Zero-inflated models: Poisson, Binomial, negative Binomial and BetaBinomial

Parametrisation

There is support two types of zero-inflated models, which we name type 0 and type 1. These are defined for both the Binomial, the Poisson, the negative Binomial and BetaBinomial likelihood. For simplicity we will describe only the Poisson as the other cases are similar.

Type 0

The (type 0) likelihood is defined as

$$\text{Prob}(y \mid \ldots) = p \times 1_{[y=0]} + (1-p) \times \text{Poisson}(y \mid y > 0)$$

where p is a hyperparameter where

$$p = \frac{\exp(\theta)}{1+\exp(\theta)}$$

and θ is the internal representation of p; meaning that the initial value and prior is given for θ. This is model is called `zeroinflatedpoisson0` (and `zeroinflatedbinomial0`).

Type 1

The (type 1) likelihood is defined as

$$\text{Prob}(y \mid \ldots) = p \times 1_{[y=0]} + (1-p) \times \text{Poisson}(y)$$

where p is a hyperparameter where

$$p = \frac{\exp(\theta)}{1+\exp(\theta)}$$

and θ is the internal representation of p; meaning that the initial value and prior is given for θ. This is model is called `zeroinflatedpoisson1` (and `zeroinflatedbinomial1`).

Link-function

As for the Poisson, the Binomial the negative Binomial and the BetaBinomial.

Hyperparameters

For Poisson and the Binomial, there is one hyperparameter; where

$$p = \frac{\exp(\theta)}{1+\exp(\theta)}$$

and the prior and initial value is is given for θ.

For the negative Binomial and BetaBinomial, there are two hyperparameters. The overdispersion parameter n for the negative Binomial is represented as

$$\theta_1 = \log(n)$$

and the prior is defined on θ_1. The zero-inflation parameter p, is represented as

$$p = \frac{\exp(\theta_2)}{1+\exp(\theta_2)}$$

and the prior and initial value is is given for θ_2. For the BetaBinomial it is similar.

Specification

- family = `zeroinflatedbinomial0`
- family = `zeroinflatedbinomial1`
- family = `zeroinflatednbinomial0`
- family = `zeroinflatednbinomial1`
- family = `zeroinflatedpoisson0`
- family = `zeroinflatedpoisson1`
- family = `zeroinflatedbetabinomial0`
- family = `zeroinflatedbetabinomial1`
- Required arguments: As for the Binomial, the negative Binomial, BetaBinomial and Poisson likelihood.

Hyperparameter spesification and default values

Zeroinflated Binomial Type 0

hyper

theta

hyperid 90001
name logit probability
short.name prob
initial -1
fixed FALSE
prior gaussian
param -1 0.2
to.theta `function(x) log(x/(1-x))`
from.theta `function(x) exp(x)/(1+exp(x))`

survival FALSE

discrete FALSE

link default logit cauchit probit cloglog loglog

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Zeroinflated Binomial Type 1

hyper

theta

hyperid 91001
name logit probability
short.name prob
initial -1

fixed FALSE
prior gaussian
param -1 0.2
to.theta `function(x) log(x/(1-x))`
from.theta `function(x) exp(x)/(1+exp(x))`

survival FALSE

discrete FALSE

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Zeroinflated NegBinomial Type 0

hyper

theta1

hyperid 95001
name log size
short.name size
initial 2.30258509299405
fixed FALSE
prior loggamma
param 1 1
to.theta `function(x) log(x)`
from.theta `function(x) exp(x)`

theta2

hyperid 95002
name logit probability
short.name prob
initial -1
fixed FALSE
prior gaussian
param -1 0.2
to.theta `function(x) log(x/(1-x))`
from.theta `function(x) exp(x)/(1+exp(x))`

survival FALSE

discrete FALSE

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Zeroinflated NegBinomial Type 1

hyper

theta1

hyperid 96001
name log size
short.name size
initial 2.30258509299405
fixed FALSE
prior loggamma
param 1 1
to.theta `function(x) log(x)`
from.theta `function(x) exp(x)`

theta2

hyperid 96002
name logit probability
short.name prob
initial -1
fixed FALSE
prior gaussian
param -1 0.2
to.theta `function(x) log(x/(1-x))`
from.theta `function(x) exp(x)/(1+exp(x))`

survival FALSE

discrete FALSE

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Zeroinflated BetaBinomial Type 0

hyper

theta1

hyperid 88001
name overdispersion
short.name rho
initial 0
fixed FALSE
prior gaussian
param 0 0.4
to.theta `function(x) log(x/(1-x))`
from.theta `function(x) exp(x)/(1+exp(x))`

theta2

hyperid 88002

name logit probability
short.name prob
initial -1
fixed FALSE
prior gaussian
param -1 0.2
to.theta `function(x) log(x/(1-x))`
from.theta `function(x) exp(x)/(1+exp(x))`

survival FALSE

discrete TRUE

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Zeroinflated BetaBinomial Type 1

hyper

theta1

hyperid 89001
name overdispersion
short.name rho
initial 0
fixed FALSE
prior gaussian
param 0 0.4
to.theta `function(x) log(x/(1-x))`
from.theta `function(x) exp(x)/(1+exp(x))`

