

**POLI 30 D: Political Inquiry**  
Professor Umberto Mignozzetti  
(Based on DSS Materials)

**Lecture 10 | Causal Inference with  
Observational Data I**

## Before we start

### Announcements:

- ▶ Quizzes and Participation: On Canvas.
- ▶ GitHub page:  
<https://github.com/umbertomig/POLI30Dpublic>
- ▶ Piazza forum: Not sure what the link is. Ask your TA!
- ▶ Note to self: Turn on the mic!

## Before we start

**Recap:** We learned:

- ▶ The definitions of theory, scientific theory, and hypotheses.
- ▶ Data, datasets, variables, and how to compute means.
- ▶ Causal effect, treatments, outcomes, and randomization.
- ▶ Sampling, descriptive statistics, and descriptive plots for one variable.
- ▶ Correlation between two continuous variables.
- ▶ Prediction of a binary and a non-binary variable.

**Great job!**

- ▶ Do you have any questions about these contents?

## Plan for Today

- Review: Causation and Randomized Experiments
- Observational Studies
- Confounding Variables or Confounders
  - Why Are Confounders a Problem?
  - Why Don't We Worry About Confounders in Randomized Experiments?
- How Can We Estimate Causal Effects with Observational Data?
  - Interpretation of  $\hat{\beta}$  When X Is the Treatment Variable and Y Is the Outcome Variable

## Review: Causation

- ▶ To measure causal effects, we need to compare the factual outcome with the counterfactual outcome
  - ▶ Fundamental problem: We can never observe counterfactual outcomes
- ▶ To estimate causal effects, we must find or create a situation in which the treatment and control groups are **comparable**.
- ▶ Only when that assumption is satisfied can we use the *factual* outcome of one group as a good *proxy* for the *counterfactual* outcome of the other.

## Review: Randomized Experiments

- ▶ In randomized experiments, we can rely on the **random assignment of treatment** to make treatment and control groups, on average, identical
- ▶ Thus, we can estimate the average treatment effect with the **difference-in-means estimator**

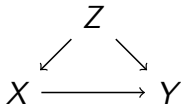
$$\overline{Y}_{\text{treatment group}} - \overline{Y}_{\text{control group}}$$

## Observational Data

- ▶ But what happens when we cannot conduct a randomized experiment and have to analyze observational data?
  - ▶ *Observational data*: data collected about naturally occurring events (i.e., researchers do not get to assign the treatment)
- ▶ We can no longer assume that treatment and control groups are comparable.
- ▶ We have to identify any relevant differences between treatment and control groups (known as confounding variables or confounders)
- ▶ Then, we have to statistically control for them so that we may claim that the two groups are comparable.

## Confounders or Confounding Variables

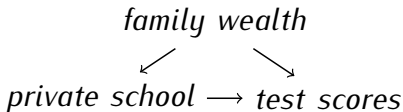
- ▶ A **confounding variable**, or **confounder**, is a variable that affects both:
  1. The likelihood of receiving the treatment  $X$ , and
  2. The outcome  $Y$
- ▶ In mathematical notation, we represent a potential confounding variable as  $Z$





## Confounders or Confounding Variables

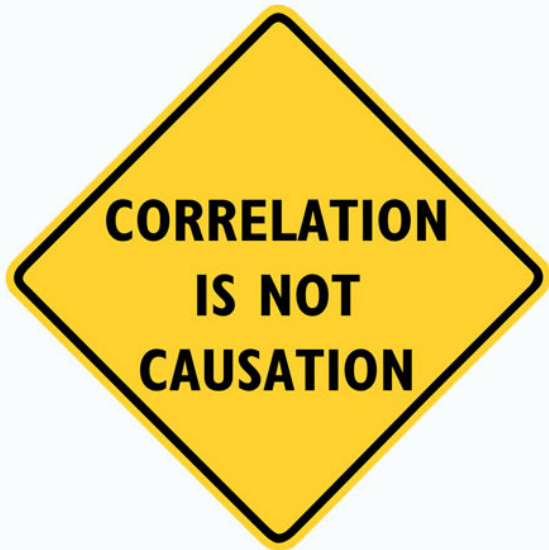
- ▶ Suppose we are interested in the average causal effect of attending a private school instead of a public one on SAT performance.
- ▶ What is the treatment variable  $X$ ?
  - ▶ What is the outcome variable  $Y$ ?
  - ▶ Can you think of a potential confounder  $Z$ ?



