

L17

# Basic Statistics for Economists

Spring 2020

Department of Statistics

# Today

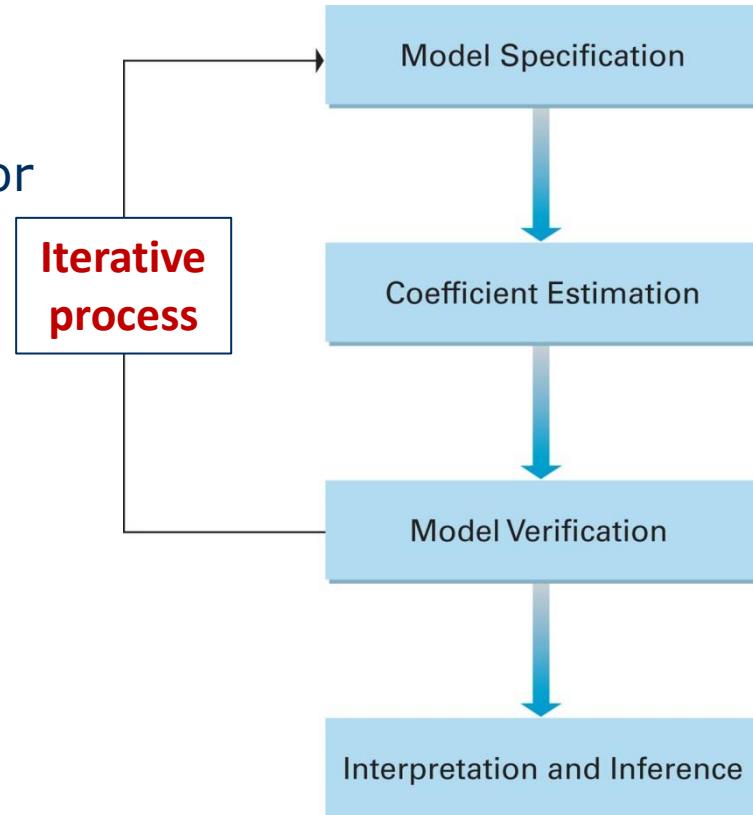
Selected topics from **12.8-9 + 13.1 + 13.4-6 NCT + review**

- Modelling
- Interpretation of *t*-test in multiple regression
  - correct statement of hypotheses
- Stepwise regression
- Dummy variables
- Specification bias, multicollinearity, heteroscedasticity
- Residual analysis

**Maybe in a different order than that of NCT**

# Analysis and the model process

- This is true any study of relationships, developing explanatory models, models for prognosis, and so on
  - Regression analysis
  - Time series analysis
  - longitudinal models (panel data)
  - ARCH/GARCH-models
  - VAR-models
  - ...



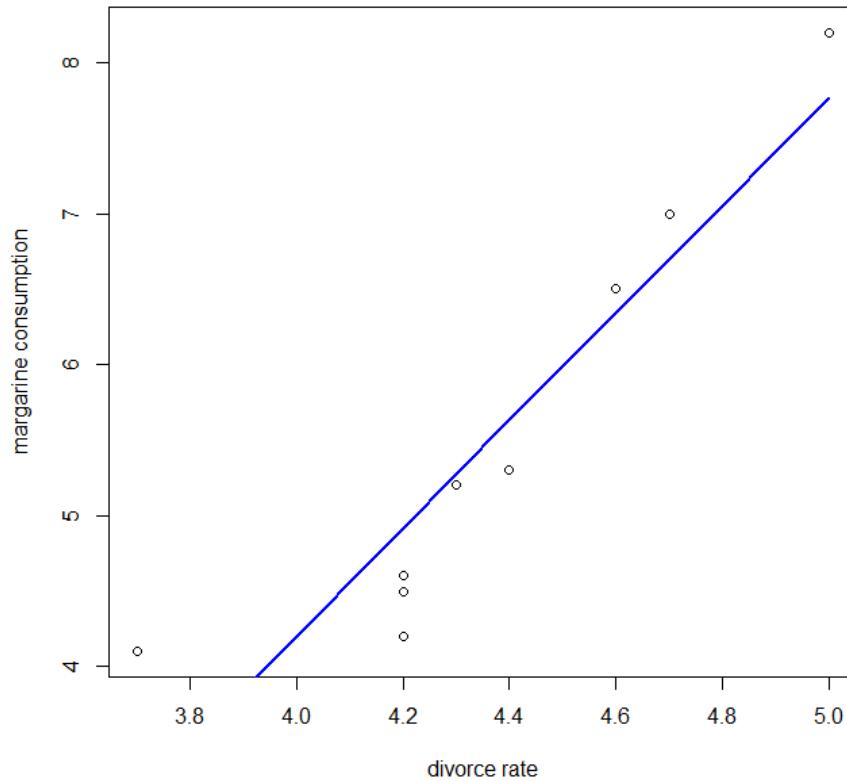
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# Model building - method

