



EN3150 Pattern Recognition Classification Part 02

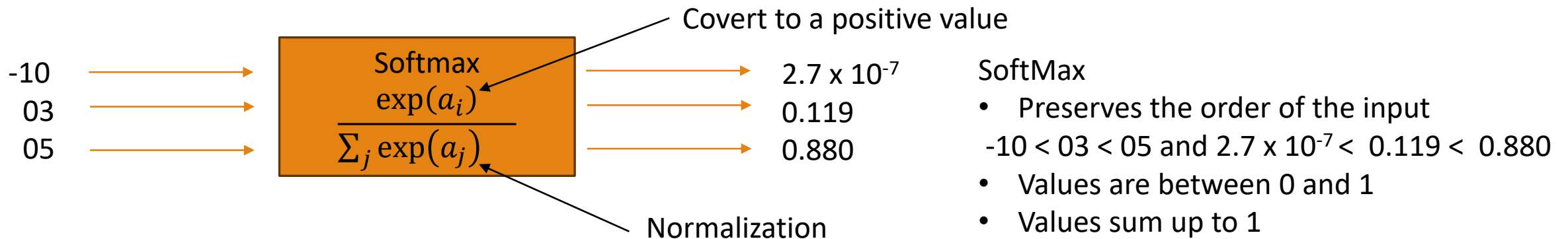
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Semester 5 – Batch 20.

Multinomial logistic regression

- Also known as multiclass logistic regression
- For multiple K classes, there are K linear functions

$$p(y_i = C_k | \mathbf{x}_i, \mathbf{W}) = \frac{\exp(w_{0k} + w_{1k}x_{1,i} + \dots + w_{Dk}x_{D,i})}{\sum_j \exp(w_{0j} + w_{1j}x_{1,i} + \dots + w_{Dj}x_{D,i})} = \frac{\exp(\mathbf{w}_k^T \mathbf{x}_i)}{\sum_j \exp(\mathbf{w}_j^T \mathbf{x}_i)} \quad \text{Softmax transformation}$$

- \mathbf{W} is K x (D+1) weight matrix and D is no of features.
- $p(y_i | x_i, \mathbf{W}) = \text{Cat}(y_i | \text{softmax}(\widetilde{\mathbf{W}}\mathbf{x}_i + \mathbf{b}))$. Here, $\mathbf{b} = [w_{01}, \dots, w_{0K}]^T$ is a K length vector*.



* \mathbf{b} can be added to first column by considering dummy feature equal to 1, $\mathbf{W} = [\mathbf{b} \quad \widetilde{\mathbf{W}}]$.

Multinomial logistic regression

$$\mathbf{c} = \begin{bmatrix} p(y_i = C_1 | x_i, \mathbf{W}) \\ p(y_i = C_2 | x_i, \mathbf{W}) \\ \vdots \\ p(y_i = C_K | x_i, \mathbf{W}) \end{bmatrix} = \begin{bmatrix} \frac{\exp(\mathbf{w}_1^T x_i)}{\sum_j \exp(\mathbf{w}_j^T x_i)} \\ \frac{\exp(\mathbf{w}_2^T x_i)}{\sum_j \exp(\mathbf{w}_j^T x_i)} \\ \vdots \\ \frac{\exp(\mathbf{w}_K^T x_i)}{\sum_j \exp(\mathbf{w}_j^T x_i)} \end{bmatrix}$$

Softmax transformation

Multinomial logistic regression

➤ How to learn weights? maximum likelihood estimation

➤ Let $a_{ik} = \mathbf{w}_k^T \mathbf{x}_i$ and $p(y_{ik} = 1 | \mathbf{x}_i, \mathbf{W}) = \tilde{y}_{ik}$

➤ Likelihood function

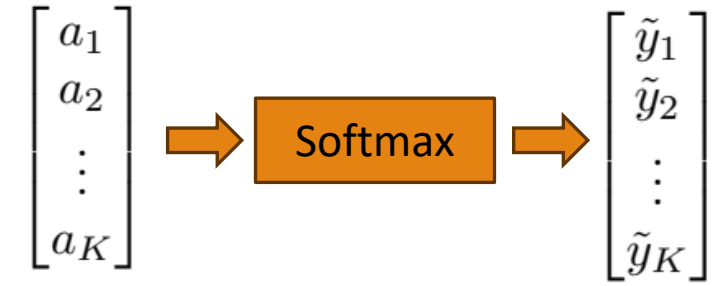
$$L(\mathbf{W}) = \prod_{i=1}^N \prod_{k=1}^K p(y_i = c_k | \mathbf{x}_i, \mathbf{W})^{c_{ik}} = \prod_{i=1}^N \prod_{k=1}^K \tilde{y}_{ik}^{c_{ik}}$$

➤ Negative likelihood function

$$NLL(\mathbf{W}) = - \sum_{i=1}^N \sum_{k=1}^K c_{ik} \log(\tilde{y}_{ik})$$

cross-entropy error function for the multiclass classification

Multinomial logistic regression



➤ Jacobian matrix

$$a_k = \mathbf{w}_k^T \mathbf{x}_i \quad \tilde{y}_k = \frac{\exp(\mathbf{w}_k^T \mathbf{x}_i)}{\sum_j \exp(\mathbf{w}_j^T \mathbf{x}_i)} = \frac{\exp(a_k)}{\sum_j \exp(a_j)} \quad \frac{\partial \tilde{y}_k}{\partial a_j} = \tilde{y}_k (\mathbb{I}(k = j) - \tilde{y}_j)$$

For 3 classes Jacobian matrix is given by

$$\mathbf{J} = \begin{bmatrix} \frac{\partial \tilde{y}_1}{\partial a_1} & \frac{\partial \tilde{y}_1}{\partial a_2} & \cdots & \frac{\partial \tilde{y}_1}{\partial a_K} \\ \frac{\partial \tilde{y}_2}{\partial a_1} & \frac{\partial \tilde{y}_2}{\partial a_2} & \cdots & \frac{\partial \tilde{y}_2}{\partial a_K} \\ \vdots & \vdots & \ddots & \vdots \\ \frac{\partial \tilde{y}_j}{\partial a_1} & \frac{\partial \tilde{y}_j}{\partial a_2} & \cdots & \frac{\partial \tilde{y}_j}{\partial a_K} \\ \vdots & \vdots & \ddots & \vdots \\ \frac{\partial \tilde{y}_K}{\partial a_1} & \frac{\partial \tilde{y}_K}{\partial a_2} & \cdots & \frac{\partial \tilde{y}_K}{\partial a_K} \end{bmatrix}$$

$$\mathbf{J} = \begin{bmatrix} \frac{\partial \tilde{y}_1}{\partial a_1} & \frac{\partial \tilde{y}_1}{\partial a_2} & \frac{\partial \tilde{y}_1}{\partial a_3} \\ \frac{\partial \tilde{y}_2}{\partial a_1} & \frac{\partial \tilde{y}_2}{\partial a_2} & \frac{\partial \tilde{y}_2}{\partial a_3} \\ \frac{\partial \tilde{y}_3}{\partial a_1} & \frac{\partial \tilde{y}_3}{\partial a_2} & \frac{\partial \tilde{y}_3}{\partial a_3} \end{bmatrix} = \begin{bmatrix} \tilde{y}_1(1 - \tilde{y}_1) & -\tilde{y}_1\tilde{y}_2 & -\tilde{y}_1\tilde{y}_3 \\ -\tilde{y}_2\tilde{y}_1 & \tilde{y}_2(1 - \tilde{y}_2) & -\tilde{y}_2\tilde{y}_3 \\ -\tilde{y}_3\tilde{y}_1 & -\tilde{y}_3\tilde{y}_2 & \tilde{y}_3(1 - \tilde{y}_3) \end{bmatrix}$$

$$\mathbf{J} = \frac{\partial \tilde{\mathbf{y}}}{\partial \mathbf{a}} = (\tilde{\mathbf{y}} \mathbf{1}^T) \odot (\mathbf{I} - \mathbf{1} \tilde{\mathbf{y}}^T)$$

