# Ultrametricity indices for the Euclidean and Boolean hypercubes

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### Introduction

#### Murtagh observed experimentally:

▶ Samples which are *sparse* and *random* in  $[0,1]^N$  or  $\mathbb{F}_2^N$  become more and more ultrametric as  $N \to \infty$ 

#### His ultrametricity coefficient:

fraction of triangles which are approximately isosceles with short base (= ultrametric)

### 1. Ultrametricity indices

Let (X, d) be a finite metric space.

### Murtagh:

$$m(X,d) := \frac{\# \text{ ultrametric } \Delta}{\# \text{ all } \Delta}$$

### topological:

$$lacksquare t(X,d) := rac{1}{\operatorname{\mathsf{diam}}(X)} \int\limits_0^{\operatorname{\mathsf{diam}}(X)} \mu(\Gamma_\epsilon) \, d\epsilon$$

Let  $\epsilon > 0$ . Vietoris-Rips graph  $\Gamma_{\epsilon}$  for (X, d):

- ► Vertices: X
- ▶ Edge: (x, y) with  $d(x, y) \le \epsilon$ .

#### Lemma

(X,d) is ultrametric  $\Leftrightarrow$  all connected components of all  $\Gamma_{\epsilon}$  are complete.

Let  $\Gamma$  be a finite graph.

- ▶  $b_0(\Gamma) := \#$ connected components of Γ
- $c(\Gamma) := \# \text{maximal cliques of } \Gamma$
- $\mu(\Gamma) := \frac{b_0(\Gamma)}{c(\Gamma)}$

$$t(X,d) = \frac{1}{\operatorname{diam}(X)} \int_{0}^{\operatorname{diam}(X)} \mu(\Gamma_{\epsilon}) d\epsilon$$
topological ultrametricity index

#### The subdominant ultrametric

▶  $\bar{d}(x,y) = \min \{ \epsilon \mid x,y \in \text{same connected component of } \Gamma_{\epsilon} \}$ 

#### Lemma

d is the subdominant ultrametric associated with d.

Let  $0 < d_0 \le \cdots \le d_n$  be the pairwise positive distances between the points of X.

#### Lemma

$$t(X,d) = \sum_{i=0}^{n-1} \alpha_i \frac{d_i}{d_n} \quad \text{with } \sum_{i=0}^{n-1} \alpha_i = 1,$$
  

$$\alpha_0 \in [0,1], \ \alpha_1, \dots, \alpha_{n-2} \in (-1,1], \ \alpha_{n-1} \in (0,1].$$

#### Proof.

- $\mu(\Gamma_{\epsilon}) = \mu_{i+1}$  is constant for  $d_i < \epsilon \le d_{i+1}$
- $\mu_0 = \mu(\Gamma_{\epsilon}) = 1$  for  $0 < \epsilon \le d_0$

$$t(X,d) = rac{1}{d_n} \left( \mu_0 d_0 + \sum_{i=1}^{n-1} \mu_i (d_i - d_{i-1}) \right)$$

$$= rac{1}{d_n} \left( \sum_{i=0}^{n-2} (\mu_i - \mu_{i+1}) d_i + \mu_{n-1} d_{n-1} \right)$$

with

$$\mu_{n-1} + \sum_{i=0}^{n-2} (\mu_i - \mu_{i+1}) = \mu_0 = 1$$



### Corollary

$$t(X, d)$$
 is scale invariant:  $t(X, d) = t(X, \sigma \cdot d)$  for  $\sigma > 0$ .

Proof.

$$t(X,d) = \sum_{i=0}^{n-1} \alpha_i \frac{d_i}{d_n} = \sum_{i=0}^{n-1} \alpha_i \frac{\sigma \cdot d_i}{\sigma \cdot d_n} = t(X, \sigma \cdot d)$$



Consider  $N \to \infty$ .

