ Population mean potassium K,  $\mu = 4.0 \text{ mEq/L}$
  - ▶ Difference in population means,  $\mu_1 \mu_2 = 0.0 \text{ mEq/L}$

# Statistical Hypotheses

- ▶ Null Hypothesis: *H*<sub>0</sub>
  - A straw man; something we hope to disprove
  - ▶ It is usually is a statement of no effects.
  - ▶ It can also be of the form  $H_0$ :  $\mu$  =constant, or  $H_0$ : probability of heads equal 1/2.
- Alternative Hypothesis: H<sub>A</sub>
  - What you expect to favor over the null
- ▶ If  $H_0$ : Mean K value = 3.5 mEq/L
  - ▶ One sided alternative hypothesis:  $H_A$ : Mean K > 3.5 mEq/L
  - ► Two-sided alternative hypothesis:  $H_A$ : Mean K  $\neq$  3.5 mEq/L (values far away from the null)

# Classical (Frequentist) Statistics

- Emphasizes hypothesis testing
- ▶ Begin by assuming  $H_0$  is true
- Examines whether data are consistent with H<sub>0</sub>
- Proof by contradiction
  - If, under H₀, the data are strange or extreme, then doubts are cast on the null.
- ► Evidence is summarized with a single statistic which captures the tendency of the data.
- ▶ The statistic is compared to the parameter value given by  $H_0$

# Classical (Frequentist) Statistics

- p-value: Under the assumption that H<sub>0</sub> is true, it is the probability of getting a statistic as or more in favor of H<sub>A</sub> over H<sub>0</sub> than was observed in the data.
- ▶ Low p-values indicate that if *H*<sub>0</sub> is true, we have observed an improbable event.
- Mount evidence against the null, and when sufficient, reject H<sub>0</sub>.
- NOTE: Failing to reject H₀ does not mean we have gathered evidence in favor of it (i.e., absence of evidence does not imply evidence of absence)
  - ► There are many reasons for not rejecting *H*<sub>0</sub> (e.g., small samples, inefficient designs, imprecise measurements, etc.)

# Classical (Frequentist) Statistics

- Clinical significance is ignored.
- ▶ Parametric statistics: assumes the data arise from a certain distribution, often a normal or Gaussian.
- Non-parametric statistics: does not assume a distribution and usually looks at ranks rather than raw values.

# Bayesian Statistics

- We can compute the probability that a statement, that is of clinical significance, is true
  - ► Given the data we observed, does medication X lower the mean cholesterol by more than 10 units?
- May be more natural than the frequentist approach, but it requires a lot more work.
- Supported by decision theory:
- ▶ Begin with a (prior) belief → learn from your data → Form a new (posterior) belief that combines the prior belief and the new data
- We can then formally integrate information accrued from other studies as well as from skeptics.
- Becoming more popular.

# Errors in Hypothesis Testing

- ► Type 1 error: Reject H<sub>0</sub> when it is true
  - Significance level (α) or Type 1 error rate: is the probability of making this type of error
  - ▶ This value is usually set to 0.05 for random reasons
- ▶ Type 2 error: Failing to reject  $H_0$  when it is false
  - The value β is the probability of a type 2 error or type 2 error rate.
- ▶ Power:  $1 \beta$ : probability of correctly rejecting  $H_0$  when it is false

	State of H <sub>0</sub>		
Decision	H₀ is true	$H_0$ is false	
Do not reject $H_0$	Correct	Type 2 error $(\beta)$	
Reject H <sub>0</sub>	Type 1 error $(\alpha)$	Correct	

└─Type 1 and type 2 errors

# Notes Regarding Hypothesis Testing

- ► Two schools of thought
  - Neyman-Pearson: Fix Type 1 error rate (say  $\alpha = 0.05$ ) and then make the binary decision, reject/do not reject
  - ► Fisher: Compute the p-value and quote the report in the publication.
  - We favor Fisher, but Neyman-Pearson is used all of the time.
- ► Fisher approach: discussion of p-values does not require discussion of type 1 and type 2 errors
  - Assume the sample was chosen randomly from a population whose parameter value is captured by  $H_0$ . The p-value is a measure of evidence against it.
- Neyman-Pearson approach: having to make a binary call (reject vs do not reject) regarding significance is arbitrary
  - ▶ There is nothing magical about 0.05
  - Statistical significance has nothing to do with clinical significance

