Grandària Mostral

Quan és prou gran?

Unitat d'Estadística I Bioinformàtica **28 de Gener de 2022**







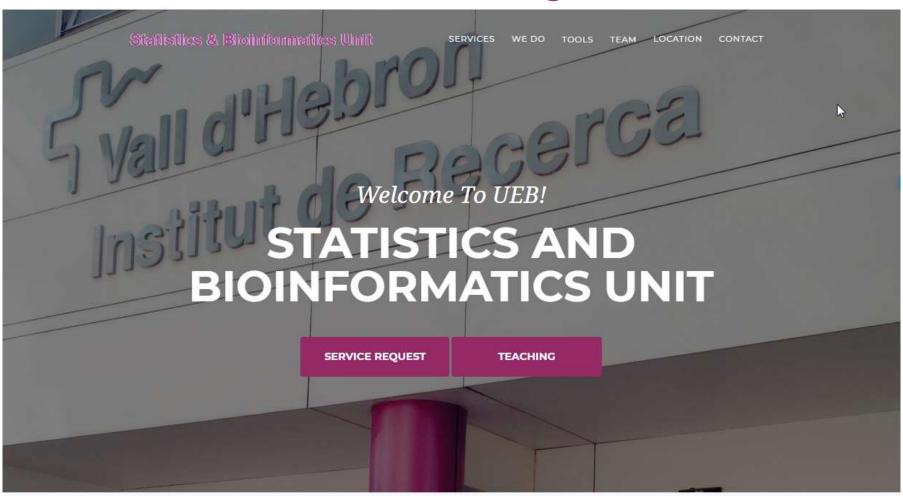
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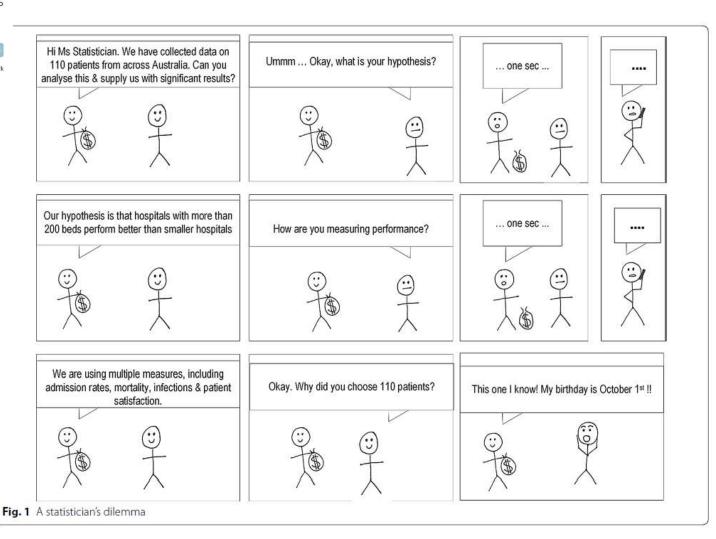
TECHNICAL NOTE

Open Access

When is enough, enough?
Understanding and solving your sample size problems in health services research

Victoria Pye*, Natalie Taylor, Robyn Clay-Williams and Jeffrey Braithwaite

The usual situation when a researcher ask a statistician about sample size for a research study

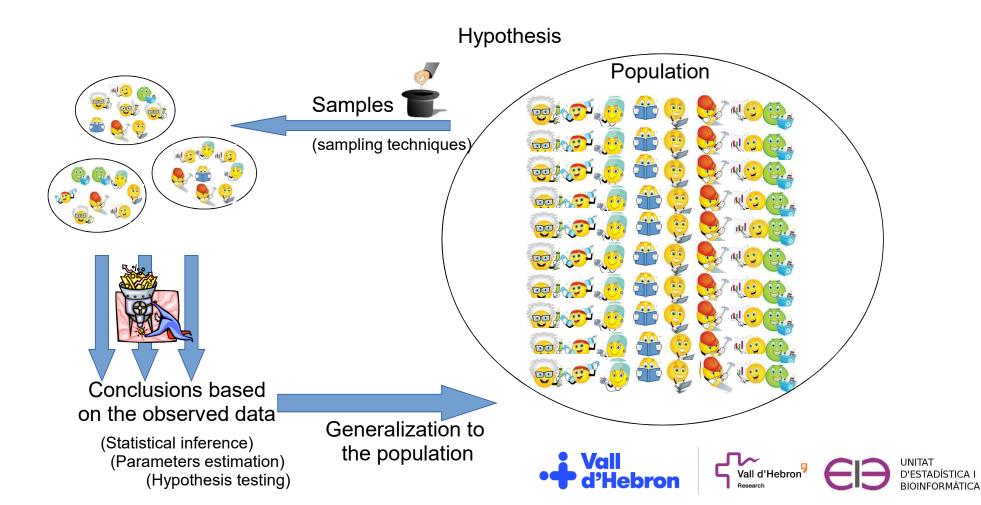








The objective of statistical inference



Billion dollar Question

How many subjects do we need for our Study









Before calculating sample size

- Have a Clear Research Question
- Selected the design of the study (Clinical trial, Observational,...)
- Identify the outcomes or end-points
- Know how variables are going to be measured(proportions, ordinal, continous)
- How many subjects you can afford
- Have a clear idea of the expected results







Too few subjects

- You will not able to answer the question
- Potentially Not ethical



- Waste of resources
- Give a harmful treatment to patients
- Identify irrelevant treatment effects "significative"
- Potentially Not ethical











Before discussing sample size calculations there are several things to keep in mind