theta2

hyperid 89002
name logit probability
short.name prob
initial -1
fixed FALSE
prior gaussian
param -1 0.2
to.theta `function(x) log(x/(1-x))`
from.theta `function(x) exp(x)/(1+exp(x))`

survival FALSE

discrete TRUE

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Zeroinflated Poisson Type 0

hyper

theta

hyperid 85001
name logit probability
short.name prob
initial -1
fixed FALSE
prior gaussian
param -1 0.2
to.theta `function(x) log(x/(1-x))`
from.theta `function(x) exp(x)/(1+exp(x))`

survival FALSE

discrete FALSE

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Zeroinflated Poisson Type 1

hyper

theta

hyperid 86001
name logit probability
short.name prob
initial -1
fixed FALSE
prior gaussian
param -1 0.2
to.theta `function(x) log(x/(1-x))`
from.theta `function(x) exp(x)/(1+exp(x))`

survival FALSE

discrete FALSE

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Example

In the following example we estimate the parameters in a simulated example for both type 0 and type 1.

Poisson

```
## type 0
n=100
a = 1
b = 1
z = rnorm(n)
eta = a + b*z
p = 0.2
E = sample(c(1,5,10,15), size=n, replace=TRUE)
lambda = E*exp(eta)

## first sample y|y>0
y = rpois(n, lambda = lambda)
is.zero = (y == 0)
while(sum(is.zero) > 0)
{
    y[is.zero] = rpois(sum(is.zero), lambda[is.zero])
    is.zero = (y == 0)
}
## then set some of these to zero
y[ rbinom(n, size=1, prob=p) == 1 ] = 0

data = list(y=y,z=z)
formula = y ~ 1+z
result0 = inla(formula, family = "zeroinflatedpoisson0", data = data, E=E)
summary(result0)

## type 1
y = rpois(n, lambda = lambda)
y[ rbinom(n, size=1, prob=p) == 1 ] = 0
data = list(y=y,z=z)
formula = y ~ 1+z
result1 = inla(formula, family = "zeroinflatedpoisson1", data = data, E=E)
summary(result1)
```

Binomial

```
## type 0
n=100
a = 1
b = 1
z = rnorm(n)
eta = a + b*z
p = 0.2
Ntrials = sample(c(1,5,10,15), size=n, replace=TRUE)
prob = exp(eta)/(1 + exp(eta))

y = rbinom(n, size = Ntrials, prob = prob)
is.zero = (y == 0)
while(sum(is.zero) > 0)
{
```

```
    y[is.zero] = rbinom(sum(is.zero), size = Ntrials[is.zero], prob = prob[is.zero])
    is.zero = (y == 0)
}
y[ rbinom(n, size=1, prob=p) == 1 ] = 0
data = list(y=y,z=z)
formula = y ~ 1+z
result0 = inla(formula, family = "zeroinflatedbinomial0", data = data, Ntrials = Ntrials)
summary(result0)

## type 1
y = rbinom(n, size = Ntrials, prob = prob)
y[ rbinom(n, size=1, prob=p) == 1 ] = 0
data = list(y=y,z=z)
formula = y ~ 1+z
result1 = inla(formula, family = "zeroinflatedbinomial1", data = data, Ntrials=Ntrials)
summary(result1)
```

Advanced example

In the following example we estimate the parameters in a simulated example for a type0 likelihood, where one linear predictor enters the zero-probability and one other linear predictor enters the non-zero Poisson for example. The same trick can be used for other models of type0. The trick is that the likelihood

$$p^* 1_{[y=0]} + (1 - p^*)P(y|y > 0)$$

can be reformulated as a Bernoulli likelihood for the "class"-variable

$$z = \begin{cases} 1, & \text{if } y = 0 \\ 0, & \text{if } y > 0. \end{cases}$$

where p^* is the probability for success, and zero-inflated type0 likelihood (with fixed $p = 0$) for those $y > 0$. Since p^* and the linear predictor in P is separated into two likelihoods, we can apply one linear predictor to each one, hence extend the basic model to cases where p^* also depends on a linear predictor. Here is a small simulated example doing this.

```
require(INLA)

n = 100
a = 0.5
b = 1.5
x1 = rnorm(n, sd = 0.5)

eta.z = -a - b*x1
z = rbinom(n, 1, inla.link.logit(eta.z, inverse=TRUE))
n.y = sum(z)

x2 = rnorm(n.y, sd = 0.5)
eta.y = a + b*x2
lambda = exp(eta.y)
y = rpois(n.y, lambda)

is.zero = (y == 0)
while(sum(is.zero) > 0)
{
   y[is.zero] = rpois(sum(is.zero), lambda[is.zero])
   is.zero = (y == 0)
}

Y = matrix(NA, n + n.y, 2)
Y[1:n, 1] = z
Y[n + 1:n.y, 2] = y

form = Y ~ 0 + mu.z + mu.y + cov.z + cov.y
ldat = list(
        Y=Y,
        mu.z=rep(1:0, c(n, n.y)),
        mu.y=rep(0:1, c(n, n.y)),
        cov.z=c(x1, rep(NA,n.y)),
        cov.y=c(rep(NA, n), x2))
```

```
res <- inla(form, data=ldat,
            family=c('binomial', 'zeroinflatedpoisson0'),
            control.family=list(
                    list(),
                    list(hyper = list(
                                 prob = list(
                                        initial = -20,
                                        fixed = TRUE)))))
round(res$summary.fix, 4)
```

Notes

None.