## One sample test for the mean

- ► Assumes the sample is drawn from a population where values are normally distributed (normality is actually not necessary)
- ▶ One sample tests for mean  $\mu = \mu_0$  (constant) don't happen very often except when data are paired (to be discussed later)
- ▶ The t-test is based on the t-statistic

$$t = \frac{\text{estimated value - hypothesized value}}{\text{standard deviation of numerator}}$$

 Standard deviation of a summary statistic is called the standard error which is the square root of the variance of the statistic

## One sample test for the mean

- ▶ Sample average:  $\overline{x} = \frac{1}{n} \sum_{i=1}^{n} x_i$ 
  - ► The estimate of the population mean based on the observed sample
- Sample variance:  $s^2 = \frac{1}{n-1} \sum_{i=1}^{n} (x_i \overline{x})^2$
- ▶ Sample standard deviation:  $s = \sqrt{s^2}$
- $H_0: \mu = \mu_0 \text{ vs. } H_A: \mu \neq \mu_0$
- One sample t-statistic

$$t = \frac{\overline{x} - \mu_0}{SE}$$

▶ Standard error of the mean,  $SE = \frac{s}{\sqrt{n}}$ 

#### One sample t-test for the mean

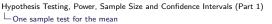
- When data come from a normal distribution and H₀ holds, the t ratio follows the t − distribution. What does that mean?
- Draw a sample from the population, conduct the study and calculate the t-statistic.
- Do it again, and calculate the t-statistic again.
- Do it again and again.
- Now look at the distribution of all of those t-statistics.
- ▶ This tells us the relative probabilities of all t-statistics if *H*<sub>0</sub> is true.

#### Example: one sample t-test for the mean

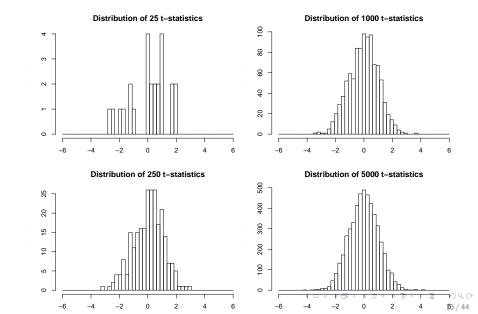
- ▶ The distribution of potassium concentrations in the target population are normally distributed with mean 4.3 and variance .1: N(4.3, .1).
- ▶  $H_0$ :  $\mu = 4.3$  vs.  $H_A$ :  $\mu \neq 4.3$ . Note that  $H_0$  is true!
- Each time the study is done,
  - Sample 100 participants
    - Calculate:

$$t = \frac{\overline{x} - 4.3}{SE}$$

Conduct the study 25 times, 250 times, 1000 times, 5000 times



☐ Hypothesis testing

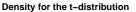


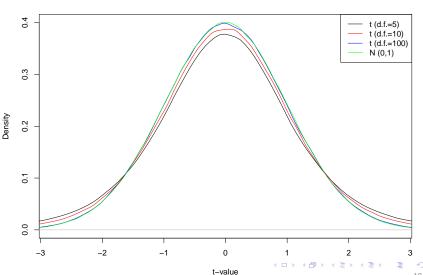
#### One sample t-test for the mean

- With very small samples (n), the t statistic can be unstable because the sample standard deviation (s) is not a precise estimate of the population standard deviation ( $\sigma$ ).
- ▶ So, the t-statistic has heavy tails for small *n*
- As n increases, the t-distribution converges to the normal distribution with mean equal to 0 and with standard deviation equal to one.
- ► The parameter defining the particular t-distribution we use (function of n) is called the degrees of freedom or d.f.
- ▶ d.f. = n number of means being estimated
- ▶ For the one-sample problem, d.f.=n-1
- $\triangleright$  Symbol is  $t_{n-1}$

Hypothesis Testing, Power, Sample Size and Confidence Intervals (Part 1)  $\cup$  One sample test for the mean

