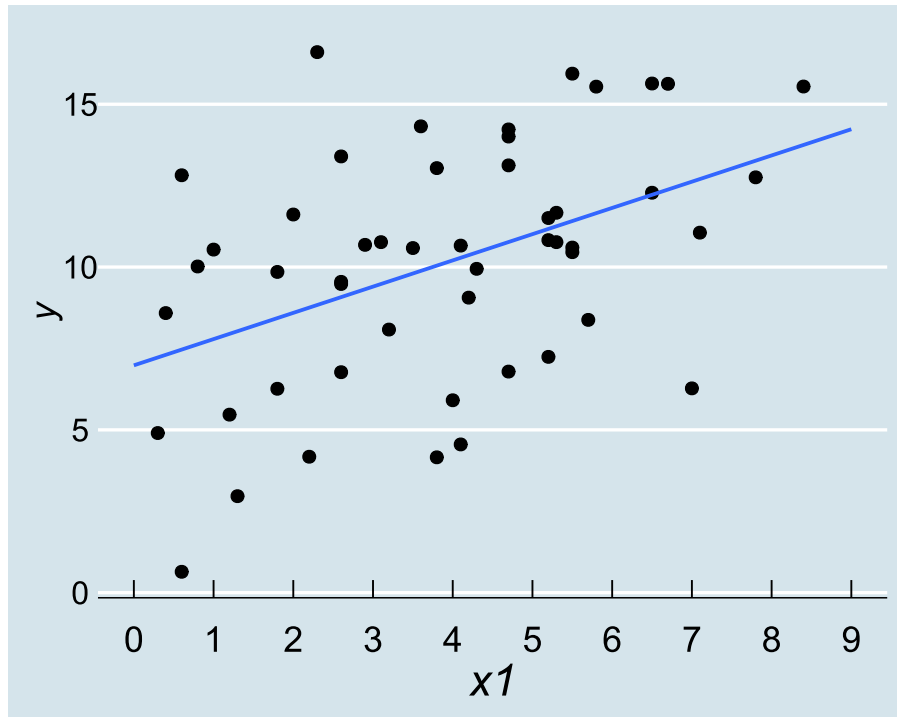
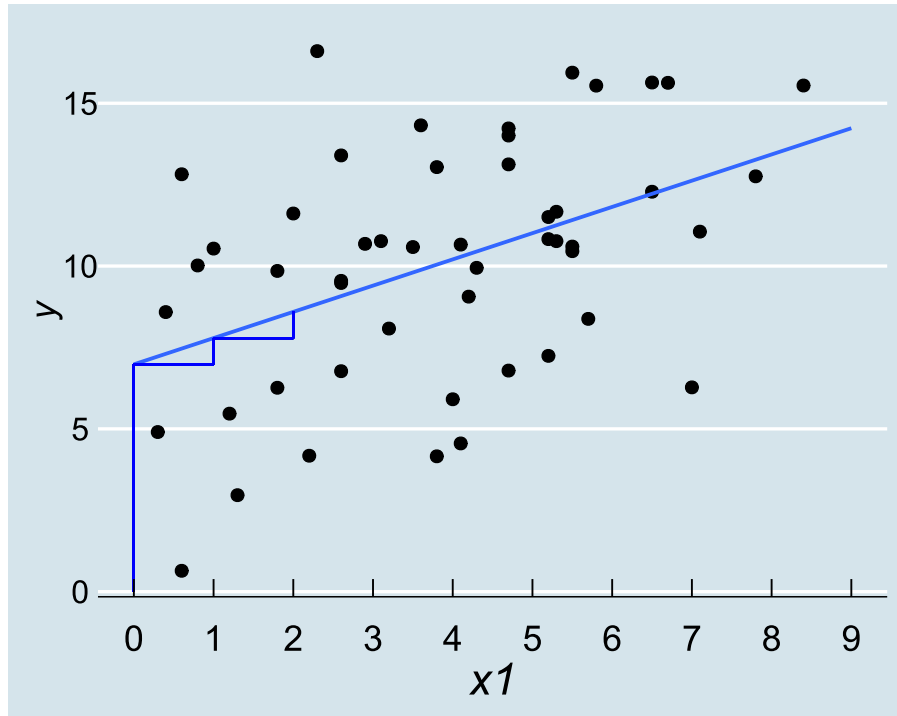