- $ightharpoonup (\mathbb{H}^N, d) := ([0, 1]^N, d_E) \text{ or } (\mathbb{F}_2^N, d_H)$
- $ightharpoonup d_E =$ Euclidean distance,  $d_H =$ Hamming distance
- ▶ X = finite random sample of  $\mathbb{H}^N$  of fixed cardinality

#### Observation.

If 
$$\frac{d_0}{d_n} \stackrel{\mathcal{P}}{\longrightarrow} 1$$
, then  $m(X,d), t(X,d) \stackrel{\mathcal{P}}{\longrightarrow} 1$ 

#### **Theorem**

Let X be uniformly distributed. Then  $\frac{d_0}{d_n} \stackrel{\mathcal{P}}{\longrightarrow} 1$ .

Proof.

Case  $\mathbb{H} = [0,1]$ . Consider for uniform iid  $x_i, y_i$ :

$$z_N = \frac{1}{N} \sum_{i=1}^{N} (x_i - y_i)^2$$

$$z_N \stackrel{\mathcal{P}}{\longrightarrow} \mathbb{E}(z_1) = \iint_{[0,1]^2} (x-y)^2 dxdy = \frac{1}{6}$$

$$\frac{d_0}{\sqrt{N}} = \min\left\{\sqrt{z_{N,0}}, \dots, \sqrt{z_{N,n}}\right\} \xrightarrow{\mathcal{P}} \frac{1}{\sqrt{6}}$$

$$\frac{d_n}{\sqrt{N}} = \max\left\{\sqrt{z_{N,0}}, \dots, \sqrt{z_{N,n}}\right\} \xrightarrow{\mathcal{P}} \frac{1}{\sqrt{6}}$$

$$\Rightarrow \quad \frac{d_0}{d_n} = \frac{d_0/\sqrt{N}}{d_n/\sqrt{N}} \xrightarrow{\mathcal{P}} \frac{1/\sqrt{6}}{1/\sqrt{6}} = 1$$

Case  $\mathbb{H} = \mathbb{F}_2$ .

- ▶ x, y uniform r.v. in  $\mathbb{F}_2^N \Rightarrow \|x\|_H \sim B(N, \frac{1}{2})$
- ► Also,  $d(x,y) = ||x + y|| \sim B(N, \frac{1}{2})$ :

$$\mathbb{P}(\|x + y\| = k, \ x \text{ fixed}) = \frac{\binom{N}{k}}{4^N}$$

$$\mathbb{P}(\|x + y\| = k) = \sum_{x} \mathbb{P}(\|x + y\| = k, \ x \text{ fixed}) = 2^N \cdot \frac{\binom{N}{k}}{4^N} = \frac{\binom{N}{k}}{2^N}$$

► For normalised positive distances  $\frac{d_H(x,y)}{N} > 0$ :

$$\mathbb{E}\left(\frac{d_{H}(x,y)}{N}\right) = \frac{1}{2\left(1 - \frac{1}{2N}\right)}$$

By Chebyshev inequality,

$$\mathbb{P}\left(\left|\frac{d_{H}(x,y)}{N} - \frac{1}{2\left(1 - \frac{1}{2^{N}}\right)}\right| > \epsilon\right) \le \frac{1}{\epsilon^{2}} \operatorname{Var}\left(\frac{d_{H}(x,y)}{N}\right) \\
= \frac{1}{\epsilon^{2}} \left(\frac{1}{4N\left(1 - \frac{1}{2^{N}}\right)} + \frac{1}{4\left(1 - \frac{1}{2^{N}}\right)} - \frac{1}{4\left(1 - \frac{1}{2^{N}}\right)^{2}}\right) \\
\to 0 \quad (N \to \infty)$$

- ▶ This means  $\frac{d_H(x,y)}{N} \stackrel{\mathcal{P}}{\longrightarrow} \frac{1}{2}$
- ► Hence,  $\frac{d_0}{N} \xrightarrow{\mathcal{P}} \frac{1}{2}$ ,  $\frac{d_n}{N} \xrightarrow{\mathcal{P}} \frac{1}{2}$
- ▶ Hence,  $\frac{d_0}{d_n} \stackrel{\mathcal{P}}{\longrightarrow} 1$



#### **Theorem**

Let  $\mathbb{H}=[0,1]$ , and  $x\in X$  with independent coordinates  $x_i\sim N(\mu_i,\sigma_i)$  such that  $\sigma_i^2\leq b$  for all i. Then  $\frac{d_0}{d_n}\stackrel{\mathcal{P}}{\longrightarrow} 1$ .