∟ Hypothesis testing





#### One sample t-test for the mean

- ▶ One sided test:  $H_0$ :  $\mu = \mu_0$  versus  $H_A$ :  $\mu > \mu_0$
- One tailed p-value:
  - ▶ Probability of getting a value from the  $t_{n-1}$  distribution that is at least as much in favor of  $H_A$  over  $H_0$  than what we had observed.
- ▶ Two-sided test:  $H_0$ :  $\mu = \mu_0$  versus  $H_A$ :  $\mu \neq \mu_0$
- Two-tailed p-value:
  - ▶ Probability of getting a value from the  $t_{n-1}$  distribution that is at least as big **in absolute value** as the one we observed.

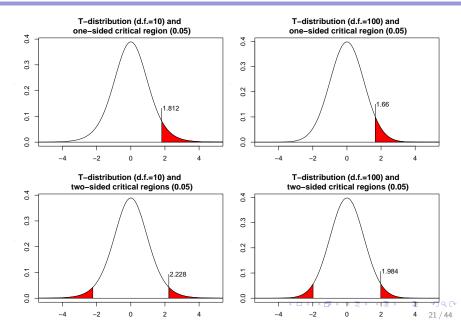
#### One sample t-test for the mean

- Computer programs can compute the p-value for a given n and t-statistic
- Critical value
  - ▶ The value in the t (or any other) distribution that, if exceeded, yields a 'statistically significant' result for type 1 error rate equal to  $\alpha$
- Critical region
  - ▶ The set of all values that are considered statistically significantly different from  $H_0$ .

Hypothesis Testing, Power, Sample Size and Confidence Intervals (Part 1)

—One sample test for the mean

☐ Hypothesis testing



## Power and Sample Size for a one sample test of means

- Power increases when
  - ▶ Type 1 error rate  $(\alpha)$  increases: type 1  $(\alpha)$  versus type 2  $(\beta)$  tradeoff
  - ▶ True  $\mu$  is very far from  $\mu_0$
  - Variance or standard deviation  $(\sigma)$  decreases (decrease noise)
  - Sample size increases
- T-statistic

$$t = \frac{\overline{x} - \mu_0}{\sigma / \sqrt{n}}$$

Power for a 2-tailed test is a function of the true mean  $\mu$ , the hypothesized mean  $\mu_0$ , and the standard deviation  $\sigma$  only through  $|\mu - \mu_0|/\sigma$ 

# Power and Sample Size for a one sample test of means

▶ Sample size to achieve  $\alpha = 0.05$ , power=0.90 is approximately

$$n = 10.51 \left(\frac{\sigma}{\mu - \mu_0}\right)^2$$

- Power calculators can be found at statpages.org/#Power
- PS is a very good power calculator (Dupont and Plummer): http://biostat.mc.vanderbilt.edu/PowerSampleSize

# Example: Power and Sample Size

- ► The mean forced expiratory volume in 1 second in a population of asthmatics is 2.5 L/sec, and the standard deviation is assumed to be 1
- ▶ How many subjects are needed to reject  $H_0: \mu = 2.5$  in favor of  $H_0: \mu \neq 2.5$  if the new drug is expected to increase the FEV to 3 L/sec with  $\alpha = 0.05$  and  $\beta = 0.1$
- $\mu_0 = 2.5, \ \mu = 3.0, \ \sigma = 1$

$$n = 10.51 \left(\frac{1}{3.0 - 2.5}\right)^2 = 42.04$$

▶ We need 43 subjects to have 90 percent power to detect a 0.5 difference from 2.5.