- **Model specification**
  - Choice of variables **grounded in area specific theory**
  - linear – non-linear (we have not talk about this, read NCT 12.7)
- **Estimation of the model (we have covered this)**
  - **Data quality** should be always be considered and assessed
- **Model verification**
  - Does it match the theory: e.g. do the estimates have correct sign?
  - Are the assumptions correct – **residual analysis**
  - A prognosis should be compared to outcomes over time
- **Interpretation, inference, prediction (usage)**
  - Confidence interval, hypothesis test, prognosis

# Example: what not to do

Divorce rate in Maine and per capita margarine consumption



source of data: <http://www.tylervigen.com/spurious-correlations>

# ***t-test for one of the coefficients, $\beta_j$***

- Hypotheses to test if a coefficient is significant (adds something to the model) ***given that all the other variables are in the model***
- Suppose that  $Y = \beta_0^\circ + \beta_1^\circ X_1 + \varepsilon$  is significant, a “good” model
- Test the extended model with an additional variable  $X_2$
- Re-estimate the model with both  $X_1$  and  $X_2$  and test:

$$H_0: \beta_2 = 0 \mid \beta_1 \neq 0 \quad \text{vs.} \quad H_1: \beta_2 \neq 0 \mid \beta_1 \neq 0$$

- This is a clearer way of stating what is really being tested: does including  $X_2$  significantly improve the model compared to just  $X_1$ 
  - remember:  $R^2$  and often also  $R_{adj}^2$  increase with every extra predictor



# Idea: Stepwise regression

- Start with the variable that has the ***strongest correlation*** with  $Y$ , lets say  $X_1$ 
  - estimate  $\beta_0$  and  $\beta_1$  and test  $H_0: \beta_1 = 0$
- Given that  $X_1$  is in the model, choose the one that ***adds the most***, e.g.  $X_2$ 
  - **NOTE!** This is not necessarily the second most one that correlates  $Y$  – it will depend on the how the variable correlates with  $X_1$
  - estimate  $\beta_0$ ,  $\beta_1$  and  $\beta_2$  and test  $H_0: \beta_2 = 0 \mid \beta_1 \neq 0$
- Continue with more variables  $X_3$ ,  $X_4$ , ..., choose the one that ***adds the most***
  - Estimate  $\beta_0$ ,  $\beta_1 - \beta_{k+1}$  and test  $H_0: \beta_{k+1} = 0 \mid \beta_1 \neq 0, \dots, \beta_k \neq 0$
- Stop when no available variable gives a significant improvement

# Stepwise regression

Important to check at every step:

- That the  $t$ -test is significant
- Changes in coefficients of the  $X$  that already are in the model
  - Large changes of the values?
  - Still significant? It happens that something that was significant now is non-significant!
  - Should you exclude a previously included variable?
- Study changes in  $R^2$ ,  $R_{adj}^2$ , and  $MSE = s_e^2$
- Check if the model can be **grounded in area specific theory**
  - You can always find the “best” model, but is it meaningful?



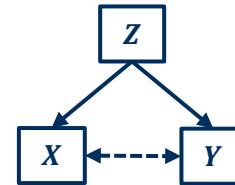
# Model building

- **Stepwise regression** is often built into statistical software packages. These functions carry out the steps **automatically**
- Different strategies:
  - **Bottom-Up**: start with an “empty” model and add  $X$  one by one.
  - **Top-Down**: start with a “full” model and reduce it by removing the worst  $X$  at every step
  - **Add-Delete strategi**: at every step, the program will do what is “best,” by adding or removing variables
- You have to be careful!

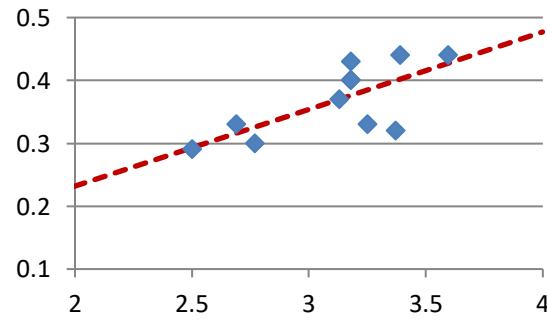
# Spurious (fake) relationships

When two variables covary strongly, but there is no explanation

- A third variable influences ... or several...
- E.g. observations are sampled over **time**; both show a **trend**

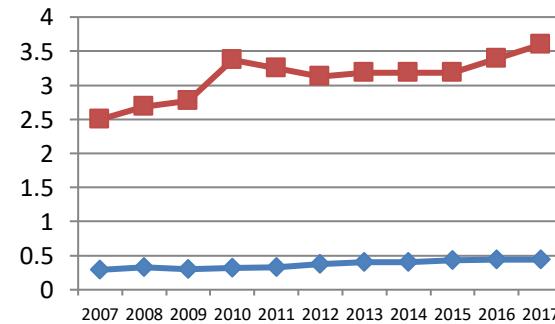


$Y = \text{Illnesses among public employees}$ ,  $X = \text{Price of household electricity, kr}$



