

COMS30035, Machine learning: Principal components analysis (PCA)

James Cussens

`james.cussens@bristol.ac.uk`

Department of Computer Science, SCEEM
University of Bristol

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Agenda

- ▶ PCA (standard presentation)

Dimensionality reduction

- Sometimes it is obvious we can throw away a dimension (i.e. a variable).

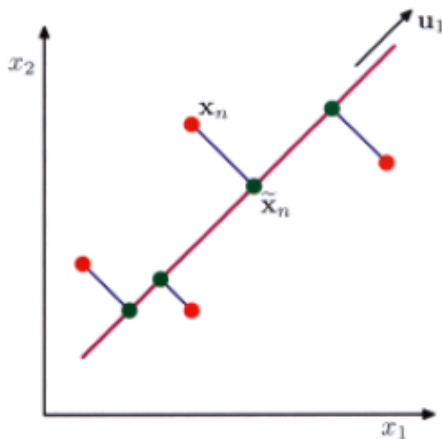
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[4.9, 3. , 1.4, 0.2, 1],  
[4.7, 3.2, 1.3, 0.2, 1],  
[4.6, 3.1, 1.5, 0.3, 1],  
[5. , 3.6, 1.4, 0.2, 1],  
. . . .
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- The idea with PCA is to rotate the data (i.e. choose a different co-ordinate system) so that we end up with dimensions with low variance ...
- ... which we can throw away without losing much information.

Motivations for PCA

- ▶ We can either view PCA as looking for projections with maximum variance [Bis06, §12.1.1],
- ▶ or looking for projections which minimise the distance from the original points to their projections [Bis06, §12.1.2].
- ▶ These are equivalent (we get the same projections)
- ▶ I will present the derivation in terms of maximising variance.

PCA in a picture (Bishop Fig 12.2)



From D dimensions to 1

- ▶ A projection from D dimensions down to 1 is defined by a D dimensional vector \mathbf{u}_1 (which we can choose to be a unit vector so $\mathbf{u}_1^T \mathbf{u}_1 = 1$).
- ▶ The projection of \mathbf{x} is simply $\mathbf{u}_1^T \mathbf{x}$.
- ▶ So which projection (which \mathbf{u}_1) is 'best'?

Eigenvector projections

Given a bunch of N data points \mathbf{x}_n , the sample covariance matrix is:

$$\mathbf{S} = \frac{1}{N} \sum_{n=1}^N (\mathbf{x}_n - \bar{\mathbf{x}})(\mathbf{x}_n - \bar{\mathbf{x}})^T$$

- ▶ The variance of the *projected data* is $\mathbf{u}_1^T \mathbf{S} \mathbf{u}_1$.
- ▶ By the usual method of differentiating (w.r.t. to \mathbf{u}_1) and setting to 0 we [Bis06, p. 562] find that

$$\mathbf{S} \mathbf{u}_1 = \lambda_1 \mathbf{u}_1 \tag{1}$$

- ▶ So \mathbf{u}_1 is an eigenvector of \mathbf{S} (with eigenvalue λ_1).
- ▶ Since $\mathbf{u}_1^T \mathbf{S} \mathbf{u}_1 = \lambda_1$, we maximise variance by setting \mathbf{u}_1 to be the eigenvector with the biggest eigenvalue.
- ▶ This eigenvector is called *the first principal component*.

And so on

- ▶ The second principal component is that direction which maximises projected variance **subject to being orthogonal to the first principal component**.
- ▶ Each subsequent principal component is chosen to maximise variance subject to being orthogonal to all previous principal components.
- ▶ It can be shown that the principal components are the eigenvectors of the covariance matrix ordered by eigenvalue.

New co-ordinates

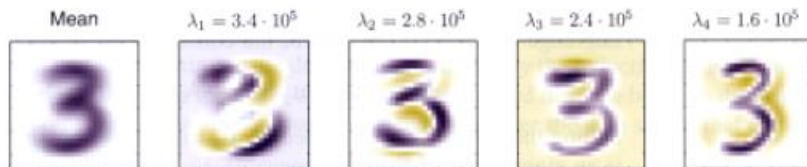
We have

$$\mathbf{x}_n = \sum_{i=1}^D (\mathbf{x}_n^T \mathbf{u}_i) \mathbf{u}_i = \sum_{i=1}^D \alpha_{ni} \mathbf{u}_i \quad (2)$$

- ▶ So each datapoint is a linear combination of principal components (= eigenvectors),
- ▶ but we (typically) only keep $M < D$ of these dimensions.
- ▶ When approximating a D -dimensional datapoint \mathbf{x}_n by an M -dimensional vector $\tilde{\mathbf{x}}_n$ the best PCA approximation accounts for the mean $\bar{\mathbf{x}}$ by adding a constant vector $\bar{\mathbf{x}} - \sum_{i=1}^M (\bar{\mathbf{x}}^T \mathbf{u}_i) \mathbf{u}_i$:

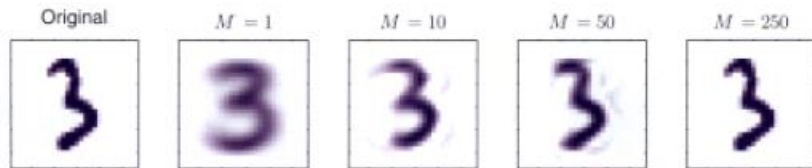
$$\begin{aligned} \tilde{\mathbf{x}}_n &= \bar{\mathbf{x}} + \sum_{i=1}^M (\mathbf{x}_n^T \mathbf{u}_i - \bar{\mathbf{x}}^T \mathbf{u}_i) \mathbf{u}_i \\ &= \sum_{i=1}^M (\mathbf{x}_n^T \mathbf{u}_i) \mathbf{u}_i + \bar{\mathbf{x}} - \sum_{i=1}^M (\bar{\mathbf{x}}^T \mathbf{u}_i) \mathbf{u}_i \end{aligned}$$

Seeing the eigenvectors (Bishop Fig 12.3)



The mean vector \bar{x} along with the first four PCA eigenvectors u_1, \dots, u_4 for the off-line digits data set, together with the corresponding eigenvalues.

Seeing PCA reconstructions (Bishop Fig 12.5)



An original example from the off-line digits data set together with its PCA reconstructions obtained by retaining M principal components for various values of M . As M increases the reconstruction becomes more accurate and would become perfect when $M = D = 28 \times 28 = 784$.

Now do the quiz!

Yes, please do the quiz for this lecture on Blackboard!



Christopher M. Bishop.

Pattern Recognition and Machine Learning.

Springer, 2006.