#### Confidence Intervals

▶ Two-sided,  $100(1-\alpha)\%$  CI for the mean  $\mu$  is given by

$$(\overline{x} - t_{n-1,1-\alpha/2} \cdot SE, \overline{x} + t_{n-1,1-\alpha/2} \cdot SE)$$

- $t_{n-1,1-\alpha/2}$  is the critical value from the t-distribution with d.f.=n-1
- ▶ For large n,  $t_{n-1,1-\alpha/2}$  is equal to 1.96 for  $\alpha = 0.05$
- ightharpoonup 1-lpha is called the confidence level or confidence coefficient

#### Confidence Intervals

- ▶  $100(1-\alpha)\%$  confidence interval (CI)
  - If we were able to repeat a study a large number of times, then  $100 \cdot (1 \alpha)$  percent of CIs would contain the true value.
- ▶ Two-sided  $100(1-\alpha)\%$  CI
  - Includes the null hypothesis  $\mu_0$  if and only if a hypothesis test  $H_0$ :  $\mu=\mu_0$  is not rejected for a 2-sided  $\alpha$  significance level test.
  - ▶ If a 95% CI does not contain  $\mu_0$ , we can reject  $H_0: \mu = \mu_0$  at the  $\alpha = 0.05$  significance level

n	$\overline{X}$	$\sigma$	p-value	95% CI
20	27.31	54.23	0.036	(1.930, 52.690)
20	27.31	59.23	0.053	(-0.410, 55.030)
20	25.31	54.23	0.051	(-0.070, 50.690)
17	27.31	54.23	0.054	(-0.572, 55.192)

Cls provide more information than p-values

# Special case: Paired data and one-sample tests

- Assume we want to study whether furosemide (or lasix) has an impact on potassium concentrations among hospitalized patients.
- ► That is, we would like to test  $H_0$ :  $\mu_{on-furo} \mu_{off-furo} = 0$  versus  $H_A$ :  $\mu_{on-furo} \mu_{off-furo} \neq 0$
- ▶ In theory, we could sample  $n_1$  participants not on furosemide and compare them to  $n_2$  participants on furosemide
- However, a very robust and efficient design to test this hypothesis is with a paired sample approach
- ▶ On n patients, measure K concentrations just prior to and 12 hours following furosemide administration.

## Special case: Paired data and one-sample tests

- ► The effect measure to test H<sub>0</sub> versus H<sub>A</sub>, is the mean, within person difference between pre and post- administration K concentrations.
- $V_i = Y_{on-furo,i} Y_{off-furo,i}$
- ▶ Note that  $\overline{W} = \overline{Y}_{on-furo} \overline{Y}_{off-furo}$ 
  - ► The average of the differences is equal to the difference between the averages
- ▶  $H_0$ :  $\mu_w = 0$  versus  $H_A$ :  $\mu_w \neq 0$  is equivalent to the above  $H_0$  and  $H_A$
- $\overline{W} = -0.075 \text{ mEq/L and } s = 0.08$

$$t_{99} = \frac{-0.075 - 0}{0.08 / \sqrt{100}} = 9.375$$

► The p-value is less than  $0.0001 \rightarrow a$  highly (!!!!) statistically significant reduction

# One Sample Methods for a Probability

- Y is binary (0/1): Its distribution is bernoulli(p) (p is the probability that Y = 1).
- **>** p is also the mean of Y and p(1-p) is the variance.
- ▶ We want to test  $H_0$  :  $p = p_0$  versus  $H_A$  :  $p \neq p_0$
- ▶ Estimate the population probability p with the sample proportion or sample average  $\hat{p}$

$$\hat{p} = \frac{1}{n} \sum_{i=1}^{n} Y_i$$

# One Sample Methods for a Probability

► A z-test is an approximate test that assumes the test statistic has a normal distribution i.e., it is a t-statistic with the d.f. very large

$$z = \frac{\hat{p} - p_0}{\sqrt{p_0(1 - p_0)/n}}$$

▶ The z-statistic has the same form as the t-statistic

$$z = \frac{\text{estimated value - hypothesized value}}{\text{standard deviation of numerator}}$$

where  $\sqrt{p_0(1-p_0)/n}$  is the standard deviation of the numerator which is the standard error assuming the  $H_0$  is true.

(see t-statistic distributions)

# One Sample test for a probability: Is our coin fair?

- $Y \sim bernoulli(p)$ :  $H_0: p = 0.5$  versus  $H_A: p \neq 0.5$
- ▶ Flip the coin 50 times. Heads (Y=1) shows up 30 times  $(\hat{p} = 0.6)$ .

