

# An Ensemble meta-prediction framework to integrate multiple regression models into a current study



Tian Gu  @GuTian\_TianGu

Department of Biostatistics, University of Michigan

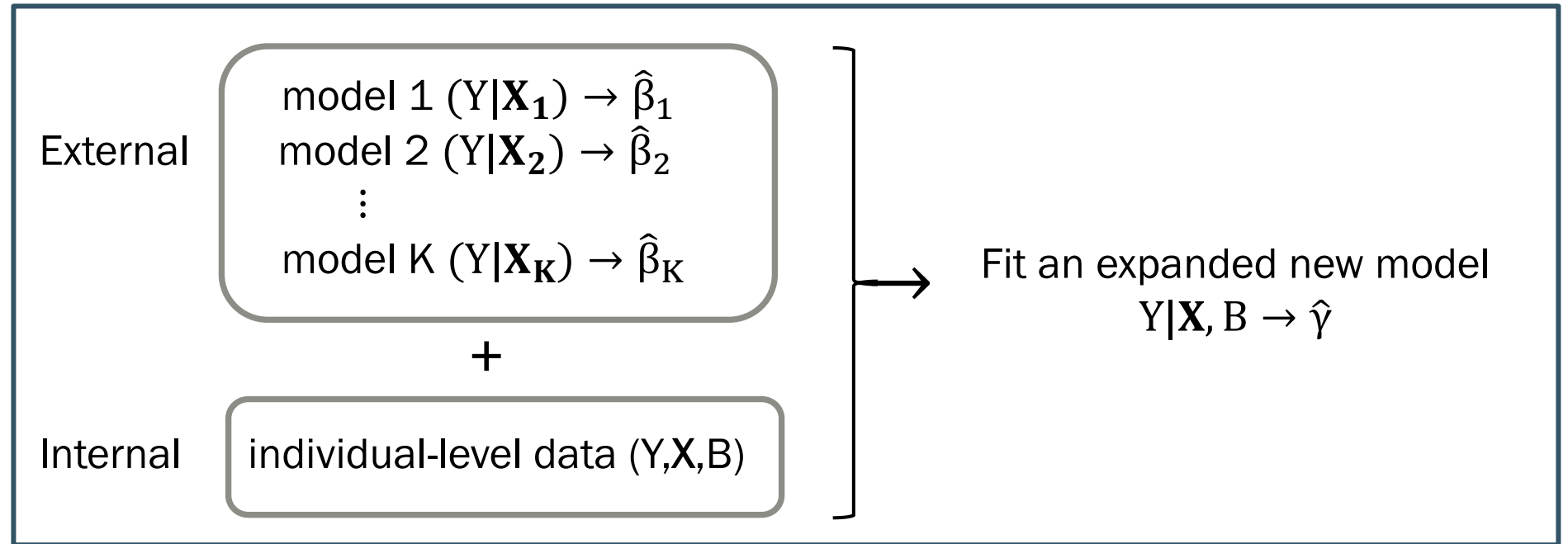


2020 JSM August 4<sup>th</sup>

# Background & Settings

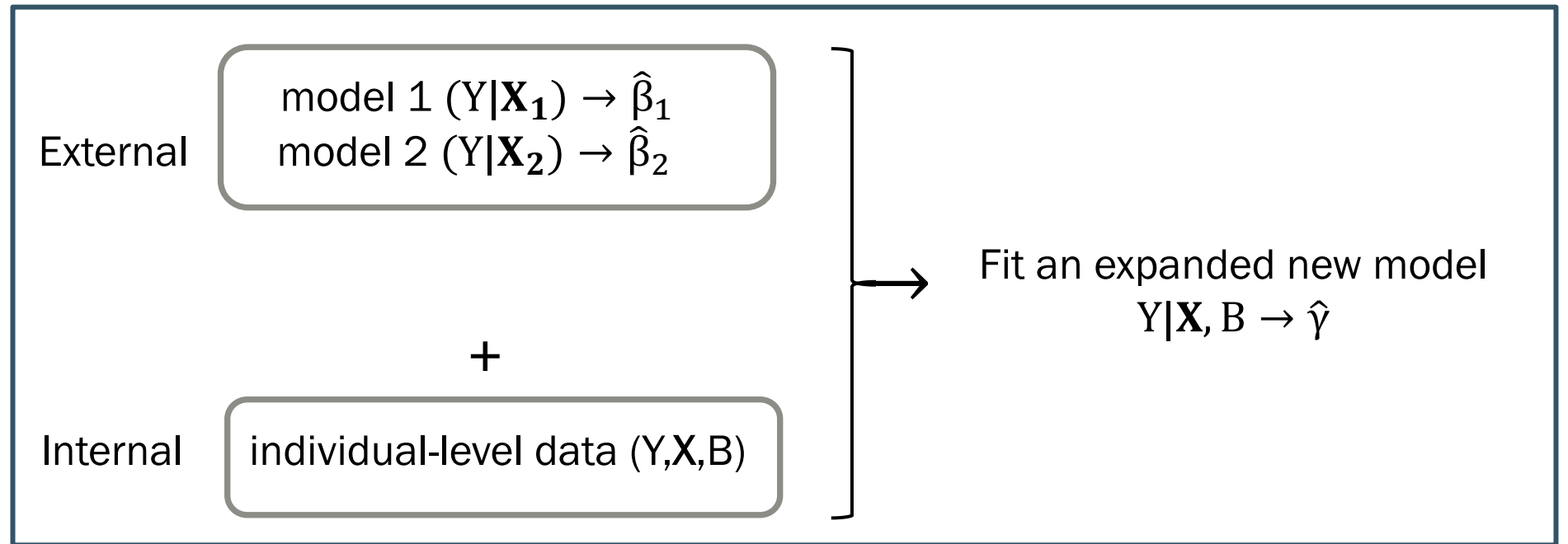
# Background in statistical language

Goal



For simplicity, all the following illustration will be using the case of 2 external models

Goal



# Proposed Framework & Estimates

# Key notations

$\hat{\mathbf{Y}}_{\mathbf{I}}$  maximum likelihood estimate (MLE) using the complete internal data

$\hat{\mathbf{Y}}_{\mathbf{CML}}$  the constrained ML (CML) estimate

$\hat{\mathbf{Y}}_{\mathbf{EB}}$  the empirical Bayes (EB) estimate

## 2-Step Framework

### Proposed Framework

Step 1:

$$\text{External } \hat{\beta}_1 \xrightarrow{\text{Chatterjee's}} \hat{\gamma}_{CML^1} \xrightarrow{\text{Estes'}} \hat{\gamma}_{EB^1} = \hat{W}_1 \hat{\gamma}_I + (I - \hat{W}_1) \hat{\gamma}_{CML^1}$$

$$\text{External } \hat{\beta}_2 \xrightarrow{\text{Chatterjee's}} \hat{\gamma}_{CML^2} \xrightarrow{\text{Estes'}} \hat{\gamma}_{EB^2} = \hat{W}_2 \hat{\gamma}_I + (I - \hat{W}_2) \hat{\gamma}_{CML^2}$$