When $k = j$, $\mathbb{I}(k = j) = 1$ else 0. $\tilde{\mathbf{y}} \mathbf{1}^T$ copies $\tilde{\mathbf{y}}$ across each column, and $\mathbf{1} \tilde{\mathbf{y}}^T$ copies $\tilde{\mathbf{y}}$ across each row. \mathbf{I} - identity matrix

Multinomial logistic regression

$$NLL(\mathbf{W}) = - \sum_{i=1}^N \sum_{k=1}^K c_{ik} \log(\tilde{y}_{ik})$$

$$\tilde{y}_{ik} = \frac{\exp(\mathbf{w}_k^T \mathbf{x}_i)}{\sum_j \exp(\mathbf{w}_j^T \mathbf{x}_i)} = \frac{\exp(a_{ik})}{\sum_j \exp(a_{ij})}$$

$$\frac{\partial \tilde{y}_k}{\partial a_j} = \tilde{y}_k (\mathbb{I}(k = j) - \tilde{y}_j)$$

$$a_{ik} = \mathbf{w}_k^T \mathbf{x}_i$$

➤ Consider i -th data sample

$$\nabla_{\mathbf{w}_k} NLL(\mathbf{W})_i = - \sum_{k=1}^K \frac{\partial NLL(\mathbf{W})_i}{\partial \tilde{y}_{ik}} \frac{\partial \tilde{y}_{ik}}{\partial a_{ik}} \frac{\partial a_{ik}}{\partial \mathbf{w}_k}$$

$$= - \sum_{k=1}^K c_{ik} \frac{\tilde{y}_{ik}}{\tilde{y}_{ik}} \tilde{y}_k (\mathbb{I}(k = i) - \tilde{y}_{ik}) \mathbf{x}_i = \sum_{k=1}^K (\tilde{y}_{ik} - c_{ik}) \mathbf{x}_i$$

➤ For all N data samples and all K classes

$$\mathbf{g}(\mathbf{W}) = \frac{1}{N} \sum_{i=1}^N \mathbf{x}_i (\tilde{\mathbf{y}}_i - \mathbf{c}_i)^T$$

$\mathbf{g}(\mathbf{W})$: $(D+1) \times K$ matrix
 \mathbf{x}_i : $(D+1) \times 1$ vector
 $(\tilde{\mathbf{y}}_i - \mathbf{c}_i)^T$: $1 \times K$ vector

Multinomial logistic regression

➤ Stochastic gradient descent

$$g(W) = \frac{1}{N} \sum_{i=1}^N \mathbf{x}_i (\tilde{\mathbf{y}}_i - \mathbf{c}_i)^T$$

➤ Update of weight (i is the sample index)

$$W_{i+1} \leftarrow W_i - \alpha \mathbf{x}_i (\tilde{\mathbf{y}}_i - \mathbf{c}_i)^T$$

➤ Batch gradient descent

$$W_{new} \leftarrow W_{old} - \alpha \frac{1}{N} \sum_{i=1}^N \mathbf{x}_i (\tilde{\mathbf{y}}_i - \mathbf{c}_i)^T$$

Multinomial logistic regression

- Hessian of the NLL for multinomial logistic regression is given by

$$H(\mathbf{W}) = \frac{1}{N} \sum_{i=1}^N (\text{diag}(\tilde{\mathbf{y}}_i) - \tilde{\mathbf{y}}_i \tilde{\mathbf{y}}_i^T) \otimes \mathbf{x}_i \mathbf{x}_i^T$$

- To develop a batch algorithm for the multiclass problem, we use the Newton-Raphson update.
- The IRLS algorithm involves evaluating the Hessian matrix

Probabilistic view of classification

- Probabilistic view of classification

- **Discriminative classifier**

- Directly fit the class posterior $p(y_i = C_k | \mathbf{x}_i, \boldsymbol{\theta})$
 - E.g., logistic regression, multi class logistic regression

- **Generative classifier**

- Model how to generate data using the conditional density $p(\mathbf{x}_i | y_i = C_k)$ and class priority $p(y_i = C_k)$. Then using Bayes rule

$$p(y = C_k | \mathbf{x}, \boldsymbol{\theta}) = \frac{p(y = C_k | \boldsymbol{\theta}) p(\mathbf{x} | y = C_k, \boldsymbol{\theta})}{\sum_{C_k'} p(y = C_k' | \boldsymbol{\theta}) p(\mathbf{x} | y = C_k', \boldsymbol{\theta})}$$

Generative models

➤ Generative classifier

generate the features \mathbf{x}
for each class

$$p(y = C_k | \mathbf{x}, \boldsymbol{\theta}) = \frac{p(y = C_k | \boldsymbol{\theta}) p(\mathbf{x} | y = C_k, \boldsymbol{\theta})}{\sum_{C_k'} p(y = C_k' | \boldsymbol{\theta}) p(\mathbf{x} | y = C_k', \boldsymbol{\theta})}$$

For two classes

$$\begin{aligned} p(y = C_1 | \mathbf{x}, \boldsymbol{\theta}) &= \frac{p(y = C_1 | \boldsymbol{\theta}) p(\mathbf{x} | y = C_1, \boldsymbol{\theta})}{p(y = C_1 | \boldsymbol{\theta}) p(\mathbf{x} | y = C_1, \boldsymbol{\theta}) + p(y = C_2 | \boldsymbol{\theta}) p(\mathbf{x} | y = C_2, \boldsymbol{\theta})} \\ &= \frac{1}{1 + p(y = C_2 | \boldsymbol{\theta}) p(\mathbf{x} | y = C_2, \boldsymbol{\theta}) / p(y = C_1 | \boldsymbol{\theta}) p(\mathbf{x} | y = C_1, \boldsymbol{\theta})} = \frac{1}{1 + \exp(-a(\mathbf{x}))} \\ &= \text{sigm}(a(\mathbf{x})) = \sigma(a(\mathbf{x})) \end{aligned}$$

$$a(\mathbf{x}) = \frac{p(y = C_1 | \boldsymbol{\theta}) p(\mathbf{x} | y = C_1, \boldsymbol{\theta})}{p(y = C_2 | \boldsymbol{\theta}) p(\mathbf{x} | y = C_2, \boldsymbol{\theta})}$$

equivalent form of the
posterior probabilities

logistic sigmoid function
squashing function

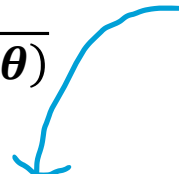
For multi-class problem it is softmax function