Source: SCB, [www.scb.se](http://www.scb.se)

$$\begin{aligned} r_{xy} &= 0,712 \\ R^2 &= 0,508 \\ p &= 0,014 \end{aligned}$$



Hard to see at this scale, but similar trends

# Dummy variables

NCT 12.8

- So far,  $X$  has mostly been **continuous** variables, but  $X$  can be **discrete**
- No problem if  $X$  is **categorical**...
- Special case:  $X$  is **binary (dichotomous)** with 0 and 1 as possible values
- If you for example want to compare group A to group B

$$X = \begin{cases} 0 & \text{if group A} \\ 1 & \text{if group B} \end{cases}$$

- $X$  is called **dummy variable** or **indicator variable**

# Dummy variable, cont.

- Model with one numerical predictor  $X_1$  and one dummy  $X_2$ :

$$\mu_{Y|X} = \beta_0 + \beta_1 X_1 + \beta_2 X_2$$

Case  $X_2 = 0$ :  $\mu_{Y|X} = \beta_0 + \beta_1 X_1 + \beta_2 \cdot 0 = \beta_0 + \beta_1 X_1$

Case  $X_2 = 1$ :  $\mu_{Y|X} = \beta_0 + \beta_1 X_1 + \beta_2 \cdot 1 = (\beta_0 + \beta_2) + \beta_1 X_1$

- Same slope  $\beta_1$  but **different intercept** for the two groups
- Test for different intercept:  $H_0: \beta_2 = 0 \mid \beta_1 \neq 0$

# Dummy variable, cont.

- Extend the model with a "new" variable =  $X_1 \cdot X_2$

$$\mu_{Y|X} = \beta_0 + \beta_1 X_1 + \beta_2 X_2 + \beta_3 (X_1 \cdot X_2)$$

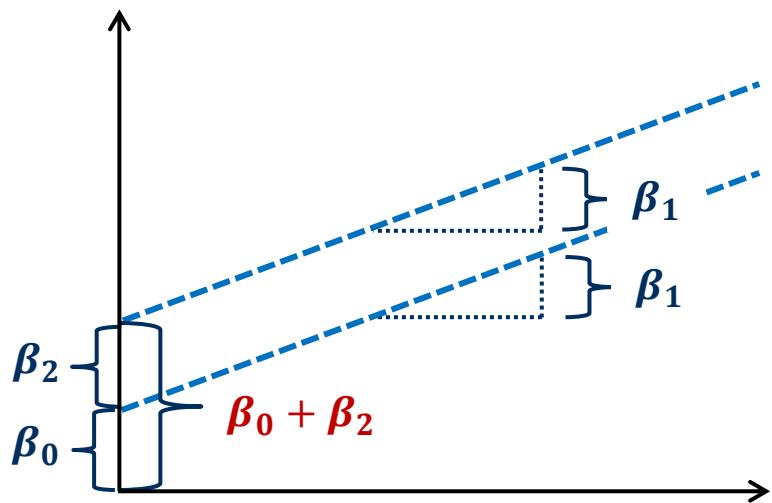
Case  $X_2 = 0$ :  $\mu_{Y|X} = \beta_0 + \beta_1 X_1$

Case  $X_2 = 1$ :  $\mu_{Y|X} = (\beta_0 + \beta_2) + (\beta_1 + \beta_3)X_1$

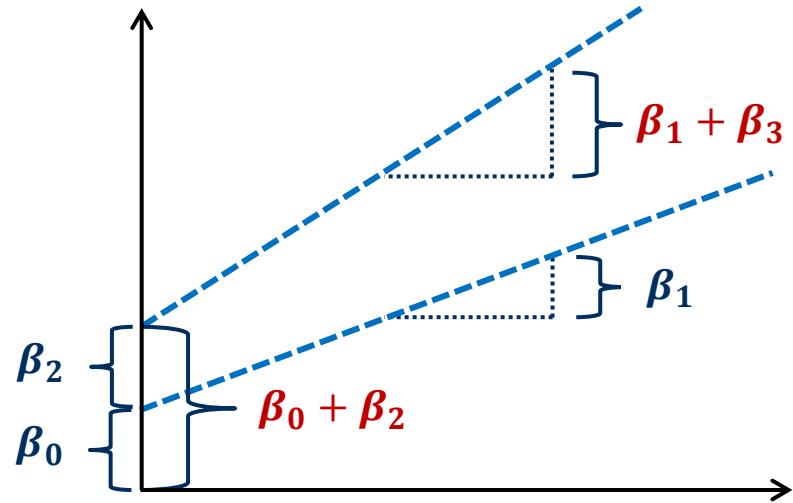
- Results in **different intercept and slope** for the two groups
- Test for different slope:  $H_0: \beta_3 = 0 \mid \beta_1 \neq 0, \beta_2 \neq 0$

# Dummy variable, graphically

- Model with different intercepts, same slope



- Model with different intercepts and slope



# Short discussion of potential problems

NCT 13.4-6

- **Specification bias**
  - Something important is missing from the model
- **Multicollinearity**
  - Predictor variables correlates
- **heteroscedasticity**
  - not constant variance



# Specification bias

## Section 13.4

Occurs if important (significant) predictor variables are left out of the model.

