Assignment 5

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https://github.com/vaibhavchhabra25/AI1103−course/blob/main/Assignment−5/main.tex

1 Problem

(UGC/MATH 2018 (June set-a)-Q.106) Let $\{X_i\}_{i\geq 1}$ be a sequence of i.i.d. random variables with $E(X_i) = 0$ and $V(X_i) = 1$. Which of the following are true?

1) $\frac{1}{n}\sum_{i=1}^{n} X_i^2 \to 0$ in probability
2) $\frac{1}{n^{3/4}}\sum_{i=1}^{n} X_i \to 0$ in probability
3) $\frac{1}{n^{1/2}}\sum_{i=1}^{n} X_i \to 0$ in probability
4) $\frac{1}{n}\sum_{i=1}^{n} X_i^2 \to 1$ in probability

2 Solution

Definition 1. *(Convergence in probability) Let $X1, X2, \dots$ be an infinite sequence of random variables, and let Y be another random variable. Then the sequence $\{X_n\}$ converges in probability to Y, if for all $\epsilon > 0$,*

$$\lim_{n\to\infty} \Pr\left(|X_n - Y| \geq \epsilon\right) = 0 \tag{2.0.1}$$

And we write as $n \to \infty$, $X_n \to Y$ in probability.

1) **Theorem 2.1.** *(Strong Law of Large Numbers) Let $X_1, X_2, \cdots$ be a sequence of i.i.d. random variables, each having finite mean $E(X_i)$. Then for any $\epsilon > 0$,*

$$\lim_{n\to\infty} \Pr\left(\left|\frac{1}{n}\sum_{i=1}^{n} X_i - E(X_i)\right| \geq \epsilon\right) = 0 \tag{2.0.2}$$

Or, $\frac{1}{n}\sum_{i=1}^{n} X_i$ converges in probability to $E(X_i)$.

Given,

$$E(X_i) = 0 \tag{2.0.3}$$
$$V(X_i) = 1 \tag{2.0.4}$$

Also, we know that,

$$E(X_i^2) = V(X_i) + (E(X_i))^2 \tag{2.0.5}$$

Putting given values, we get,

$$E(X_i^2) = 1 + 0^2 \tag{2.0.6}$$
$$\implies E(X_i^2) = 1 \tag{2.0.7}$$

So, $E(X_i^2)$ is finite.
Let $F_{X_i}(x)$ be the c.d.f. for the random variable X_i. As $\{X_i\}$ is sequence of i.i.d. random variables, it follows the following conditions $\forall x, x_i \in \mathbb{R}$:

a) $F_{X_1}(x) = F_{X_2}(x) = \cdots = F_{X_n}(x) = F(x)$
b) $F_{X_1,X_2\dots X_n}(x_1, x_2 \dots x_n) = F(x_1)F(x_2)\dots F(x_n)$

Let $Y_i = X_i^2$. Then for $y \geq 0$,

$$F_{Y_i}(y) = \Pr\left(Y_i \leq y\right) \tag{2.0.8}$$
$$\implies F_{Y_i}(y) = \Pr\left(X_i^2 \leq y\right) \tag{2.0.9}$$
$$\implies F_{Y_i}(y) = \Pr\left(-\sqrt{y} \leq X_i \leq \sqrt{y}\right) \tag{2.0.10}$$
$$\implies F_{Y_i}(y) = \Pr\left(X_i \leq \sqrt{y}\right) - \Pr\left(X_i \leq -\sqrt{y}\right) \tag{2.0.11}$$
$$\implies F_{Y_i}(y) = F_{X_i}(\sqrt{y}) - F_{X_i}(-\sqrt{y}) \tag{2.0.12}$$

Using (a),

$$F_{Y_i}(y) = F(\sqrt{y}) - F(-\sqrt{y}) \tag{2.0.13}$$

So, by (2.0.13),

$$F_{X_1^2}(y) = F_{X_2^2}(y) = \cdots = F_{X_n^2}(y) = F_1(y) \tag{2.0.14}$$

where $F_1(y)$ is the c.d.f. of $Y_i = X_i^2$.
Now, for $y_i \geq 0$, consider

$$F_{Y_1,Y_2,\dots,Y_n}(y_1, y_2, \dots, y_n) = \Pr\left(Y_1 \leq y_1, Y_2 \leq y_2, \dots, Y_n \leq y_n\right) \tag{2.0.15}$$

$$= \Pr\left(X_1^2 \le y_1, X_2^2 \le y_2, \dots, X_n^2 \le y_n\right) \quad (2.0.16)$$

$$= \Pr\left(-\sqrt{y_1} \le X_1 \le \sqrt{y_1}, -\sqrt{y_2} \le X_2 \le \sqrt{y_2}, \dots, -\sqrt{y_n} \le X_n \le \sqrt{y_n}\right) \quad (2.0.17)$$