Describing a pattern with a line



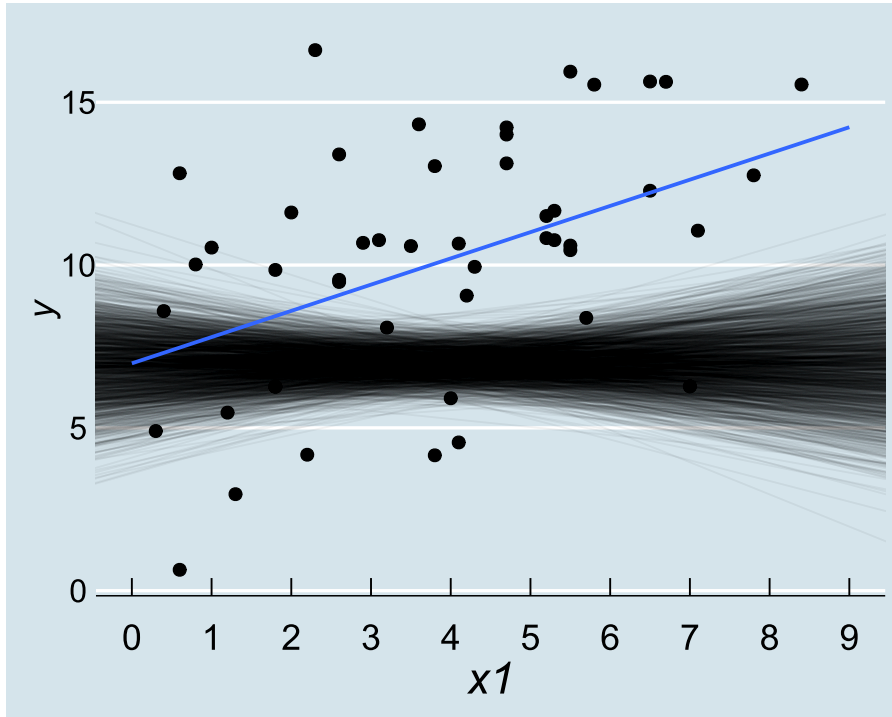
```
ggplot(df, aes(x = x1, y = y)) +  
  geom_point()
```