- This will make the **OLS estimates** of the coefficients of the included variables **biased**. Shown on 572, but not part of the course!
- The **statistical inference** (hypothesis test and CI) can in this case be **misleading**.
- The effect of the missing variables are instead captured by the residual variance ( $s_e^2$ ) which typically becomes much greater, i.e. **worse precision**.



# Multicollinearity

## Section 13.5

- When two or more of the explanatory variables are strongly correlated with each other  $\Rightarrow$  serious problems for the entire model!
- Illustration (constructed example):

### Variables

- $Y$ : Sales (mkr)
- $X_1$ : Print advertising (100 tkr)
- $X_2$ : Other advertising (100 tkr)
- $n = 7$

$Y$	$X_1$	$X_2$	$X_3$
7	4	1	5
9	7	2	9
16	9	5	14
19	12	8	20
25	15	10	25
26	17	14	31
33	20	17	37

# Multicollinearity, cont.

Regression Analysis: Y versus X1

The regression equation is

$$Y = -0,31 + 1,63 X1$$

Predictor	Coef	SE Coef	T	P
Constant	-0,306	1,388	-0,22	0,834
X1	1,6327	0,1058	15,42	0,000

$$S = 1,48186 \quad R-Sq = 97,9\% \quad R-Sq(adj) = 97,5\%$$

$$r_{x_1,y} = 0,990$$

**NOTE:**  $R^2$  close to 1 and significant coefficients ( $p = 0$ ) for both models

Regression Analysis: Y versus X2

The regression equation is

$$Y = 6,68 + 1,55 X2$$

Predictor	Coef	SE Coef	T	P
Constant	6,676	1,282	5,21	0,003
X2	1,5485	0,1302	11,89	0,000

$$S = 1,90830 \quad R-Sq = 96,6\% \quad R-Sq(adj) = 95,9\%$$

$$r_{x_2,y} = 0,983$$

# Multicollinearity, cont.

Regression Analysis: Y versus X1; X2

The regression equation is

$$Y = 0,89 + 1,34 X_1 + 0,279 X_2$$

Comments: see next page

Predictor	Coef	SE Coef	T	P
Constant	0,890	3,594	0,25	0,817
X1	1,3436	0,7948	1,69	0,166
X2	0,2791	0,7592	0,37	0,732

$$S = 1,62948$$

$$R-Sq = 98,0\%$$

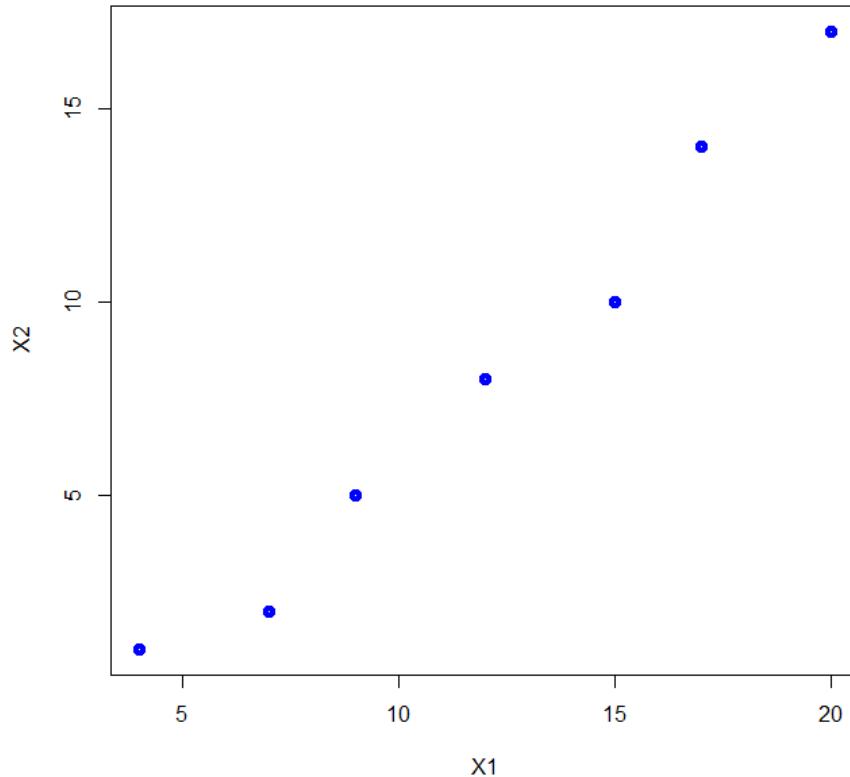
$$R-Sq(adj) = 97,0\%$$

Analysis of Variance

Source	DF	SS	MS	F	P
Regression	2	522,81	261,40	98,45	0,000
Residual Error	4	10,62	2,66		
Total	6	533,43			



# Plot: Print ads vs. Other ads



# Multicollinearity, cont.

- Large  $R^2$  but marginally greater than the previous models
  - but  $R^2_{adj}$  has decreased compared to model with just  $X_1$
- Test  $H_0: \beta_1 = \beta_2 = 0$  gives significant results ( $F$ -test,  $p = 0$ )
  - But neither  $b_1$  or  $b_2$  are significantly different from zero with  $p = 0,166$  and  $0,732$  respectively
- Large change in the estimated  $X_2$  coefficient, slightly smaller change of the  $X_1$  coefficient
- Standard errors  $s_{b_1}$  and  $s_{b_2}$  are approximately 7,5 and 6 times greater compared to the model with only  $X_1$ , the residual standard error  $s_e$  has also increased
  - We have also lost a degree of freedom



# Multicollinearity, cont.

- $X_1$  and  $X_2$  covary so strongly that we cannot discern their respective effects on  $Y$ 
  - Their correlation is  $r_{x_1,x_2} = 0,989$
- Do try a model with all three explanatory variables  $X_1$ ,  $X_2$ , and  $X_3$  and see what happens!
  - The problems will be even more serious. Do you see why?