Generative models

➤ Generative classifier

$$p(y = C_k | \mathbf{x}, \boldsymbol{\theta}) = \frac{p(y = C_k | \boldsymbol{\theta}) p(\mathbf{x} | y = C_k, \boldsymbol{\theta})}{\sum_{C_k'} p(y = C_k' | \boldsymbol{\theta}) p(\mathbf{x} | y = C_k', \boldsymbol{\theta})}$$

generate the features \mathbf{x} for each class



➤ For two classes

$$p(y = C_1 | \mathbf{x}, \boldsymbol{\theta}) = \frac{p(y = C_1 | \boldsymbol{\theta}) p(\mathbf{x} | y = C_1, \boldsymbol{\theta})}{p(y = C_1 | \boldsymbol{\theta}) p(\mathbf{x} | y = C_1, \boldsymbol{\theta}) + p(y = C_2 | \boldsymbol{\theta}) p(\mathbf{x} | y = C_2, \boldsymbol{\theta})}$$

➤ Linear discriminant analysis* $\log p(\mathbf{x} | y = C_k, \boldsymbol{\theta}) = \tilde{\mathbf{w}}^T \mathbf{x} + \text{const}$ (linear function of \mathbf{x})

*Although names says discriminant this is a generative model

Generative models: Continuous inputs

- Class-conditional densities are Gaussian distributed

$$p(\mathbf{x}|y = Ck, \boldsymbol{\theta}) = \mathcal{N}(\mathbf{x}|\boldsymbol{\mu}_k, \Sigma_k) \triangleq \frac{1}{(2\pi)^{D/2} |\Sigma_k|^{1/2}} \exp \left[-\frac{1}{2} (\mathbf{x} - \boldsymbol{\mu}_k)^T \Sigma_k^{-1} (\mathbf{x} - \boldsymbol{\mu}_k) \right]$$

Multi-Variate Gaussian
 $\boldsymbol{\mu}_k$ = mean vector
 Σ_k = covariance matrix

- Assume that **same covariance matrix** is shared with all classes

$$p(\mathbf{x}|y = Ck, \boldsymbol{\theta}) = \mathcal{N}(\mathbf{x}|\boldsymbol{\mu}_k, \Sigma) \triangleq \frac{1}{(2\pi)^{D/2} |\Sigma|^{1/2}} \exp \left[-\frac{1}{2} (\mathbf{x} - \boldsymbol{\mu}_k)^T \Sigma^{-1} (\mathbf{x} - \boldsymbol{\mu}_k) \right]$$

Generative models: Continuous inputs

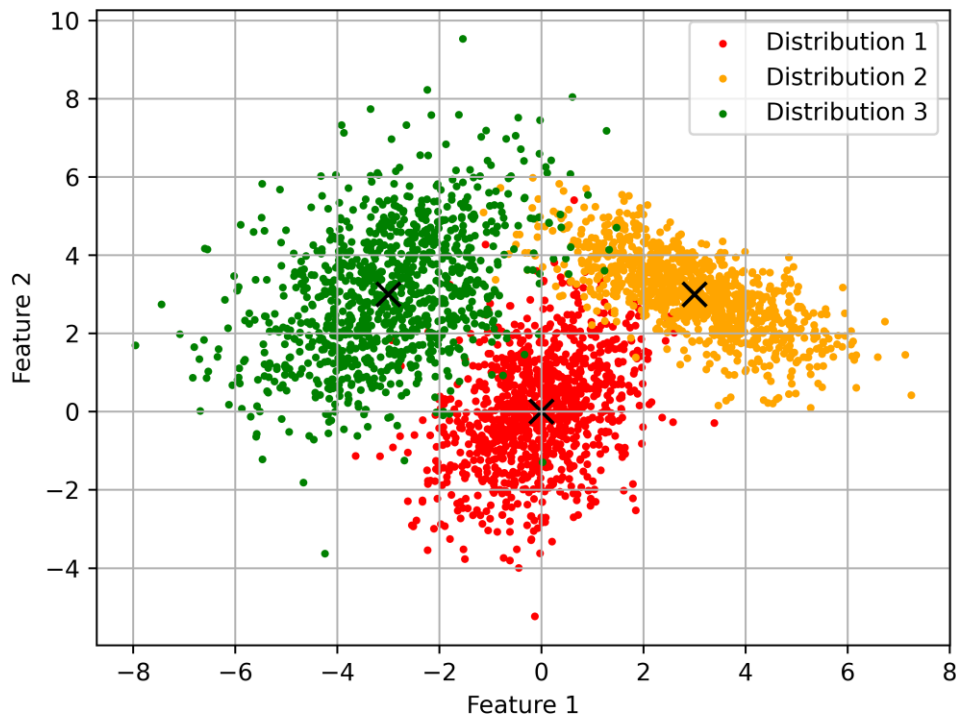
- Class-conditional densities are Gaussian distributed

$$p(\mathbf{x}|y = C_k, \boldsymbol{\theta}) = \mathcal{N}(\mathbf{x}|\boldsymbol{\mu}_k, \Sigma) \triangleq \frac{1}{(2\pi)^{D/2} |\Sigma|^{1/2}} \exp \left[-\frac{1}{2} (\mathbf{x} - \boldsymbol{\mu}_k)^T \Sigma^{-1} (\mathbf{x} - \boldsymbol{\mu}_k) \right]$$