Coefficients describe the line



$$y_i = b_0 + b_1(x_{1i}) + \varepsilon_i$$

We can test the coefficients

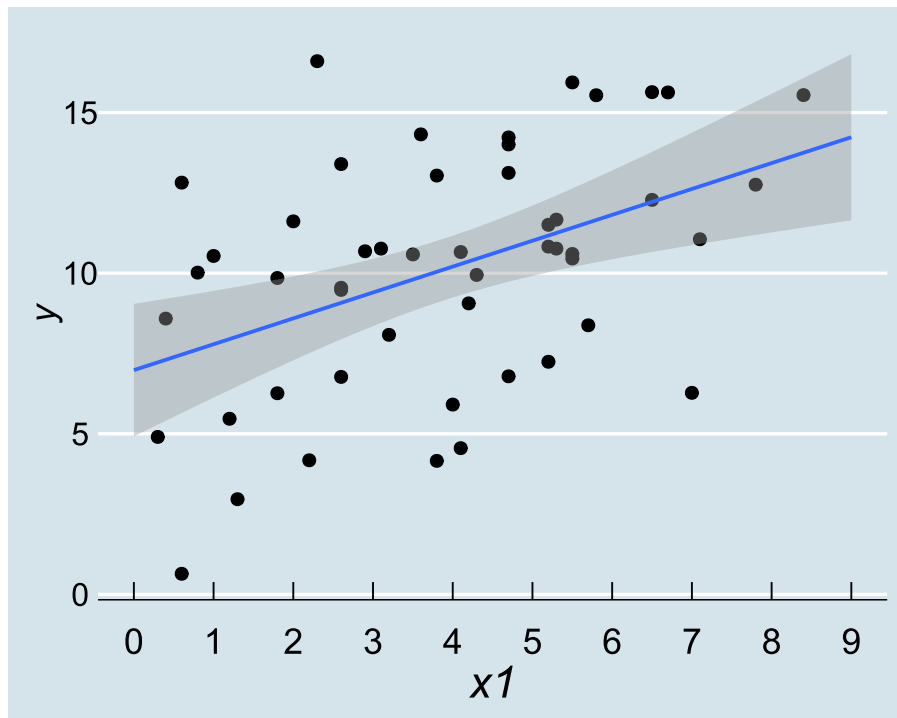


$$y_i = b_0 + b_1(x_{1i}) + \varepsilon_i$$

In the "null universe" where $b_1 = 0$, when sampling this many people, what is the probability that we will find a relationship at least as extreme as the one we *have* found?

```
