Statistics

-	Stal	utics	12 10 10 10 10 10 10 10 10 10 10 10 10 10		oniket.
	The	es of data	<u> </u>	119 - 220 110 - 11	XX SVA
	Categorical or Qualitative	lata	Numarical	or Quan	titative data
	Nominal Ordin				Continuos
		order	Discrete age/nank		weight / height
	No heirarchy good	order as a	35,23		59.5, 72.7
	Return Dalmar Control	550 19 97.6 273 135-60577	1343	N I AV	100
	Population mean u	= Exi	[N=no of	· done	in Dogwester]
	ALMOST PRODUCT THE TOTAL CONTRACTOR OF THE PRODUCT TO THE PRODUCT	^/		1.50	
	Sample mean X	= Exi	[n = no. of	tems 1	n cample?
	2 2 3 2 14	n	J	0.05	in samp a
2	Mean is posone to ou	tliers m	edian can sol	AR puttion	oma Üsus.
	mode can work best				
5.455	Weighted mean we als	im Some	weights 8 trumin	med mean	we Hemone
	fraction of values for	on hoth	the sides To you	men a Out	Hiore.
U	James James F		are such to the	Thouse our	XXXX
L	Variance - 2 = 5	(xi-x)2	used in finding	dispersion	in of data
		AX	12 M 2 15 2 17	5.3.4.8	
	Mean absolute deviation	MAD -	ZIZI KI	21	
-	Ct. 1. I I	CA - F	14 : 14	10 - 10 - 00	really to
	Standard deviation	SD = Vo	this Rolds The	ame uni	t as data
	Coefficient of variation	CV= Sta	ndond deviat x 100	Trgu	ies the idea of compare
			mean	spread of	colums & compare
	EURO COMPANIA			41.03	
L	Categorical - Foregue Numerical - Bins history	te solun	uns - Categoric	al & Nu	reducal
a)	Categorical - Foregue	mcy bara	Carts, relative	frequency	, cumulative for
B	Numerical - Bing histor	mams, no	rmal Skew, bin	nodel, c	netarm, no pater
		•	P. Limber	- 10° v	0 , ,
2	graph for Sivoriate	Column :	C&C, NON	C&N	1 Polass
a	Categorical & C - Continge	my table	or owsstab	Survived	1 2 3
В	Numerical & N :- Scar	terplots	7 P V V	0	12 14 50
C) Numerical & N : Scar) C&N : Bor charts we	th aggregat	af" or ows to	268 P	ivot tables