## Why Are Confounders a Problem?

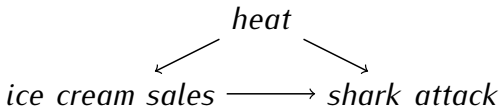
- ▶ They obscure the causal relationship between  $X$  and  $Y$ !
- ▶ In the example above, if we observed that, on average, private school students perform better than public school students, we would not know whether it is:
  - ▶ Because they attended a private school or
  - ▶ Because they came from wealthier families
- ▶ We would not know what portion of the observed differences in SAT performance (the difference-in-means estimator), if any, could be attributed to:
  - ▶ Attending a private school versus
  - ▶ Coming from a wealthy family.

## Why Are Confounders a Problem?



## Why Are Confounders a Problem?

- ▶ In the presence of confounders, correlation does not necessarily imply causation.
- ▶ Just because two variables are highly correlated, it does **not** mean that one causes the other:
  - ▶ There could be a third variable that causes both!
- ▶ Ice cream sales and shark attacks are highly correlated.
  - ▶ Does this mean that eating ice cream increases the probability that a shark will attack you?



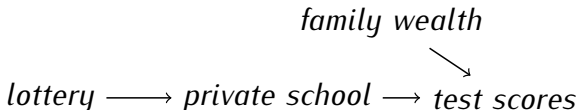
## Why Are Confounders a Problem?

### IN THE PRESENCE OF CONFOUNDERS

- Correlation does not imply causation.
- The treatment and control groups are not **comparable**.
- The difference-in-means estimator does **NOT** provide a valid estimate of the average causal effect!

## Why Don't We Worry About Confounders in Randomized Experiments?

- ▶ Randomization of treatment assignment eliminates all potential confounders.
  - ▶ That is this is the **gold standard** for causal inference.
- ▶ Ensures that treatment and control are comparable by breaking the link between any potential confounder.
- ▶ If we have a lottery to randomly determine who will attend the private school, then we break the wealth link.



## How Can We Estimate Causal Effects with Observational Data?

- ▶ We cannot rely on random treatment assignments to eliminate potential confounders.
- ▶ We must identify all potential confounders and statistically control for them using a multiple linear regression model.
- ▶ Before we learn how to do that, we will fit a linear regression model to find the *difference-in-means estimator*.

## Using the Linear Regression to Compute the Difference-in-Means Estimator

When  $X$  is the treatment variable, and  $Y$  is the outcome variable of interest, the estimated slope coefficient ( $\hat{\beta}_1$ ) is **equivalent** to the *difference-in-means estimator*.

- ▶ Let us return to a beloved example: *Does Social Pressure Affect Turnout?*
- ▶ Registered voters were randomly assigned to either:
  - a. receive a message designed to induce social pressure or
  - b. receive nothing



## Does Social Pressure Affect Turnout?



(Based on Alan S. Gerber, Donald P. Green, and Christopher W. Larimer. 2008. "Social Pressure and Voter Turnout: Evidence from a Large-Scale Field Experiment." *American Political Science Review*, 102 (1): 33-48.)

## Does Social Pressure Affect Turnout?

1. Load and look at the data:

```
voting <- read.csv("https://raw.githubusercontent.com/umber  
head(voting) # shows first six observations  
##    birth message voted  
## 1   1981      no     0  
## 2   1959      no     1  
## 3   1956      no     1  
## 4   1939     yes     1  
## 5   1968      no     0  
## 6   1967      no     0
```

2. Creating the treatment variable:

```
voting$pressure <- ifelse(voting$message=="yes", 1, 0)
```

## Does Social Pressure Affect Turnout?

3. Compute difference-in-means estimator directly

```
mean(voting$voted[voting$pressure==1]) -  
  mean(voting$voted[voting$pressure==0])  
## [1] 0.08130991
```