Extentions

There are some extentions available which currently is only implemented for the cases where its needed/requested.

Type 2 Is like Type 1 but where (for the Poisson)

$$p = 1 - \left(\frac{E\exp(x)}{1 + E\exp(x)}\right)^{\alpha}$$

where $\alpha > 0$ is the hyperparameter instead of p (and $E\exp(x)$ is the mean). Available for Poisson as `zeroinflatedpoisson2`, for binomial as `zeroinflatedbinomial2` and for the negative binomial as `zeroinflatednbinomial2`.

The internal representation is $\theta = \log(\alpha)$ and prior is defined on $\log(\alpha)$.

Zeroinflated Poisson Type 2

hyper

theta

- **hyperid** 87001
- **name** log alpha
- **short.name** a
- **initial** 0.693147180559945
- **fixed** FALSE
- **prior** gaussian
- **param** 0.693147180559945 1
- **to.theta** `function(x) log(x)`
- **from.theta** `function(x) exp(x)`

survival FALSE

discrete FALSE

link default log

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Zeroinflated Binomial Type 2

hyper

- **theta**
  - **hyperid** 92001
  - **name** alpha
  - **short.name** alpha
  - **initial** -1
  - **fixed** FALSE
  - **prior** gaussian
  - **param** -1 0.2
  - **to.theta** `function(x) log(x)`
  - **from.theta** `function(x) exp(x)`

survival FALSE

discrete FALSE

link default logit cauchit probit cloglog loglog

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Zeroinflated Negative Binomial Type 2

hyper

- **theta1**
  - **hyperid** 99001
  - **name** log size
  - **short.name** size
  - **initial** 2.30258509299405
  - **fixed** FALSE
  - **prior** loggamma
  - **param** 1 1
  - **to.theta** `function(x) log(x)`
  - **from.theta** `function(x) exp(x)`
- **theta2**
  - **hyperid** 99002
  - **name** log alpha
  - **short.name** a
  - **initial** 0.693147180559945
  - **fixed** FALSE
  - **prior** gaussian
  - **param** 2 1
  - **to.theta** `function(x) log(x)`
  - **from.theta** `function(x) exp(x)`

survival FALSE

discrete FALSE

link default log

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0.0.1 Zero and N-inflated Binomial likelihood: type 3

This is the case where

$$\begin{aligned}\text{Prob}(y|\ldots) &= p_0 \times 1_{[y=0]} + \\ &\quad p_N \times 1_{[y=N]} + \\ &\quad (1-p_0-p_N) \times \text{binomial}(y, N, p)\end{aligned}$$

where:

$$p = \frac{\exp(\eta)}{1+\exp(\eta)} \qquad p_0 = \frac{p^{\alpha_0}}{1+p^{\alpha_0}+(1-p)^{\alpha_N}} \qquad p_N = \frac{(1-p)^{\alpha_N}}{1+p^{\alpha_0}+(1-p)^{\alpha_N}}$$

There are 2 hyperparameters, α_0 and α_N, governing zero-inflation where: The zero-inflation parameters α_0 and α_N are represented as $\theta_0 = \log(\alpha_0)$; $\theta_N = \log(\alpha_N)$ and the prior and initial value is given for θ_0 and θ_N respectively.

Here is an example

```
nsim<-10000
x<-rnorm(nsim)
alpha0<-1.5
alphaN<-2.0
p = exp(x)/(1+exp(x))
p0 = p^alpha0 / (1 + p^alpha0 + (1-p)^alphaN)
pN = (1-p)^alphaN / (1 + p^alpha0 + (1-p)^alphaN)
P<-cbind(p0, pN, (1-p0 -pN))
N<-rpois(nsim,20)
y<-rep(0,nsim)
for(i in 1:nsim)
    y[i]<-sum(rmultinom(1,size = 1,P[i,])*c(0,N[i],rbinom(1,N[i],p[i])))
formula = y ~1 + x
r = inla(formula, family = "zeroninflatedbinomial3",  Ntrials = N, verbose = TRUE,
           data = data.frame(y, x))
```

and the default settings

hyper

theta1

hyperid 93101
name alpha0
short.name alpha0
initial 2
fixed FALSE
prior gaussian
param 4 1
to.theta `function(x) log(x)`
from.theta `function(x) exp(x)`

theta2

hyperid 93102
name alphaN
short.name alphaN

initial 2
fixed FALSE
prior gaussian
param 4 1
to.theta `function(x) log(x)`
from.theta `function(x) exp(x)`

survival FALSE

discrete FALSE

link default logit cauchit probit cloglog loglog

pdf zeroinflated