## Coefficients:
##               Estimate Std. Error t value Pr(>|t|)
## (Intercept)    6.982      1.026     6.81  1.5e-08 ***
## x1              0.806      0.234     3.45  0.0012 **
```

We can test the coefficients



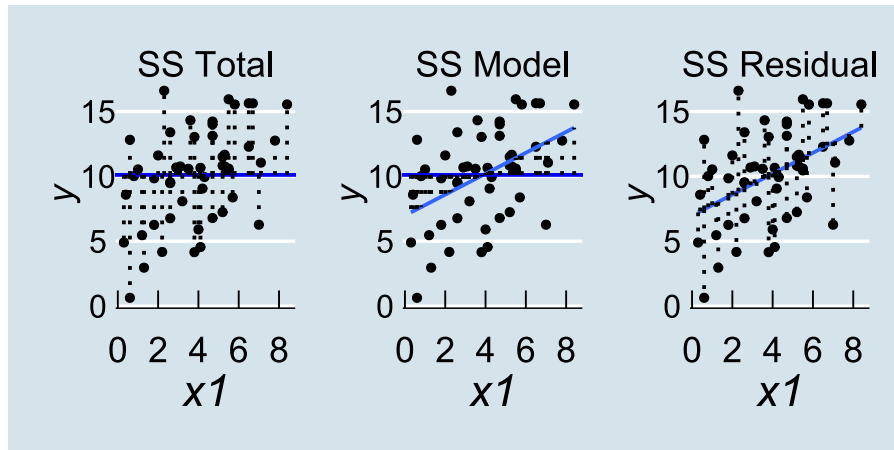
$$y_i = b_0 + b_1(x_{1i}) + \varepsilon_i$$

Plausible range of values for b_1 :

```
confint(model)
```

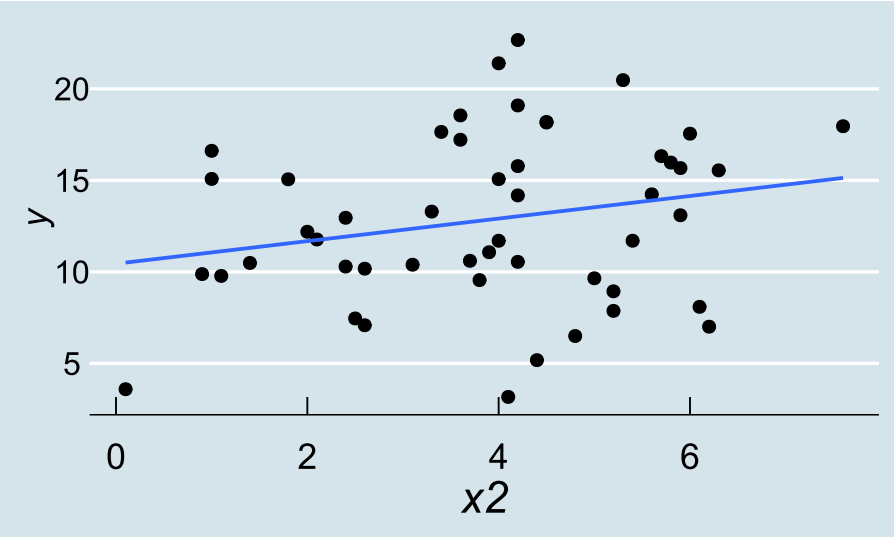
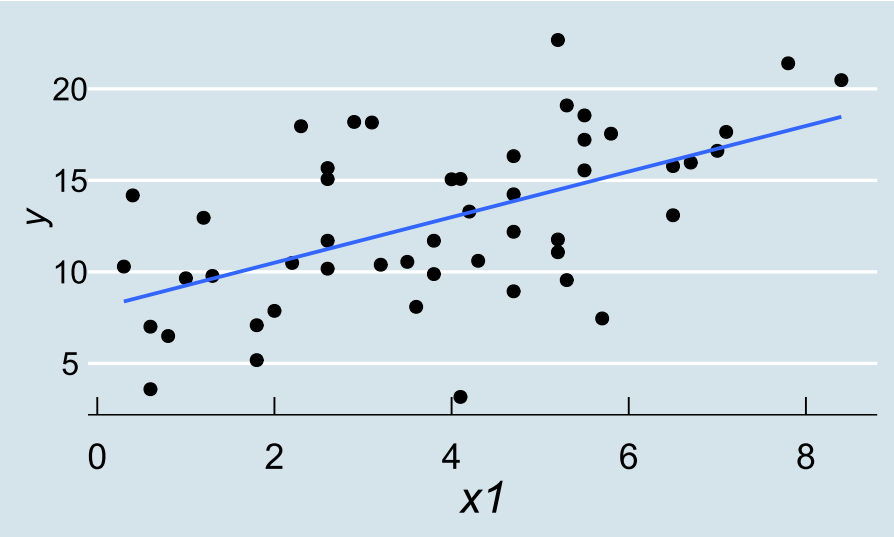
##	2.5 %	97.5 %
## (Intercept)	4.919	9.044
## x1	0.336	1.275