What can we do?

- Specify a different model!



# Absolute multicollinearity

- Anta  $Y$  = shipping cost,  $X_1$  = distance in kilometers,  $X_2$  = distance in Swedish mil (10 km = 1 mil).
- Model including both:

$$\begin{aligned}\mu_{Y|X} &= \beta_0 + \beta_1 X_1 + \beta_2 X_2 = \beta_0 + \beta_1 X_1 + \beta_2 \cdot 10X_1 \\ &= \beta_0 + (\beta_1 + 10\beta_2)X_1 = \beta_0 + \alpha_1 X_1\end{aligned}$$

- This is really two variables,  $X_1$  and  $Y$ , but we are trying to estimate this using 3 coefficients
- Mathematically, this will not work!

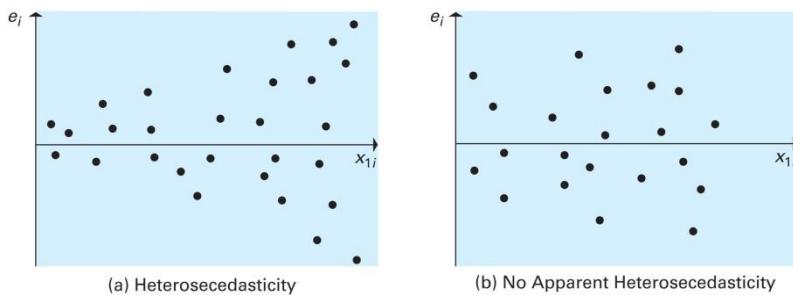
The relationships between the independent variables do not need to be obvious. Strong correlations can also occur in more complex situations. Be careful!



# heteroscedasticity

## Section 13.6

- Model assumptions: the error terms are normally distributed with expected value = 0 and **constant variance**  $\sigma_\varepsilon^2$
- This is often not a good assumption



Plot the residuals ( $e_i$ ) against each of the explanatory variables ( $x_{ki}$ )

- E.g. Larger companies are subject to more factors which may affect variation of  $Y$ , compared to smaller companies.

# heteroscedasticity, cont.

Problem:

- OLS estimates are **not efficient** (not best precision)
- Hypothesis test and confidence interval for the coefficients are **not correctly defined** – e.g. you can not be sure that your CI or  $p$ -values are even approximately correct.
- **NCT** describes a test for homoscedasticity – **skip it!**
- Do plot the residuals  $e_i$  against the explanatory variables  $x_{ki}$  and also against the predicted values  $\hat{y}_i$ !



# Verification of the model: Residual analysis

NCT 12.9 p. 534-537

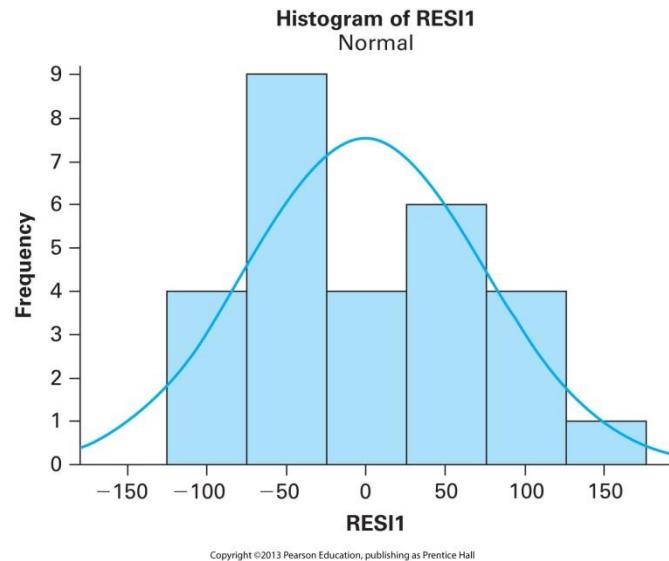
- The model assumption must be verified!
  1.  $Y$  is a **linjeär** function of  $X_1, \dots, X_K$  + error term  $\varepsilon_i$
  2.  $X_1, \dots, X_K$  are all **independent** of the error terms  $\varepsilon_i$ .
  3. The error terms are  $\varepsilon_i$  **normally distributed** r.v. with expected value  $E(\varepsilon_i) = 0$  and **constant variance**  $\sigma_\varepsilon^2$
  4. The error terms are pairwise **independent** of each other
- The residuals  $e_i$  are proxies of  $\varepsilon_i$   $\Rightarrow$  **analyze the residuals**
  - We have just talked about the assumption of constant variance and how we can check this