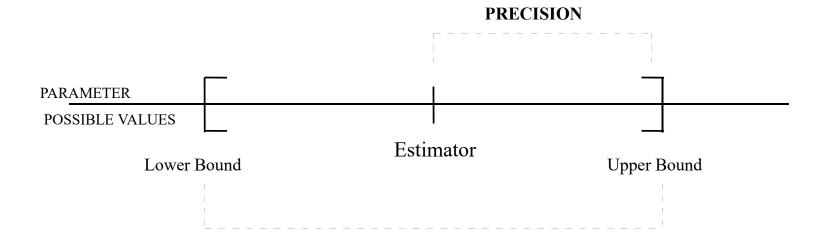
- Type of calculations depend on study goal.
 - Estimation
 - Testing
- Preliminary concepts to be used
 - Standard error of an estimator
 - Confidence interval
 - Hypothesis testing
 - Type I error
 - Power of a test (1-β)







Sample size based on Confidence Intervals



Values in which we are confident that real population parameter is inside With a prefixed confidence level (Usually 95%)

Estimator \pm Coef._{1- $\alpha/2$} x Standard Error







Standard error of the mean

A measure of how variable is the sample mean when computed in all samples of size n

Standard deviation of the distribution of sample mean

standard error =
$$\frac{\sigma}{\sqrt{n}} \cong \frac{s}{\sqrt{n}}$$







Standard error of a proportion

The standard error of a proportion is computed similarly to the SEM.

Instead of the standard deviation it uses the population proportion in the formula If p is not known the estimated proportion is used

standard error =
$$\sqrt{\frac{p \cdot (1-p)}{n}} \cong \sqrt{\frac{\hat{p} \cdot (1-\hat{p})}{n}}$$







Formulas for confidence intervals

Data normally distributed Population variance known (unrealistic assumption)

Population variance unkown, estimated by sample variance

Data: Counts of presence or absence of an event Sample must be "big enough"

 $Z_{1-\alpha/2}$ are quantiles of standard Normal N(0,1) distribution

$$\overline{X}_n - z_{1-\alpha/2} \frac{\sigma}{\sqrt{n}} \le \mu \le \overline{X}_n + z_{1-\alpha/2} \frac{\sigma}{\sqrt{n}}$$

$$\overline{X}_n - t_{1-\alpha/2} \frac{S}{\sqrt{n}} \le \mu \le \overline{X}_n + t_{1-\alpha/2} \frac{S}{\sqrt{n}}$$

$$\hat{p} \pm z_{1-\alpha/2} \sqrt{\frac{\hat{p}\hat{q}}{n}}; \quad n \ge 30, n\hat{p} \ge 5, n\hat{q} \ge 5$$

1- α	0,90	0,95	0,99
Z _{1-α/2}	1,64	1,96	2,58







Sample SIZE for mean

Precision =
$$z_{1-\alpha/2}$$
 * $ee = z_{1-\alpha/2}$ * $\frac{\sigma}{\sqrt{n}}$

$$n = \frac{z_{1-\alpha/2}^2 \sigma^2}{precision^2}$$

- If interval range is 10 (precision =10/2=5)
- Confidence level is 95%
- Standard deviation is 20

$$n = \frac{1.96^2 20^2}{5^2} = 62$$















Sample size for proportion

Precision =
$$z_{1-\alpha/2}$$
 * $ee = 1.96 * \sqrt{\frac{\hat{p}(1-\hat{p})}{n}}$

$$n = \frac{1.96^2 \, \hat{p}(1-\hat{p})}{precision^2} =$$

- Assume precision is 5% (Interval = p±.05)
- Confidence level is 95%
- If it is known that p is around 12.5%

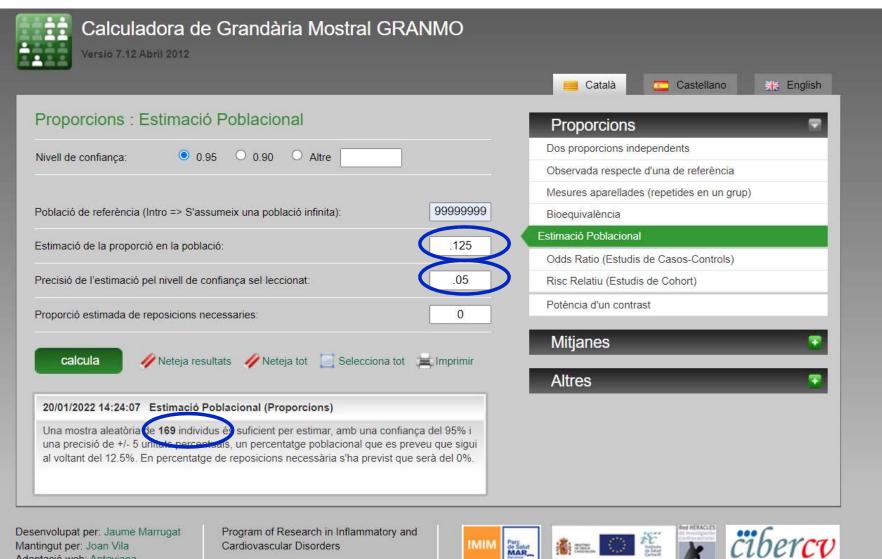
$$n = \frac{1.96^2.125(1 - .125)}{.05^2} = 168$$

$$n = \frac{1.96^2.50(1 - .50)}{.05^2} = 384$$









Adaptació web: Antaviana

Els autors no es fan responsables de les consequencies del seu ús.

