





Lecture 2.3 - Maximum Likelihood

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(Bishop 1.2.3 - 1.2.5)

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Maximum Likelihood Principle

- ▶ Dataset $D = (x_1, x_2, ..., x_N)$ of N independent observations.
- Likelihood of the dataset:
- Maximum likelihood principle: the most likely "explanation" of D is given by \mathbf{w}_{ML} which maximizes the likelihood function

$$\mathbf{w}_{\mathrm{ML}} = \operatorname{arg\,max} \rho(\rho(\omega))$$

• i.i.d. assumption: each $x_i \in D$ is independently distributed according to the same distribution, conditioned on **w**.

 $x \sim p(x | w)$

If i.i.d., joint distribution

$$p(D|\mathbf{w}) = p(x_1, x_2, ..., x_N|\mathbf{w}) = \prod_{i=1}^{N} p(\mathbf{x}_i | \mathbf{w})$$

Maximum Likelihood Estimation

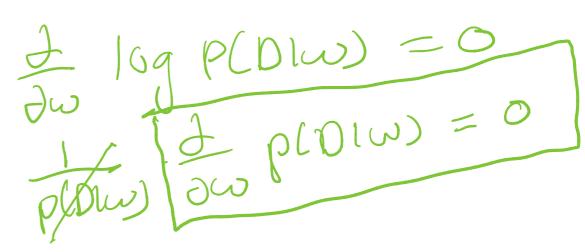


Maximum likelihood estimation w_{ML}

$$\mathbf{w}_{\mathrm{ML}} = \arg\max_{\mathbf{w}} p(D|\mathbf{w}) = \arg\max_{\mathbf{w}} \prod_{i=1}^{N} p(x_i|\mathbf{w})$$

numerical underflow/overflow

▶ How do we maximize?



Maximize log-likelihood instead:

$$\mathbf{w}_{\mathrm{ML}} = \arg\max_{\mathbf{w}} \prod_{i=1}^{N} p(x_i | \mathbf{w}) = \arg\max_{i=1}^{N} \log \prod_{i=1}^{N} \omega$$

$$= \arg\max_{\mathbf{w}} \sum_{i=1}^{N} p(x_i | \mathbf{w}) = \arg\max_{i=1}^{N} \log \min_{\mathbf{w}} \omega$$

From function: $E(D; \mathbf{w}) = -\log p(D|\mathbf{w}) = -\sum_{i=1}^{N} \log p(x_i|\mathbf{w})$

• i.i.d. Gaussian distributed real variables $D = (x_1, x_2, ..., x_N)$

$$p(x|\mathbf{w}) = \mathcal{N}(x|\mu, \sigma^2) \qquad p(D|\mu, \sigma^2) = \frac{1}{(2\pi\sigma^2)^{N/2}} \prod_{i=1}^{N} \exp\left[-\frac{1}{2\sigma^2}(x_i - \mu)^2\right]$$

$$p(\chi_i, \chi_2, \dots, \chi_N) M_i \sigma^2) = p(\chi_i | M_i \sigma^2) p(\chi_2 | \mu_i, \sigma^2) . \dots$$

$$\log \chi^{a} \cdot \text{Log likelihood} \qquad \mathcal{N}$$

$$\log p(D|\mu, \sigma^2) = \log \left(2\pi\sigma^2\right) + \sum_{i=1}^{N} \log p(D|\mu, \sigma^2) = \log \left(2\pi\sigma^2\right) + \sum_{i=1}^{N} \log p(\chi_i - \mu_i)^2$$

$$= -\frac{\mathcal{N}}{2} \left(\log \left(2\pi\sigma^2\right) + \sum_{i=1}^{N} -\frac{1}{2\sigma^2} \left(\chi_i - \mu_i\right)^2\right)$$

Estimate model parameters: μ_{ML} , $\sigma_{ML}^2 = \underset{\mu,\sigma^2}{\operatorname{argmax}} \log p(D \mid \mu, \sigma^2)$

$$\frac{\partial}{\partial u} \log p(D/M, \sigma^2) = 0$$
Solve for $u \rightarrow MmL$

log likelihood:

$$\log p(D|\mu,\sigma^2) = -\frac{N}{2}\log 2\pi\sigma^2 - \frac{1}{2\sigma^2}\sum_{i=1}^N (x_i-\mu)^2$$
 Maximum Likelihood solution for μ