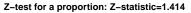
$$z = \frac{0.6 - 0.5}{\sqrt{(0.5)(0.5)/50}} = 1.414$$

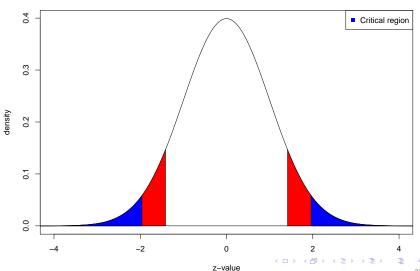
- ▶ The p-value associated with Z is  $2 \times$  the area under the normal curve to the right of z=1.414 (e.g. the area to the right of 1.414 plus the area to the left of -1.414)
- ▶ The critical value for a 2-sided  $\alpha = 0.05$  significance level test is 1.96
- ▶ The p-value associated with this test is approximately 0.16
- Note that if p is very small or very large or if n is small, use exact methods (e.g. Fishers exact test or permutation test)

 $\hbox{Hypothesis Testing, Power, Sample Size and Confidence Intervals (Part 1) } \\$ 

One sample methods for a probability

Hypothesis testing





One sample methods for a probability

Power, confidence intervals, and sample size

#### Power and confidence intervals

- Power increases when
  - n increases
  - p departs from p<sub>0</sub>
  - $ightharpoonup p_0$  departs from 0.5

$$z = \frac{\hat{p} - p_0}{\sqrt{p_0(1 - p_0)/n}}$$

- Confidence interval
  - ▶ 95%CI:  $(\hat{p} 1.96 \cdot \sqrt{\hat{p}(1 \hat{p}/n)}, \hat{p} 1.96 \cdot \sqrt{\hat{p}(1 \hat{p}/n)})$
- ▶ For the coin flipping example:  $\hat{p} = 0.6$  and the 95% CI is given by

$$0.6 \pm 1.96 \cdot \sqrt{0.6 \times 0.4/50} = (0.464, 0.736)$$

which is consistent with the 0.16 p-value that we had observed for  $H_0$ : p = 0.5.

☐ Hypothesis tests

## Two sample test for means

- Two groups of patients (not paired)
- ▶ These are much more common than 1 sample tests
- We assume data come from a normal distribution (although this is not completely necessary)
- For now, assume the two groups have equal variability in response distribution
- Test whether population means are equal
- Example: All patient in population 1 are treated with clonidine after limb amputation and all patients in population 2 are treated with standard therapy.
- Scientific question:
  - ► What is the difference in the mean pain scale scores at 6 months following the amputation?

## Two sample test for means

- ▶  $H_0$ :  $\mu_1 = \mu_2$  which can be generalized to  $H_0$ :  $\mu_1 \mu_2 = 0$  or  $H_0$ :  $\mu_1 \mu_2 = \delta$
- ▶ The quantity of interest (QOI) is  $\mu_1 \mu_2$
- ▶ If we want to test  $H_0$ :  $\mu_1 \mu_2 = 0$  and if we assume the two populations have equal variances, then the t- statistic is given by:

$$t = \frac{\text{point estimate of the QOI} - 0}{\text{standard error of the numerator}}$$

▶ The estimate of the QOI:  $\overline{x}_1 - \overline{x}_2$ 

## Two sample test for means

- For two independent samples variance of the sum or of differences in means is equal to the sum of the variances
- ► The variance of the QOI is then given by  $\frac{\sigma^2}{n_1} + \frac{\sigma^2}{n_2}$
- We need to estimate a single  $\sigma^2$  from the two samples
- We use a weighted average of the two sample variances

$$s^{2} = \frac{(n_{1} - 1)s_{1}^{2} + (n_{2} - 1)s_{2}^{2}}{n_{1} + n_{2} - 2}$$

► The true standard error of the difference in sample means:  $\sigma \sqrt{\frac{1}{n_1} + \frac{1}{n_2}}$ 

• Estimate with  $s\sqrt{\frac{1}{n_1}+\frac{1}{n_2}}$ 

## Two sample test for means

The t-statistic is given by,

$$t = \frac{\overline{x}_1 - \overline{x}_2}{s\sqrt{\frac{1}{n_1} + \frac{1}{n_2}}}$$

- ▶ Under  $H_0$  t, has a t-distribution with  $n_1 + n_2 2$  degrees of freedom.
- ► The -2 comes from the fact that we had to estimate the center of 2 distributions