Institut Municipal d'Investigació Mèdica, Barcelona, Spain









TAT STADÍSTICA I NFORMÀTICA

Type I Error α

Probability of rejecting the Null Hypothesis when its true (5%)

Power 1- β (Type II Error)

Probability of rejecting the Null Hypothesis when its false (80%,90%)

Variability of the data σ^2

Variance of the data

Effect Size δ

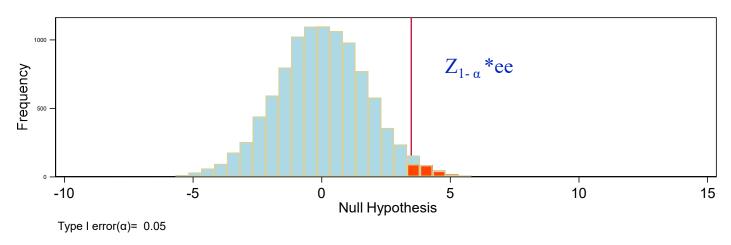
Minimum detectable difference betweeen the two groups to compare

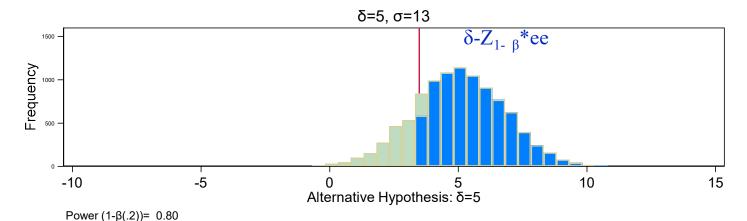






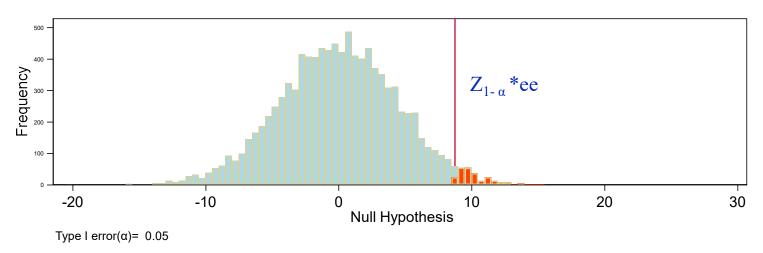
N=107 por grupo

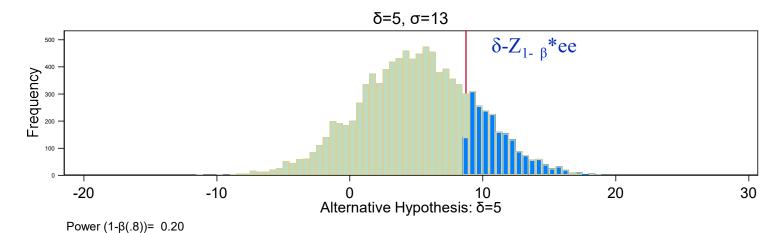






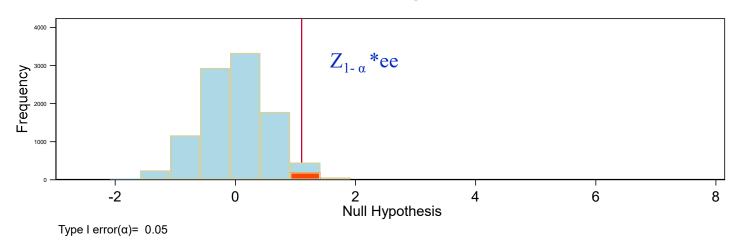
N=17 por grupo

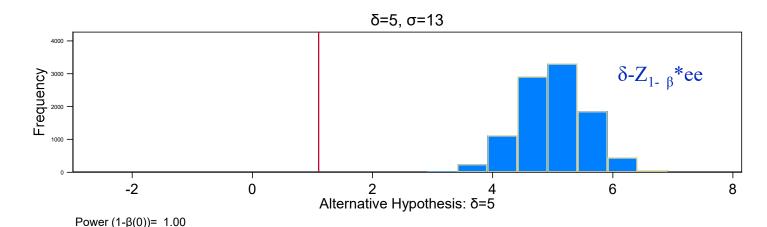






N=1070 por grupo







We want

$$Z_{1-\alpha}$$
*ee= δ - $Z_{1-\beta}$ *ee

We know α,β and $ee=\sigma/\sqrt{n}$

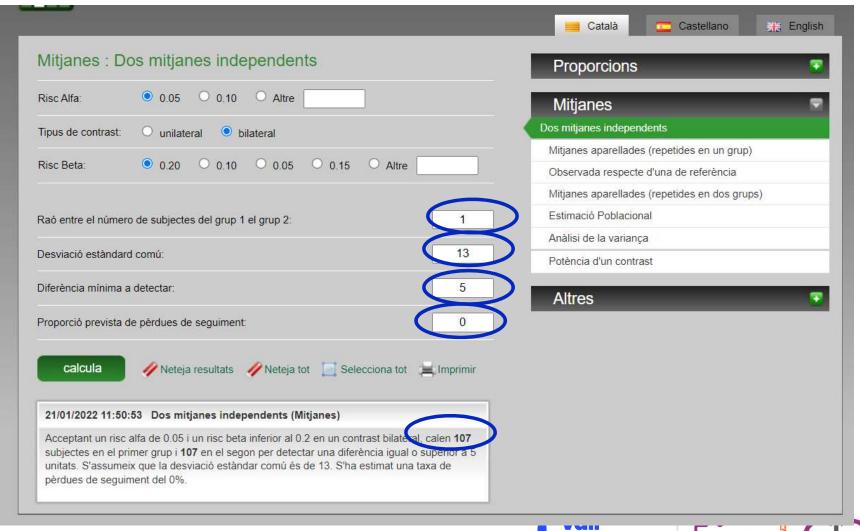
$$n = \frac{2\sigma^{2}(z_{1-\alpha} + z_{1-\beta})^{2}}{\delta^{2}}$$