4	Quantiles : Divide a set of numeric data into equal sized group
	Quartiles: Q1 (25 percentile), 82 (50 percentile), 93 (75 perce)
	Deciles (10th percentite) 01 D9
dil	Parcontile: Q1.D2 Pgg
6363	Quintile - 5 equal parts [minimum, Q1, Q2, Q3, man] of describe ()
**	18 43 - 47 - Menuality Hange Bon of Bonger
L	PI = P (N+1) PI= does nowcentile value location
	N= no. of obs. P= percentile scant
En	find 75 percentile score from data
**	78 82 84 88 91 92 94 96 as 99 > Solt the values
NA	PL= 75 (10+1)= 33 = 8.25 : 98 - 98 : 96+ (98-96) 1=60 W
	100 4 8 + 9 = 96.5 42
	- Mason is much to the Billion can solve Entries is
West L	Percentile of a value:
\$-38m	Percentile grant = x + 0.5 y x > no of values below quen
	n y > no of values equal to que
	1 2 3 n + total no- of values
Gen	78,82,84,88,91,93,44,96,98,99
	.: 3+0.8(1) = 0.35 = 35%
	Maan about depation, MAD - Elxo XI - 01
En	How to create hamplots?
te	6 213 241 260 281 290 314 321 350 1500
dea H	0=====================================
Sieny Its	92-30 (11)-3.3 - 288.3
	01 = 25(11) = 2.75 = 213+4.75(2115212) = 10 tokuler
70	01 = a5(11) - 2.75: 213+0.75(241-213) = 234 whater
t pulaling t	THE RESERVE OF THE PROPERTY OF
in pri	83=75'x11/100=8.25=328.25
-	max = 93 + 1.5 (IQR) = 469 693 234 285 400
	(30,0 502 01 5 378 = 224 501)
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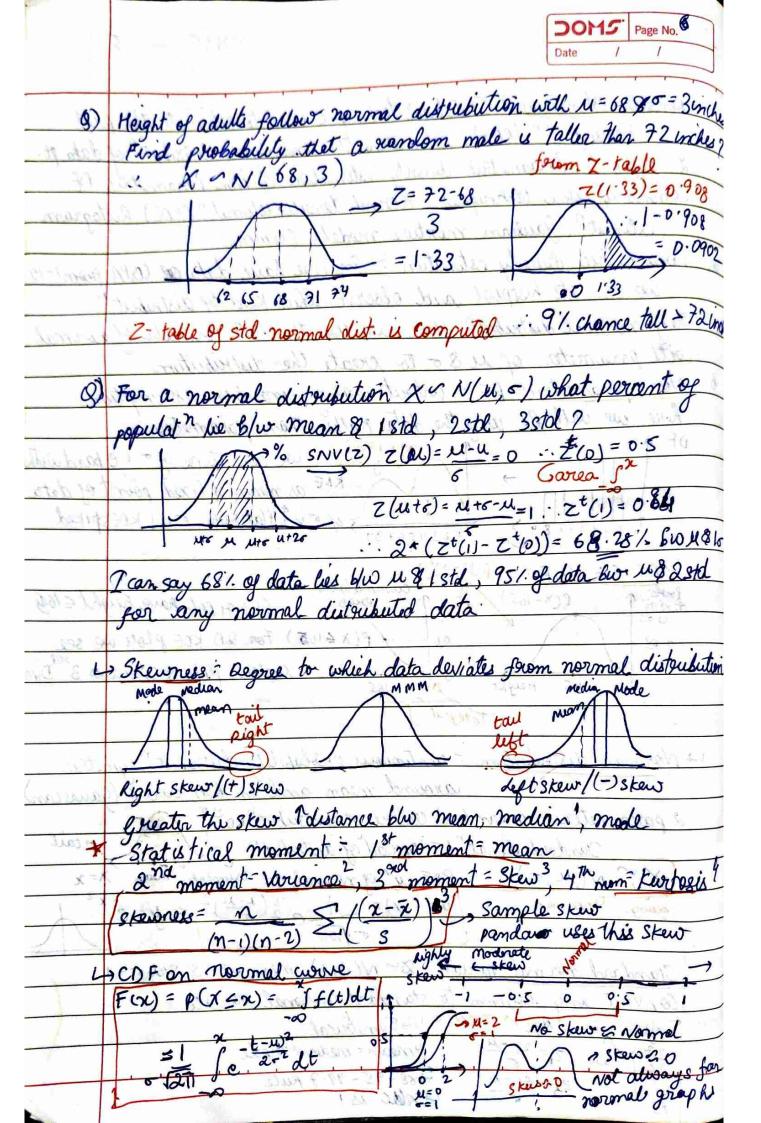
Variables change toge	my ocetated measures how much &
	thull take a on I other I or I
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May KV Syx 1 Co	TXXX STUXXX
→	X1××××××××××××××××××××××××××××××××××××
Covariance x is (+) we co	inarance istare covariance = 50
their confidences with	e easter wast along with
variane formula -	LIPT WILL TELL
Population	Sample
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N	2 - 1 - 1 - 1 - x
V - Value Al X V in appoint	A 11 11 11 11 11 11 11 11 11 11 11 11 11
2, My > population mean	E, y→ sample mean
2, soy > population / memor	n > Same
(-> no of observation.	× 13.00 × 30 0 2013 A
2	(701) Deep plane persons
Lis advantage of covarian	nio:
ioes not tell us about stru	ength of relat n b/w 2 variable as magnitude
of covariance is affected	d by the scale of variable
Covariance tells us abou	ut the spread of data it gets scaled up
J. B. W. L. I 4 2 3	- Mt es Part Burbles Be Bucker
ornelation - Refore to sta	atistical relat " b/w variables it measure
the doored to which 2	variable are related & how they Change inse (x,y) Invor
GOVERNO E COVERNO	inge (X, y) Invorn
C On	* of y = (2) x 1 = (292 m)
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standard d	wat of t
if I things are correlated	does not mean it is constin or course
Caucation refers to a ca	we & effect relat" Ww & variables
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legachs for multivariate	E Colums:
graphs for multivariate Scatterplats 3D, Barplats	is with Hul, Facetgoids, Jointplots,

şl.