Step 2:

$$\begin{bmatrix} \hat{\gamma}_{EB^1} \\ \hat{\gamma}_{EB^2} \end{bmatrix} \sim N(\gamma_0, V),$$

$\hat{\gamma}_0$  = A weighted average of  $\hat{\gamma}_{EB^1}$  and  $\hat{\gamma}_{EB^2}$

# Constrained Maximum Likelihood Estimation for Model Calibration Using Summary-Level Information From External Big Data Sources

Nilanjan CHATTERJEE, Yi-Hau CHEN, Paige MAAS, and Raymond J. CARROLL

Proposed  
Framework

Step 1:

$$\text{External } \hat{\beta}_1 \xrightarrow{\text{Chatterjee's}} \hat{\gamma}_{CML^1} \xrightarrow{\text{Estes'}} \hat{\gamma}_{EB^1} = \hat{W}_1 \hat{\gamma}_I + (I - \hat{W}_1) \hat{\gamma}_{CML^1}$$

$$\text{External } \hat{\beta}_2 \xrightarrow{\text{Chatterjee's}} \hat{\gamma}_{CML^2} \xrightarrow{\text{Estes'}} \hat{\gamma}_{EB^2} = \hat{W}_2 \hat{\gamma}_I + (I - \hat{W}_2) \hat{\gamma}_{CML^2}$$

Step 2:

$$\begin{bmatrix} \hat{\gamma}_{EB^1} \\ \hat{\gamma}_{EB^2} \end{bmatrix} \sim N(\gamma_0, V),$$

$\hat{\gamma}_0$  = A weighted average of  $\hat{\gamma}_{EB^1}$  and  $\hat{\gamma}_{EB^2}$