Sample size for mean differences

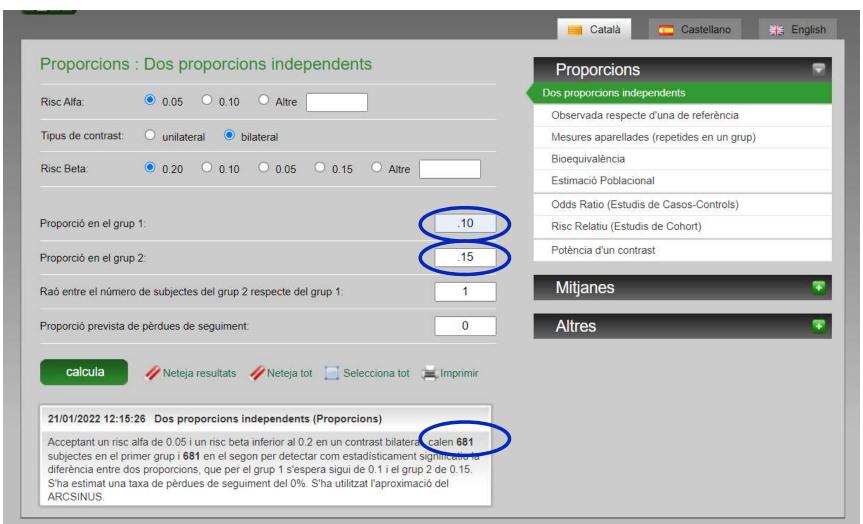








Sample size for differences in proportions



UNITAT D'ESTADÍSTICA I BIOINFORMÀTICA

Some formulas to calculate sample size

Problema	Datos necesarios	Tamaño muestral
Estimación de una proporción	p=Proporción esperada o 0.50 d= precisión(mitad amplitud I.C.) e= error porcentual sobre verdadero parámetro z _{1-α/2} = percentil N(0,1) para significación α	$n = \frac{z_{1-\alpha/2}^{2} p(1-p)}{d^{2}}$ $n = \frac{z_{1-\alpha/2}^{2} (1-p)}{e^{2}}$
Estimación de una media	σ ² =varianza esperada d= precisión(mitad amplitud I.C.) z _{1-α/2} = percentil N(0,1) para significación α	$n = \frac{z_{1-\alpha/2}^2 \sigma^2}{d^2}$
Diferencia de proporciones	p ₁ =proporción grupo 1 p ₂ = proporción grupo 2 D=p ₁ -p ₂ z _{1-α/2} = percentil N(0,1) para significación α z _{1-β} = percentil N(0,1) para poder β p= (p ₁ +p ₂)/2	$n = \frac{\left[z_{1-\alpha/2}\sqrt{2p(1-p)} + z_{1-\beta}\sqrt{p_1(1-p_1) + p_2(1-p_2)}\right]}{D^2}$ $n = \frac{\left[z_{1-\alpha/2} + z_{1-\beta}\right]^2}{2(\arcsin\sqrt{p_2} - \arcsin\sqrt{p_1})^2}$ para enfermedades de ocurrencia rara
Diferencia de medias	σ12= varianza grupo 1 $ σ22= varianza grupo 2 $ $ z1-α/2= percentil N(0,1) para $ significación α $ z1-β= percentil N(0,1) para poder β D= diferencia de medias$	$n = \frac{\left[z_{1-\alpha/2} + z_{1-\beta}\right]^2 (\sigma_1^2 + \sigma_2^2)}{D^2}$ Vali d'Hebron







Some formulas to calculate sample size

Problema	Datos necesarios	Tamaño muestral
	OR= Odds ratio que se estima p ₂ = Proporción de expuestos en los controles	$p_1 = \frac{ORp_2}{ORp_2 + (1 - p_2)}$
Estimación de una OR	p ₁ = Proporción de expuestos en los casos e=amplitud relativa del C.I. z _{1-α/2} = percentil N(0,1) para significación α	$n = z_{1-\alpha/2}^2 \frac{\gamma_{p_1(1-p_1)} + \gamma_{p_2(1-p_2)}}{\ln(1-e)^2}$
	RR= Riesgo relativo que se estima p ₂ = Proporción de casos en los no expuestos	$p_1 = RRp_2$
$ \begin{array}{ccc} \textbf{Estimación de un} & p_1 = \Pr \\ \textbf{RR} & expues \\ e = amp \\ z_{1-\alpha/2} = \\ signifi \end{array} $	p ₁ = Proporción de casos en los expuestos e=amplitud relativa del C.I. z _{1-α/2} = percentil N(0,1) para significación α z ₁₋₈ = percentil N(0,1) para poder β	$n = z_{1-\alpha/2}^2 \frac{\frac{(1-p_1)}{p_1} + \frac{(1-p_2)}{p_2}}{\ln(1-e)^2}$
	OR= Odds ratio que se estima p ₂ = Proporción de expuestos en los controles p ₁ = Proporción de expuestos en los casos	$p_1 = \frac{ORp_2}{ORp_2 + (1 - p_2)}$
Contraste OR>1	$z_{1-\alpha/2}$ = percentil N(0,1) para significación α $z_{1-\beta}$ = percentil N(0,1) para poder β	$n = \frac{\left[z_{1-\alpha/2}\sqrt{2p_2(1-p_2)} + z_{1-\beta}\sqrt{p_1(1-p_1) + p_2(1-p_2)}\right]^2}{(p_1 - p_2)^2}$
	RR= Riesgo relativo que se estima p ₂ = Proporción de casos en los no expuestos	$p_1 = RRp_2$
Contraste RR >1	p ₁ = Proporción de casos en los expuestos z _{1-α/2} = percentil N(0,1) para significación α z _{1-β} = percentil N(0,1) para poder β	$n = \frac{\left[z_{1-\alpha/2}\sqrt{2p_2(1-p_2)} + z_{1-\beta}\sqrt{p_1(1-p_1) + p_2(1-p_2)}\right]^2}{(p_1 - p_2)^2}$