# Analysis of residuals, cont.

**Normally distributed** error terms  $\varepsilon_i$

- There are test, e.g. Jarque-Bera test. (not covered)
- **Histogram** of  $e_i$  or "Probability plots" (not covered)



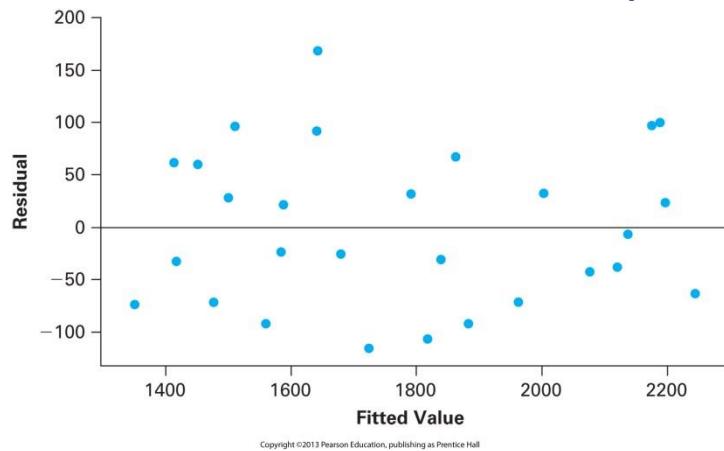
Does not have to be very normal like. If it looks "almost" normal, then that is ok, too.

The other assumptions are more critical: independence, constant variance, linearity

# Analysis of residuals, cont.

$X_1, \dots, X_K$  **independent** of error terms  $\varepsilon_i$

- Plot residuals  $e_i$  against each of  $X_1, \dots, X_K$
- Also plot against the predicted values  $\hat{y}$  (= linj.komb. av  $X_1, \dots, X_K$ )
- Should be void of clear patterns and should look random



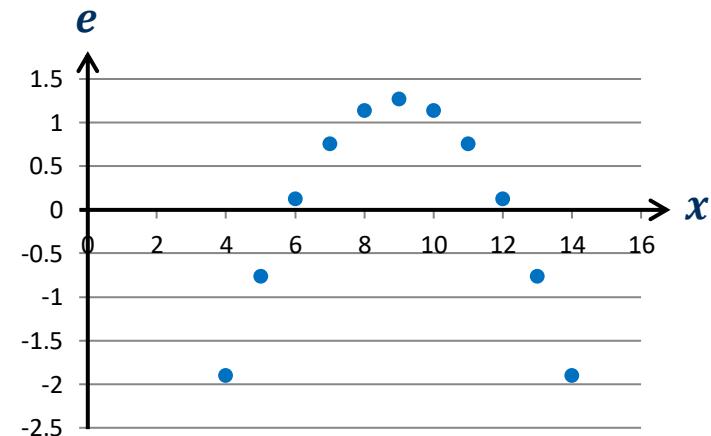
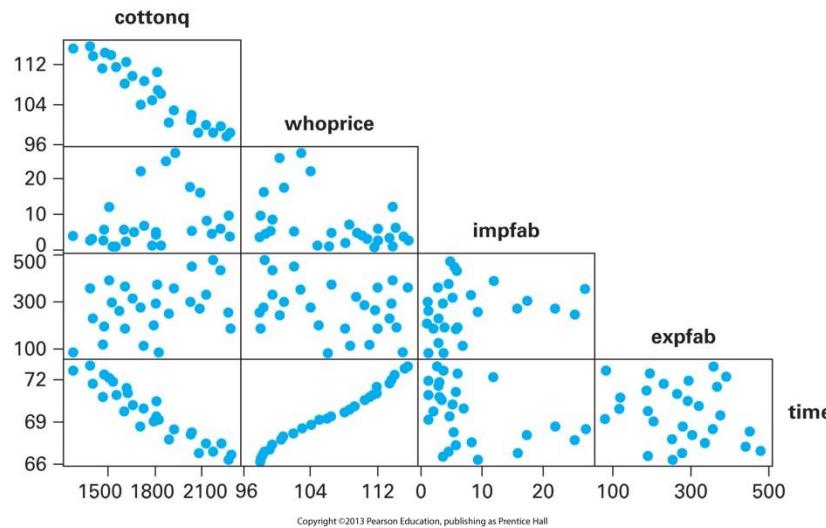
A "clear pattern" could be any kind of pattern. To understand what causes the pattern requires more thinking and discussion. But dependence, non-linearity causes typical patterns (see next page).