10	Random variables - Set of possible values from a grandom lxp. Ex coin foss H&T & = & 0, 13 H=0, T=1
Н	Kandom Variables. Set of possible values just
n de	Ex coin toss H&T = a , 15 reportal letters?
	Random variables [Capital letters]
	Types of Random variables: Discrete RV = 21,2,3,4,8,63016 Continuous RV = 21-103 CGPA
	Types of Random variables: Discrete RV = 21 - 103 CGPA
L	Probability distribut "- dist of all of the possible outcome of a vandom variable along with their propability values (our toss (10) T(1)
	Tropicing disputer along with their propability values
	Conten (410) T(1)
	Parst 11 1/2
	Peroposition 1/2 1/2
	we can derive mathematical for to model relator b/w outcome & fort
	X > NUSCOMO 1/- 54 1 0
	Probability mass funct (PDF) Probability density funct (PDF) Commutative Distribut fr (CDF)
-	Probability mass funct" (PMF) , Cummulature Distribut (CDF)
1 12	Parobability density funct (PDF)
4	1111 - Prospering distribution of aister garages to
gritiga	Sum of prob is equals to 1
The second	:. y = 5 16 if X & 21, 2, 3, 4, 5, 6 }
AN IN	as Bar 2 of of otherwise of head to still as more than
	EMF of pmf describes the probability of X found at = x
12.4121	First 5 Plax 42) Teles Sensition of water and stocker a
-2X 4,5 -3	En for dice role P(4) = 1/6
T Carps	$CMF(4) = P(X \le 4) = P(1) + P(2) + P(3) + P(4) = 4/6$
	$CMF(6) = P(X \le 6) = 1 \land np. cumsum(s)$
	Control x 10 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1
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	So we take small step of I to find pouble of density of density of
1	getting marks blu 4-4.00 , that value is on or front
de de	9 300 bt
A M	y-01/45 x (mark
	P(06x410)=1=Axea