Since $X_1, X_2, \dots, X_n$ are independent,

$$F_{Y_1,Y_2,\dots,Y_n}(y_1, y_2, \dots, y_n) = \Pr\left(-\sqrt{y_1} \le X_1 \le \sqrt{y_1}\right)\Pr\left(-\sqrt{y_2} \le X_2 \le \sqrt{y_2}\right) \dots \Pr\left(-\sqrt{y_n} \le X_n \le \sqrt{y_n}\right) \quad (2.0.18)$$

Using (2.0.10) and (2.0.14),

$$F_{Y_1,Y_2,\dots,Y_n}(y_1, y_2, \dots, y_n) = F_{Y_1}(y_1)F_{Y_2}(y_2)\dots F_{Y_n}(y_n) \quad (2.0.19)$$

$$= F_1(y_1)F_1(y_2)\dots F_1(y_n) \quad (2.0.20)$$

So,

$$F_{X_1^2,X_2^2,\dots,X_n^2}(y_1, y_2, \dots, y_n) = F_1(y_1)F_1(y_1)\dots F_1(y_n) \quad (2.0.21)$$

By (2.0.14) and (2.0.21), $\{X_i^2\}$ must also be a sequence of i.i.d. random variables.
So, we can apply S.L.L.N. to this sequence.
Then, $\frac{1}{n}\sum_{i=1}^{n} X_i^2$ converges in probability to $E(X_i^2)$.
Or, $\frac{1}{n}\sum_{i=1}^{n} X_i^2 \to 1$ in probability.
Thus, option 1 is wrong.

2) Now, we define

$$Y_n = \frac{1}{n^{3/4}}\sum_{i=1}^{n} X_i \quad (2.0.22)$$

Then,

$$E(Y_n) = \frac{1}{n^{3/4}}E\left(\sum_{i=1}^{n} X_i\right) \quad (2.0.23)$$

$$\implies E(Y_n) = \frac{1}{n^{3/4}}\left(E(X_1) + E(X_2)\dots E(X_n)\right) \quad (2.0.24)$$

Using (2.0.3)

$$E(Y_n) = \frac{1}{n^{3/4}}(0) = 0 \quad (2.0.25)$$

Now,

$$V(Y_n) = V\left(\frac{1}{n^{3/4}}\sum_{i=1}^{n} X_i\right) \quad (2.0.26)$$

$$\implies V(Y_n) = \frac{1}{n^{3/2}}V\left(X_1 + X_2 + \dots + X_n\right) \quad (2.0.27)$$

As $X_1, X_2, \cdots X_n$ are independent of each other,

$$V(Y_n) = \frac{1}{n^{3/2}}\left(V(X_1) + V(X_2) + \dots + V(X_n)\right) \quad (2.0.28)$$

Using (2.0.4)

$$V(Y_n) = \frac{1}{n^{3/2}}(1 + 1 + \dots + 1) = \frac{1}{n^{3/2}} \times n = \frac{1}{n^{1/2}} \quad (2.0.29)$$

Now for any $\epsilon > 0$, consider the probability

$$\Pr\left(|Y_n - 0| \ge \epsilon\right) = \Pr\left(|Y_n - E(Y_n)| \ge \epsilon\right) \quad (2.0.30)$$

Applying Chebyschev's inequality here, we get,

$$\Pr\left(|Y_n - 0| \ge \epsilon\right) \le \frac{V(Y_n)}{\epsilon^2} = \frac{1}{n^{1/2}\epsilon^2} \quad (2.0.31)$$

So,

$$\lim_{n\to\infty}\Pr\left(|Y_n - 0| \ge \epsilon\right) \le \lim_{n\to\infty}\frac{1}{n^{1/2}\epsilon^2} = 0 \quad (2.0.32)$$

$$\implies \lim_{n\to\infty}\Pr\left(\left|\frac{1}{n^{3/4}}\sum_{i=1}^{n} X_i - 0\right| \ge \epsilon\right) = 0 \quad (2.0.33)$$

So, $\frac{1}{n^{3/4}}\sum_{i=1}^{n} X_i \to 0$ in probability.
Thus, option 2 is correct.

3)

4) As proved for option (1), $\frac{1}{n}\sum_{i=1}^{n} X_i^2 \to 1$ in probability. So option 4 is correct.

So the answer must be options 2,4.