# Analysis of residuals, cont.

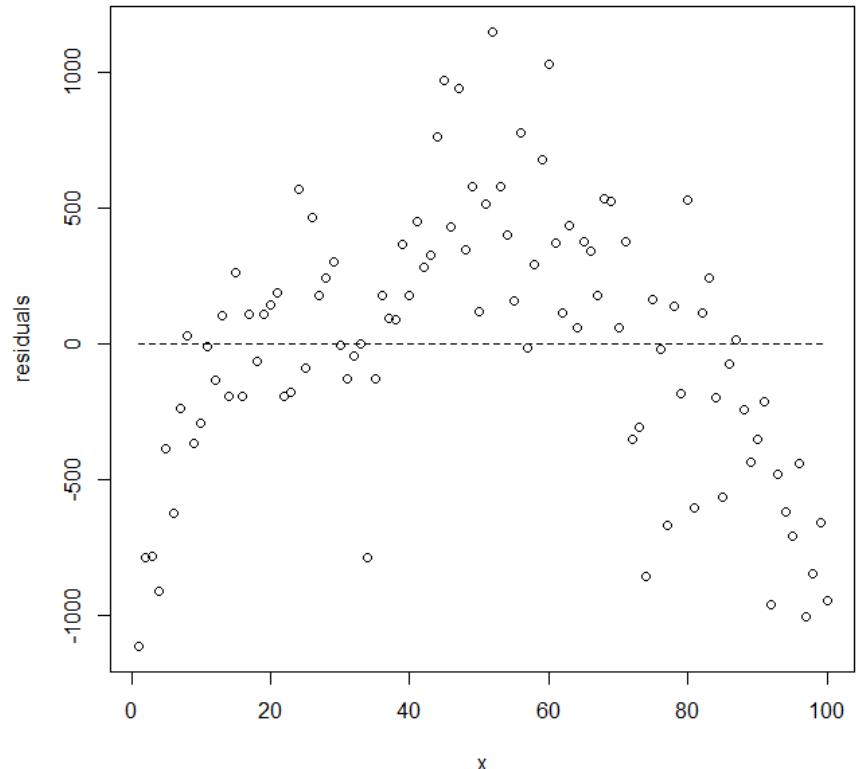
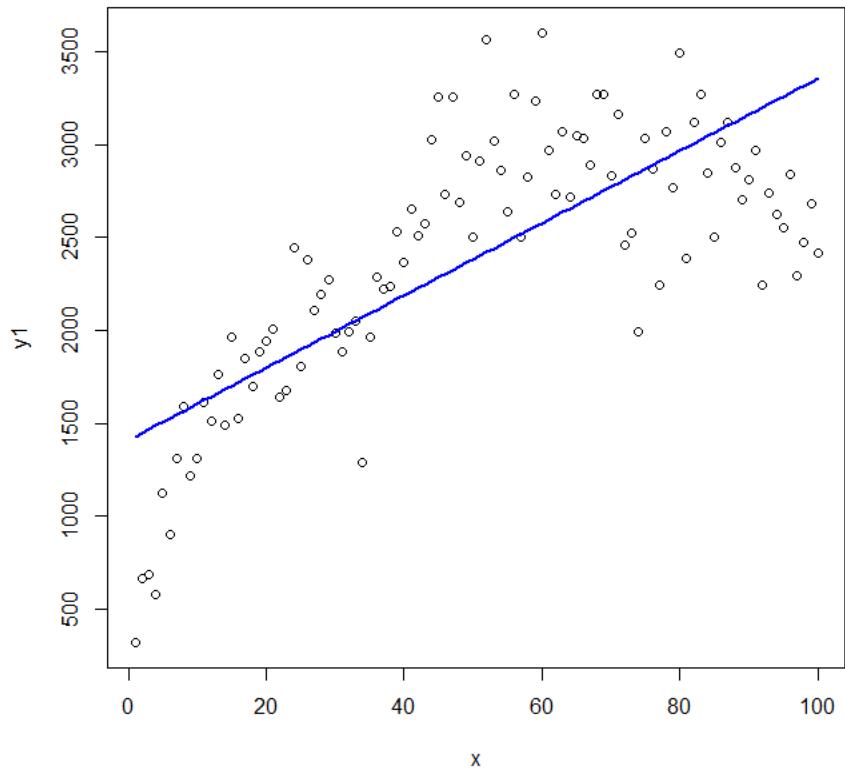
## Linear correlation

- Plot  $Y$  against each of  $X_1, \dots, X_K$ ; matrix plots
- Also plots with  $e_i$  against each of  $X_1, \dots, X_K$  and  $\hat{y}$  can indicate non-linearity