Some tips

- Sample size goes up
 - \circ for smaller α
 - For higher β
 - For smaller δ
 - For higher σ
 - For p closer to 50%
- Sample size is higher for proportions than means
- Sample size must be calculated a priori. Is not sensible to calculate power after
- SD can be calculated from 95% CI
- Upper-Lower limit of a CI is about 4 Standard Error and $SE=s/\sqrt{n}$
- Some % of survivors can be obtained from Kaplan-Meier survival curves and can be used for calculations
- Sample size is not an exact science and must be the product of calculations and reality







Effect size δ

Is the number measuring the strength of the relation between two (o more) group of comparison. It can be calculated from the data.

- Mean difference of a quantitative measure between two populations
- Difference of proportions of a dichotomous measure in two populations
- Correlation coefficient between two quantitative variatons
- Regression coeficient from a multivariable regression model
- Risk difference, Relative Risk, Odds Ratio, Hazard Ratio

It is preferred that the effect size measurement is standardized

$$d=\frac{\bar{x}_1-\bar{x}_2}{s}.$$

$$w = \sqrt{\sum_{i=1}^m rac{(p_{1i} - p_{0i})^2}{p_{0i}}}$$

$$f^2 = \frac{R^2}{1 - R^2}$$

• Eta-squared
$$(\eta^2)$$
 $\eta^2 = \frac{SS_{\text{Treatment}}}{SS_{\text{Total}}}$.

$$\eta^2 = rac{SS_{ ext{Treatment}}}{SS_{ ext{Total}}}.$$

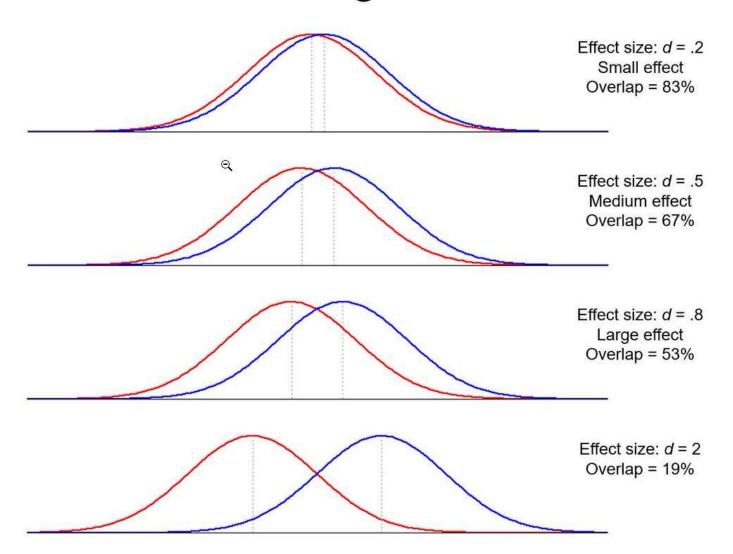
Table I	Values of	Effect Sizes	and Their	Interpretation
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Kind of Effect Size	Small	Medium	Large
r	.10	.30	.50
d	0.20	0.50	0.80
$\eta^2_{\ p}$.01	.06	.14
f^2	.02	.15	.35

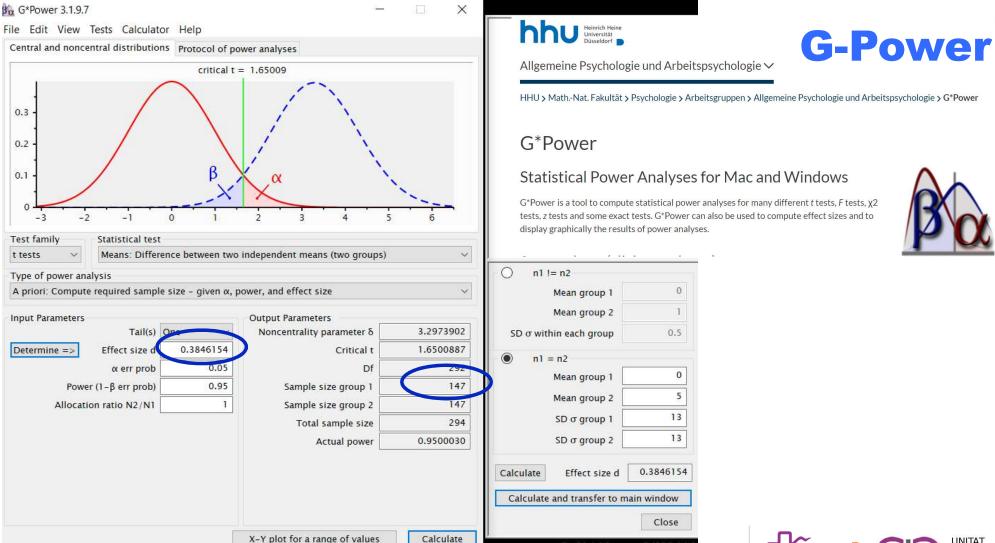


Source: Cohen, J. (1992). A power primer. Psychological Bulletin, 112, 155-159. doi:10.1037/0033-2909.112.1.155

