Weekly Report 8

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Topics Covered

1 PRINCIPAL COMPONENTS ANALYSIS

It is a linear projection method. In this approach, we seek to find a mapping from the d – dimensional space to a k – dimensional space (k < d). This mapping must ensure that there is a minimum loss of information. Given a vector x and directional vector w, the projection of x on w is given as

$$z = w^T x, z \in \mathbb{R}, w, x \in \mathbb{R}^d$$

PCA is unsupervised learning algorithm. In this approach, the only thing we care about is variance. The principal component, denoted as w_1 is defined as direction for which the projection $z_1=w_1^Tx$ has maximum variance. Assuming x to be a random vector (composed of random variables) i.e. $x=(x_1,x_2,...,x_d)$. We define covariance matrix of x to be Σ , where $\Sigma[i,j]=cov(x_i,x_j)$. It is known fact that covariance matrix is a random vector is a positive semi-definite matrix i.e. all the eigen values of covariance matrix are non-zero. In this setting, the variance of projection z_1 is given as $var(z_1)=w_1^T\Sigma w_1$. To find the principal component, we have to maximize this variance under constraint that $\|w_1\|=1$. We transform this into Lagrange problem.

$$w_1 = \operatorname*{argmax}_{w} (w^T \Sigma w - \alpha (w^T w - 1))$$

After differentiating and putting the derivative to zero, we get

$$\Sigma w_1 = \alpha w_1$$
$$var(z_1) = w_1^T \Sigma w_1 = \alpha w_1^T w_1 = \alpha$$

The equations obtained implies that the principal component is an eigenvector of the covariance matrix Σ corresponding to eigenvalue α , which is also the variance of the projection. Hence, as principal component gives the direction of maximum variance, it can be concluded that the eigenvector of covariance matrix Σ corresponding to largest eigenvalue $\lambda_1 = \alpha$ is the principal component.

The second principal component w_2 must be the direction of maximum variance given that it is orthogonal to the first principal component w_1 . Therefore, transforming this to Lagrange problem we have,

$$w_2 = \operatorname*{argmax}_w \left(w^T \Sigma w - \alpha (w^T w - 1) - \beta (w^T w_1 - 0) \right)$$

Differentiating wrt to w and putting the derivative to zero, we get

$$2\Sigma w_2 \,-\, 2\alpha w_2 \,-\, \beta w_1 \,=\, 0$$

Pre-multiplying with \boldsymbol{w}_1^T , we have,

$$2w_1^T \Sigma w_2 - 2\alpha w_1^T w_2 - \beta w_1^T w_1 = 0$$

Now, $w_1^T w_2 = 0$ as both the directions are orthogonal. Consider $s = w_1^T \Sigma w_2$. As s is a scalar, we have

$$w_1^T \Sigma w_2 = (w_1^T \Sigma w_2)^T = w_2^T \Sigma w_1 = \lambda_1 w_2^T w_1 = 0$$

This implies $\beta = 0$. Hence, we have

$$\Sigma w_2 = \alpha w_2$$
$$var(z_2) = w_2^T \Sigma w_2 = \alpha w_2^T w_2 = \alpha$$

Therefore, it can be concluded that the second principal component is the eigenvector of covariance matrix Σ corresponding to second largest eigenvalue $\lambda_2 = \alpha$. It is guaranteed that w_1 and w_2 will be orthogonal as they are eigenvectors of two distinct eigenvalue of a

symmetric matrix Σ . Similarly, we can find other principal components. If the input dimensions are uncorrelated already, then the rank of covariance matrix will be d. Therefore, we will have d distinct eigenvalues hence eigenvectors. In such a scenario, using PCA is wasteful. However, if it is known that the input dimensions are correlated to an extent (like in image and speech processing tasks), then we will have only k (< d) principal components and thus a reduction in dimensionality is achieved. We can also choose number of principal components we want depending on how much variance we want to lose. Suppose we have k non-zero eigenvalues (sorted in decreasing order), if we select the first m eigenvalues (hence first m principal components) only, then fraction of variance retained is given as

$$\frac{\lambda_1 + \lambda_2 + \dots + \lambda_m}{\lambda_1 + \lambda_2 + \dots + \lambda_m + \dots + \lambda_k}$$

