90% Confidence Level Upper Bound

Brief discussion of Feldman & Cousins

Outline

Parameter Estimataion	3
Confidence Interval (CI)	5
Standard CIs	12
Feldman & Cousins CI	15
F&C CI construction	22
Bayesian	29

Outline

Outline 1
Parameter Estimataion 3
Poisson Distribution 4
Confidence Interval (CI) 5
CI Definition 6
CI Construction: Confidence Belt 8
Standard CIs
Acceptance region? 13
Example CIs14
Feldman & Cousins CI
Question: which one to use? 16
Problem: "flip-flop" 17
Standard Poisson (90% Upper) CI 18
Comparison: look at small n
Standard Poisson (90% central) CI
Comparison: look at large n21
F&C CI construction
Maximum Likelihood
Likelihood Ratio25

Example acceptance region for $\mu=0.5$	27
Bayesian	29
Bayes Theorem	30
Summary: Bayesian Credible Interval	32

Parameter Estimataion

• Expect an average of (real) $\mu \ge 0$ neutrinos per time

Poisson Distribution

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$$\Pr(n|\mu) = \frac{e^{-\mu}\mu^n}{n!}$$

• Prob. of data n, given parameter μ ; aka the *likelihood*.

Poisson Distribution

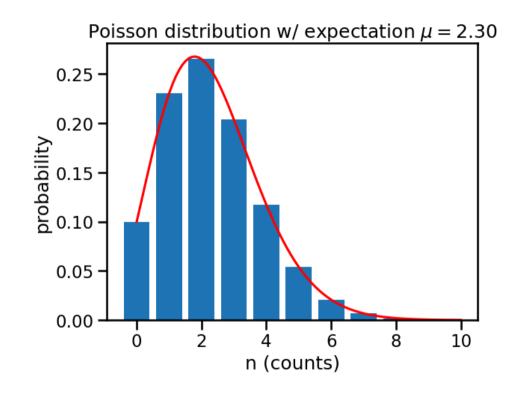
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• Example:



Confidence Interval (CI)

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- Repeat experiment; get outcome $x_1 \to \text{construct} [\mu_l(x_1), \mu_u(x_1)]$
- More experiments; get a bunch of intervals. *i.e.* we get a set

$$C \equiv \{ [\mu_l(x_0), \mu_u(x_0)], [\mu_l(x_1), \mu_u(x_1)], [\mu_l(x_2), \mu_u(x_2)] \dots \}$$

- $C \equiv \{ [\mu_l, \mu_u], [\mu_l, \mu_u], [\mu_l, \mu_u] \dots \}$
- The set C has the property that

$$P([\mu_l, \mu_u] \ni \mu_t) = \alpha\%$$

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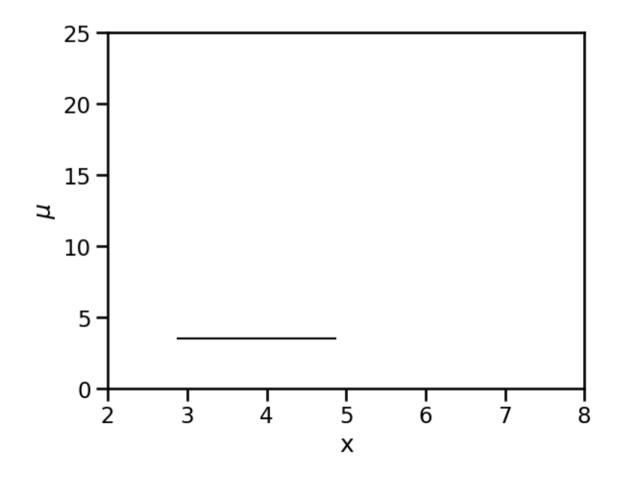
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- ullet The members of C are called $confidence\ intervals.$

Confidence Interval (CI)