Understanding Effect Sizes

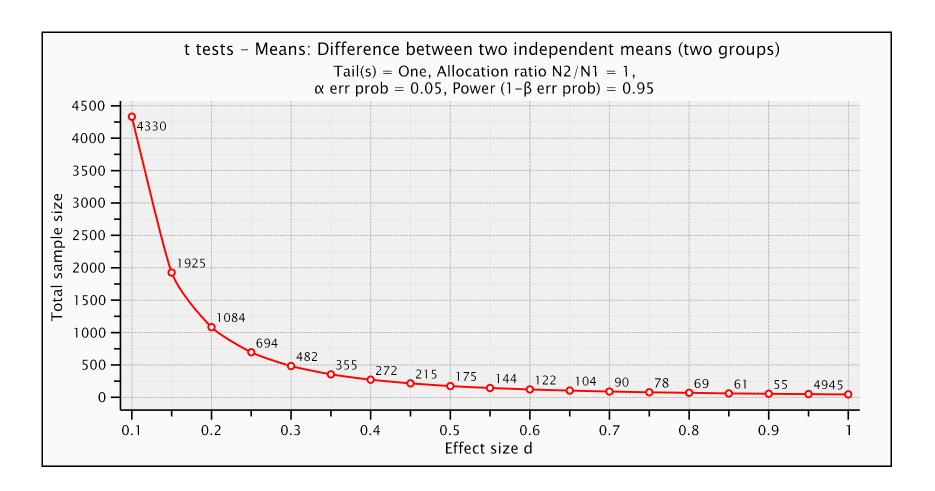








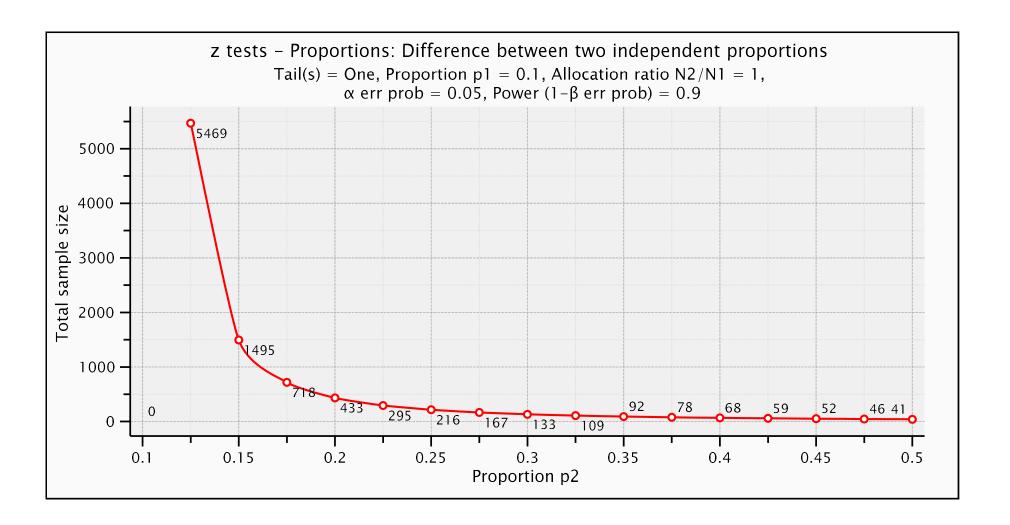










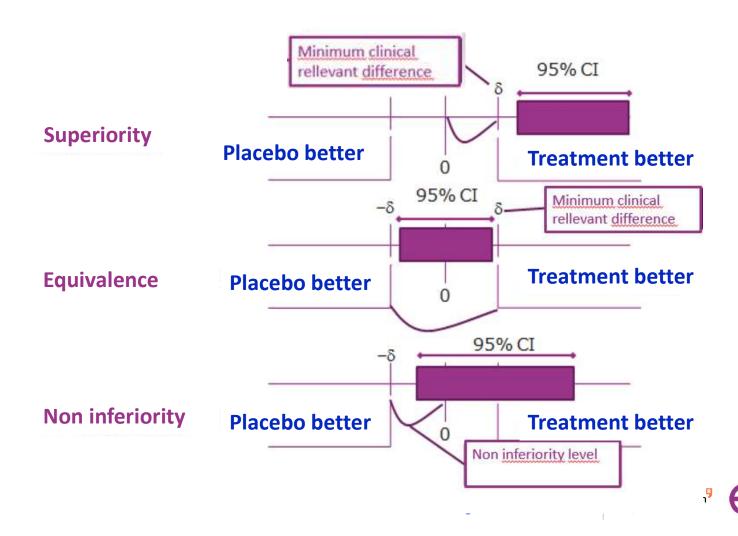








Equivalence and Non inferiority Trials



D'ESTADÍSTICA I

Practical Biostatistics

How to Calculate Sample Size in Randomized Controlled Trial?