### Assumptions

- 1 Known form of  $g_\beta(Y|X)$
- 2  $f_\gamma(Y|X, B)$  is correctly-specified
- 3  $F(X, B)$  is the same in internal and external studies

Untestable  
Transportability  
assumptions may  
lead to bias

## Proposed Framework

### Construct constraints

- 1  $p_i \equiv Pr(X = X_i, B = B_i), \quad \sum_{i=1}^n p_i = 1$
- 2 Connect external score function to internal model

$$\begin{aligned}
 0 &= E_{Y,X,B} \left( \frac{\partial l_\beta(Y|X)}{\partial \beta} \right) = E_{Y,X,B} (U_\beta(Y|X)) \\
 &= E_{X,B} \{ E_{Y|X,B} (U_\beta(Y|X) | X, B) \} \\
 &= \int_{X,B} \int_{Y|X,B} U_\beta(Y|X) f_\gamma(Y|X, B) dY dF(X, B) \\
 &= \sum_{i=1}^n \int_{Y|X,B} U_\beta(Y|X) f_\gamma(Y|X, B) dY p_i
 \end{aligned}$$

- Maximize the likelihood  $\prod_{i=1}^n f_\gamma(Y_i|X_i, B_i) p_i$  w.r.t  $\gamma$  and  $p_i$  subject to two constraints  
 $\sum_{i=1}^n p_i = 1$  and  $\sum_{i=1}^n \int U_\beta(Y|X) f_\gamma(Y|X, B) dY p_i = 0$

$$\begin{aligned}
 \hat{\gamma}_{CML} &= \arg \max_{\gamma, p_i} \left\{ \prod_{i=1}^n f_\gamma(Y_i|X_i, B_i) p_i \right. \\
 &\quad \left. + \lambda_1 \left( \sum_{i=1}^n p_i - 1 \right) \right. \\
 &\quad \left. + \lambda_2 \sum_{i=1}^n \int U_\beta(Y|X) f_\gamma(Y|X, B) dY p_i \right\}
 \end{aligned}$$

Step 1:

$$\text{External } \hat{\beta}_1 \xrightarrow{\text{Chatterjee's}} \hat{\gamma}_{CML^1} \xrightarrow{\text{Estes'}} \hat{\gamma}_{EB^1} = \hat{W}_1 \hat{\gamma}_I + (I - \hat{W}_1) \hat{\gamma}_{CML^1}$$

$$\text{External } \hat{\beta}_2 \xrightarrow{\text{Chatterjee's}} \hat{\gamma}_{CML^2} \xrightarrow{\text{Estes'}} \hat{\gamma}_{EB^2} = \hat{W}_2 \hat{\gamma}_I + (I - \hat{W}_2) \hat{\gamma}_{CML^2}$$

Step 2:

$$\begin{bmatrix} \hat{\gamma}_{EB^1} \\ \hat{\gamma}_{EB^2} \end{bmatrix} \sim N(\gamma_0, V),$$

$\hat{\gamma}_0 =$  A weighted average of  $\hat{\gamma}_{EB^1}$  and  $\hat{\gamma}_{EB^2}$

$$\text{Var}([\hat{\gamma}_{CML^1}, \hat{\gamma}_{CML^2}, \hat{\gamma}_I]^T)$$

$$= \begin{bmatrix} \text{Var}(\hat{\gamma}_{CML^1}) & \text{Cov}(\hat{\gamma}_{CML^1}, \hat{\gamma}_{CML^2}) & \text{Cov}(\hat{\gamma}_{CML^1}, \hat{\gamma}_I) \\ & \text{Var}(\hat{\gamma}_{CML^2}) & \text{Cov}(\hat{\gamma}_{CML^2}, \hat{\gamma}_I) \\ & & \text{Var}(\hat{\gamma}_I) \end{bmatrix}$$

We derived the closed-form covariance terms

## Proposed Framework

Step 1:

$$\text{External } \hat{\beta}_1 \xrightarrow{\text{Chatterjee's}} \hat{\gamma}_{CML^1} \xrightarrow{\text{Estes'}} \hat{\gamma}_{EB^1} = \hat{W}_1 \hat{\gamma}_I + (I - \hat{W}_1) \hat{\gamma}_{CML^1}$$


$$\text{External } \hat{\beta}_2 \xrightarrow{\text{Chatterjee's}} \hat{\gamma}_{CML^2} \xrightarrow{\text{Estes'}} \hat{\gamma}_{EB^2} = \hat{W}_2 \hat{\gamma}_I + (I - \hat{W}_2) \hat{\gamma}_{CML^2}$$