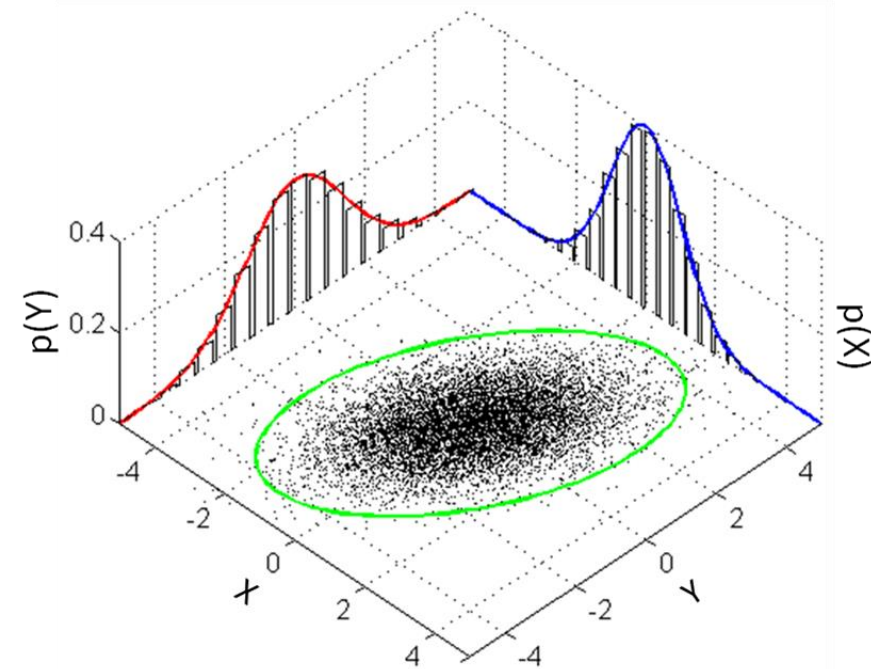
Multi-Variate Gaussian

$\boldsymbol{\mu}_k$ = mean vector

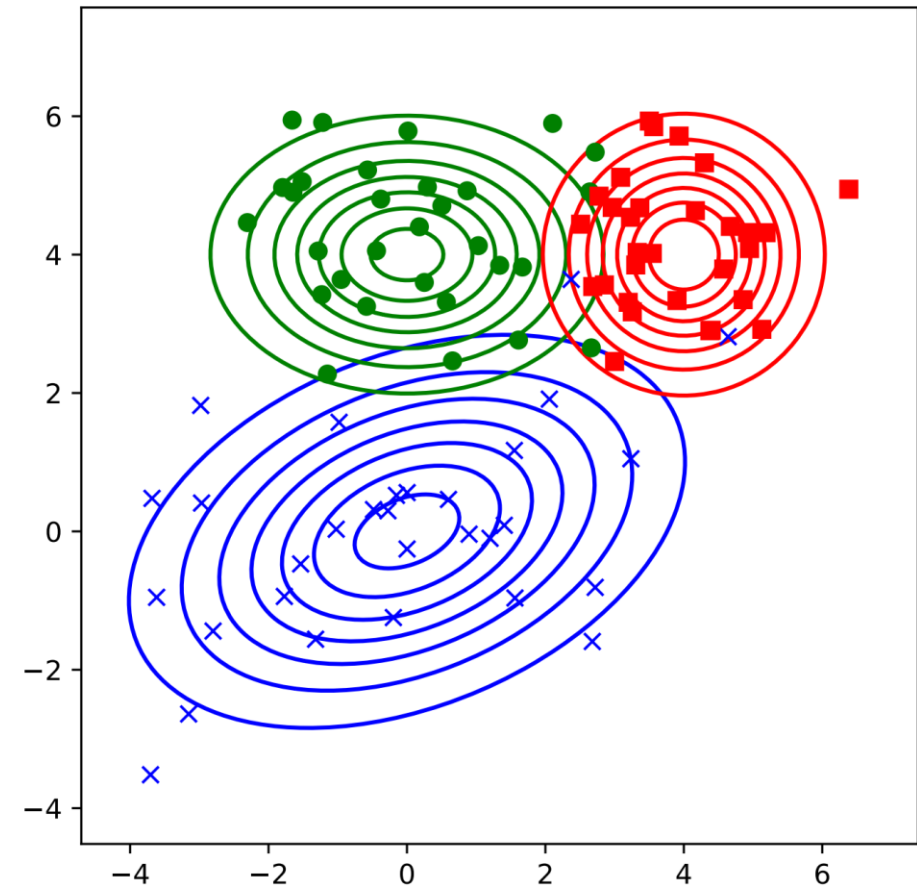
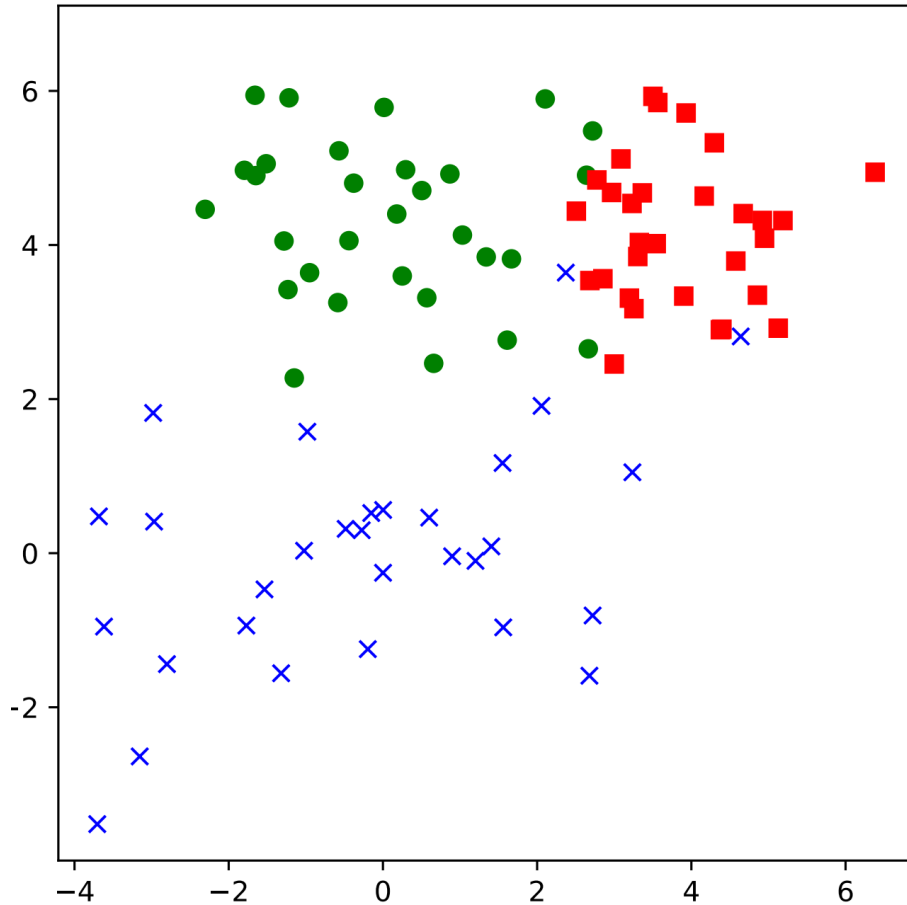
Σ = covariance matrix



$$\begin{aligned} \mu_1 &= \begin{bmatrix} 0 \\ 0 \end{bmatrix}, C_1 = \begin{bmatrix} 1 & 0.5 \\ 0.5 & 2 \end{bmatrix} \\ \mu_2 &= \begin{bmatrix} 3 \\ 3 \end{bmatrix}, C_2 = \begin{bmatrix} 2 & -1 \\ -1 & 1 \end{bmatrix} \\ \mu_3 &= \begin{bmatrix} -3 \\ 3 \end{bmatrix}, C_3 = \begin{bmatrix} 2 & 1 \\ 1 & 3 \end{bmatrix} \end{aligned}$$



Generative models: Continuous inputs



Generative models: Continuous inputs

- Class-conditional densities are Gaussian distributed

$$p(\mathbf{x}|y = C_k, \boldsymbol{\theta}) = \mathcal{N}(\mathbf{x}|\boldsymbol{\mu}_k, \Sigma) \triangleq \frac{1}{(2\pi)^{D/2}|\Sigma|^{1/2}} \exp\left[-\frac{1}{2}(\mathbf{x} - \boldsymbol{\mu}_k)^T \Sigma^{-1}(\mathbf{x} - \boldsymbol{\mu}_k)\right]$$