We can test the model



```
##  
## Multiple R-squared:  0.199,    Adjusted R-squared:  0.182  
## F-statistic: 11.9 on 1 and 48 DF,  p-value: 0.00118
```

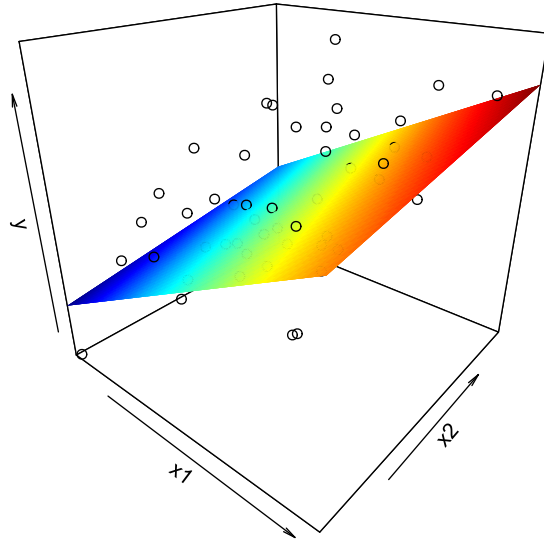
More predictors



More predictors (2)

$$y_i = b_0 + b_1(x_{1i}) + b_2(x_{2i}) + \varepsilon_i$$

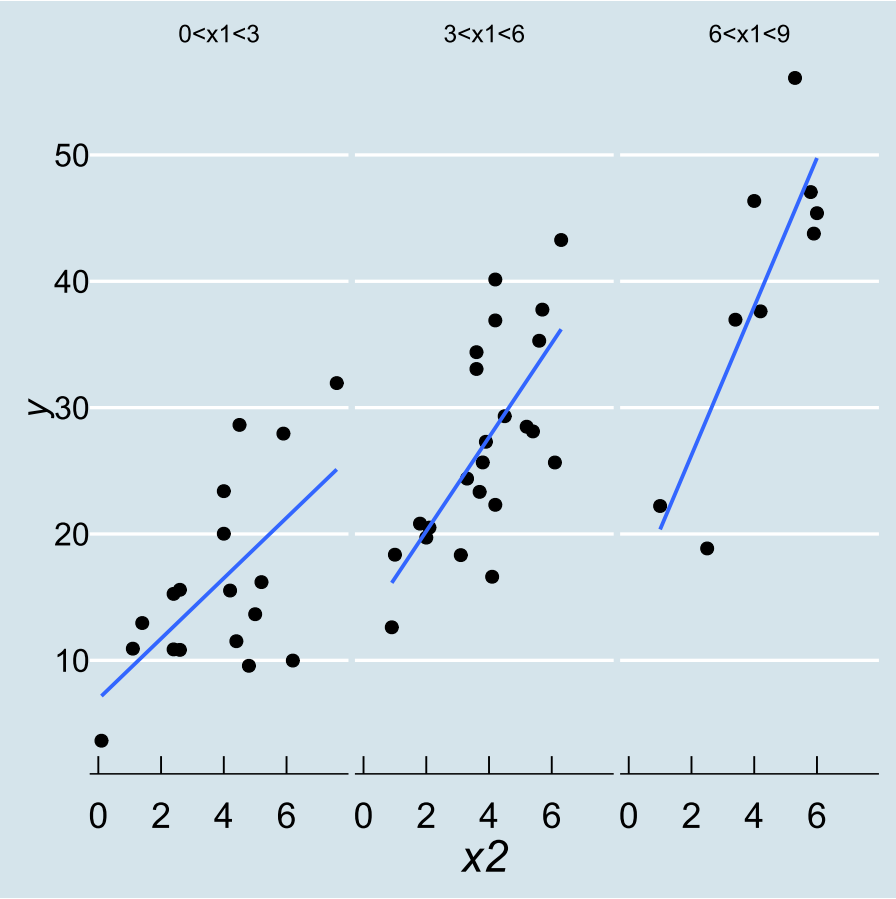
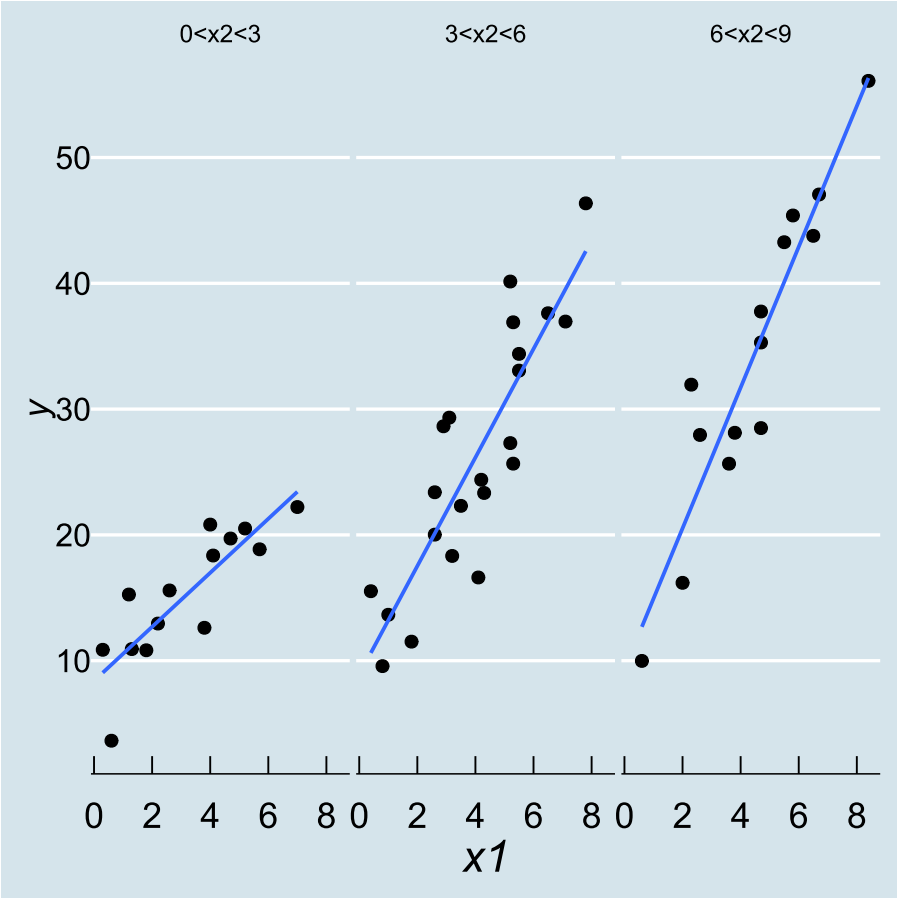
More predictors (3)



Even more predictors...