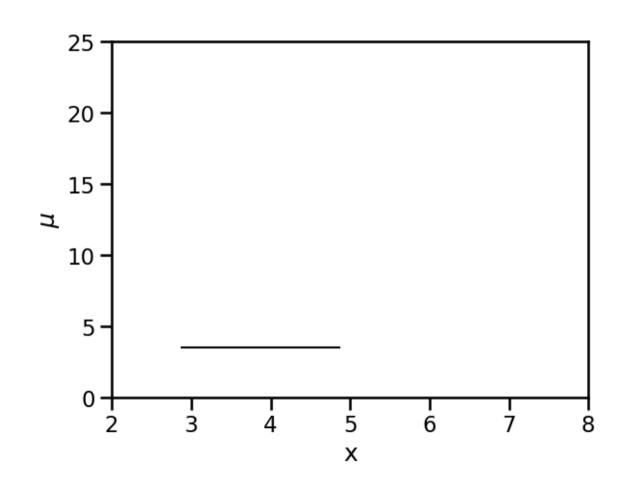
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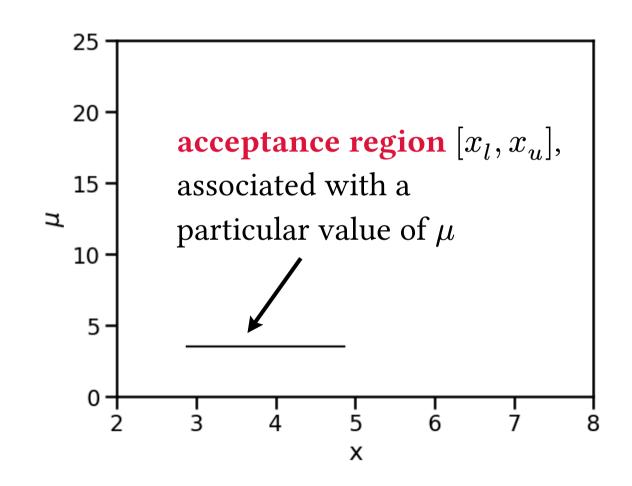
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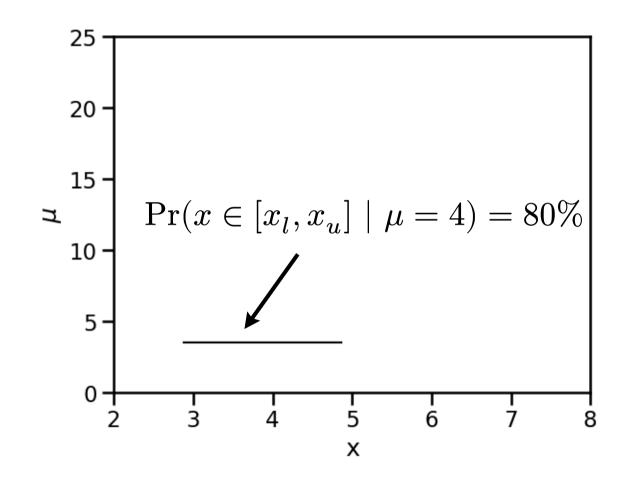
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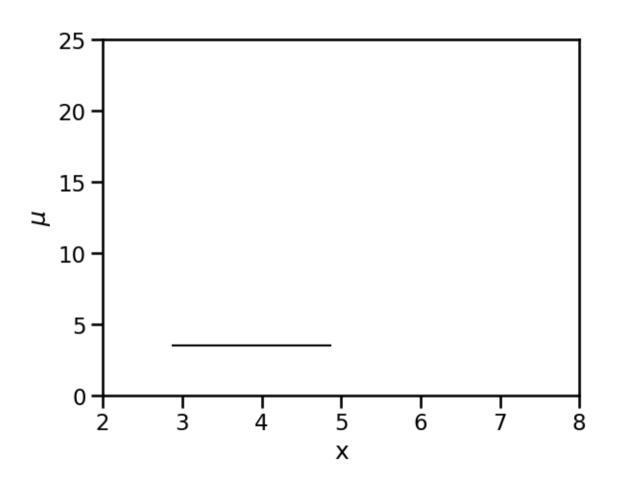


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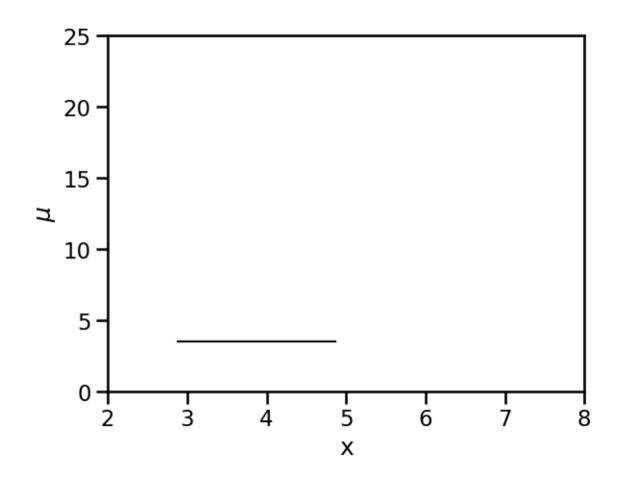


Confidence Interval (CI)

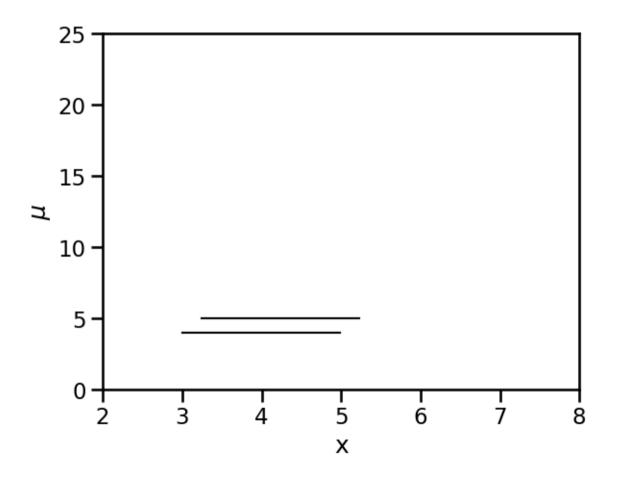
• But of course μ_t is unknown.