Step 2:

$$\begin{bmatrix} \hat{\gamma}_{EB^1} \\ \hat{\gamma}_{EB^2} \end{bmatrix} \sim N(\gamma_0, V),$$

$$\hat{\gamma}_0 = \text{A weighted average of } \hat{\gamma}_{EB^1} \text{ and } \hat{\gamma}_{EB^2}$$

# Empirical Bayes Estimation and Prediction Using Summary-Level Information From External Big Data Sources Adjusting for Violations of Transportability

Jason P. Estes<sup>1</sup>  · Bhramar Mukherjee<sup>1</sup> · Jeremy M. G. Taylor<sup>1</sup>

Proposed  
Framework

Step 1:

$$\begin{aligned} \text{External } \hat{\beta}_1 &\xrightarrow{\text{Chatterjee's}} \hat{\gamma}_{CML^1} \xrightarrow{\text{Estes'}} \hat{\gamma}_{EB^1} = \hat{W}_1 \hat{\gamma}_I + (I - \hat{W}_1) \hat{\gamma}_{CML^1} \\ \text{External } \hat{\beta}_2 &\xrightarrow{\text{Chatterjee's}} \hat{\gamma}_{CML^2} \xrightarrow{\text{Estes'}} \hat{\gamma}_{EB^2} = \hat{W}_2 \hat{\gamma}_I + (I - \hat{W}_2) \hat{\gamma}_{CML^2} \end{aligned}$$

Step 2:

$$\begin{bmatrix} \hat{\gamma}_{EB^1} \\ \hat{\gamma}_{EB^2} \end{bmatrix} \sim N(\gamma_0, V),$$

$\hat{\gamma}_0$  = A weighted average of  $\hat{\gamma}_{EB^1}$  and  $\hat{\gamma}_{EB^2}$

To correct the potential bias due to incomparability: the larger the difference between internal & external estimates, the more  $\hat{\gamma}_{EB}$  shrinks toward  $\hat{\gamma}_I$

$$\begin{cases} \gamma \sim N(\gamma_0, A) \\ \hat{\gamma}_I | \gamma \sim N(\gamma, \Sigma) \end{cases} \Rightarrow \hat{\gamma}_{posterior} = A(\Sigma + A)^{-1} \hat{\gamma}_I + \Sigma(\Sigma + A)^{-1} \gamma_0$$

$$\begin{aligned} \hat{\gamma}_{EB} &= \hat{A}(\hat{\Sigma} + \hat{A})^{-1} \hat{\gamma}_I + \hat{\Sigma}(\hat{\Sigma} + \hat{A})^{-1} \hat{\gamma}_{CML} \\ &= \hat{W} \hat{\gamma}_I + (I - \hat{W}) \hat{\gamma}_{CML} \end{aligned} \quad \begin{cases} \hat{A} &= (\hat{\gamma}_I - \hat{\gamma}_{CML})(\hat{\gamma}_I - \hat{\gamma}_{CML})^T \\ \hat{\Sigma} &= \hat{Var}(\hat{\gamma}_I) \end{cases}$$

## Proposed Framework

Step 1:

$$\text{External } \hat{\beta}_1 \xrightarrow{\text{Chatterjee's}} \hat{\gamma}_{CML1} \xrightarrow{\text{Estes'}} \hat{\gamma}_{EB1} = \hat{W}_1 \hat{\gamma}_I + (I - \hat{W}_1) \hat{\gamma}_{CML1}$$

$$\text{External } \hat{\beta}_2 \xrightarrow{\text{Chatterjee's}} \hat{\gamma}_{CML2} \xrightarrow{\text{Estes'}} \hat{\gamma}_{EB2} = \hat{W}_2 \hat{\gamma}_I + (I - \hat{W}_2) \hat{\gamma}_{CML2}$$

Step 2:

$$\begin{bmatrix} \hat{\gamma}_{EB1} \\ \hat{\gamma}_{EB2} \end{bmatrix} \sim N(\gamma_0, V),$$