4. Alternatively, we can fit a linear model where X is the treatment variable and Y is the outcome variable.

## Does Social Pressure Affect Turnout?

- Recall: the R function to fit a linear model is `lm()`

```
lm(voted ~ pressure, data=voting)
##
## Call:
## lm(formula = voted ~ pressure, data = voting)
##
## Coefficients:
## (Intercept)      pressure
##      0.29664         0.08131
```

- Fitted model:  $\widehat{voted} = 0.30 + 0.08 \text{ pressure}$
- Note that  $\widehat{\beta}_1$  has the same value as the difference-in-means estimator above (both equal 0.08)

## Interpretation of $\hat{\beta}_1$ When X Is the Treatment Variable, and Y Is the Outcome Variable

- ▶ Start the same as in predictive models:
  - ▶ Definition:  $\hat{\beta}$  is the  $\Delta \hat{Y}$  associated with  $\Delta X=1$ 
    - ▶  $\hat{\beta} = 0.08$  is the  $\Delta \widehat{\text{voted}}$  associated with  $\Delta \text{pressure}=1$
  - ▶ Receiving a social-pressure inducing message is associated with a predicted increase in the probability of voting of 8 p.p., on average
- ▶ Unit of measurement of  $\hat{\beta}_1$ ? same as  $\Delta \bar{Y}$ .
  - ▶ Since Y is binary,  $\Delta \bar{Y}$  is measured in p.p., and so is  $\hat{\beta}$  (after x 100)

## Interpretation of $\hat{\beta}_1$ When $X$ Is the Treatment Variable, and $Y$ Is the Outcome Variable

- ▶ Since  $X$  is the treatment variable and  $Y$  is the outcome variable,  $\hat{\beta}_1$  is equivalent to the difference-in-means estimator
- ▶ As a result, we can interpret  $\hat{\beta}_1$  using **causal language**
- ▶ **Predictive language:** We estimate that receiving the message inducing social pressure *is associated with a predicted increase* in the probability of voting of 8 p.p., on average
- ▶ **Causal language:** We estimate that receiving the message inducing social pressure *increases* the probability of voting by 8 p.p., on average

## Interpretation of $\hat{\beta}_1$ When X Is the Treatment Variable, and Y Is the Outcome Variable

- ▶ This should be a valid estimate of the average treatment effect if there are no confounding variables present:
  - ▶ If registered voters who received the message are comparable to those who did not.
- ▶ Since the data come from a randomized experiment, there should be no confounding variables (why?)
- ▶ And thus, the difference-in-means estimator should produce a valid estimate of the average treatment effect

## Interpretation of $\hat{\beta}_1$ When X Is the Treatment Variable, and Y Is the Outcome Variable

- ▶ **Conclusion:** A message inducing social pressure increases the probability of voting by eight p.p., on average.
  - ▶ Valid estimate of the ATE if registered voters who received the message are comparable to those who did not.
  - ▶ This is a reasonable assumption, given that the data come from a randomized experiment.
- ▶ Note that this is the same conclusion we arrived at in a previous lecture.



## Interpretation of $\hat{\beta}_1$ When $X$ Is the Treatment Variable, and $Y$ Is the Outcome Variable

### INTERPRETATION OF THE ESTIMATED SLOPE COEFFICIENT IN THE SIMPLE LINEAR MODEL:

- ▶ By default, we interpret  $\hat{\beta}_1$  using predictive language: It is the  $\Delta\hat{Y}$  *associated with*  $\Delta X=1$ .
- ▶ When  $X$  is the treatment variable, then  $\hat{\beta}_1$  is equivalent to the difference-in-means estimator and, thus, we interpret  $\hat{\beta}_1$  using causal language: It is the  $\Delta\hat{Y}$  *caused by*  $\Delta X=1$ . This causal interpretation is valid if no confounding variables exist: the treatment and control groups are comparable.

## Summary

### ► Today's Class:

- Observational Studies
- Confounding Variables or Confounders
  - Why Are Confounders a Problem?
  - Why Don't We Worry About Confounders in Randomized Experiments?
- How Can We Estimate Causal Effects with Observational Data?
- Interpretation of  $\hat{\beta}_1$  when X is the Treatment Variable, and Y Is the Outcome Variable.

### ► Next class:

- More Causality with Observational Data:
  - We will use *Multiple Regression* models to control for confounders.

Questions?

See you in the next class!