Multi-Variate Gaussian

$\boldsymbol{\mu}_k$ = mean vector

Σ = covariance matrix

- For two classes

$$p(y = C_1|\mathbf{x}, \boldsymbol{\theta}) = \frac{p(y = C_1|\boldsymbol{\theta})p(\mathbf{x}|y = C_1, \boldsymbol{\theta})}{p(y = C_1|\boldsymbol{\theta})p(\mathbf{x}|y = C_1, \boldsymbol{\theta}) + p(y = C_2|\boldsymbol{\theta})p(\mathbf{x}|y = C_2, \boldsymbol{\theta})}$$

independent of C_1

prior probability of class C_1

$$p(y = C_1|\boldsymbol{\theta}) = \pi_1$$

$$\log(p(y = C_1|\mathbf{x}, \boldsymbol{\theta})) = \log p(y = C_1|\boldsymbol{\theta}) + \log(p(\mathbf{x}|y = C_1, \boldsymbol{\theta})) + \text{constant}$$

$$= \log p(y = C_1|\boldsymbol{\theta}) + \log(\mathcal{N}(\mathbf{x}|\boldsymbol{\mu}_1, \Sigma)) + \text{constant}$$

$$= \log p(y = C_1|\boldsymbol{\theta}) + -\frac{1}{2}\log(|\Sigma|) - \frac{D}{2}\log(2\pi) \left[-\frac{1}{2}(\mathbf{x} - \boldsymbol{\mu}_1)^T \Sigma^{-1}(\mathbf{x} - \boldsymbol{\mu}_1)\right] + \text{constant}$$

$$= \log \pi_1 - \frac{1}{2}\log(|\Sigma|) - \frac{D}{2}\log(2\pi) \left[-\frac{1}{2}(\mathbf{x} - \boldsymbol{\mu}_1)^T \Sigma^{-1}(\mathbf{x} - \boldsymbol{\mu}_1)\right] + \text{constant}$$

Generative models: Continuous inputs

- Class-conditional densities are Gaussian distributed

$$p(\mathbf{x}|y = C_k, \boldsymbol{\theta}) = \mathcal{N}(\mathbf{x}|\boldsymbol{\mu}_k, \Sigma) \triangleq \frac{1}{(2\pi)^{D/2}|\Sigma|^{1/2}} \exp \left[-\frac{1}{2}(\mathbf{x} - \boldsymbol{\mu}_k)^T \Sigma^{-1}(\mathbf{x} - \boldsymbol{\mu}_k) \right]$$

Multi-Variate Gaussian

$\boldsymbol{\mu}_k$ = mean vector

Σ = covariance matrix

$$\log(p(y = C_k|\mathbf{x}, \boldsymbol{\theta})) = \log p(y = C_k|\boldsymbol{\theta}) + \log(p(\mathbf{x}|y = C_k, \boldsymbol{\theta})) + \text{constant}$$

$$= \log \pi_k - \frac{1}{2} \log(|\Sigma|) - \frac{D}{2} \log(2\pi) \left[-\frac{1}{2}(\mathbf{x} - \boldsymbol{\mu}_k)^T \Sigma^{-1}(\mathbf{x} - \boldsymbol{\mu}_k) \right] + \text{constant}$$

$$= \log \pi_k - \frac{1}{2} \boldsymbol{\mu}_k^T \Sigma^{-1} \boldsymbol{\mu}_k + \mathbf{x}^T \Sigma^{-1} \boldsymbol{\mu}_k + \text{constant} - \frac{1}{2} \mathbf{x}^T \Sigma^{-1} \mathbf{x}$$

$$= \boxed{\log \pi_k - \frac{1}{2} \boldsymbol{\mu}_k^T \Sigma^{-1} \boldsymbol{\mu}_k} + \mathbf{x}^T \boxed{\Sigma^{-1} \boldsymbol{\mu}_k} + \boxed{\text{constant} - \frac{1}{2} \mathbf{x}^T \Sigma^{-1} \mathbf{x}}$$

$$p(y = C_1|\mathbf{x}, \boldsymbol{\theta}) = \sigma(a(\mathbf{x})) \Rightarrow \log(p(y = C_k|\mathbf{x}, \boldsymbol{\theta})) = (\tilde{\mathbf{w}}^T \mathbf{x} + w_0)$$

$$\tilde{\mathbf{w}} = \Sigma^{-1} \boldsymbol{\mu}_k$$

$$w_0 = \log \pi_k - \frac{1}{2} \boldsymbol{\mu}_k^T \Sigma^{-1} \boldsymbol{\mu}_k$$

Due to shared covariance matrix assumption, the quadratic part $\mathbf{x}^T \Sigma^{-1} \mathbf{x}$ cancels off and $\log(p(y = C_k|\mathbf{x}, \boldsymbol{\theta}))$ is a linear function of \mathbf{x} . This is called linear discriminant analysis.

Generative models: Continuous inputs

➤ Maximum likelihood estimation

$$p(\mathbf{x}|y = C_k, \boldsymbol{\theta}) = \mathcal{N}(\mathbf{x}|\boldsymbol{\mu}_k, \boldsymbol{\Sigma}) \triangleq \frac{1}{(2\pi)^{D/2} |\boldsymbol{\Sigma}|^{1/2}} \exp \left[-\frac{1}{2} (\mathbf{x} - \boldsymbol{\mu}_k)^T \boldsymbol{\Sigma}^{-1} (\mathbf{x} - \boldsymbol{\mu}_k) \right]$$

$$p(y = C_k|\mathbf{x}, \boldsymbol{\theta}) = \frac{p(y = C_k|\boldsymbol{\theta})p(\mathbf{x}|y = C_k, \boldsymbol{\theta})}{\sum_{C_k'} p(y = C_k'|\boldsymbol{\theta})p(\mathbf{x}|y = C_k', \boldsymbol{\theta})}$$

$$p(y = C_k|\mathbf{x}, \boldsymbol{\theta}) \propto p(y = C_k|\boldsymbol{\theta})p(\mathbf{x}|y = C_k, \boldsymbol{\theta})$$

➤ For two classes

$$p(y = C_1|\mathbf{x}, \boldsymbol{\theta}) \propto p(y = C_1|\boldsymbol{\theta})p(\mathbf{x}|y = C_1, \boldsymbol{\theta})$$

$$p(y = C_2|\mathbf{x}, \boldsymbol{\theta}) \propto p(y = C_2|\boldsymbol{\theta})p(\mathbf{x}|y = C_2, \boldsymbol{\theta})$$

$$p(y = C_1|\mathbf{x}, \boldsymbol{\theta}) \propto \pi_1 p(\mathbf{x}|y = C_1, \boldsymbol{\theta})$$

$$p(y = C_2|\mathbf{x}, \boldsymbol{\theta}) \propto (1 - \pi_1) p(\mathbf{x}|y = C_2, \boldsymbol{\theta})$$

$$p(y = C_1|\mathbf{x}, \boldsymbol{\theta}) \propto \pi_1 \mathcal{N}(\mathbf{x}|\boldsymbol{\mu}_1, \boldsymbol{\Sigma})$$

$$p(y = C_2|\mathbf{x}, \boldsymbol{\theta}) \propto (1 - \pi_1) \mathcal{N}(\mathbf{x}|\boldsymbol{\mu}_2, \boldsymbol{\Sigma})$$

➤ likelihood function

$$p(\mathbf{c}|\pi_1, \boldsymbol{\mu}_2, \boldsymbol{\mu}_1, \boldsymbol{\Sigma}) = \mathcal{L} = \prod_i [p(y = C_1|\mathbf{x}_i, \boldsymbol{\theta})]^{c_i} [p(y = C_2|\mathbf{x}_i, \boldsymbol{\theta})]^{1-c_i} = \prod_i [\pi_1 \mathcal{N}(\mathbf{x}_i|\boldsymbol{\mu}_1, \boldsymbol{\Sigma})]^{c_i} [(1 - \pi_1) \mathcal{N}(\mathbf{x}_i|\boldsymbol{\mu}_2, \boldsymbol{\Sigma})]^{1-c_i}$$

