

Assignment 5

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Download all latex codes from

<https://github.com/vaibhavchhabra25/AI1103-course/blob/main/Assignment-5/main.tex>

Given,

$$E(X_i) = 0 \quad (2.0.3)$$

$$V(X_i) = 1 \quad (2.0.4)$$

Also, we know that,

$$E(X_i^2) = V(X_i) + (E(X_i))^2 \quad (2.0.5)$$

Putting given values, we get,

$$E(X_i^2) = 1 + 0^2 = 1 \quad (2.0.6)$$

$$\implies E(X_i^2) = 1 \quad (2.0.7)$$

So, $E(X_i^2)$ is finite.

As $\{X_i\}$ is sequence of i.i.d. random variables, it follows the following conditions $\forall x, x_i \in \mathbb{R}$:

- 1) $\Pr(X_1 \leq x) = \Pr(X_k \leq x) \quad \forall k \in \{1, 2, 3 \dots n\}$
- 2) $\Pr(X_1 \leq x_1, X_2 \leq x_2 \dots X_n \leq x_n) = \Pr(X_1 \leq x_1) \cdot \Pr(X_2 \leq x_2) \dots \Pr(X_n \leq x_n)$

We can rewrite (1) as

$$\Pr(X_1^2 \leq x^2) = \Pr(X_k^2 \leq x^2) \quad (2.0.8)$$

Putting $x^2 = y$, we get,

$$\Pr(X_1^2 \leq y) = \Pr(X_k^2 \leq y) \quad (2.0.9)$$

Also, we can rewrite (2) as

$$\begin{aligned} \Pr(X_1^2 \leq x_1^2, X_2^2 \leq x_2^2 \dots X_n^2 \leq x_n^2) = \\ \Pr(X_1^2 \leq x_1^2) \cdot \Pr(X_2^2 \leq x_2^2) \dots \Pr(X_n^2 \leq x_n^2) \end{aligned} \quad (2.0.10)$$

Putting $x_i^2 = y_i$, we get,

$$\begin{aligned} \Pr(X_1^2 \leq y_1, X_2^2 \leq y_2 \dots X_n^2 \leq y_n) = \\ \Pr(X_1^2 \leq y_1) \cdot \Pr(X_2^2 \leq y_2) \dots \Pr(X_n^2 \leq y_n) \end{aligned} \quad (2.0.11)$$

By (2.0.9) and (2.0.11), $\{X_i^2\}$ must also be a sequence of i.i.d. random variables.

So, we can apply S.L.L.N. to this sequence.

Then, $\frac{1}{n} \sum_{i=1}^n X_i^2$ converges in probability to $E(X_i^2)$.

1 PROBLEM

(UGC/MATH 2018 (June set-a)-Q.106) Let $\{X_i\}_{i \geq 1}$ be a sequence of i.i.d. random variables with $E(X_i) = 0$ and $V(X_i) = 1$. Which of the following are true?

- 1) $\frac{1}{n} \sum_{i=1}^n X_i^2 \rightarrow 0$ in probability
- 2) $\frac{1}{n^{3/4}} \sum_{i=1}^n X_i \rightarrow 0$ in probability
- 3) $\frac{1}{n^{1/2}} \sum_{i=1}^n X_i \rightarrow 0$ in probability
- 4) $\frac{1}{n} \sum_{i=1}^n X_i^2 \rightarrow 1$ in probability

2 SOLUTION

Definition 1. (Convergence in probability)

Let X_1, X_2, \dots be an infinite sequence of random variables, and let Y be another random variable. Then the sequence $\{X_n\}$ converges in probability to Y , if for all $\epsilon > 0$,

$$\lim_{n \rightarrow \infty} \Pr(|X_n - Y| \geq \epsilon) = 0 \quad (2.0.1)$$

And we write as $n \rightarrow \infty$, $X_n \rightarrow Y$ in probability.

Theorem 2.1. (Strong Law of Large Numbers)

Let X_1, X_2, \dots be a sequence of i.i.d. random variables, each having finite mean $E(X_i)$. Then for any $\epsilon > 0$,

$$\lim_{n \rightarrow \infty} \Pr\left(\left|\frac{1}{n} \sum_{i=1}^n X_i - E(X_i)\right| \geq \epsilon\right) = 0 \quad (2.0.2)$$

Or, $\frac{1}{n} \sum_{i=1}^n X_i$ converges in probability to $E(X_i)$.

Or, $\frac{1}{n} \sum_{i=1}^n X_i^2 \rightarrow 1$ in probability.
Thus, option 1 is wrong and option 4 is correct.

Thus, option 2 is also correct.
So the answer must be options 2,4.

Now, we define

$$Y_n = \frac{1}{n^{3/4}} \sum_{i=1}^n X_i \quad (2.0.12)$$

Then,

$$E(Y_n) = \frac{1}{n^{3/4}} E\left(\sum_{i=1}^n X_i\right) \quad (2.0.13)$$

$$\Rightarrow E(Y_n) = \frac{1}{n^{3/4}} (E(X_1) + E(X_2) + \cdots + E(X_n)) \quad (2.0.14)$$

Using (2.0.3)

$$E(Y_n) = \frac{1}{n^{3/4}} (0) = 0 \quad (2.0.15)$$

Now,

$$V(Y_n) = V\left(\frac{1}{n^{3/4}} \sum_{i=1}^n X_i\right) \quad (2.0.16)$$

$$\Rightarrow V(Y_n) = \frac{1}{n^{3/2}} V(X_1 + X_2 + \cdots + X_n) \quad (2.0.17)$$

As X_1, X_2, \dots, X_n are independent of each other,

$$V(Y_n) = \frac{1}{n^{3/2}} (V(X_1) + V(X_2) + \cdots + V(X_n)) \quad (2.0.18)$$

Using (2.0.4)

$$V(Y_n) = \frac{1}{n^{3/2}} (1 + 1 + \cdots + 1) = \frac{1}{n^{3/2}} \times n = \frac{1}{n^{1/2}} \quad (2.0.19)$$

Now for any $\epsilon > 0$, consider the probability

$$\Pr(|Y_n - 0| \geq \epsilon) = \Pr(|Y_n - E(Y_n)| \geq \epsilon) \quad (2.0.20)$$

Applying Chebyshev's inequality here, we get,

$$\Pr(|Y_n - 0| \geq \epsilon) \leq \frac{V(Y_n)}{\epsilon^2} = \frac{1}{n^{1/2} \epsilon^2} \quad (2.0.21)$$

So,

$$\lim_{n \rightarrow \infty} \Pr(|Y_n - 0| \geq \epsilon) \leq \lim_{n \rightarrow \infty} \frac{1}{n^{1/2} \epsilon^2} = 0 \quad (2.0.22)$$

$$\Rightarrow \lim_{n \rightarrow \infty} \Pr\left(\left|\frac{1}{n^{3/4}} \sum_{i=1}^n X_i - 0\right| \geq \epsilon\right) = 0 \quad (2.0.23)$$

So, $\frac{1}{n^{3/4}} \sum_{i=1}^n X_i \rightarrow 0$ in probability.