2 LINEAR DISCRIMINANT ANALYSIS

It is a linear projection, supervised method for classification problem. For sake of simplicity, let's have only two classes C_1 and C_2 . We are interested in the direction w such that after projection in this direction, the instances of the two classes are as well separated as possible. In this case, we have a dimensionality reduction from d to 1. Let w be the required direction. Then, projection of instance x is given as $z = w^T x$. Let μ_1, m_1 be mean of instances belonging to class C_1 before and after the projection and μ_2, m_2 for class C_2 . Let the training set be given as $\{x^{(i)}, y^{(i)}\}$ where $y^{(i)} = 1$ if $x^{(i)}$ belongs to class C_1 and $y^{(i)} = 0$, if $x^{(i)}$ belongs to C_2 .

$$\begin{split} m_1 &= \frac{\sum_{i=1}^N w^T x^{(i)} y^{(i)}}{\sum_{i=1}^N y^{(i)}} = w^T \mu_1 \\ m_2 &= \frac{\sum_{i=1}^N w^T x^{(i)} (1 - y^{(i)})}{\sum_{i=1}^N (1 - y^{(i)})} = w^T \mu_2 \end{split}$$

The scatter of sample (representing the spread) from \mathcal{C}_1 and \mathcal{C}_2 after the projection is defined as

$$\begin{split} s_1 &= \sum_{i=1}^N \left(\left(\boldsymbol{w}^T \boldsymbol{x}^{(i)} - \boldsymbol{m}_1 \right)^2 \cdot \boldsymbol{y}^{(i)} \right) = \sum_{i=1}^N \left(\boldsymbol{w}^T \big(\boldsymbol{x}^{(i)} - \boldsymbol{m}_1 \big) \big(\boldsymbol{x}^{(i)} - \boldsymbol{m}_1 \big)^T \boldsymbol{y}^{(i)} \boldsymbol{w} \right) = \boldsymbol{w}^T \boldsymbol{S}_1 \boldsymbol{w} \\ s_2 &= \sum_{i=1}^N \left(\left(\boldsymbol{w}^T \boldsymbol{x}^{(i)} - \boldsymbol{m}_2 \right)^2 \cdot (1 - \boldsymbol{y}^{(i)}) \right) = \sum_{i=1}^N \left(\boldsymbol{w}^T \big(\boldsymbol{x}^{(i)} - \boldsymbol{m}_1 \big) \big(\boldsymbol{x}^{(i)} - \boldsymbol{m}_1 \big)^T \big(1 - \boldsymbol{y}^{(i)} \big) \boldsymbol{w} \right) = \boldsymbol{w}^T \boldsymbol{S}_2 \boldsymbol{w} \end{split}$$

Here, S_1 and S_2 are the within-class scatter matrix for C_1, C_2 respectively. Now, several implementations of LDA exists depending on how we define the criterion of being well-separated. One variation is the *Fisher's linear discriminant* in which we expect mean of the classes to be as far as possible and scatter of each class to be as less as possible. Therefore, we maximize the following function.

$$J(w) = \frac{(m_1 - m_2)^2}{s_1^2 + s_2^2}$$

We have, $(m_1-m_2)^2=(w^T\mu_1-w^T\mu_2)^2=w^T(\mu_1-\mu_2)(\mu_1-\mu_2)^Tw=w^TS_Bw$, where $S_B=(\mu_1-\mu_2)(\mu_1-\mu_2)^T$ is the between-class scatter matrix. For the denominator part, we have $s_1^2+s_2^2=w^TS_1w+w^TS_2w=w^T(S_1+S_2)w=w^TS_Ww$. Hence, the objective function reduces to

$$J(w) = \frac{w^T S_B w}{w^T S_W w} = \frac{|w^T (\mu_1 - \mu_2)|^2}{w^T S_W w}$$

Differentiating wrt w and putting the derivative to 0, we get $w = S_w^{-1}(\mu_1 - \mu_2)$.

Novel Ideas

In my research project on 'Classification of Autism Disorder using Deep Learning', I used both *PCA* and *LDA*: *PCA* for identifying important modules of the brain and *LDA* for classification purpose. In that, we got test accuracy of about 59.43% which was not the best (best test accuracy was observed for *SVM* classifier). I think this is because in LDA, the dimensions are reduced to unity and hence model becomes quite simpler. You can see the project at https://github.com/utkarsh512/fMRI-classification-of-ASD/