Generative models: Continuous inputs

➤ Likelihood function

$$p(\mathbf{c}|\pi_1, \mu_2, \mu_1, \Sigma) = \mathcal{L} = \prod_i [p(y = C_1|\mathbf{x}_i, \boldsymbol{\theta})]^{c_i} [p(y = C_2|\mathbf{x}_i, \boldsymbol{\theta})]^{1-c_i} = \prod_i [\pi_1 \mathcal{N}(\mathbf{x}_i|\mu_1, \Sigma)]^{c_i} [(1-\pi_1) \mathcal{N}(\mathbf{x}_i|\mu_2, \Sigma)]^{1-c_i}$$

➤ Log likelihood function

$$\sum_i c_i \log [\pi_1 \mathcal{N}(\mathbf{x}_i|\mu_1, \Sigma)] + (1 - c_i) \log [(1-\pi_1) \mathcal{N}(\mathbf{x}_i|\mu_2, \Sigma)]$$

➤ First maximize with respect to π_1

$$\mathcal{L}_{\pi_1} = \sum_i c_i \log [\pi_1] + (1 - c_i) \log [(1-\pi_1)]$$

$$\frac{\partial \mathcal{L}_{\pi_1}}{\partial \pi_1} = 0 \quad \pi_1 = \frac{N_1}{N} = \frac{N_1}{N_1 + N_2}$$

Generative models: Continuous inputs

➤ Likelihood function

$$p(\mathbf{c}|\pi_1, \boldsymbol{\mu}_2, \boldsymbol{\mu}_1, \boldsymbol{\Sigma}) = \mathcal{L} = \prod_i [p(y = C_1|\mathbf{x}_i, \boldsymbol{\theta})]^{c_i} [p(y = C_2|\mathbf{x}_i, \boldsymbol{\theta})]^{1-c_i} = \prod_i [\pi_1 \mathcal{N}(\mathbf{x}_i|\boldsymbol{\mu}_1, \boldsymbol{\Sigma})]^{c_i} [(1-\pi_1)\mathcal{N}(\mathbf{x}_i|\boldsymbol{\mu}_2, \boldsymbol{\Sigma})]^{1-c_i}$$

➤ Log likelihood function

$$\sum_i c_i \log [\pi_1 \mathcal{N}(\mathbf{x}_i|\boldsymbol{\mu}_1, \boldsymbol{\Sigma})] + (1 - c_i) \log [(1 - \pi_1) \mathcal{N}(\mathbf{x}_i|\boldsymbol{\mu}_2, \boldsymbol{\Sigma})]$$

➤ First maximize with respect to $\boldsymbol{\mu}_1$

$$\mathcal{L}_{\boldsymbol{\mu}_1} = \sum_i c_i \log [\pi_1 \mathcal{N}(\mathbf{x}_i|\boldsymbol{\mu}_1, \boldsymbol{\Sigma})] = \sum_i c_i \left[-\frac{1}{2} (\mathbf{x}_i - \boldsymbol{\mu}_1)^T \boldsymbol{\Sigma}^{-1} (\mathbf{x}_i - \boldsymbol{\mu}_1) \right] + \text{const}$$

$$\frac{\partial \mathcal{L}_{\boldsymbol{\mu}_1}}{\partial \boldsymbol{\mu}_1} = 0 \quad \boldsymbol{\mu}_1 = \frac{1}{N_1} \sum_{i=1}^N c_i \mathbf{x}_i \quad \text{Similarly } \boldsymbol{\mu}_2 = \frac{1}{N_2} \sum_{i=1}^N (1 - c_i) \mathbf{x}_i$$

Generative models: Continuous inputs

➤ Likelihood function

$$p(\mathbf{c}|\pi_1, \boldsymbol{\mu}_2, \boldsymbol{\mu}_1, \boldsymbol{\Sigma}) = \mathcal{L} = \prod_i [p(y = c_1 | \mathbf{x}_i, \boldsymbol{\theta})]^{c_i} [p(y = c_2 | \mathbf{x}_i, \boldsymbol{\theta})]^{1-c_i} = \prod_i [\pi_1 \mathcal{N}(\mathbf{x}_i | \boldsymbol{\mu}_1, \boldsymbol{\Sigma})]^{c_i} [(1 - \pi_1) \mathcal{N}(\mathbf{x}_i | \boldsymbol{\mu}_2, \boldsymbol{\Sigma})]^{1-c_i}$$

➤ Log likelihood function

$$\sum_i c_i \log [\pi_1 \mathcal{N}(\mathbf{x}_i | \boldsymbol{\mu}_1, \boldsymbol{\Sigma})] + (1 - c_i) \log [(1 - \pi_1) \mathcal{N}(\mathbf{x}_i | \boldsymbol{\mu}_2, \boldsymbol{\Sigma})]$$

➤ First maximize with respect to $\boldsymbol{\Sigma}$

$$\mathcal{L}_{\boldsymbol{\Sigma}} = \sum_i c_i \log [\pi_1 \mathcal{N}(\mathbf{x}_i | \boldsymbol{\mu}_1, \boldsymbol{\Sigma})] = \sum_i -\frac{1}{2} c_i \log(|\boldsymbol{\Sigma}|) - \left[\frac{1}{2} (\mathbf{x}_i - \boldsymbol{\mu}_1)^T \boldsymbol{\Sigma}^{-1} (\mathbf{x}_i - \boldsymbol{\mu}_1) \right] + \text{const}$$

$$\frac{\partial \mathcal{L}_{\boldsymbol{\Sigma}}}{\partial \boldsymbol{\Sigma}} = 0 \quad \boldsymbol{\Sigma} = \frac{1}{N} \sum_{i=1}^N (\mathbf{x}_i - \bar{\mathbf{x}})(\mathbf{x}_i - \bar{\mathbf{x}})^T = \frac{1}{N} \left(\sum_{i=1}^N \mathbf{x}_i \mathbf{x}_i^T \right) - \bar{\mathbf{x}} \bar{\mathbf{x}}^T$$

Generative models: Continuous inputs

➤ Maximum likelihood estimation

$$\hat{\pi}_k = \frac{N_k}{N} \quad \hat{\boldsymbol{\mu}}_k = \frac{1}{N_k} \sum_{i: y_i=k}^N \mathbf{x}_i \quad \hat{\boldsymbol{\Sigma}}_k = \frac{1}{N_k} \sum_{i: y_i=k}^N (\mathbf{x}_i - \boldsymbol{\mu}_k) (\mathbf{x}_i - \boldsymbol{\mu}_k)^T$$

For tied variance ($\boldsymbol{\Sigma} = \boldsymbol{\Sigma}_k$)

$$\hat{\boldsymbol{\Sigma}} = \frac{1}{N} \sum_{i=1}^N (\mathbf{x}_i - \bar{\mathbf{x}}) (\mathbf{x}_i - \bar{\mathbf{x}})^T = \frac{1}{N} \left(\sum_{i=1}^N \mathbf{x}_i \mathbf{x}_i^T \right) - \bar{\mathbf{x}} \bar{\mathbf{x}}^T$$