Proof.

$$z_N = \frac{1}{N} \sum_{i=1}^N (x_i - y_i)^2$$

$$\mathbb{E}(z_N) = \frac{2}{N} \sum_{i=1}^N \sigma_i^2 \le \frac{2N}{N} b = 2b$$

$$\text{Var } z_N = \frac{1}{N^2} \sum_{i=1}^N \mathbb{E}\left((x_i - y_i)^2\right) = \frac{1}{N} \mathbb{E}(z_N)$$

Chebyshev inequality:

$$\mathbb{P}(|z_N - \mathbb{E}(z_N)| > \epsilon) \leq \frac{\mathsf{Var}\,z_N}{\epsilon^2} = \frac{1}{N\epsilon^2}\mathbb{E}(z_N) \leq \frac{2b}{N\epsilon^2} \to 0$$

▶ 
$$\mathbb{E}(z_N) > 0$$
 is increasing and bounded  $\Rightarrow \mathbb{E}(z_N) \to \zeta > 0$ 

$$\Rightarrow z_N \xrightarrow{\mathcal{P}} \zeta > 0$$

$$ightharpoonup 
ightharpoonup rac{d_0}{d_n} \stackrel{\mathcal{P}}{\longrightarrow} rac{\sqrt{\zeta}}{\sqrt{\zeta}} = 1$$



Categorial data.

X in complete disjunctive form

- elementary vector x<sub>i</sub> has precisely one 1-entry.
- $d(x_i, y_i) = 2\delta_{x_i, y_i}$
- $d(x,y) = \sum_{i=1}^{\ell} d(x_i,y_i)$
- ▶  $\mathbb{P}(d(x_i, y_i) = 2) = 1 \frac{1}{k_i}$
- $\mathbb{E}\left(\frac{d(x,y)}{\ell}\right) = \frac{2}{\ell} \sum_{i=1}^{\ell} \left(1 \frac{1}{k_i}\right) = 2\left(1 \frac{1}{\ell} \sum_{i=1}^{\ell} \frac{1}{k_i}\right)$
- ► Var  $\left(\frac{d(x,y)}{\ell}\right) = \frac{1}{\ell} \sum_{i=1}^{\ell} \frac{4}{\ell} \left(\frac{1}{k_i} \frac{1}{k_i^2}\right)$

Categorial data.

#### Theorem

Let 
$$k_i \geq 2$$
. If  $\frac{1}{\ell} \sum_{i=1}^{\ell} \frac{1}{k_i}$  converges for  $\ell \to \infty$ , then  $\frac{d_0}{d_n} \stackrel{\mathcal{P}}{\longrightarrow} 1$ .

#### Proof.

Chebyshev inequality:

$$\mathbb{P}\left(\left|\frac{d(x,y)}{\ell} - \mathbb{E}\frac{d(x,y)}{\ell}\right| > \epsilon\right) \le \frac{\operatorname{Var}\frac{d(x,y)}{\ell}}{\epsilon^2}$$

$$= \frac{1}{\epsilon^2} \cdot \frac{1}{\ell} \sum_{i=1}^{\ell} \frac{4}{\ell} \left(\frac{1}{k_i} - \frac{1}{k_i^2}\right) \to 0 \quad \text{(Cesàro means)}$$

If 
$$\lim_{\ell \to \infty} \frac{1}{\ell} \sum_{i=1}^{\ell} \frac{1}{k_i} = C$$
, then  $C \leq \frac{1}{2}$  and

$$\frac{d(x,y)}{\ell} \xrightarrow{\mathcal{P}} 2(1-C) > 0$$

$$\Rightarrow \frac{d_0}{d_0} \xrightarrow{\mathcal{P}} \frac{2(1-C)}{2(1-C)} = 1$$