4) Den	sety estimation - Statistical technique used to estimate the prop.
103/105	density for (PDF), estimating the underlying distribut of data pt.
2	types - Parametric density reportion of all - carements of
Com	monte used techniques - Kound down estimation & Non-parametric OF
1/6.0-	monly used techniques - Kounel density estimat " (KOE), Ristogram
al Par	estimet, Gausian mixture models (GMMs)
4	ametric density estination - En use have data of CCPA from 1-10
5.	se see its histplot and observe which kind of distrubit"
Jogy	normal distribut then we use the PDF formula of normal
(OCT	parameters of u & o to create the distribution
O) Nor	- parametric DE - without making underlying assumptions
her	e we will be using the data rather than parameters
DF	e we will be using the data reather than parameters add points we can turn of a c bandwith **KDE* as mean is fixed point of data if = in flat KDE; = in KDESpiked
	> KOE as mean is fixed point of data
13	of out flat KDE; out KDESpiked
Tel 8 1,000	123456781
4 C1	OF for PDF: pifferentiat" Cummulation (x = 165) dx power 60% of ppl have height < 166
dense	by 1 (x = 163) (x) pros. 60% of ppc name registre = 109
0.1	of P(XZ165) FOR UD KOE PLOTE WE SEE
woulden	Constitute of the constitution of the constitu
	165 Height 7 165
1	Integrat S
LA NO	ermal distoubilion - continuous probability distrubut symetric
	around mean and in bell curve (Gaussian)
20	arameters: mean (4) centre of distoubut " PO) () () tail
p- 1	Standard deviat (5) spread of distribut
1 44	(N(U, o) PDF Eq " of normal dut asymptotic mean >= x
be	longs informed y= f(x)= 1 = 2 (2-11) = y= e
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St	andard normal variate 20 N(O, 1) mean u=08 Std. 5=1
Co	nvent and column to standard normal variety
NGW 6	ge - Age > Z~N(O,1) & Symetrical
2	2 34 7= 27-4 4 mean = median mode
41 6	23=61-14 468-95-99.7 Rule
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limit theorem, assumpt of ML algo like CaMM & Linear Reg (everar	sü
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Discrete non-gaussian	
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Kvoctosi is 4th Statistical moment, it measures tailedness of PD (Nortosis = n × (n+1) x 2 (x-1) 4 - 3 × (n-1) 2 xurtosis 1 = 1 = 1 = 1 = 1 = 1 = 1 = 1 = 1 = 1	
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outlers of	
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four nominal it is Take a column X and oceates its	
1st quantile, 2nd 99th quantile and do the same for a norma	Lda
Now draw a scatterplot and if it fets in same line then	
date is normal and gaussian in nature	Wgy
It can use any other distribution and can compare how	_ 1
Similar it is to our data	
in taken his like break the entrivale into smaller a well-	
is uniform distribution: All outcomes are equal within nange	
Uniform - Diecreti	
Denoted as (X & U(a, b) s parameters a= 1 lower b= 6 higher	
Denoted as (be a b) b = 6 higher	
4 PDF & CDF for uniform distrubition (skewness=0)	
45 PDF & CDF for uniform distrubution (Skewness = 0)	
over b-a PDF of tognormal	2.
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a b x 0 a b x 120 12TI	1
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	•
Log normal distribut - Heavy tailed continue PD whose logs is	4
normally distributed. * Distribut " siska log is normal log(x) ~ N(11, ='))
11.10.1	•
is called as log-normal (u, 5	
say if we have just stew data => ln(x) ~ N(11,0	17

		and Catalinus
100	Pareto distrukution - Exhibe	g power - law behavious
g branch	Power law where x & y ar	dota Auc Ex wealth distribution of sculer, Not always true, a=0
Charlet J.	y = k x 2 80% of	data Auc excurrent such a = 0
	or 1 reat 1 tail & 100:2	D same C3 100 mm () = 3
	al peak + tail 1	
194	exampled of son ha xm 20%	2
	PDF = y = xxm	eg(s) parata distrusi
(8)	your to detect distribut "is	pareto? > log(2)
7.	Take a greath of log(y) VI	pareto? -> log(a) s log(a) is straight line take our reference as pareto plot
2 826 1	We can also are ad plot	take our reference as parelo plot
ts s	a a returned to and country	in to transfer end
i-I mai di	> Transformation : To conve	townsform, power (Sq, Bgot);
f"transf-	Log triansform, successional	transform, power (Sq, 8 gpt);
power tre-	Box-cox transform, Yea-	johnson townsform.
Tu)	Lition and and congare le	The come lede any other deline
(4)	dog transform: Take log	of values, if data is light skew
-	On Taking log we break to	Le interval into smaller ones.
3900	Lang log matter	Can't use log on the value
- 20	112081213	Marinima - Consultation
-	unear scale	DOMEN IN X WILL E FOOD
	ark.	
	Reciprocal (1/2) (c) S	Juane(x2) (a) Squt (52)
	(de	It skewed)
e)	Box-con transform: We	Examine values of 2 forom -5 to 5
-4-23	and we chose the optim	at value the close to normal dut
-		Not for negative numbers.
	$\chi_{i}^{\lambda} = \int \dot{\chi}_{i}^{\lambda} - 1 \lambda \neq 0$	719
22 44 -	$\int_{0}^{\pi} \ln(x_{i}) \lambda = 0$	ust in tuder in homers against the
1 1 1 1	l ln(xi) $\lambda=0$	attended to the sold all
		Limited 112 to 129 to 189 to 1
, -7 i -	1 maral and 1	is called it lay thought