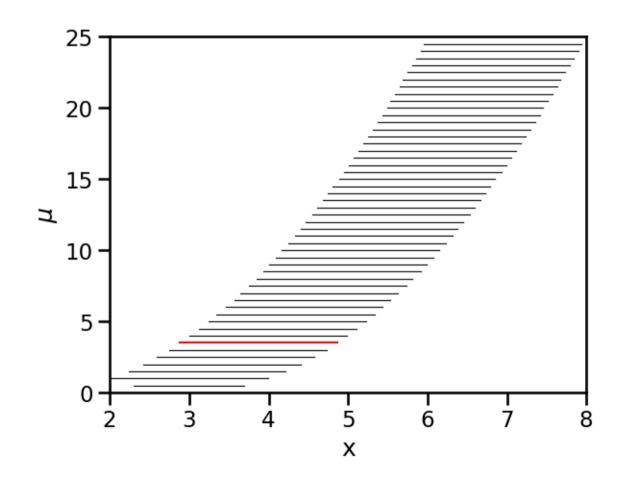
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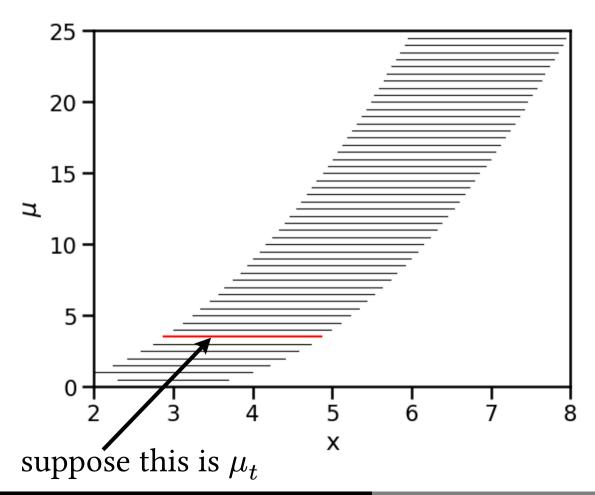
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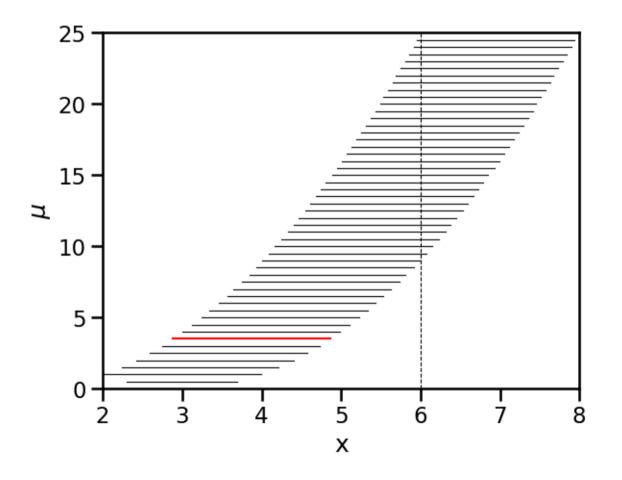


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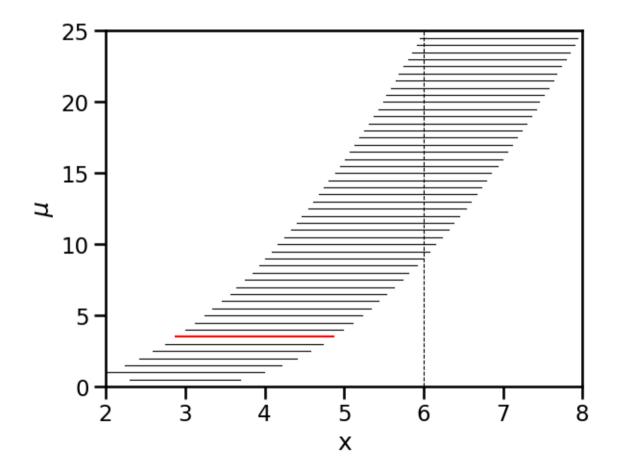


Confidence Interval (CI)