- $ightharpoonup m_N := m(\mathbb{F}_2^N, d_H)$
- $> t_N := t(\mathbb{F}_2^N, d_H)$

#### **Theorem**

$$\frac{1}{N} < t_N < m_N < \frac{C}{\sqrt{N}}$$

for N >> 0 with C > 0. In particular

$$\lim_{N\to\infty}t_N=\lim_{N\to\infty}m_N=0$$

## Proof that $t_N < \frac{2}{N}$ .

- ▶ For  $k \le \epsilon < k+1$  each k-face of  $\mathbb{F}_2^N$  is a maximal clique of  $\Gamma_\epsilon$
- ▶  $\Gamma_{\epsilon}$  is connected for  $k \ge 1$
- #k-faces of  $\mathbb{F}_2^N = 2^{N-k} \binom{N}{k}$

$$t_{N} \leq \frac{1}{N} \left( 1 + \sum_{k=1}^{N-1} \frac{1}{2^{N-k} {N \choose k}} \right) \leq \frac{1}{N} + \frac{1}{N} \sum_{k=1}^{N-1} \frac{1}{2^{N-k}}$$
$$= \frac{1}{N} + \frac{1}{N} \left( 1 - \frac{1}{2^{N-1}} \right) = \frac{2}{N} \left( 1 - \frac{1}{2^{N}} \right)$$
$$< \frac{2}{N}$$

Proof that  $t_N > \frac{1}{N}$ .

$$t_{N} = \frac{1}{N} \left( 1 + \sum_{k=1}^{N-1} \frac{1}{c(\Gamma_{k})} \right) > \frac{1}{N}$$



- ullet  $\mathcal{U}_{\mathcal{N}}:=\left\{ \mathsf{ultrametric}\ \Delta\ \mathsf{in}\ \mathbb{F}_{2}^{\mathcal{N}}
  ight\}$
- $ightharpoonup \mathbb{F}_2^N$  acts without fixed points on  $\mathcal{U}_N$  via translations
- $ightharpoonup 
  ightharpoonup u_N := rac{|\mathcal{U}_N|}{2^N} \in \mathbb{N}$

### Proposition

$$u_{N} = \sum_{k=3}^{N} \sum_{\left\lceil \frac{2k}{3} \right\rceil \le i \le k} {N \choose k} {k \choose i} {i \choose k-i}$$

#### Proof.

- ▶ Via translation:  $\Delta$  has form (0, a, b)
- ► Side lengths: ||a||, ||b||, ||a + b||
- ▶ I := supp(a),  $J = \text{supp}(b) \Rightarrow \text{supp}(a+b) = I \triangle J$
- ▶ Assume triangle is in a k-face, but not inside one of its faces:  $\Rightarrow |I \cup J| = k$
- ▶  $\Delta$  ultrametric  $\Rightarrow |I| = |J|$  and  $|I \triangle J| \le |I|$
- with  $\ell = |I \cap J|$  this means:

$$2|I| - \ell = k$$
$$2|I| - 2\ell \le i$$

▶ solution:  $|I| \ge \left\lceil \frac{2k}{3} \right\rceil$ 

number of such triangles:

$$\sum_{\left\lceil \frac{2k}{3}\right\rceil \le i \le k} \binom{k}{i} \binom{i}{\ell} = \sum_{\left\lceil \frac{2k}{3}\right\rceil \le i \le k} \binom{k}{i} \binom{i}{2i-k}$$
$$= \sum_{\left\lceil \frac{2k}{3}\right\rceil \le i \le k} \binom{k}{i} \binom{i}{k-i}$$

▶ All ultrametric  $\Delta$  of form (0, a, b):

$$u_{N} = \sum_{k=3}^{N} {N \choose k} \sum_{\left\lceil \frac{2k}{3} \right\rceil \le i \le k} {k \choose i} {i \choose k-i}$$