$\hat{\gamma}_0 =$  A weighted average of  $\hat{\gamma}_{EB1}$  and  $\hat{\gamma}_{EB2}$

$$\begin{bmatrix} \hat{\gamma}_0^{EB1} \\ \hat{\gamma}_1^{EB1} \\ \vdots \\ \hat{\gamma}_p^{EB1} \end{bmatrix} + \begin{bmatrix} \hat{\gamma}_0^{EB2} \\ \hat{\gamma}_1^{EB2} \\ \vdots \\ \hat{\gamma}_p^{EB2} \end{bmatrix} \xrightarrow{?} \hat{\gamma}_0$$

Target evaluation metric

$$\sum_{i=1}^n \hat{Var}(X_i \hat{\gamma}_0)$$

$\downarrow$   $1 \times (PH)$        $\downarrow$   $(PH) \times 1$

Goal: combine 2  $\hat{\gamma}_{EB}$ 's to minimize the expected prediction variance, which reduce the p-dimension optimization to a scalar

Proposed Framework

Step 1:

$$\begin{aligned} \text{External } \hat{\beta}_1 &\xrightarrow{\text{Chatterjee's}} \hat{\gamma}_{CML1} \xrightarrow{\text{Estes'}} \hat{\gamma}_{EB1} = \hat{W}_1 \hat{\gamma}_I + (I - \hat{W}_1) \hat{\gamma}_{CML1} \\ \text{External } \hat{\beta}_2 &\xrightarrow{\text{Chatterjee's}} \hat{\gamma}_{CML2} \xrightarrow{\text{Estes'}} \hat{\gamma}_{EB2} = \hat{W}_2 \hat{\gamma}_I + (I - \hat{W}_2) \hat{\gamma}_{CML2} \end{aligned}$$

Step 2:

$$\begin{bmatrix} \hat{\gamma}_{EB1} \\ \hat{\gamma}_{EB2} \end{bmatrix} \sim N(\gamma_0, V),$$

$\hat{\gamma}_0 =$  A weighted average of  $\hat{\gamma}_{EB1}$  and  $\hat{\gamma}_{EB2}$

$$\begin{bmatrix} \hat{\gamma}_0^{EB1} \\ \hat{\gamma}_1^{EB1} \\ \vdots \\ \hat{\gamma}_p^{EB1} \end{bmatrix} + \begin{bmatrix} \hat{\gamma}_0^{EB2} \\ \hat{\gamma}_1^{EB2} \\ \vdots \\ \hat{\gamma}_p^{EB2} \end{bmatrix} \xrightarrow{?} \hat{\gamma}_0$$

Naïve approach:

Inverse variance weighed estimate,  
ignoring the correlation between  $\hat{\gamma}_{EB}$ 's

$$V = \text{Var} \begin{bmatrix} \hat{\gamma}_{EB1} \\ \hat{\gamma}_{EB2} \end{bmatrix} = \begin{bmatrix} \text{Var}(\hat{\gamma}_{EB1}) & \text{Cov}(\hat{\gamma}_{EB1}, \hat{\gamma}_{EB2}) \\ \text{Cov}(\hat{\gamma}_{EB2}, \hat{\gamma}_{EB1}) & \text{Var}(\hat{\gamma}_{EB2}) \end{bmatrix}$$

Proposed  
Framework

Step 1:

$$\begin{aligned} \text{External } \hat{\beta}_1 &\xrightarrow{\text{Chatterjee's}} \hat{\gamma}_{CML1} \xrightarrow{\text{Estes'}} \hat{\gamma}_{EB1} = \hat{W}_1 \hat{\gamma}_I + (I - \hat{W}_1) \hat{\gamma}_{CML1} \\ \text{External } \hat{\beta}_2 &\xrightarrow{\text{Chatterjee's}} \hat{\gamma}_{CML2} \xrightarrow{\text{Estes'}} \hat{\gamma}_{EB2} = \hat{W}_2 \hat{\gamma}_I + (I - \hat{W}_2) \hat{\gamma}_{CML2} \end{aligned}$$

Step 2:

$$\begin{bmatrix} \hat{\gamma}_{EB1} \\ \hat{\gamma}_{EB2} \end{bmatrix} \sim N(\gamma_0, V),$$

$\hat{\gamma}_0 =$  A weighted average of  $\hat{\gamma}_{EB1}$  and  $\hat{\gamma}_{EB2}$

## Proposed Estimates

Optimal Covariance  
Weighted Estimate

$$\text{weight 1} \begin{bmatrix} \hat{\gamma}_0^{EB1} \\ \hat{\gamma}_1^{EB1} \\ \vdots \\ \hat{\gamma}_p^{EB1} \end{bmatrix} + (1 - \text{weight 1}) \begin{bmatrix} \hat{\gamma}_0^{EB2} \\ \hat{\gamma}_1^{EB2} \\ \vdots \\ \hat{\gamma}_p^{EB2} \end{bmatrix}$$