Baoliang Zhong, MD

Sample size goes up

Table 1 Hypothesis testing of different design

design	null hypothesis	alternative hypothesis	test statistics
non-inferiority	H0:T-S=-δ	Ha: T-S>-δ	Z=(d+δ)/sd
equivalence	H10:T-S=-δ	H1a: T-S>-δ	$Z1=(d+\delta)/sd$
	H20:T-S=δ	H2a: T-S<δ	$Z2=(\delta -d)/sd$
statistical superiority	H0:T-S=0	Ha: T-S>0	Z=d/sd
clinical superiority	H0:T-S=δ	Ha: T-S>δ	$Z=(d-\delta)/sd$

T: new treatment; S: standard treatment; δ: clinically admissible margin of non-inferiority/equivalence/ superiority; d: the effectiveness difference between T and S; sd: the standard error of d; Z: Z obeys standard normal distribution.









REVIEW Open Access

Through the looking glass: understanding non-inferiority

Jennifer Schumi* and Janet T Wittes

Sample size goes up

Table 2 Approximate sample sizes required for non-inferiority comparison of proportions

True proportion in active control	Non-inferiority bound using 10% margin	Approximate sample size per group assuming 1:1 randomization to new treatment and control required under:		
		Equal effects 5% benefit 10% bene		
0.1	0.09	19,200	8,725	5,050
0.2	0.18	8,500	3,900	2,250
0.3	0.27	4,970	2,260	1,300
0.4	0.36	3,200	1,450	825
0.5	0.45	2,100	1,000	550
0.6	0.54	1,440	640	360
0.7	0.63	930	405	225

Sample sizes calculated using Pass 2008 methods for non-inferiority tests of two independent proportions, using the Z statistic with continuity correction and pooled variance, with a target power of 90% and α level of 0.025.







Interim analysis

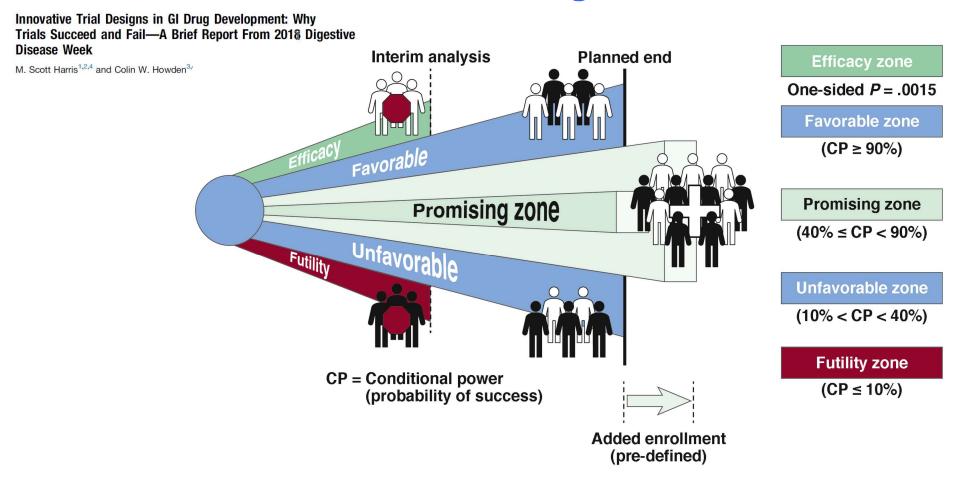
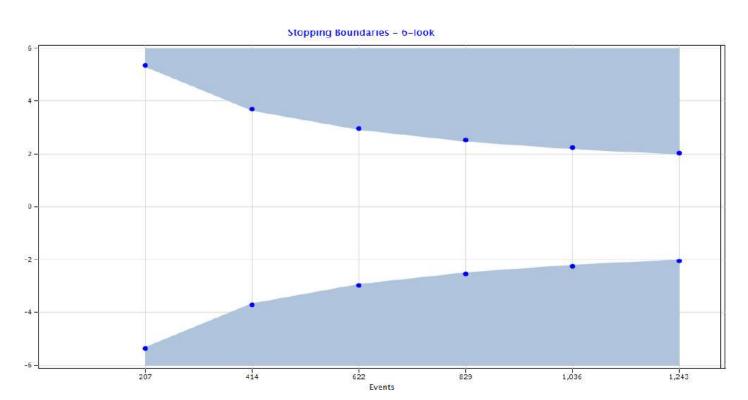


Figure 1In blinded sample size re-estimation, the study commences with an optimistic sample size, assuming a large treatment effect. Patients can be added at IA if the conditional power (CP) falls in the promising zone, the example showing $CP \ge 40\%$ and $\le 90\%$. A favorable ($\ge 90\%$) or unfavorable (10% to 10%) CP would result in no change in enrollment. The trial could also be stopped at IA for success (10%) or futility (CP of 10%). The starting sample size of the blinded sample size reestimate is typically smaller than grouped sequential design because of the flexibility of adding patients at the interim analysis (IA). The study is conducted as a same phase trial, because unlike a phase II/III 2-stage design, no change in dose, treatment, subpopulation, or study phase is implemented at the IA. Modified with permission from Mehta 10%

Six look sequential desing









The present and the future of sample size: **SIMULATION**

- Simulation: Computing intensive approach that mimics the data generating process to evalute properties of a test statistic or confidence interval (Power)
- Model the population parameters of the study (Survival function, Regression models, etc. Lost of follow up, Atrition rate, Changes in treatments, etc.
- Sampling strategy and simulate sampling
- Analytics methods and analysis model :
- Performance of the model (power and precision).
- Calculated Distribution of the effects measurements
- Select sample size







N=10/ por grupo 10 15 Null Hypothesis Type I error(α)= 0.05 δ =5, σ =13 Frequency -10 -5 Alternative Hypothesis: δ=5 Power (1-β(.2))= 0.80 .3 -.2 150 250 350 Potencia calculada Potencia simulada