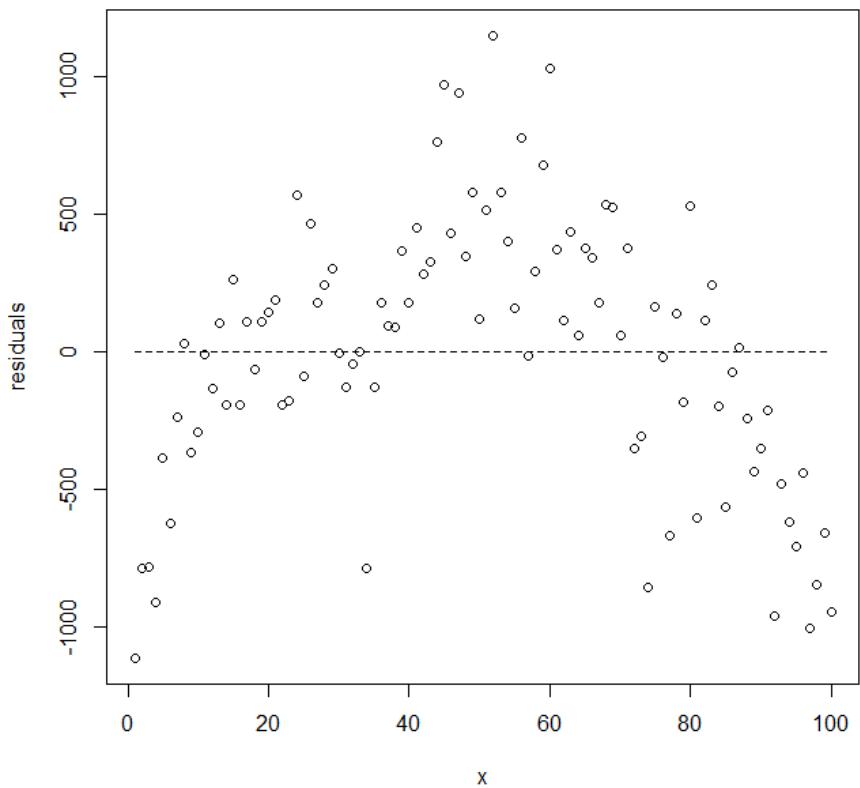
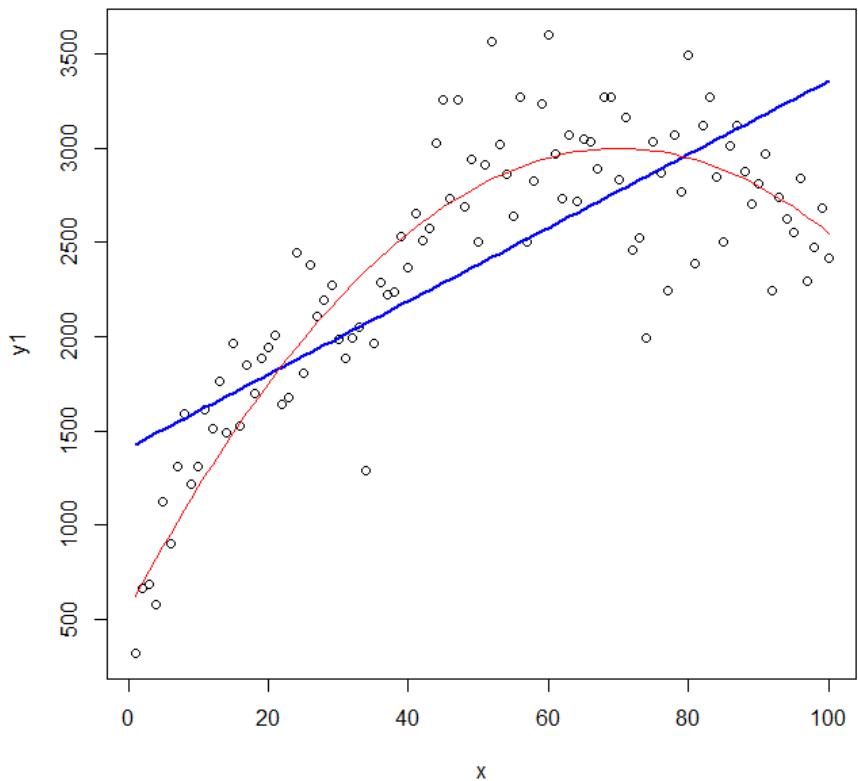


Clear pattern because of non-linearity!

# Example: non-linear relationship



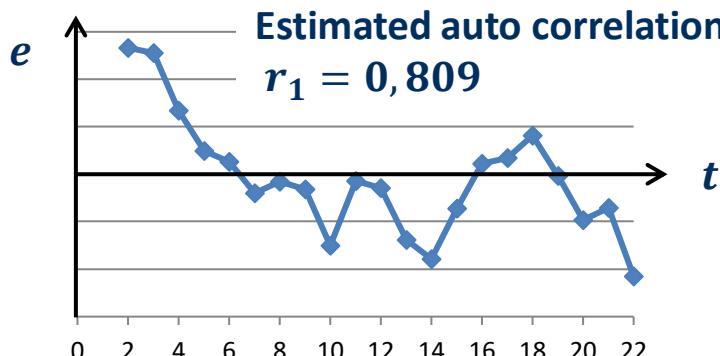
# Example: non-linear relationship, cont.



# Residual analysis, cont.

The error terms  $\varepsilon_i$  are **dependent** of each other

- Common vid repeated measures over time
  - The value of today influences the value of tomorrow
- Let  $\varepsilon_t$  denote the error term at time  $t$
- **Auto correlation**, serial correlation:  $\rho_1 = \text{Corr}(\varepsilon_t, \varepsilon_{t-1})$



Note! Described in section 12.9, but not part of reading instructions!

Can be difficult to notice if the error terms are dependent in some other way.

Area specific knowledge required!

# Next time

- Another kind of test:  **$\chi^2$ -test**
- Used for **categorical data**
  - Goodness-of-fit test
  - Independence test/ Homogeneity test