$$y_i = b_0 + b_1(x_{1i}) + b_2(x_{2i}) + \dots + b_k(x_{ki}) + \varepsilon_i$$

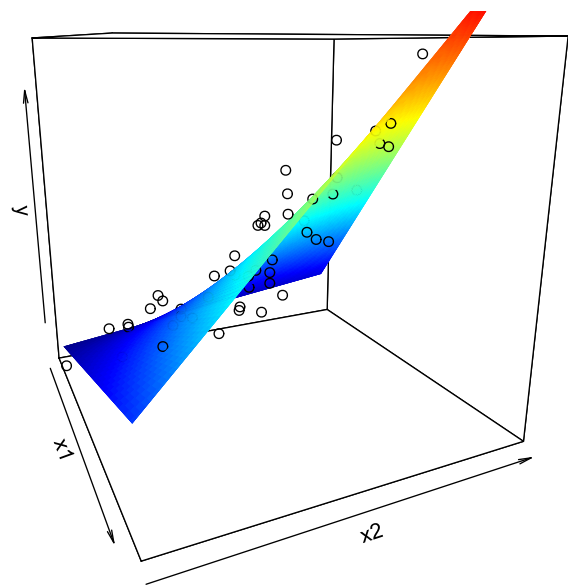
associations that depend on other things



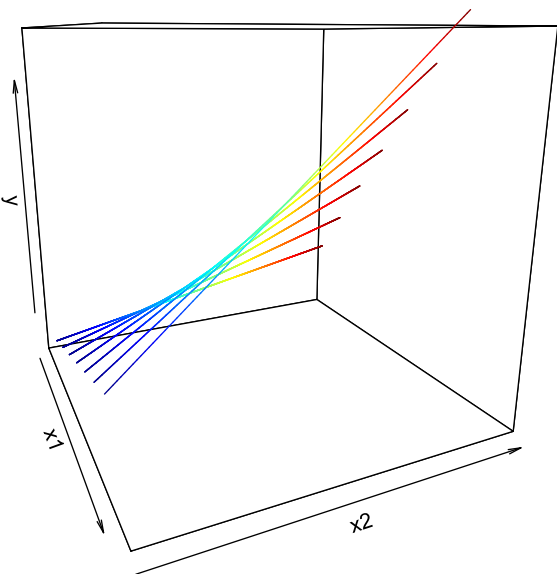
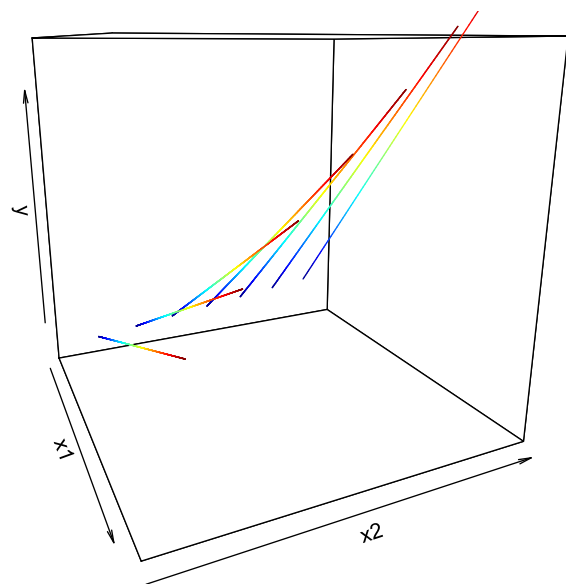
interactions

$$y_i = b_0 + b_1(x_{1i}) + b_2(x_{2i}) + b_3(x_{1i} \cdot x_{2i}) + \varepsilon_i$$

interactions

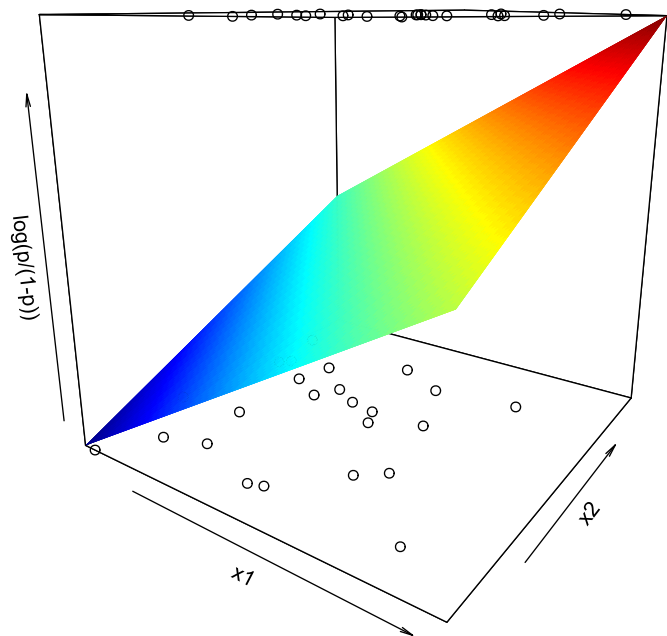


interactions (2)