Example Simulation mean difference

1. Select Parameters

$$\mu_0$$
- μ_0 =0
 μ_1 - μ_0 = δ =5
 σ =13

- 2. Select series of sample size n=10,20...,490,500
- 3. Calculate ee=Standard error= $\sigma \sqrt{\frac{1}{n} + \frac{1}{n}}$
- 4. Sampling 1000 from Normal(δ ,ee)
- 5. Calculate Critical value= $cv=z_{1-\alpha/2}\sigma\sqrt{\frac{1}{n}+\frac{1}{n}}$
- 6. Calculate power as % of observations over cv
- 7. Plot sample size against power







Power Calculations for Survival Analyses Via Monte Carlo Estimation

David B. Richardson, PhD'

- Period follow up 20 years
- Two groups(x) equally distributed(1:1)
 (sampling from binomial p=0.50)
- Generate I cohorts of size N
- Effect size φ were hazard ratio=exp(φ)
- Baseline risk exponential $\exp(\delta)$
- Disease status= $log(h) = \delta + \phi x$
- Calculate proportional hazard
- Power= proportion fof pvalues less than .05 for a big number of repetitions.

Example Simulation Clinical Trial

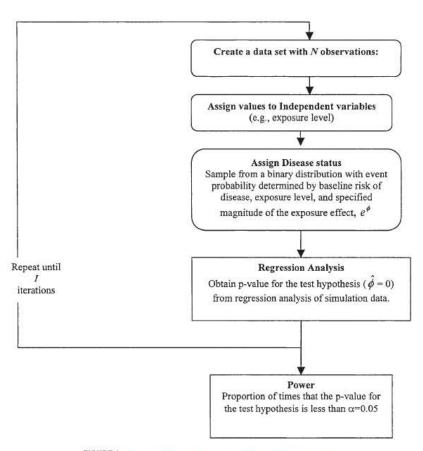


FIGURE 1. Flowchart of the process followed for Monte Carlo estimation of statistical power.







AMERICAN JOURNAL OF INDUSTRIAL MEDICINE 44:532-539 (2003)

Power Calculations for Survival Analyses Via Monte Carlo Estimation

David B. Richardson, PhD*

Parameter	Example 1	
Type Ferror probability (a)	0.05	
Number of iterations (/)	1,000	
Number of persons in study cohort (N)	2,500, 5,000, 10,000	
Exposure effect (e ^d)	1.01, 1.05, 1.10, 1.25	
Length of follow-up (in years)	20	
Exposure (x)	B (1,0.5)	
Age at entry (in years)	n.a.	
Length of employment (in years)	n.a.	
Baseline annual mortality rate (e^{δ_0}) (e^{δ_0} in Example 2—4)	0.0084	

TABLE II. Estimates of Statistical Power by Total Study Size and Exposure Effect Size for a Cohort Study Comparing Survival in Exposed Versus Unexposed Groups of Equal Size (Example 1)

Study size	Relative risk (hazard ratio)			
	1.01	1.05	1.10	1.25
2,500	0.04ª	0.07ª	0.15 ^a	0.62 ^a
	0.03 ^b	0.07 ^b	0.16 ^b	0.63 ^b
5,000	0.05 ^a	0.10 ^a	0.25 ^a	0.91°
	0.03 ^b	0.10 ^b	0.27 ^b	0.90 ^b
10,000	0.05 ^a	0.16ª	0.50 ^a	1.00°
	0.04 ^b	0.16 ^b	0.48 ^b	1.00 ^b

[&]quot;Statistical power estimated by Monte Carlo approach under conditions specified in Table I using the program in Appendix 1. (Values shown include estimates of power derived via Monte Carlo simulation.)







^bStatistical power calculated using the SSP internet based computer program for power and sample size calculations under conditions specified in Table I. (Values shown include estimates of power derived via the method described by Lachin and Foulkes.)

Summary

- Important step in designing a study
- Although for simple situations is easy to calcultate sample size, when having several outcomes and covariates performing a sample size can be complicated
- Because sample size errors and costs related to performing a RCT, is recomended to take caution and ask for statistical advice
- Simulation is the actual option to perform sample size before starting a RCT







BMC Research Notes

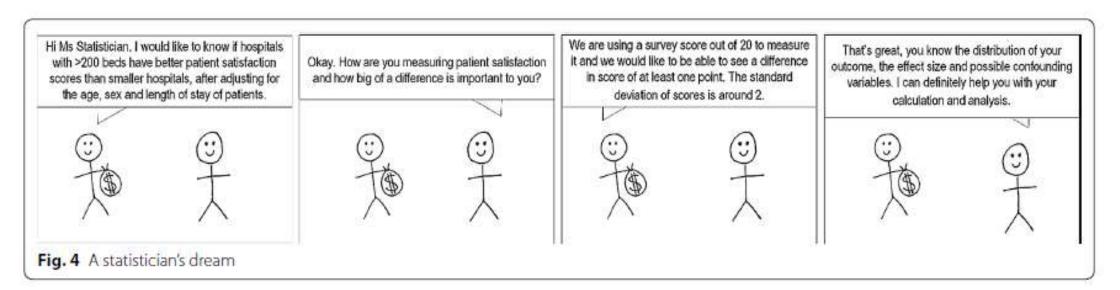
TECHNICAL NOTE

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When is enough, enough?
Understanding and solving your sample size problems in health services research

Victoria Pye*, Natalie Taylor, Robyn Clay-Williams and Jeffrey Braithwaite

Next time you want to calculate a sample size do not forget to know the Distribution of the Outcome, the Effect Size, and the possible Confounding variables









Gràcies \ Gracias \ Thank you