Generative models: Continuous inputs

➤ Quadrature discriminant analysis

➤ We drop the shared covariance matrix assumption

$$p(\mathbf{x}|y = C_k, \boldsymbol{\theta}) = \mathcal{N}(\mathbf{x}|\boldsymbol{\mu}_k, \Sigma_k) \triangleq \frac{1}{(2\pi)^{D/2}|\Sigma_k|^{1/2}} \exp\left[-\frac{1}{2}(\mathbf{x} - \boldsymbol{\mu}_k)^T \Sigma_k^{-1}(\mathbf{x} - \boldsymbol{\mu}_k)\right]$$

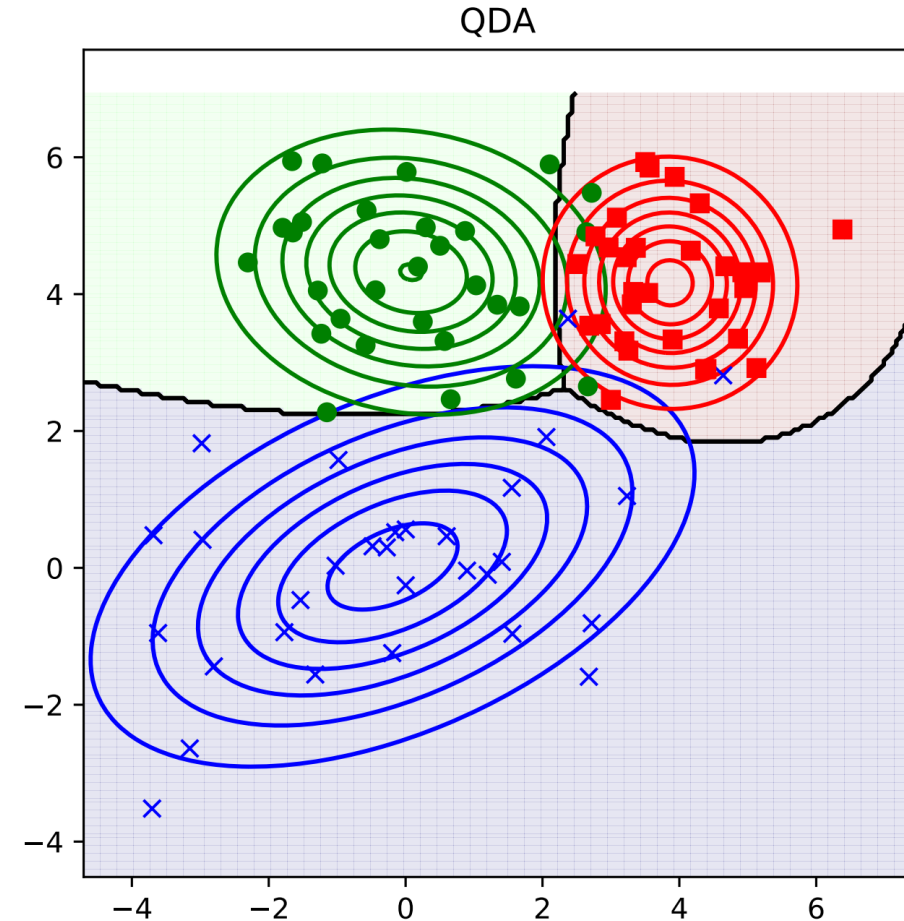
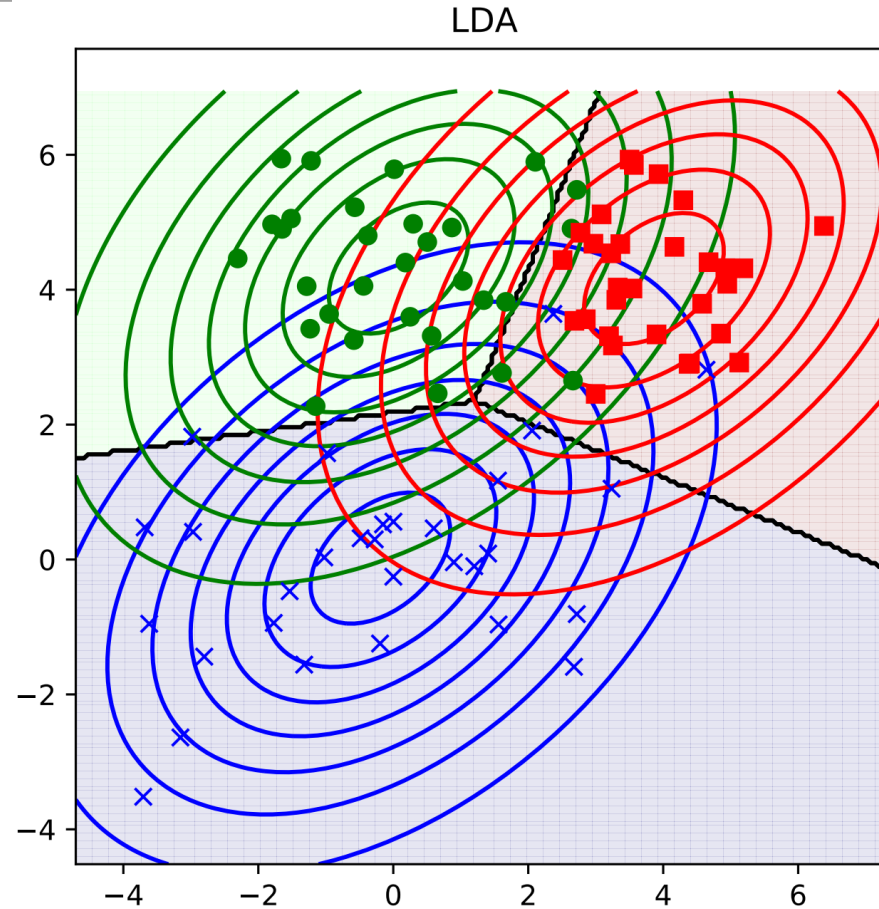
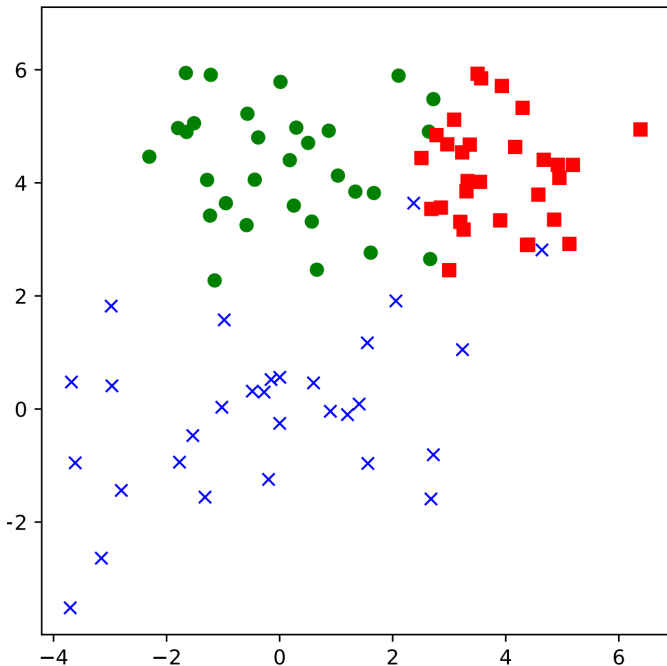
Multi-Variate Gaussian
 $\boldsymbol{\mu}_k$ = mean vector
 Σ_k = covariance matrix

$$p(y = C_k|\mathbf{x}, \boldsymbol{\theta}) = \frac{p(y = C_k|\boldsymbol{\theta})p(\mathbf{x}|y = C_k, \boldsymbol{\theta})}{\sum_{C_{k'}} p(y = C_{k'}|\boldsymbol{\theta})p(\mathbf{x}|y = C_{k'}, \boldsymbol{\theta})}$$