Maximum Likelihood solution for μ

$$\frac{\partial}{\partial \mu} \log p(D|\mu, \sigma^2) = \frac{1}{2\sigma^2} \sum_{b=1}^{N} 2/(\chi_b - \mu) = 0$$

$$\sum_{b=1}^{N} (\chi_b - \mu) = 0$$

 $\frac{N}{\sum_{i \geq 1} M} = \sum_{i \geq 2}^{N} X_i$

sample mean

log likelihood:

$$\log \text{ likelihood:}$$

$$\log p(D|\mu,\sigma^2) = -\frac{N}{2}\log 2\pi\sigma^2 - \frac{1}{2\sigma^2}\sum_{i=1}^N(x_i-\mu)^2$$

$$\text{Maximum Likelihood solution for }\sigma^2$$

Maximum Likelihood solution for
$$\sigma^2$$

$$\frac{\partial}{\partial \sigma^2} \log p(D|\mu, \sigma^2) = -\frac{N}{2} \frac{1}{2\pi 6^2} \cdot 2\pi + \frac{1}{264} \sum_{c=0}^{N} (\chi_c - \mu)^2$$

$$\frac{26^4}{2} - N \sigma^2 + \sum_{c=1}^{N} (\chi_c - \mu)^2 = 0$$

$$\frac{1}{2} = \frac{1}{N} \sum_{i=1}^{N} (x_i - y_i)^2$$
 sample variance

How well do the ML estimators represent the true

parameters?
$$p(D|\mu, \sigma^2) = \frac{1}{(2\pi\sigma^2)^{N/2}} \prod_{i=1}^{N} \exp\left[-\frac{1}{2\sigma^2} (x_i - \mu)^2\right]$$

• If I draw multiple datasets, what is the expected value of $\mu_{\rm ML}$?

$$\mathbb{E}_{D \sim p(D|\mu,\sigma^{2})}[\mu_{\mathrm{ML}}] = \mathbb{E}\left[\frac{1}{N}\sum_{i=1}^{N}x_{i}\right] = \frac{1}{N}\sum_{i=1}^{N}\mathcal{F}_{O \sim p(D|\mu,\sigma^{2})}[\chi_{i}]$$

$$= \frac{1}{N}\sum_{i=1}^{N}\mathcal{F}_{\chi_{i}} - p(\chi_{i}|\mu,\sigma^{2})[\chi_{i}] = \frac{1}{N}\sum_{i=1}^{N}\mathcal{F}_{O \sim p(D|\mu,\sigma^{2})}[\chi_{i}]$$

$$= \frac{1}{N}\sum_{i=1}^{N}\mathcal{F}_{\chi_{i}} - p(\chi_{i}|\mu,\sigma^{2})[\chi_{i}] = \frac{1}{N}\sum_{i=1}^{N}\mathcal{F}_{O \sim p(D|\mu,\sigma^{2})}[\chi_{i}]$$

Bias of estimator:

$$\mathbb{E}[\mu_{\mathrm{ML}}] - \mu = 0$$

ML estimate of the variance:

$$\mathbb{E}_{D \sim p(D|\mu, \sigma^2)}[\sigma_{ML}^2] = \mathbb{E}\left[\frac{1}{N} \sum_{i=1}^{N} \left(x_i - \frac{1}{N} \sum_{n=1}^{N} x_n\right)^2\right]$$

ML estimate of the variance:

$$\mathbb{E}_{D \sim p(D|\mu,\sigma^2)}[\sigma_{ML}^2] = \mathbb{E}\left[\frac{1}{N} \sum_{i=1}^{N} \left(x_i - \frac{1}{N} \sum_{n=1}^{N} x_n\right)^2\right]$$

$$= \frac{1}{N} \sum_{i=1}^{N} \mathbb{E} \left[(xi - \frac{1}{N} \sum_{n=1}^{N} x_n)^2 \right]$$

ML estimate of the variance:

