

概统
第八章

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8.1.

$$\begin{aligned} 1. (1) \quad E X &= \int_{-\infty}^{+\infty} f(x; \theta) x dx \\ &= \int_{-\theta}^{+\theta} \frac{2}{\theta^2} (\theta - x) x dx \\ &= \frac{1}{3} \theta \end{aligned}$$

$$\therefore \frac{1}{3} \hat{\theta} = \frac{1}{n} \sum_{i=1}^n X_i$$

$$\therefore \hat{\theta} = \frac{3}{n} \sum_{i=1}^n X_i$$

$$\begin{aligned} 3. \quad E X &= (1-p)^{1-p} \cdot \dots \cdot n(1-p)^{n-1} p \\ &= p \left[\frac{1-p}{1-(1-p)} \right]' p \frac{d(1-p)}{dp} \\ &= \frac{1}{p} \end{aligned}$$

$$\therefore \frac{1}{\hat{p}} = \frac{1}{n} \sum_{i=1}^n X_i$$

$$\therefore \hat{p} = \frac{n}{\sum_{i=1}^n X_i}$$

$$7. \quad L(\theta) = \prod_{i=1}^n \frac{1}{\sigma \sqrt{2\pi}} e^{-\frac{X_i^2}{2\sigma^2}} =$$

$$L(\theta) = \frac{1}{\sigma^n} \frac{1}{\sqrt{2\pi}}^n \exp\left(-\frac{\sum_{i=1}^n X_i^2}{2\sigma^2}\right)$$

$$= \frac{2^{-n/2}}{\sigma^n} \exp\left(-\frac{\sum_{i=1}^n X_i^2}{2\sigma^2}\right)$$

$$\ln L(\theta) = -n \ln 2 - \frac{n}{2} \ln \pi - \frac{n}{2} \ln \sigma^2 - \frac{\sum_{i=1}^n X_i^2}{2\sigma^2}$$

$$\frac{d \ln L(\theta)}{d \sigma} = -\frac{n}{\sigma} - \frac{\sum_{i=1}^n X_i^2}{\sigma^3}$$

$$\text{令 } -\frac{n}{\sigma} - \frac{\sum_{i=1}^n X_i^2}{\sigma^3} = 0$$

$$\text{得 } \hat{\sigma}^2 = \frac{\sum_{i=1}^n X_i^2}{n}$$

8. (2) $L(\theta) = \prod_{i=1}^n f(x_i; \theta_1, \theta_2)$

$$L(\theta) = \prod_{i=1}^n \frac{1}{\theta_2 - \theta_1} f(x_i; \theta_1, \theta_2)$$

且

$$L(\hat{\theta}) = \max_{\theta \in \Theta} L(\theta)$$

$$\therefore \theta_2 = \max \{x_1, \dots, x_n\}$$

$$\theta_1 = \min \{x_1, \dots, x_n\}$$

$$L(\hat{\theta}) = \left(\frac{1}{\max \{x_1, \dots, x_n\} - \min \{x_1, \dots, x_n\}} \right)^n$$

$$11. \quad L(p) = \prod_{j=1}^{100} C_{10}^j p^j (1-p)^{10-j} = \prod_{j=1}^{100} C_{10}^j p^{\sum_{j=1}^{100} j} (1-p)^{1000 - \sum_{j=1}^{100} j}$$

$$\frac{d L(p)}{d p} = \frac{\sum_{j=1}^{100} j}{p} - \frac{1000 - \sum_{j=1}^{100} j}{1-p}$$

$$\text{令 } \frac{d \ln L(p)}{d p} = 0 \text{ 得}$$

$$p = 0.499$$

8.2

$$\begin{aligned} 2. 1) E(X) &= \int_{-\infty}^{+\infty} f(x) x dx \\ &= \int_0^{\theta} \frac{3x^2}{\theta^3} dx \\ &= \frac{3}{4}\theta \end{aligned}$$

$$E(T_1) = \frac{2}{3}(E(X_1) + E(X_2)) = \theta$$

$$E(T_2) = \frac{7}{8} E(\max\{X_1, X_2\}) = \frac{7}{8} \times \frac{6}{7} \theta = \frac{3}{4}\theta$$

$\therefore T_1, T_2$ 均为无偏估计

$$(2) \text{ 令 } Z = \max\{X_1, X_2\}$$

$$f(z) = \int_{-\infty}^z \int_{-\infty}^z \frac{3x^2}{\theta^3} \cdot \frac{3y^2}{\theta^3} dx dy = \frac{z^6}{\theta^6}$$

$$f(z) = \frac{dF(z)}{dz} = \frac{6z^5}{\theta^6}$$

$$EZ = \int_{-\infty}^{+\infty} z f(z) dz = \frac{6}{7}\theta$$

$$ET_2 = \frac{7}{8}EZ = \theta$$

$\therefore T_1, T_2$ 均为无偏估计

$$\begin{aligned} (3) E(T_1^2) &= \frac{4}{9} [E(X_1^2) + E(X_2^2) + 2E(X_1 X_2)] \\ &= \frac{4}{9} \left(\frac{6}{5}\theta^2 + \frac{6}{5}\theta^2 \right) \\ &= \frac{32}{30}\theta^2 \end{aligned}$$

$$\begin{aligned} E(T_2^2) &= \frac{49}{36} \int_0^{\theta} z^2 f(z) dz \\ &= \frac{3}{4}\theta^2 \cdot \frac{49}{36} \\ &= \frac{49}{48}\theta^2 \end{aligned}$$

$$D(T_1) = E(T_1^2) - (ET_1)^2 = \frac{1}{30}\theta^2$$

$$D(T_2) = E(T_2^2) - (ET_2)^2 = \frac{1}{48}\theta^2$$

$$D(T_1) > D(T_2)$$

$$4. E(\theta) = \theta$$

$$E(\theta^2) = D(\theta) + E(\theta)^2 = \frac{1}{6}\theta^2 + D(\theta) > \theta^2$$