Gaussian discriminant analysis or GDA

$$p(y = C_k|\mathbf{x}, \boldsymbol{\theta}) = \frac{\frac{1}{(2\pi)^{D/2}|\Sigma_k|^{1/2}} \exp\left[-\frac{1}{2}(\mathbf{x} - \boldsymbol{\mu}_k)^T \Sigma_k^{-1}(\mathbf{x} - \boldsymbol{\mu}_k)\right]}{\sum_{C_{k'}} \pi_{k'} |2\pi \Sigma_{k'}|^{-1/2} \exp\left[-\frac{1}{2}(\mathbf{x} - \boldsymbol{\mu}_{k'})^T \Sigma_{k'}^{-1}(\mathbf{x} - \boldsymbol{\mu}_{k'})\right]}$$

QDA & LDA

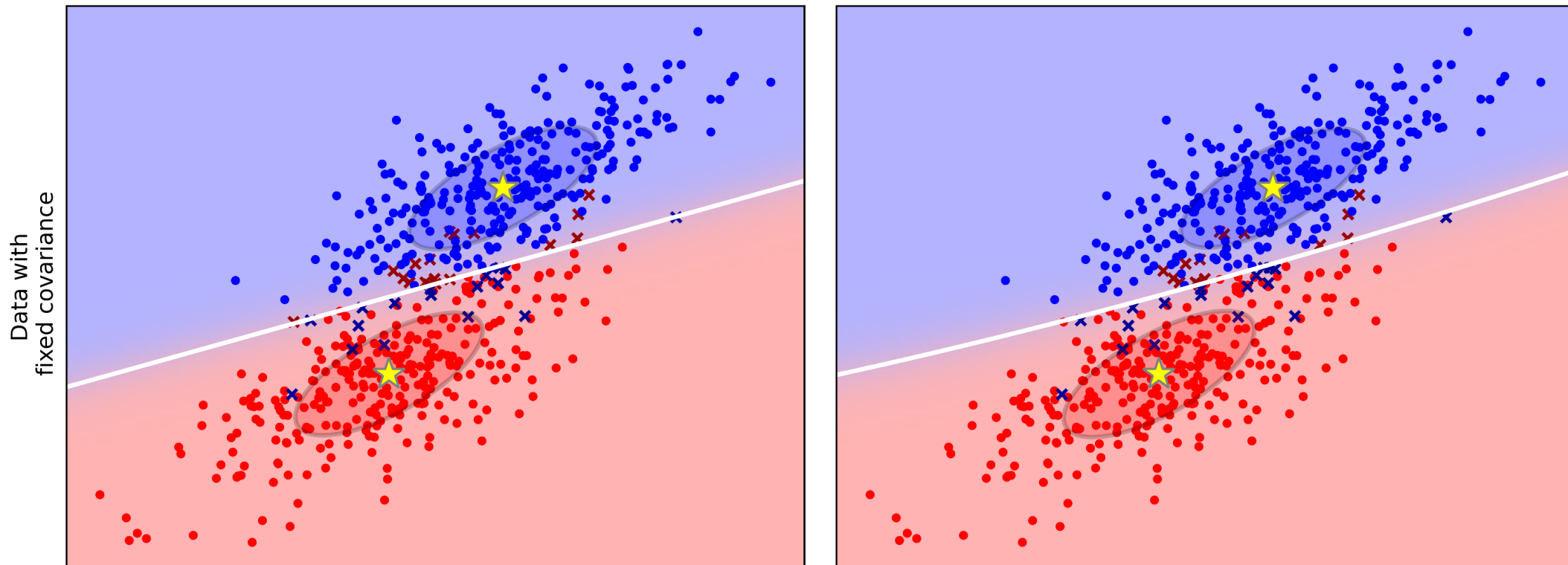


QDA & LDA

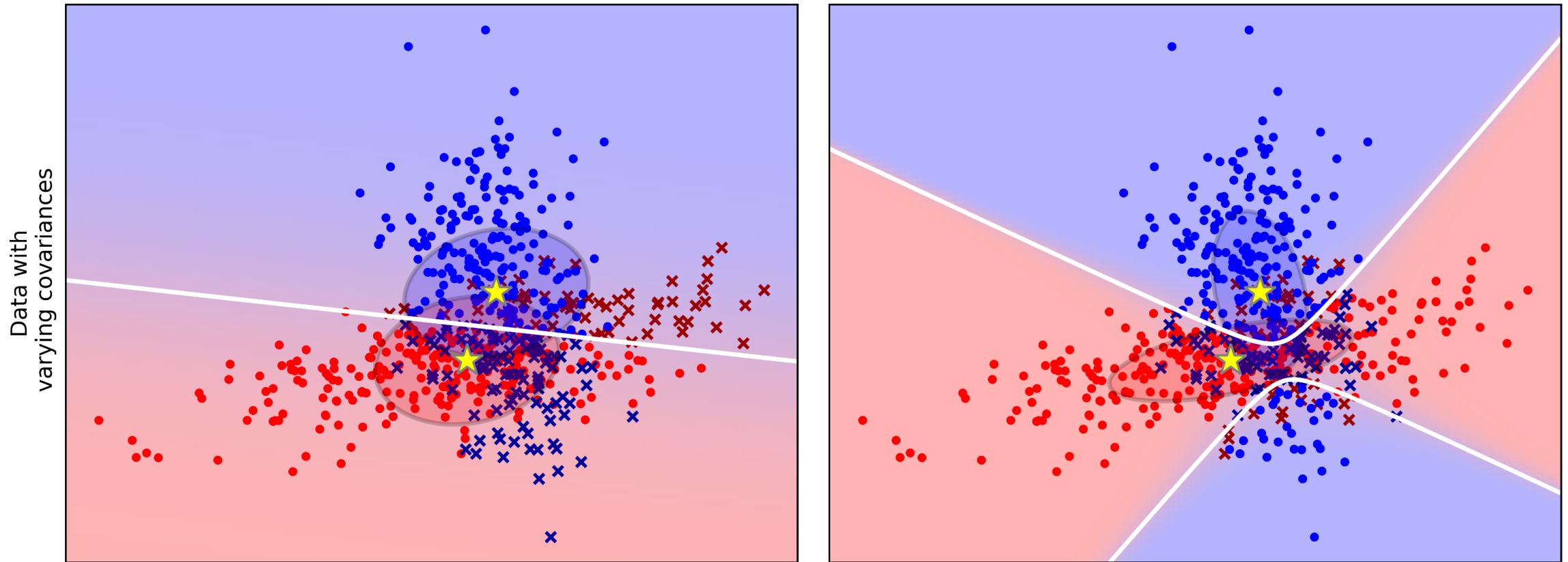
Linear Discriminant Analysis vs Quadratic Discriminant Analysis

Linear Discriminant Analysis

Quadratic Discriminant Analysis



QDA & LDA



Advantages of generative classifiers

- Ease of fitting
- Handling missing features
- Class-specific learning
- Can handle unlabeled data
- Robustness to spurious features



Thank You

Q & A