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$$= \frac{1}{N} \sum_{i=1}^{N} \mathbb{E}\left[(x_i - \frac{1}{N} \sum_{n=1}^{N} x_n)^2\right] = \frac{1}{N} \sum_{i=1}^{N} \mathbb{E}\left[x_i^2 - \frac{2x_i}{N} \sum_{n=1}^{N} x_n + \frac{1}{N^2} \sum_{m=1}^{N} \sum_{n=1}^{N} x_m x_n\right]$$

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$$= \frac{1}{N}\sum_{i=1}^N \mathbb{E}\left[(x_i - \frac{1}{N}\sum_{n=1}^N x_n)^2\right] = \frac{1}{N}\sum_{i=1}^N \mathbb{E}\left[x_i^2 - \frac{2x_i}{N}\sum_{n=1}^N x_n + \frac{1}{N^2}\sum_{m=1}^N \sum_{n=1}^N x_m x_n\right]$$

$$= \frac{1}{N}\sum_{i=1}^N \left\{\mathbb{E}\left[x_i^2\right] - \frac{2}{N}\sum_{n=1}^N \mathbb{E}\left[x_i x_n\right] + \frac{1}{N^2}\sum_{m=1}^N \sum_{n=1}^N \mathbb{E}\left[x_m x_n\right]\right\}$$

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$$\mathbb{E}[x_i x_j] = \begin{cases} n^2 + \sigma^2 & \text{if } i = j \quad \text{Covex.}, x_i J = |E(x_i)J - |E(x_i)J|^2 = \sigma^2 \\ n^2 & \text{if } i \neq j_{COve(x_i)}, x_j J = |E(x_i x_j)J - |E(x_i)J|E(x_j)J = \sigma \end{cases}$$

For data generated from

$$p(D|\mu,\sigma^2) = \frac{1}{(2\pi\sigma^2)^{N/2}} \prod_{i=1}^{N} \exp\left[-\frac{1}{2\sigma^2} (x_i - \mu)^2\right]$$

ML gives biased estimator

$$\mathbb{E}[\sigma_{\mathrm{ML}}^2] = \frac{N-1}{N} = 2$$

Unbiased variance estimator:

$$\tilde{\sigma}^2 = \frac{N}{N-1} \sigma_{NL}^2 = \frac{1}{N-1} \sum_{i=1}^{N} (\chi_i - \mu_i)^2$$

$$1E[\tilde{\sigma}^2] = \sigma^2$$

Biased Maximum Likelihood Estimator

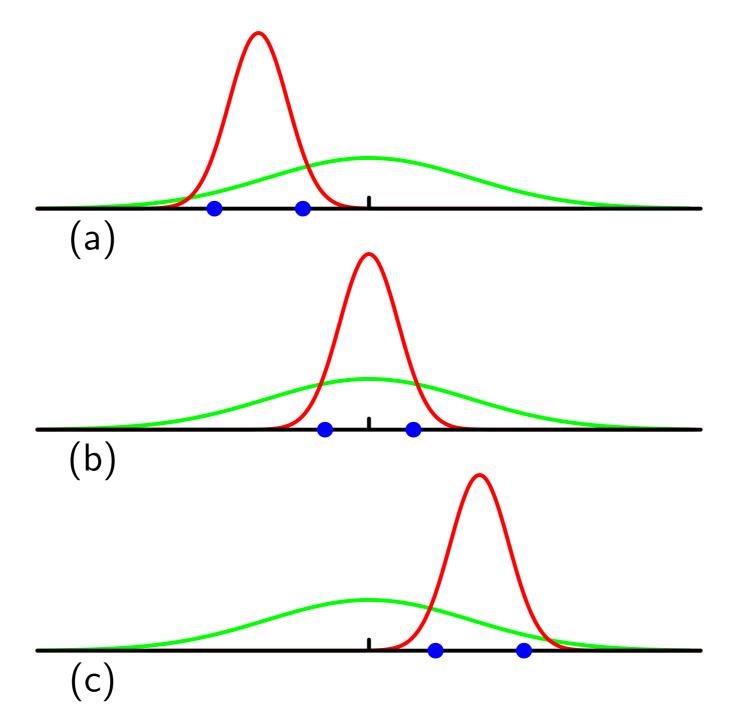


Figure: Bias in ML estimator for variance (Bishop 1.15)