# Example: two sample test for means

▶ 
$$n_1 = 8$$
,  $n_2 = 21$ ,  $s_1 = 15.34$ ,  $s_2 = 18.23$ ,  $\overline{x}_1 = 132.86$ ,  $\overline{x}_2 = 127.44$ 

$$s^2 = \frac{7(15.34)^2 + 20(18.23)^2}{7 + 20} = 307.18$$

$$s = \sqrt{307.18} = 17.527$$

$$se = 17.527\sqrt{\frac{1}{8} + \frac{1}{21}} = 7.282$$

$$t = \frac{5.42}{7.282} = 0.74$$

on 27 d.f.

## Example: two sample test for means

- ▶ The two-sided p-value is 0.466
  - ▶ You many verify with the surfstat.org t-distribution calculator
- ▶ The chance of getting a difference in means as large or larger than 5.42 if the two populations have the same mean in 0.466.
- No evidence to suggest that the population means are different.

Power, confidence intervals, and sample size

## Power and sample size: two sample test for means

- Power increases when
  - $ightharpoonup \Delta = \mid \mu_1 \mu_2 \mid \text{increases}$
  - $ightharpoonup n_1$  or  $n_2$  increases
  - $ightharpoonup n_1$  and  $n_2$  are close
  - $ightharpoonup \sigma$  decreases
  - $ightharpoonup \alpha$  increases
- ▶ Power depends on  $n_1$ ,  $n_2$ ,  $\mu_1$ ,  $\mu_2$ , and  $\sigma$  approximately through

$$\frac{\Delta}{\sigma\sqrt{\frac{1}{n_1}+\frac{1}{n_2}}}$$

- When using software to calculate power you can put in 0 for  $\mu_1$  and  $\Delta$  for  $\mu_2$  since all that matters is their difference
- $ightharpoonup \sigma$  is often estimated from pilot data

Power, confidence intervals, and sample size

# Power and sample size: two sample test for means

- Example
  - From available data, ascertain a best guess of  $\sigma$ : assume it is 16.847.
  - Assume  $\Delta$ =5,  $n_1$  = 100,  $n_2$  = 100,  $\alpha$  = 0.05
  - ▶ The surfstat software computes a power of 0.555
- The required sample size decreases with
  - $k = \frac{n_2}{n_1} \rightarrow 1$
  - ▶ ∆ large
  - $ightharpoonup \sigma$  small
  - ightharpoonup lpha large
  - Lower power requirements

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Two sample tests for means

Power, confidence intervals, and sample size

# Power and sample size: two sample test for means

▶ An approximate formula for required sample sizes to achieve power=0.9 with  $\alpha = 0.05$  is

$$n_1 = \frac{10.51\sigma^2(1 + \frac{1}{k})}{\Delta^2}$$

$$n_2 = \frac{10.51\sigma^2(1 + k)}{\Delta^2}$$

$\sigma$	Δ	K	$n_1$	$n_2$	n
16.847	5	1.0	239	239	478
16.847	5	1.5	199	299	498
16.847	5	2.0	177	358	537
16.847	5	3.0	160	478	638

▶ Usually, websites are recommended for these calculations.

Two sample tests for means

Power, confidence intervals, and sample size

# Confidence interval: two sample test for means

Confidence interval

$$[(\overline{x}_1 - \overline{x}_2) - t_{n_1+n_2-2,1-\alpha/2} \times s \times \sqrt{\frac{1}{n_1} + \frac{1}{n_2}}, \\ (\overline{x}_1 - \overline{x}_2) + t_{n_1+n_2-2,1-\alpha/2} \times s \times \sqrt{\frac{1}{n_1} + \frac{1}{n_2}}]$$

Δ	S	$n_1$	n <sub>2</sub>	LCI	UCI
5	16.847	100	100	3.01	6.99
5	16.847	75	125	2.95	7.05
5	16.847	50	150	2.70	7.30

Power, confidence intervals, and sample size

# Summary

- Hypothesis testing, power, sample size, and confidence intervals
  - One sample test for the mean
  - One sample test for a probability
  - ► Two sample test for the mean