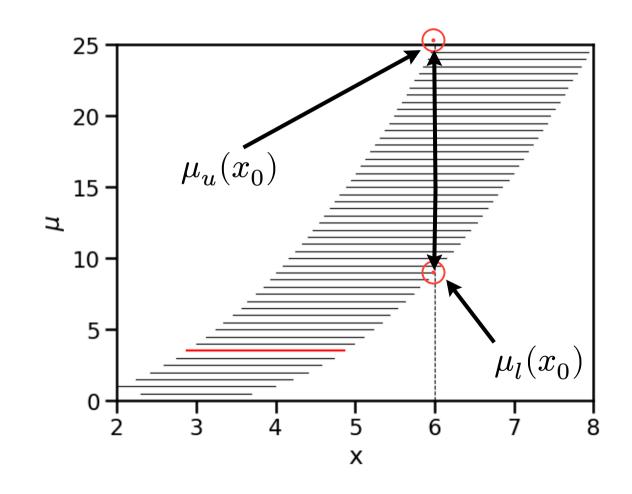
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- The probability of x_0 falling in the acceptance region (red) is 80%, by construction

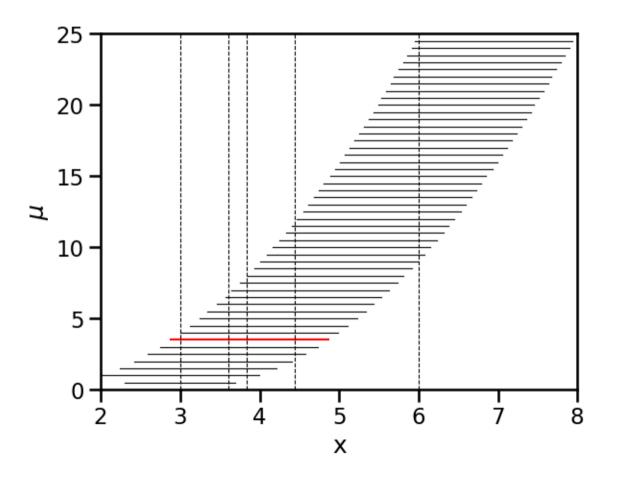


- Make a measurement, get result $x_0 = 6$
- The probability of x_0 falling in the acceptance region (red) is 80%, by construction
- The confidence interval $[\mu_l, \mu_u]$ from this experiment is the vertical intercept.



Confidence Interval (CI)

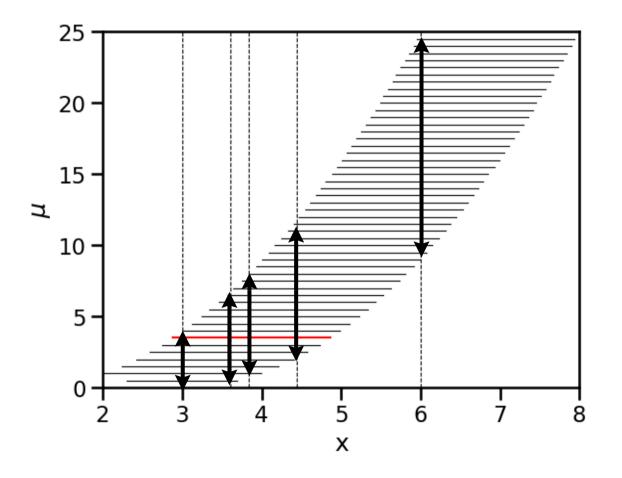
• Make some more measurements



CI Construction: Confidence Belt

Confidence Interval (CI)

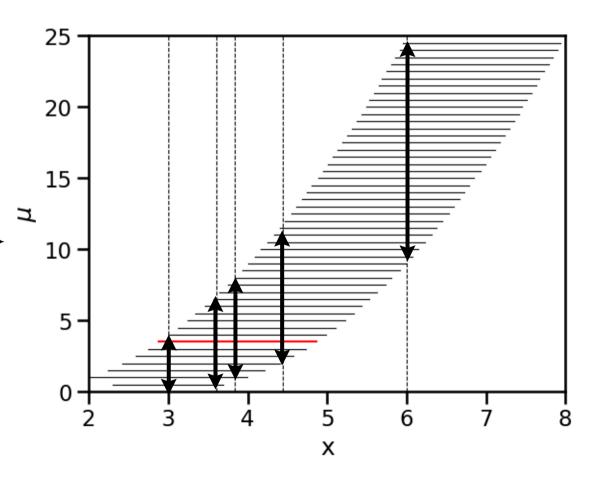
- Make some more measurements
- Get some more confidence intervals.



- Make some more measurements
- Get some more confidence intervals.
- Have a set

$$C = \{\operatorname{CI}_1, \operatorname{CI}_2, \operatorname{CI}_3, \operatorname{CI}_4, \operatorname{CI}_5\}$$

• 80% of this set would cover the true value, μ_t .



Standard CIs

Acceptance region?

• Recall acceptance region:

$$\Pr(n \in [n_1, n_2] \mid \mu_{\text{fixed}}) = 90\%$$

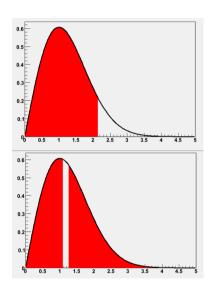
¹See Feldman & Cousins Section II.B

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• <u>Complete freedom</u> in choosing a 90% range.

