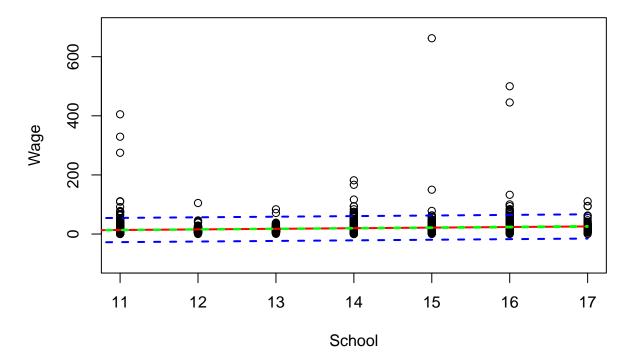
FTAP Homework 5

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Problem 1

 \mathbf{a}

```
wgs <- read.xls(xls = "censuswage.xls")</pre>
lm.out <- lm(Wage ~ School, wgs[wgs$School >= 10,])
confint(lm.out)
##
                    2.5 %
                              97.5 %
## (Intercept) -12.256742 -4.875641
## School
                 1.709791 2.321660
b
plot(Wage ~ School, wgs[wgs$School > 10, ], ylim=c(-100, 700))
pred.int <- predict(lm.out, newdata = data.frame(School=c(10:17)), interval = "predict")</pre>
conf.int <- predict(lm.out, newdata = data.frame(School=c(10:17)), interval = "confidence")</pre>
lines(c(10:17), pred.int[,1], col="red", lwd = 2)
lines(c(10:17), pred.int[,2], col="blue", type="l", lty=2, lwd = 2)
lines(c(10:17), pred.int[,3], col="blue", type="1", lty=2, lwd = 2)
lines(c(10:17), conf.int[,2], col="green", type="l", lty=2, lwd = 2)
lines(c(10:17), conf.int[,3], col="green", type="1", lty=2, lwd = 2)
```



Problem 2

```
ceo <- read.xls(xls = "ceosalary.xls")</pre>
lm.out <- lm(salary ~ comten + ceoten + sales, ceo)</pre>
print(lm.out) # Point Estimates
##
## Call:
## lm(formula = salary ~ comten + ceoten + sales, data = ceo)
##
## Coefficients:
## (Intercept)
                     comten
                                   ceoten
                                                 sales
     674.17896
                   -3.05712
                                 15.62693
                                               0.03858
summary(lm.out)$coefficients[,"Std. Error"] # Std Errors (You can also see this from sqrt(diag(vcov(lm.
  (Intercept)
##
                      comten
                                    ceoten
                                                  sales
## 89.434836362 3.504800534 6.006818282
                                            0.006732074
confint(lm.out) # Confidence intervals
                      2.5 %
                                   97.5 %
## (Intercept) 497.65504252 850.70287558
                -9.97479541
## comten
                               3.86055426
## ceoten
                 3.77084090
                             27.48301240
## sales
                 0.02529341
                              0.05186856
summary(lm.out)$coefficients[,"Pr(>|t|)"]
## (Intercept)
                      comten
                                    ceoten
                                                  sales
## 2.562366e-12 3.842719e-01 1.008491e-02 4.351557e-08
```

 \mathbf{a}

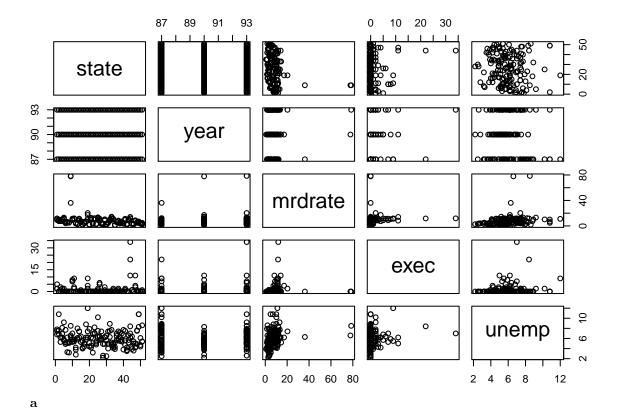
Because the confidence interval for the slope coefficient of company tenure on ceo salary includes 0. We could hypothesise that company tenure is unrelated to ceo salary. However, company tenure is probably highly correlated with ceo tenure. Therefore, the large covariance between the two variables may be causing the counter intuative relationship between company tenure and salary.

b

Because the confidence interval for sales does not include 0 and because the p-value for the statistical significance for sales as a determininent of sales is far below the 99% confidence level we can be confident that theory two is incorrect.

Probelm 3

```
redrum <- read.xls(xls = "murder.xls")
plot(redrum)</pre>
```



redrum\$CapPun <- as.numeric(redrum\$exec > 0)
lm.out <- lm(mrdrate ~ unemp + CapPun, redrum)
lm.out</pre>

```
##
## Call:
## lm(formula = mrdrate ~ unemp + CapPun, data = redrum)
##
## Coefficients:
## (Intercept) unemp CapPun
## 0.116 1.253 1.753
```

b

The coefficient on the capital punishment variable means that for a given level of unemployment, whether a community has capital punishment or not is associated with 1.753% higher murder rate.