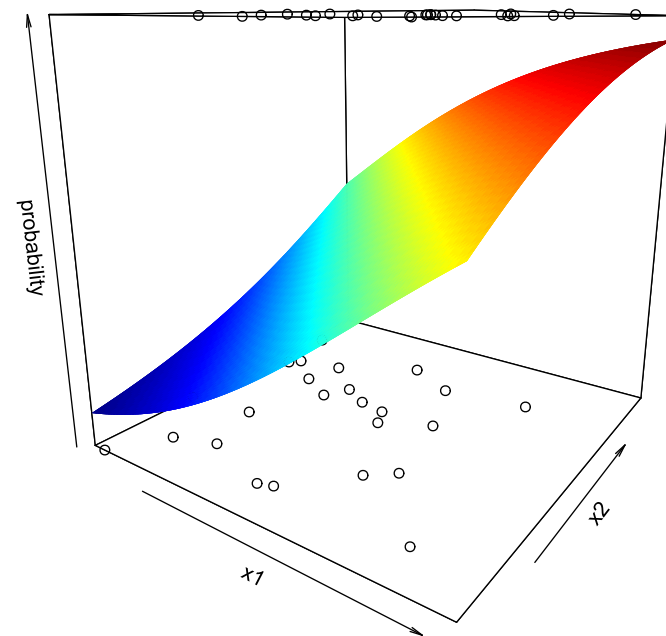


other outcomes

$$\ln\left(\frac{p}{1-p}\right) = b_0 + b_1(x_{1i}) + b_2(x_{2i})$$



$$\ln\left(\frac{p}{1-p}\right) \Rightarrow \frac{p}{1-p} \Rightarrow p$$



Checking Assumptions: Linear Models

required

- linearity of relationship
- for the *residuals*:
 - normality
 - homogeneity of variance
 - independence

desirable

- uncorrelated predictors
- no "bad" (overly influential) observations

Checking Assumptions: Logit Models

required

- linearity of relationship between IVs and log-odds
- for the *residuals*:
 - normality
 - homogeneity of variance
 - independence

desirable

- uncorrelated predictors
- no "bad" (overly influential) observations
- large samples (due to maximum likelihood fitting)

Part 2

Common Tests as linear models

```
usmr <- read_csv("https://uoepsy.github.io/data/usmr2022.csv")
```

lm vs correlation

regression, continuous predictor

```
summary(lm(sleeprating ~ height, data = usmr))
```

```
##
## Call:
## lm(formula = sleeprating ~ height, data = usmr)
##
## Residuals:
##      Min       1Q   Median       3Q      Max
## -65.42 -11.66   5.52  16.96  36.16
##
## Coefficients:
##              Estimate Std. Error t value Pr(>|t|)
## (Intercept)   18.785     46.879    0.4    0.69
## height         0.279      0.278    1.0    0.32
##
## Residual standard error: 22.6 on 76 degrees of freedom
## (2 observations deleted due to missingness)
## Multiple R-squared:  0.013,    Adjusted R-squared:  4.73e-05
## F-statistic:    1 on 1 and 76 DF,  p-value: 0.32
```

Correlation

```
cor.test(usmr$height, usmr$sleeprating)
```

```
##
##      Pearson's product-moment correlation
##
## data:  usmr$height and usmr$sleeprating
## t = 1, df = 76, p-value = 0.3
## alternative hypothesis: true correlation is not equal to 0
## 95 percent confidence interval:
##  -0.1112  0.3284
## sample estimates:
##      cor
## 0.1142
```

lm vs t.test

regression, intercept

```
summary(lm(height ~ 1, data = usmr))
```

```
##
## Call:
## lm(formula = height ~ 1, data = usmr)
##
## Residuals:
##      Min       1Q   Median       3Q      Max
## -15.407  -7.607  -0.107   7.523  20.893
##
## Coefficients:
##              Estimate Std. Error t value Pr(>|t|)
## (Intercept)    168.11         1.04    162   <2e-16 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 9.22 on 78 degrees of freedom
## (1 observation deleted due to missingness)
```