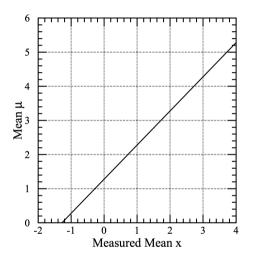


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- *A* choice¹ leads to the **upper limit**.



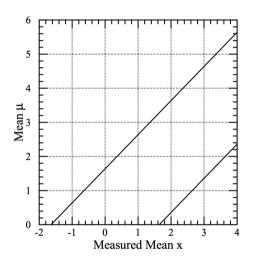
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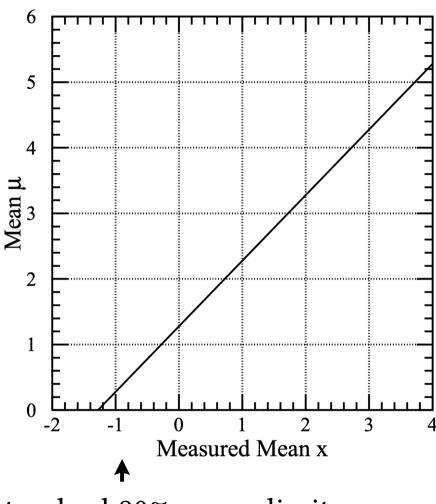
$$\Pr(n \in [n_1, n_2] \mid \mu_{\text{fixed}}) = 90\%$$

- Complete freedom in choosing a 90% range.
- *A* choice¹ leads to the **upper limit**.
- Another common choice leads to the two-sided central interval.



¹See Feldman & Cousins Section II.B

Example CIs

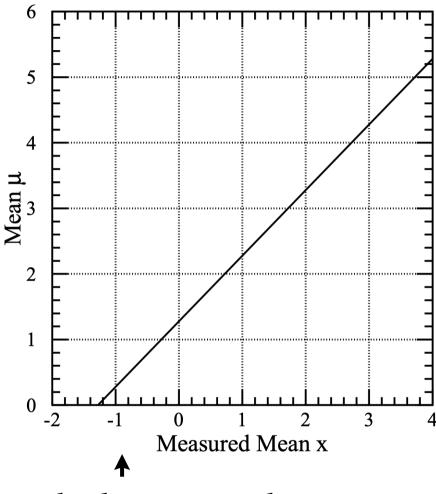


standard 90% upper limit

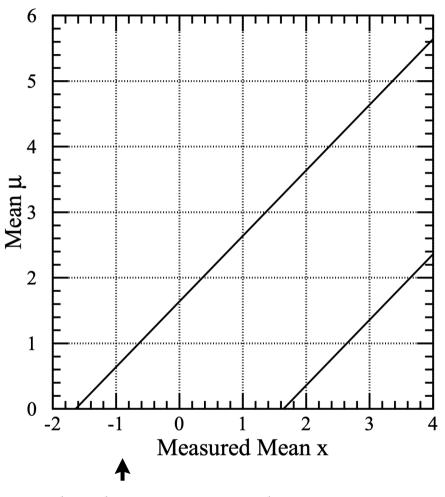
Gaussian likelihood with known $\sigma = 1$