 \mathbf{c}

summary(lm.out)

```
##
## Call:
## lm(formula = mrdrate ~ unemp + CapPun, data = redrum)
##
## Residuals:
## Min 1Q Median 3Q Max
## -8.892 -3.549 -1.159 0.933 69.414
```

```
##
## Coefficients:
               Estimate Std. Error t value Pr(>|t|)
##
                 0.1160
                            2.6812
                                      0.043 0.96554
## (Intercept)
## unemp
                 1.2531
                            0.4357
                                      2.876 0.00462 **
## CapPun
                            1.6480
                                      1.064 0.28917
                 1.7530
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 8.949 on 150 degrees of freedom
## Multiple R-squared: 0.06479,
                                     Adjusted R-squared:
## F-statistic: 5.196 on 2 and 150 DF, p-value: 0.00658
Null Hypothesis: H_0: \beta_2 = 0
Alternative: H_1: \beta_2 \neq 0
```

Because the p-value for β_2 is >5% we fail to reject the null hypothesis that whether or not a community has capital punishment influences the murder rate.

Controlling for the effects of unemployment, the presence of capital punishment we cannot say that capital punishment is correlated to the murder rate.

 \mathbf{d}

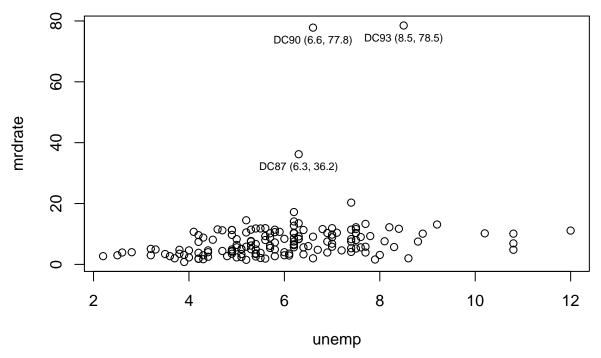
```
predict(lm.out, newdata = data.frame(unemp=c(6,6), CapPun=c(1,0)), interval = "predict")

## fit lwr upr
## 1 9.387374 -8.512036 27.28678
## 2 7.634422 -10.127569 25.39641
```

These intervals are very large! Both intervals include 0 also both intervals include values with are both >2x and < x/2 the fit (spot estimate) value.

 \mathbf{e}

```
plot(mrdrate ~ unemp, redrum)
dcUnemp <- redrum[redrum$state == 9,]$unemp
dcMrdrate <- redrum[redrum$state == 9,]$mrdrate
text(x = dcUnemp, y = dcMrdrate, labels = c("DC87 (6.3, 36.2)", "DC90 (6.6, 77.8)", "DC93 (8.5, 78.5)")</pre>
```



```
redrumNoDC <- redrum[redrum$state != 9,]
lm.out <- lm(mrdrate ~ unemp + CapPun, redrumNoDC)
lm.out

##
## Call:
## lm(formula = mrdrate ~ unemp + CapPun, data = redrumNoDC)
##
## Coefficients:
## (Intercept) unemp CapPun
## 2.1320 0.6443 3.5957</pre>
```

Based on the new regression which excludes Washington D.C. for an given level of unemployment, whether or not a state has capital punishment is associated with a 3.6% higher murder rate.

```
summary(lm.out)
```

```
##
## Call:
## lm(formula = mrdrate ~ unemp + CapPun, data = redrumNoDC)
##
## Residuals:
##
       Min
                1Q Median
                                 3Q
                                        Max
   -6.5515 -2.0081 -0.3782
                            1.4475
##
## Coefficients:
               Estimate Std. Error t value Pr(>|t|)
##
## (Intercept)
                 2.1320
                             0.9315
                                      2.289
                                               0.0235 *
## unemp
                                      4.235 4.01e-05 ***
                 0.6443
                             0.1522
## CapPun
                 3.5957
                             0.5729
                                      6.276 3.68e-09 ***
##
```

```
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 3.095 on 147 degrees of freedom
## Multiple R-squared: 0.3106, Adjusted R-squared: 0.3012
## F-statistic: 33.11 on 2 and 147 DF, p-value: 1.344e-12
```

Null Hypothesis: $H_0: \beta_2 = 0$

Alternative: $H_1: \beta_2 \neq 0$

Because the p-value for the β_2 coefficient is less than .001% we are able to be more than 99% confident in rejecting the null hypothesis.

Discussion:

We know that the t-statistic for multiple varaiable regressions depend the standard error of the coefficient we are examining.

$$t = \frac{b_j - \beta_j}{s_{b_j}}$$

Furthermore, we know that the standard error for the coefficient depends on \bar{X} :

$$\sqrt{(\sigma^2 * [1 + \frac{1}{n} + \frac{(X_f - \bar{X})}{(n-1) * s_x^2}])}$$

Therefore, due to removing the large