Selective Coefficient Learner

$$\begin{bmatrix} \text{weight 1} \hat{\gamma}_0^{EB1} + (1 - \text{weight 1}) \hat{\gamma}_0^{EB2} \\ \text{weight 2} \hat{\gamma}_1^{EB1} + (1 - \text{weight 2}) \hat{\gamma}_1^{EB2} \\ \vdots \\ \text{weight } p+1 \hat{\gamma}_p^{EB1} + (1 - \text{weight } p+1) \hat{\gamma}_p^{EB2} \end{bmatrix}$$

## Proposed Framework

Step 1:

$$\text{External } \hat{\beta}_1 \xrightarrow{\text{Chatterjee's}} \hat{\gamma}_{CML1} \xrightarrow{\text{Estes'}} \hat{\gamma}_{EB1} = \hat{W}_1 \hat{\gamma}_I + (I - \hat{W}_1) \hat{\gamma}_{CML1}$$

$$\text{External } \hat{\beta}_2 \xrightarrow{\text{Chatterjee's}} \hat{\gamma}_{CML2} \xrightarrow{\text{Estes'}} \hat{\gamma}_{EB2} = \hat{W}_2 \hat{\gamma}_I + (I - \hat{W}_2) \hat{\gamma}_{CML2}$$

Step 2:

$$\begin{bmatrix} \hat{\gamma}_{EB1} \\ \hat{\gamma}_{EB2} \end{bmatrix} \sim N(\gamma_0, V),$$

$$\hat{\gamma}_0 = \text{A weighted average of } \hat{\gamma}_{EB1} \text{ and } \hat{\gamma}_{EB2}$$

## Proposed Estimates

Weighting across different  $\hat{\gamma}_{EB}$ 's: same weight for all coefficient within one  $\hat{\gamma}_{EB}$

## Optimal Covariance Weighted Estimate

$$\text{Weight 1} \begin{bmatrix} \hat{\gamma}_0^{EB1} \\ \hat{\gamma}_1^{EB1} \\ \vdots \\ \hat{\gamma}_p^{EB1} \end{bmatrix} + (1 - \text{weight 1}) \begin{bmatrix} \hat{\gamma}_0^{EB2} \\ \hat{\gamma}_1^{EB2} \\ \vdots \\ \hat{\gamma}_p^{EB2} \end{bmatrix}$$

## Selective Coefficient Learner

$$\begin{bmatrix} \text{weight 1} \hat{\gamma}_0^{EB1} + (1 - \text{weight 1}) \hat{\gamma}_0^{EB2} \\ \text{weight 2} \hat{\gamma}_1^{EB1} + (1 - \text{weight 2}) \hat{\gamma}_1^{EB2} \\ \vdots \\ \text{weight } p+1 \hat{\gamma}_p^{EB1} + (1 - \text{weight } p+1) \hat{\gamma}_p^{EB2} \end{bmatrix}$$

## Proposed Framework

Step 1:

$$\text{External } \hat{\beta}_1 \xrightarrow{\text{Chatterjee's}} \hat{\gamma}_{CML1} \xrightarrow{\text{Estes'}} \hat{\gamma}_{EB1} = \hat{W}_1 \hat{\gamma}_I + (I - \hat{W}_1) \hat{\gamma}_{CML1}$$

$$\text{External } \hat{\beta}_2 \xrightarrow{\text{Chatterjee's}} \hat{\gamma}_{CML2} \xrightarrow{\text{Estes'}} \hat{\gamma}_{EB2} = \hat{W}_2 \hat{\gamma}_I + (I - \hat{W}_2) \hat{\gamma}_{CML2}$$

Weighting across each coefficient: different weight for coefficients within one  $\hat{\gamma}_{EB}$

Step 2:

$$\begin{bmatrix} \hat{\gamma}_{EB1} \\ \hat{\gamma}_{EB2} \end{bmatrix} \sim N(\gamma_0, V),$$

$$\hat{\gamma}_0 = \text{A weighted average of } \hat{\gamma}_{EB1} \text{ and } \hat{\gamma}_{EB2}$$