	A Tri 15	2MOC	Page No. 9
		Date /	,
f)	Yeo-Johnson transform - Adjusts box cox $ \begin{array}{ccccccccccccccccccccccccccccccccccc$	& can	poly on -) Ve num
	$C[(x;+1)^{\lambda}-17]/\lambda \qquad \text{if } \lambda \neq 0.$	2;≥0	The state of the s
144	$\chi_i^2 = \lim_{x \to +\infty} \chi_i + \lim_{x \to +\infty} \chi_i = 0$	7: 20	escolles.
	1-[(-xi+1)2-2-17/(2-2) 1 = 2 = 2	2: (0	Leen a
	$\left(-\ln(-x_i+1)\right) \text{if } \lambda=2.$	1:40	1
.*	And the Market of the Market o	1.7	13
4	Cornoulli Distribution - Binary outcome,	oithon S	ucas or failure
4.15	$PMF = P(X = x) = P^{2}(1-p)^{1-2} = P = Probable$	1688	54
2.30	P(X=0)=0.58 b(x=1)=0.8 1-0=00000 001	niture.	5.2-
	$fm\bar{f} = P(x=x) = p^{x}(1-p)^{1-x}$; $p = prob of support of sup$	- DALBATÓ	+0 11)
*	for bevery classificat "like spam/kam, fram	dete fall	s in Rernoulli
	messia com 3 times 1	indipende	at events)
L	Binomial distribution - Number of suces	esina	fixed number
incl	Bunomial distrukution: Number of sucess pandent bernoulli touals, with 2 possible of	utcomes	(suces & failure
-	where probability of suces is constant for	each to	ual.
- 353	It has 2 parameters n= no of truals &	p → proto	bility of sucess.
	-		
8)	No-one watches out of 3 people = 1/8 (NNN) I out of 3 watch = 3/8 [YNN, NNY, NYN] B	() - 2	-22=3=8 outon
198 C	No-one watches out of 3 people = 1/8 (NNN)) ///	. 4
ni a	I out of 3 watch = 3/8 [YNN, NNY, NYN] B	ut what	if we can't get
1	Lout of 3 watch = 3/8 [YYN, N, YY, YNY]	he sam	ple space
1	Bout of 3 watch 1/8 LYYY We wil	l use th	e formula
1	POP of Binomial = P(X=x) = "CxP" (1-f		n=no of truels
	Jour of 3 watch = 3/8 [YNN, NNY, NYN] B Lout of 3 watch = 3/8 [YYN, N, YY, YNY] To Bout of 3 watch = 1/8 [YYY] We will POP of Binomial = P(X = x) = MC_2Px(1-16) 3(2)(2)'= 3/8	anak sor	= descred result
	2(2) (2) 8	WALLE TO	(1) 1217
. ,	large productional results in the state of t	ti ch	lateil mount
	Sampling distrubution - It is PD that describe of a Sample statistic (such as mean, varia	the sta	Miles
0	of a Sample statistic Court as mean, varia	rices jo	trick 16
	independent samples of same site of proper	→ ×, V	wristelity of a
A Service	independent samples of Same Size of property Take So samples x x x x x x x x x x x x x x x x x x x	→ ×2	Sample statistic
Mexico	Salary noth		
305	Salary 100th Somple Sample Sample of Sample mean Take	XIDO	1
1.2	1. (1,12/3 1/100 of sample mean Take	2 means	lavance

Central limit theorem: CLT states distrubition of sample means of large no of independent & identical distributed random Variable will approach to normal distrubution, sugardless of underlying distribution of the variables.