one sample t.test

```
t.test(usmr$height, mu=0)
```

```
##
##      One Sample t-test
##
## data:  usmr$height
## t = 162, df = 78, p-value <2e-16
## alternative hypothesis: true mean is not equal to 0
## 95 percent confidence interval:
##  166.0 170.2
## sample estimates:
## mean of x
##      168.1
```

lm vs t.test (2)

regression, binary predictor

```
summary(lm(height ~ catdog, data = usmr))
```

```
##
## Call:
## lm(formula = height ~ catdog, data = usmr)
##
## Residuals:
##      Min       1Q   Median       3Q      Max
## -13.655  -7.501   0.499   8.399  19.499
##
## Coefficients:
##              Estimate Std. Error t value Pr(>|t|)
## (Intercept)    166.36       1.55  107.60  <2e-16 ***
## catdogdog         3.15       2.07   1.52   0.13
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 9.15 on 77 degrees of freedom
## (1 observation deleted due to missingness)
## Multiple R-squared:  0.0291,    Adjusted R-squared:  0.0165
## F-statistic: 2.31 on 1 and 77 DF,  p-value: 0.133
```

two sample t.test

```
t.test(height ~ catdog, data = usmr,
        var.equal = TRUE)
```

```
##
##      Two Sample t-test
##
## data:  height by catdog
## t = -1.5, df = 77, p-value = 0.1
## alternative hypothesis: true difference in means between group cat and group
## 95 percent confidence interval:
##  -7.2706  0.9796
## sample estimates:
## mean in group cat mean in group dog
##              166.4              169.5
```

lm vs Traditional ANOVA

regression, binary predictor

```
summary(lm(height ~ eyecolour, data = usmr))
```

```
##
## Call:
## lm(formula = height ~ eyecolour, data = usmr)
##
## Residuals:
##      Min       1Q   Median       3Q      Max
## -17.55  -5.77   0.00   6.79  17.75
##
## Coefficients:
##              Estimate Std. Error t value Pr(>|t|)
## (Intercept)    170.254     2.183   77.99  <2e-16 ***
## eyecolourbrown    -3.682     2.609   -1.41    0.16
## eyecolourgreen     2.717     4.125    0.66    0.51
## eyecolourgrey    -0.254     9.515   -0.03    0.98
## eyecolourhazel   -2.404     3.935   -0.61    0.54
## eyecolourother   -4.834     5.775   -0.84    0.41
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 9.26 on 73 degrees of freedom
## (1 observation deleted due to missingness)
## Multiple R-squared:  0.0563,    Adjusted R-squared:  -0.00837
## F-statistic: 0.871 on 5 and 73 DF,  p-value: 0.505
```

anova

```
summary(aov(height ~ eyecolour, data = usmr))
```

```
##
##              Df Sum Sq Mean Sq F value Pr(>F)
## eyecolour      5     373    74.7    0.87  0.51
## Residuals     73    6261    85.8
## 1 observation deleted due to missingness
```


lm vs Traditional ANOVA

If you should say to a mathematical statistician that you have discovered that linear multiple regression and the analysis of variance (and covariance) are identical systems, he would mutter something like "Of course—general linear model," and you might have trouble maintaining his attention. If you should say this to a typical psychologist, you would be met with incredulity, or worse. Yet it is true, and in its truth lie possibilities for more relevant and therefore more powerful research data.

Cohen (1968)

History

Multiple Regression

- introduced c. 1900 in biological and behavioural sciences
 - aligned to "natural variation" in observations
 - tells us that means (\bar{y}) are related to groups (g_1, g_2, \dots, g_n)
-
- both produce F -ratios, discussed in different language, but identical

ANOVA

- introduced c. 1920 in agricultural research
- aligned to experimentation and manipulation
- tells us that groups (g_1, g_2, \dots, g_n) have different means (\bar{y})

Why Teach LM/Regression?

- LM has less restrictive assumptions
 - especially true for unbalanced designs/missing data
- LM is far better at dealing with covariates
 - can arbitrarily mix continuous and discrete predictors
- LM is the gateway to other powerful tools
 - mixed models and factor analysis (→ MSMR)
 - structural equation models

Goodbye!