## Proposed Estimates

### Optimal Covariance Weighted Estimate

$$\hat{\gamma}_0 = \text{Weight}_1 \hat{\gamma}_{EB^1} + \text{Weight}_2 \hat{\gamma}_{EB^2}$$

$$\text{Weight} = \arg \min_{\text{Weight}} \text{Var}(X^T \hat{\gamma}_0)$$

$$\text{s.t.} \begin{cases} \text{Weight}^T = 1 \\ \text{Weight} \in (0, 1) \end{cases}$$

### Selective Coefficient Learner

$$\hat{\gamma}_{X_j}^* = \sum_{k \in E_j} \frac{\hat{\text{Var}}(\hat{\gamma}_{X_j^k})}{\sum_{k \in E_j} \hat{\text{Var}}(\hat{\gamma}_{X_j^k})} \hat{\gamma}_{X_j^k}$$

$$\hat{\gamma}_{\text{SC-Learner}} = [\hat{\gamma}_{X_0}^*, \hat{\gamma}_{X_1}^*, \dots, \hat{\gamma}_{X_p}^*, \hat{\gamma}_B]^T$$

## Proposed Framework

Step 1:

$$\text{External } \hat{\beta}_1 \xrightarrow{\text{Chatterjee's}} \hat{\gamma}_{CML^1} \xrightarrow{\text{Estes'}} \hat{\gamma}_{EB^1} = \hat{W}_1 \hat{\gamma}_I + (I - \hat{W}_1) \hat{\gamma}_{CML^1}$$

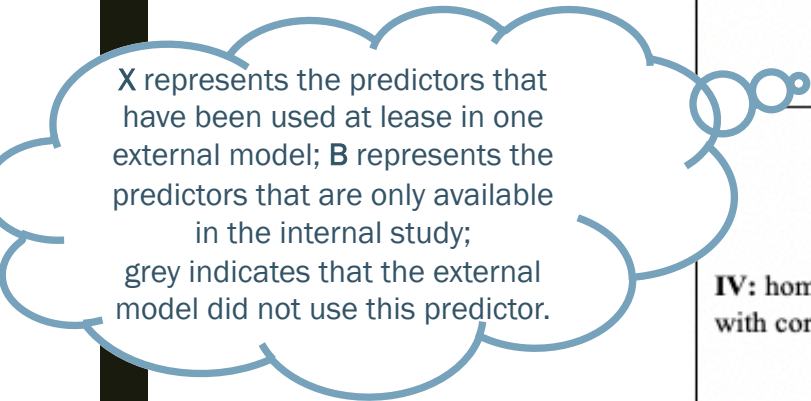
$$\text{External } \hat{\beta}_2 \xrightarrow{\text{Chatterjee's}} \hat{\gamma}_{CML^2} \xrightarrow{\text{Estes'}} \hat{\gamma}_{EB^2} = \hat{W}_2 \hat{\gamma}_I + (I - \hat{W}_2) \hat{\gamma}_{CML^2}$$

Step 2:

$$\begin{bmatrix} \hat{\gamma}_{EB^1} \\ \hat{\gamma}_{EB^2} \end{bmatrix} \sim N(\gamma_0, V),$$

$$\hat{\gamma}_0 = \text{A weighted average of } \hat{\gamma}_{EB^1} \text{ and } \hat{\gamma}_{EB^2}$$

# **Simulation/Real Data Analysis Results and Conclusion**



Simulation Setting Snapshot												
Overall performance: $\bar{V}(Xi\hat{\gamma})$ , SSE, Brier Score												
Covariate-wise performance: variance decrease for single coefficient												
I: homogeneous populations with correct external models		X1	X2	X3	X4	B						
	Internal											
	External 1											
	External 2											
	External 3											
II: homogeneous populations with uncertainties in external models		X1	X2	X3	X4	B						
	Internal											
	External 1 (small sample size)											
	External 2											
	External 3											
III: heterogenous X distribution	1		X1	X2	X3	X4	B					
		Internal										
		External 1										
		External 2 (different B X)										
		External 3										
	2		X1	X2	X3	X4	B					
		Internal										
		External 1										
		External 2										
		External 3 (different Y X, B)										
IV: homogeneous populations with correct external models	1		X1	X2	X3	X4	X5	X6	X7	X8	X9	B
		Internal										
		External 1										
		External 2										
		External 3										
	2		X1	X2	X3	X4	B1	B2	B3	B4	B5	
		Internal										
		External 1										
		External 2										
		External 3										