$$\Pr(x \mid \mu) = \frac{1}{\sqrt{2\pi}} \cdot e^{-(x-\mu)^2/2}$$

Example CIs

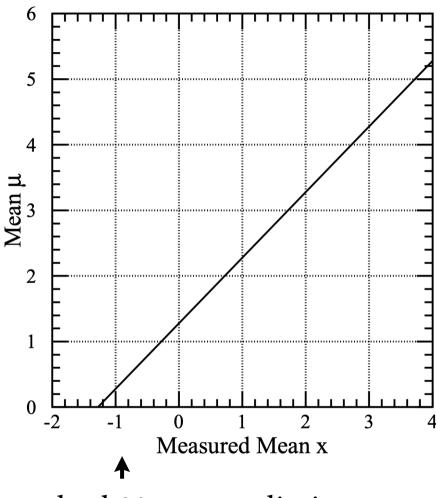


standard 90% upper limit

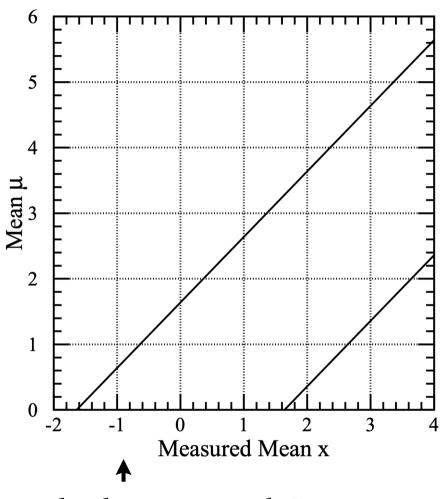


standard 90% central CI

Feldman & Cousins CI



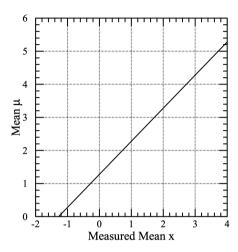
standard 90% upper limit

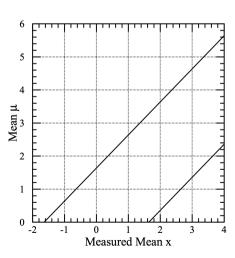


standard 90% central CI

Feldman & Cousins CI

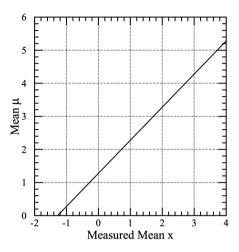
Problem: "flip-flop"

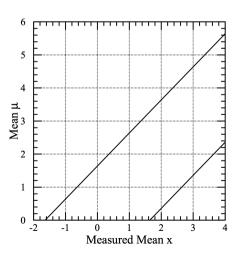




• One **cannot** pick the type of interval afterthe-fact. The Choice is made before we perform the experiment. Otherwise we would be "flip-flopping".

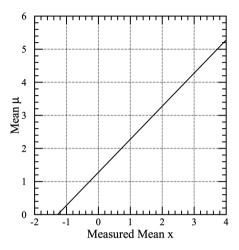
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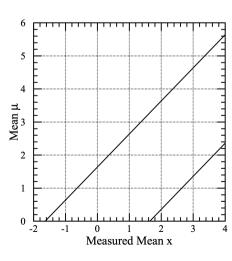




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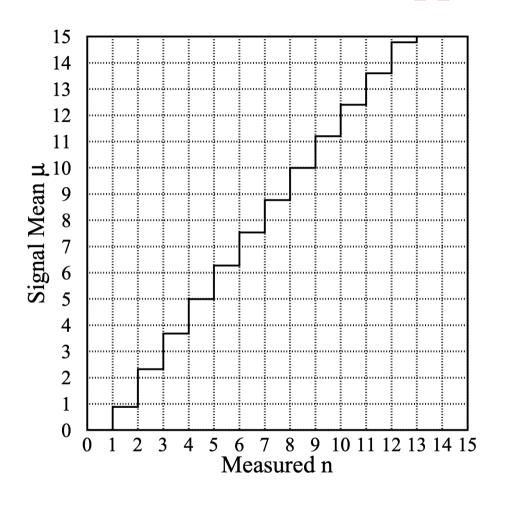




- One **cannot** pick the type of interval afterthe-fact. The Choice is made before we perform the experiment. Otherwise we would be "flip-flopping".
- It turns out that flip-flopping leads to invalid intervals
- Feldman & Cousins's approach removes the possibility of (or motivation to commit) flip-flopping.