Take sample distribut " & x x x x ... x 30 samples

of sample mean x, , x2, x3 · × 100 ← 100 times Now if we plot \$1, \$2 : X00 > It will be norm isouspecture of original distrubut, skew, Cognormal, foretrit The mean will be the same, the vocance will reduce by a factor of n i e no of samples of sample n > 30] Wow we cannot say that is of sample mean is same or appron to population mean. We can our confidence interval and say we are 95 % Swel that our mean will lies between [u-25, u+25] as This holds 95% of our value so we are sure about mean it ways to calcutal CI+ > 2-procedure [pop-std available = + t-procedure [pop -> Stol Xavailable not Confidence intervals is created for parameters & not statistion Statutus help us get the confidence interval for a parameter -> Find confidence interval using 2- procedure: Condit 1 - , 2- procedure formula used where data is randomly selected, Standard deviation is known for a populat , data is normal CT = point letimate = Zaz = 1-a=Confidence level 95/ confidence o = Std populat n n = sample size =) say 100 * 4 t- procedure: It follows students T distribut a Normal parameter is DOF= n-1; if n 1 Student & Normal It has fatter tails, as s'is also uncertain so grange increases of CI; tx > Zx But n1 tx = Zx Point est sample mean

Whether the data at Rand supports sufficiently support a particular hypothesis. It allows us to make probabilistic statements about population parameter. Say pumple so will there sales 1, what is the prop of this hypothesis Ho 4 Null hypothesis (Status que) - It is a statement which assumes no significant change It serves as starting point. It is the assumption of no effect until proven otherwise ie lawyer foroves that he is the guilty one & null hypothesis (Ho) ear he is onot quity / not committed owne. The purpose of hypothesis and prove in the favour of atternative hypothesis (Hi or Ha) OR Hat > Alternative hypothesis (Kesearch hypothesis H, or Ha) - Pt contradition null hypothesis and claims significant negult. We need to contradict and prove the man is accused. Always any one of them is touce either null or alternative hypothesis We try to gather evidence to reject the null hypothesis is toure. We couldn't gather much evidence to prove him accused That does not mean he is completely not quilty Rejection region approach: 1. Formulate a Null and alternate hypothesis 2. Select significant We (0.05 or 0.01) of sugesting mill hypothesis
3. Check assumptions (example distribution) Normal, laynormal, category 4. Recide apperaporate test (Z-test, I-test, chi-quare test, ANOVA 5. State the relevant test statistic It () pop son categorical 7. Reject or not reject mill hypothesis. 8. Interpret the result :

- 8	4 (0 000	and a dist	tu
8)	Sugare company has training program. Avg. pro	i di	J
	Suppose company has training program . A vg. prus before program is sounite with a std. of s	unus	
month	TANK TO THE TANK THE		
	a mar ad the est incomplete the		1
74	Significantly productively has increased ?	- CC	
uster Sel	Ho = 4 = 50 Nutl hypothesy x = 53	m= 30	
1/	Significantly productively has implianted. Ho = 11 = 50 ~ Nutl hypothesis 12 = 50 Ha = 11 > 50 ~ Alternate hypothesis 12 = 50 2) d = 0.05 ~ 5%.	tool	
	7-	- Lest	
NA:	3) Normality is valid as sample > 30/ pop std. (0)	u Kny	wn
4 / 4	1) Z - best	Francis	2
	2 - test 2 - 50 - 2.28 / 2	D.03	2-tollo
	5) X= X Y = 55/30	4 7	
Std 1	pormal variet of on reject region 0 Usin	dZ(1.	65)
100	4) Z - test 5) Z = x - y = 53 - 50 = 3.28 Normal varieb \(\tau \) / \(\tau \)	earon	80
OH CO	165/3.28 up can great the null hypothes	es wi	1695/
4.	(3) Reject Ho	(B	nfidence
(0)	1 = C0	at twee	Cl.
3/	n = 80, $n = 70$, $n = 70$, $n = 10$, $n =$	11 hypot	tesu
PHE	1) Ho-12=30, Ha - M730 2) A=0.05	The state of	000
	2) d=0.05 -0.025	No.	50,00
	3) Normality V Cisio	77	
<u>b</u>	4) Ztest, 2-49-50 =-1.58 21-1.58 4/540 -20/2-1.58	2/1/2	11
September 1	7/090 -1.96	1.96	7
1	The 1 als Tune 2 Garage - Mootheris MOX ->		
Late Land		s Train CY	4.1
	Type 1 is false positive; rejects Howken Hoxx	Ho tous	Ha towe
(B)	Type ? u talse negative, willege to whim I would	Type I	1
	Pouls of committing Type IT would is B Rejet to		TypeTi
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