\therefore 不是无偏估计

6.1)

$$\textcircled{1} E(\hat{\sigma}^2) = \frac{1}{n} \sum_{i=1}^n E[(X_i - \mu_0)^2] = \frac{1}{n} \sum_{i=1}^n \sigma^2 = \sigma^2$$

\therefore 是无偏估计

$$\textcircled{2} P\left\{ \left| \frac{1}{n} \sum_{i=1}^n E(X_i - \mu_0)^2 - \sigma^2 \right| < \varepsilon \right\}$$

$$= P\left\{ |S^2 - ES^2| < \varepsilon \right\}$$

$$> 1 - \frac{DS^2}{\varepsilon^2}$$

$$DS^2 = \frac{\sigma^2}{(n-1)^2} \cdot D\left[\frac{(n-1)S^2}{\sigma^2}\right] = \frac{2\sigma^4}{n-1}$$

$$\therefore P > 1 - \frac{2\sigma^4}{\varepsilon^2}$$

$$\therefore P \rightarrow 1$$

\therefore 是一致性估计

$$(2) X_i - \bar{X} \sim (0, \frac{n-1}{n} \sigma^2)$$

$$E(\hat{\sigma}_1^2) = \frac{1}{n-1} \sum_{i=1}^n \frac{n-1}{n} \sigma^2 = \sigma^2$$

$$E(\hat{\sigma}_2^2) = \frac{1}{n} \sum_{i=1}^n \frac{n-1}{n} \sigma^2 = \frac{n-1}{n} \sigma^2$$

$$E(\hat{\sigma}_3^2) = \frac{1}{n-1} \sum_{i=1}^n \frac{n-1}{n} \sigma^2 = \frac{n-1}{n} \sigma^2$$

$$E(\hat{\sigma}_4^2) = \frac{1}{n} \sum_{i=1}^n \sigma^2 = \sigma^2$$

$\therefore \hat{\sigma}_1^2, \hat{\sigma}_4^2$ 为无偏估计

$$D(\hat{\sigma}_1^2) = D(S^2) = \frac{2\sigma^4}{n-1}$$

$$D(\hat{\sigma}_4^2) = D\left(\frac{\sigma^2}{n}\right) = \frac{\frac{1}{n^2} \sum_{i=1}^n E[(X_i - \mu_0)^4]}{(\sigma^2)^2} = \frac{3\sigma^4}{n}$$

$$\therefore D(\hat{\sigma}_1^2) > D(\hat{\sigma}_4^2)$$

$\therefore \hat{\sigma}_4^2$ 优良性更佳

概统(2)
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$\{m, n\}$
 $\frac{m}{n}$
 G_b
 S_1
 F

8.4

$$3. P\left\{\bar{X} - z_{1-\frac{\alpha}{2}} \frac{\sigma}{\sqrt{n}} \leq \mu \leq \bar{X} + z_{1-\frac{\alpha}{2}} \frac{\sigma}{\sqrt{n}}\right\} = 1 - \alpha = 99\%$$

$\therefore \alpha$ 取 0.01

查表得 $-1.17 \leq \mu \leq 6.57$

\therefore 置信区间为 $[-1.17, 6.57]$

$$(2) P\left\{\bar{X} - t_{1-\frac{\alpha}{2}}(3) \cdot \frac{S}{\sqrt{2}} \leq \mu \leq \bar{X} + t_{1-\frac{\alpha}{2}}(3) \cdot \frac{S}{\sqrt{2}}\right\} = 0.95$$

$$S^2 = 2.277$$

\therefore 置信区间为 $[-0.923, 6.224]$

6. 两侧分位点分别为 $-z_{1-\frac{\lambda}{2}}$ 和 $-z_{1-\frac{\gamma}{2}}$

$$(\lambda + \gamma = \alpha)$$

$$\text{即 } \varphi(b) - \varphi(a) = 1 - \alpha$$

求 a 和 $b - a$ 最小

$$\text{令 } g(a, b) = b - a + \lambda [\varphi(b) - \varphi(a) - 1 + \alpha]$$

$$\frac{dg}{da} = -1 + \lambda \left(\frac{e^{-\frac{a^2}{2}}}{\sqrt{2\pi}} \right) = 0$$

$$\frac{dg}{db} = 1 + \lambda \frac{e^{-\frac{b^2}{2}}}{\sqrt{2\pi}} = 0$$

$$\text{得 } b = -a$$

$$\text{即 } |z_{1-\lambda}| = |z_{1-\gamma}|$$

$$\therefore \lambda = \frac{\alpha}{2} = \gamma$$

$\therefore [z_{1-\frac{\alpha}{2}}, z_{1-\frac{\alpha}{2}}]$ 为最优区间