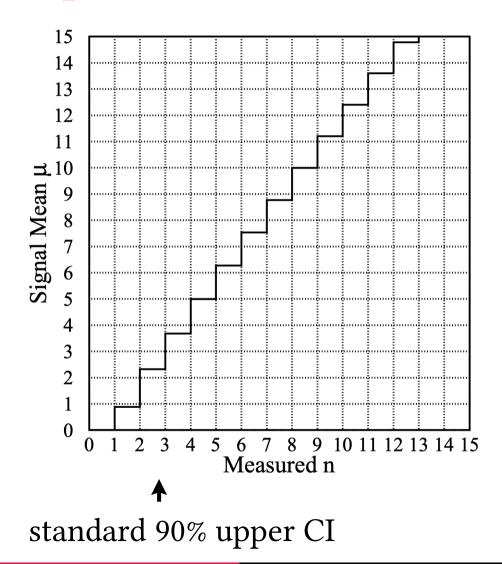
Standard Poisson (90% Upper) CI

Feldman & Cousins CI



Poisson likelihood with background

$$\mathcal{L} \equiv \Pr(n \mid \mu) = \frac{(\mu + b)^n e^{-(\mu + b)}}{n!}$$

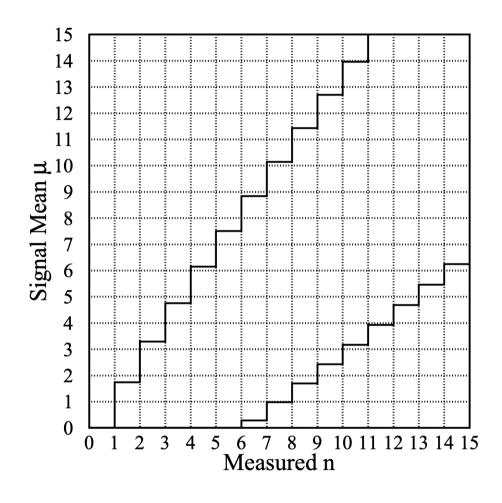


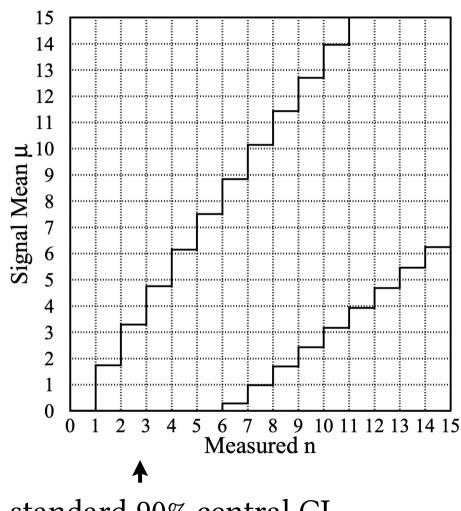
110 Signal Mean 10 11 12 13 14 15 Measured n

Feldman & Cousins

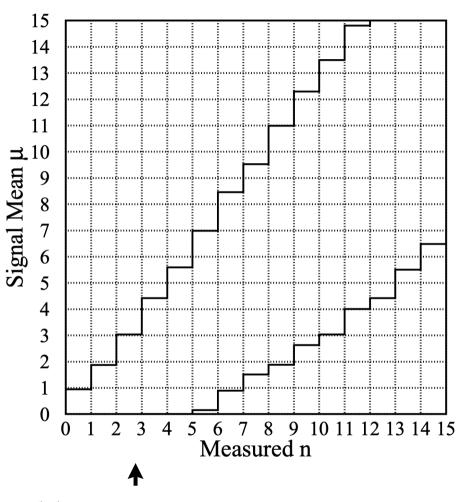
Standard Poisson (90% central) CI

Feldman & Cousins CI

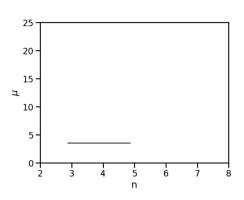




standard 90% central CI



Feldman & Cousins



• Recall acceptance region:

$$\Pr(n \in [n_1, n_2] \mid \mu_{\text{fixed}}) = 90\%$$

- <u>Complete freedom</u> in choosing how to construct the acceptance regions.
- Consider likelihood: Poisson with background *b*:

$$\mathcal{L} \equiv \Pr(n \mid \mu) = \frac{(\mu + b)^n e^{-(\mu + b)}}{n!}$$

- F&C propose to compute a likelihood ratio R
 - This needs a "best fit" $\mu_{\text{best}} \equiv \max(0, n b)$

Derivation

• Likelihood is a Poisson in this case.

$$\mathcal{L} \equiv \Pr(n \mid \mu) = \frac{(\mu + b)^n e^{-(\mu + b)}}{n!}$$

• Find maximum (fixing n, vary μ):

$$\frac{\mathrm{d}\mathcal{L}}{\mathrm{d}\mu}\bigg|_{\mu=\mu_{\mathrm{best}}} = 0$$

- Result: "best fit" $\mu = \mu_{\text{best}} = n b$
- Require physical $\mu \ge 0 \Rightarrow \mu_{\text{best}} = \max(0, n-b)$

- Do this for representative values of μ ; say we start with $\mu=0.5$
 - As an example background, b = 3
 - $ightharpoonup \Rightarrow \mu_{\mathrm{best}} \equiv \max(0, n-b) = \max(0, n-3)$
- Procedure:

 \blacktriangleright

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>

n	$\Pr(n \mu)$	$\mu_{ m best}$	$\Pr(n \mu_{ ext{best}})$	R	rank
0					

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- Procedure:
 - For n = 0, compute $Pr(n \mid \mu = 0.5)$

▶

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n	$\Pr(n \mu)$	$\mu_{ m best}$	$\Pr(n \mu_{ ext{best}})$	R	rank
0	0.03				

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 - For n=0, compute $\mu_{\text{best}}=0$

▶

\boldsymbol{n}	$\Pr(n \mu)$	$\mu_{ m best}$	$\Pr(n \mu_{ ext{best}})$	R	rank
0	0.03	0			

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 - For n=0, compute $\mu_{\text{best}}=0$
 - For n = 0, compute $Pr(n \mid \mu = \mu_{best})$

•

$oldsymbol{n}$	$\Pr(n \mu)$	$\mu_{ m best}$	$\Pr(n \mu_{ ext{best}})$	R	rank
0	0.03	0	0.05		

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- Procedure:
 - For n = 0, compute $Pr(n \mid \mu = 0.5)$
 - For n=0, compute $\mu_{\text{best}}=0$
 - For n = 0, compute $Pr(n \mid \mu = \mu_{best})$
 - ightharpoonup Divide likelihoods to get R.