#### **Proof that** $m_N \rightarrow 0$ .

Gosper's approximation:

$$n! = \sqrt{2\pi} \left(\frac{n}{e}\right)^n \sqrt{n + \frac{1}{6}} \cdot Q(n), \quad \lim_{n \to \infty} Q(n) = 1$$

$$m_N \sim \frac{6u_N}{4^N} = \frac{6}{4^N} \sum_{k=3}^N \sum_{\left\lceil \frac{2k}{2} \right\rceil \le i \le k} \frac{N!}{(N-k)!(k-i)!^2(2i-k)!}$$

$$<\frac{6Q_{\max}}{(2\pi)^{\frac{3}{2}}N^{\frac{3}{2}}}\sum\sum\frac{\left(1+\frac{1}{6N}\right)^{\frac{1}{2}}}{\left(1-\frac{k}{N}+\frac{1}{6N}\right)^{\frac{1}{2}}\left(\frac{k}{N}-\frac{i}{N}+\frac{1}{6N}\right)\left(\frac{2i}{N}-\frac{k}{N}+\frac{1}{6N}\right)^{\frac{1}{2}}}}{\cdot\left(\frac{1}{4\left(1-\frac{k}{N}\right)^{1-\frac{k}{N}}\left(\frac{k}{N}-\frac{i}{N}\right)^{2\left(\frac{k}{N}-\frac{i}{N}\right)}\left(\frac{2i}{N}-\frac{k}{N}\right)^{\frac{2i}{N}-\frac{k}{N}}}\right)^{N}}\right)^{N}}$$

$$\sim\frac{6Q_{\max}N^{\frac{1}{2}}}{(2\pi)^{\frac{3}{2}}}\int\limits_{\frac{3}{N}}^{1}\int\limits_{\frac{3}{2}x}^{x}h_{N}(x,y)e^{Nf(x,y)}\,dydx$$

$$\begin{split} Q_{\text{max}} &= \max \left\{ \frac{Q(N)}{Q(N-k)Q(k-i)^2Q(2i-k)} \right\} \\ h_N(x,y) &= \frac{1}{\left(1-x+\frac{1}{6N}\right)^{\frac{1}{2}}\left(x-y+\frac{1}{6N}\right)\left(2y-x+\frac{1}{6N}\right)^{\frac{1}{2}}} \\ f(x,y) &= -\log 4 - (1-x\log(1-x)-2(x-y)\log(x-y) \\ &- (2y-x)\log(2y-x) \end{split}$$

- $(\frac{3}{4}, \frac{1}{2})$  is unique global maximum of f(x, y)
- $f(\frac{3}{4},\frac{1}{2})=0$
- ▶ Hessian matrix  $H(f(\frac{3}{4}, \frac{1}{2}))$  is negative definite



⇒ Laplace method yields:

$$egin{aligned} m_{N} &\lesssim rac{6Q_{\mathsf{max}}N^{rac{1}{2}}}{(2\pi)^{rac{3}{2}}} \left(rac{2\pi}{N}
ight) \cdot \det\left(H(f\left(rac{3}{4},rac{1}{2}
ight))
ight)^{-rac{1}{2}} h_{N}\left(rac{3}{4},rac{1}{2}
ight) e^{Nf\left(rac{3}{4},rac{1}{2}
ight)} \ &pprox rac{2.5651}{N^{rac{1}{2}}} \end{aligned}$$

#### Lower bound for $m_N$ .

▶ Replace < by > and  $Q_{max}$  by

$$Q_{\min} = \min \left\{ \frac{Q(N)}{Q(N-k)Q(k-i)^2Q(2i-k)} \right\}$$

$$\Rightarrow m_{N} \gtrsim \frac{6Q_{\min}}{(2\pi)^{\frac{1}{2}}N^{\frac{1}{2}}} \approx \frac{2.4}{N^{\frac{1}{2}}}$$

### **Conclusion**

- ► For random samples of fixed size, the ultrametricity indices tend to one if dimension tends to infinity.
- ► For the discrete hypercube, the ultrametricity indices tend to zero as dimension tends to infinity.
- In particular, the fraction of ultrametric triangles becomes negligible in the discrete hypercube, as dimension tends to infinity.
- Randomness and sparsity pick precisely these, as dimension tends to infinity.