Abbreviation:  $\bar{V}(Xi\hat{\gamma}) = \frac{1}{N} \sum_{i=1}^N \hat{V}(Xi\hat{\gamma})$ , estimated mean variance of predicted log odds over observations in the validation dataset;

$SSE = \sum_{i=1}^N (p_{i0} - \hat{p}_i)^2$ , sum of squared error over observations in the validation dataset;

Scaled Brier score =  $\sum_{i=1}^N (Y_i - \hat{p}_i)^2 / \sum_{i=1}^N (Y_i - \bar{Y})^2$ , scaled Brier score over observations in the validation dataset

# Summary of simulation & prostate cancer data analysis

- Incorporating 3 correct external model information
  - Overall: ↓ ~25%
  - Covariate-wise: ↓ ~30% each
- Incorporating 2 correct & 1 incorrect external model information
  - Low-weight for the wrong external model information
  - Overall: ↓ 15-25% depending on dimension
  - Covariate-wise: ↓ 15-30% each
- Regarding both overall and covariate-wise performance, **Optimal Covariate Weighted Estimate (OCWE)** and **Selective Covariate Learner (SC-Learner)** are better than
  - Direct regression using internal data only
  - EB estimate solely (one-time shrinkage)
  - Inverse variance weighted estimate
- OCWE vs. SC-Learner
  - SC-Learner has better covariate-wise performance when the dimension difference of external calculators was large (i.e. certain external calculators used more X's than other calculators)

e.g.

[illegible]

# In Conclusion

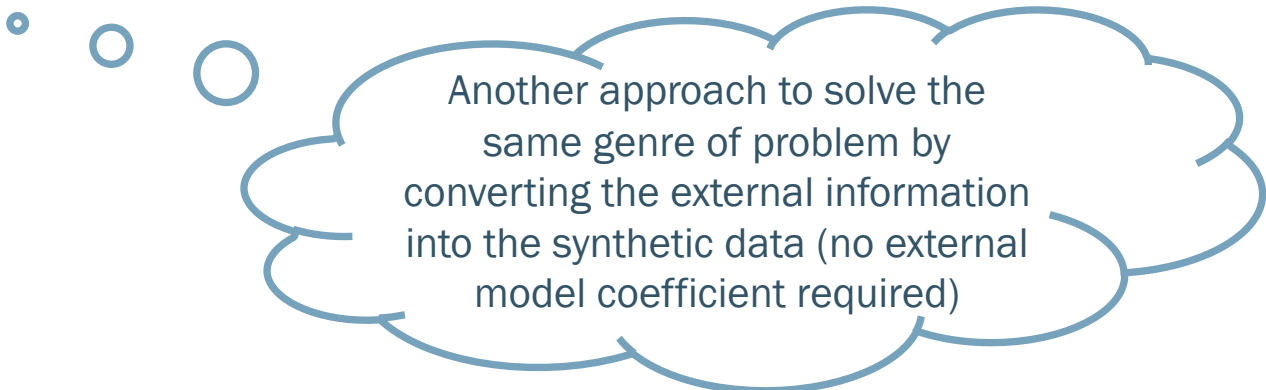
- The proposed framework is flexible and robust
  - It can incorporate external models that use a slightly different set of covariates
  - It can identify the most relevant external information and diminish the influence of information that is less compatible with the internal data
  - It nicely balances the bias-variance tradeoff while preserving the most efficiency gain
- The 2 proposed estimates--Optimal Covariate Weighted Estimate (OCWE) and Selective Covariate Learner (SC-Learner) are efficient
  - More efficient than the naïve analysis of the internal data and other naïve combinations of external estimators
  - More efficient than incorporating single external information

# Reference

Chatterjee, N., Chen, Y.-H., Maas, P. & Carroll, R. J. (2016). Constrained maximum likelihood estimation for model calibration using summary-level information from external big data sources. *Journal of the American Statistical Association*, 111(513), 107–117.

Estes, J.P. & Mukherjee, B. & Taylor, J.M.G. (2018). Empirical Bayes estimation and prediction using summary-Level information from external big data sources adjusting for violations of transportability. *Statistics in Biosciences*, 10(3), 568–586.

Gu, T. & Taylor, J.M.G. & Cheng, W. & Mukherjee, B. (2019). Synthetic data method to incorporate external information into a current study. *Canadian Journal of Statistics*, 47(4), 580-603.



Another approach to solve the same genre of problem by converting the external information into the synthetic data (no external model coefficient required)

**Thank you!**