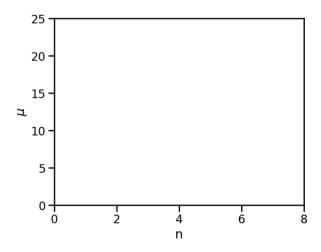
\boldsymbol{n}	$\Pr(n \mu)$	$\mu_{ m best}$	$\Pr(n \mu_{ ext{best}})$	R	rank
0	0.03	0	0.05	0.607	

F&C CI construction

• As a reminder, this is still just for $\mu = 0.5$ (and example b = 3)

n	$\Pr(n \mu)$	$\mu_{ m best}$	$\Pr(n \mu_{ ext{best}})$	R	rank
0	0.03	0	0.05	0.607	

• We will see that the acceptance region for $\mu = 0.5$ is this:

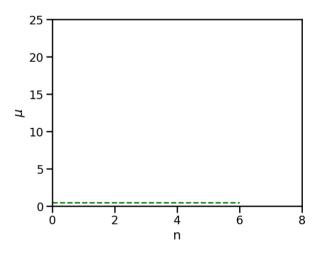


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n	$\Pr(n \mu)$	$\mu_{ m best}$	$\Pr(n \mu_{ ext{best}})$	R	rank
0	0.03	0	0.05	0.607	

• To construct the region, make a new row for n=1

n	$\Pr(n \mu)$	$\mu_{ m best}$	$\Pr(n \mu_{ ext{best}})$	R	rank
0	0.030	0	0.050	0.607	
1	0.106	0	0.149	0.708	

F&C CI construction

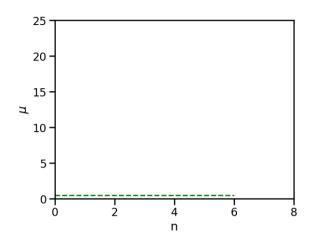
And then for a bunch of other n.

\boldsymbol{n}	$\Pr(n \mu)$	$\mu_{ m best}$	$\Pr(n \mu_{ ext{best}})$	R	rank
0	0.030	0	0.050	0.607	
1	0.106	0	0.149	0.708	
2	0.185	0	0.224	0.826	
3	0.216	0	0.224	0.963	
4	0.189	1	0.195	0.966	
5	0.132	2	0.175	0.753	
6	0.077	3	0.161	0.480	
7	0.039	4	0.149	0.259	

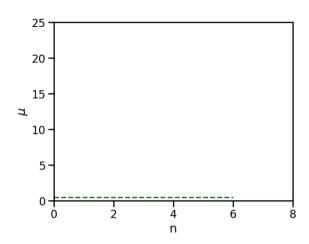
\boldsymbol{n}	$\Pr(n \mu)$	$\mu_{ m best}$	$\Pr(n \mu_{ ext{best}})$	R	rank
0	0.030	0	0.050	0.607	6
1	0.106	0	0.149	0.708	5
2	0.185	0	0.224	0.826	3
3	0.216	0	0.224	0.963	2
4	0.189	1	0.195	0.966	1
5	0.132	2	0.175	0.753	4
6	0.077	3	0.161	0.480	7
7	0.039	4	0.149	0.259	

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- Next, construct the acceptance region for other μ as well.



Bayesian

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Bayesian

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Bayesian

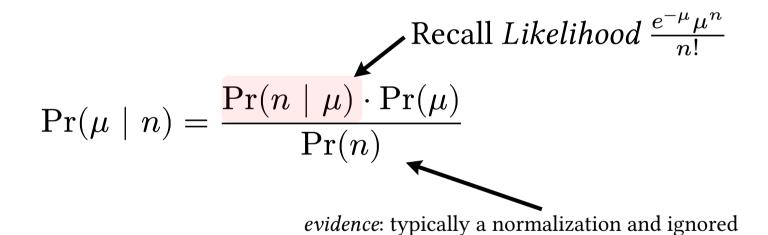
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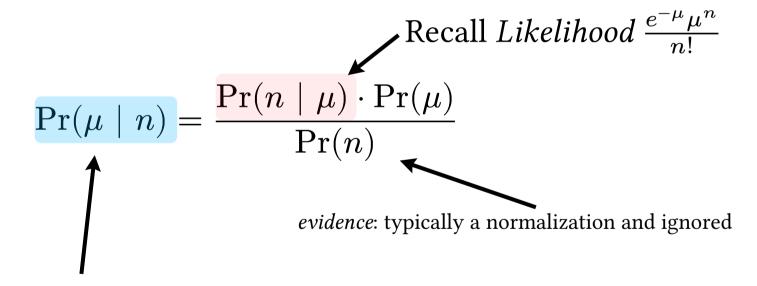
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Bayes Theorem



Posterior, encodes possible values of μ_t as a probability dist.

The "truth" μ_t is **not** fixed from the Bayesian perspective

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- Suppose we measure n=0 event, then the posterior is

$$\Pr(\mu \mid n = 0) = \Pr(n = 0 | \mu) = \frac{e^{-\mu} \cdot \mu^n}{n!} = e^{-\mu}$$

Bayesian

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• so we "estemate with 90% confidence that $\mu \leq 2.3$ " base on a non-detection.

