Biostatistics (MATH11230)

Measures of disease-exposure association

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- For simplicity, we will start with a risk factor that can only take two values (e.g., exposed and unexposed).

Relative risk

$$\mathsf{RR} = \frac{\mathsf{Pr}(D \mid E)}{\mathsf{Pr}(D \mid \bar{E})} = \frac{\mathsf{Pr}(D \mid E)}{\mathsf{Pr}(D \mid \mathsf{not}\; E)}.$$

- \hookrightarrow A RR = 1, the so-called null value, implies that $\Pr(D \mid E) = \Pr(D \mid \text{not } E)$, which is equivalent to saying that D and E are independent (no association between the exposure to the risk factor E and the outcome D).
- → A RR > 1 indicates that there is a greater risk or probability of *D* when exposed than when
 not exposed (positive association between the exposure to the risk factor *E* and the
 outcome *D*).
- → When RR < 1 there is a reduced risk or probability of D when exposed than when not exposed (negative association between the exposure to the risk factor E and the outcome D).
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Relative risk

- \hookrightarrow We shall note that the relative risk has an implicit upper bound.
- \hookrightarrow This is because the maximum possible value for a risk and, in particular for $\Pr(D \mid E)$ is one, and therefore $RR \le \frac{1}{\Pr(D|\text{not }E)}$.
- \hookrightarrow This restriction on the range of the relative risk is only problematic if the disease outcome is common, as in this case $\Pr(D \mid \text{not } E)$ may become large.

Relative risk

- \hookrightarrow We shall note that the relative risk is not symmetric in the role of the two factors *D* and *E*.
- → That is, the relative risk for E associated with D is a different measure of association than
 the relative risk for D associated with E, i.e.,

$$\frac{\Pr(\textit{E} \mid \textit{D})}{\Pr(\textit{E} \mid \mathsf{not} \; \textit{D})} \neq \frac{\Pr(\textit{D} \mid \textit{E})}{\Pr(\textit{D} \mid \mathsf{not} \; \textit{E})}.$$

Relative risk

- Let us illustrate the relative risk calculation with data from the table below (Jewell, 2003, p 22), which lists the vital status of all births in the United States in 1991, one year after date of birth, and categorised by the marital status of the mother at the birth and the birth weight of the infant.
- → Low birth weight was defined as a birth where the newborn's weight is less than 2.5kg.

| | Mother's M | Iarital Status | | |
|----------------------|----------------------------|---------------------|-----------|--|
| Infant Mortality | Unmarried | Married | Total | |
| Death | 16,712 | 18,784 | 35,496 | |
| Live at 1 year | 1,197,142 | 2,878,421 | 4,075,563 | |
| Total | 1,213,854 | 2,897,205 | 4,111,059 | |
| Birth weight | | | | |
| Infant Mortality | Low Birth weight | Normal Birth weight | Total | |
| Death | 21,054 | 14,442 | 35,496 | |
| Live at 1 year | 271,269 | 3,804,294 | 4,075,563 | |
| Total | 292,323 | 3,818,736 | 4,111,059 | |
| Source: National Cer | iter for Health Statistics | | | |

Relative risk

 Following the definition, the relative risk for infant mortality in the USA in 1991, associated with the mother being unmarried, is

$$\frac{\text{Pr(death} \mid \text{unmarried mother})}{\text{Pr(death} \mid \text{married mother})} = \frac{16712/1213854}{18784/2897205} = 2.12.$$

- This result indicates that there is a positive association between infant mortality and mother being unmarried.
- The risk of an infant death with an unmarried mother is about twice the risk of an infant death with a married mother.

Relative risk

→ In turn, the relative risk for infant mortality in the USA in 1991, associated with a low birth weight is given by

$$\frac{\text{Pr(death | low birth weight)}}{\text{Pr(death | normal birth weight)}} = \frac{21054/292323}{14442/3818736} = 19.0.$$

→ There is thus a much greather effect of birth weight on infant mortality than marital status.

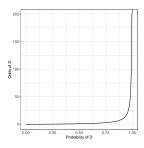
Odds of disease

- \hookrightarrow The relative risk measures the risk of the outcome *D* through Pr(D).
- → Another quantity is the **odds** of disease, given by

$$\frac{Pr(D)}{Pr(\text{not }D)} = \frac{Pr(D)}{1 - Pr(D)},$$

and which gives the likelihood of *D* occurring relative to it not occurring.

 \hookrightarrow The odds of D provide the same information as $\Pr(D)$ since knowing one of the quantities immediately determines the other.





Odds ratio

- \hookrightarrow The odds ratio for D associated with E is therefore defined by

$$OR = \frac{Pr(D \mid E)}{Pr(\text{not } D \mid E)} / \frac{Pr(D \mid \text{not } E)}{Pr(\text{not } D \mid \text{not } E)}.$$

- \hookrightarrow The odds ratio is a number between 0 and ∞ .
- → Because there is no upper limits for the odds of D, the odds ratio, by opposition to the relative risk, has no implicit upper bound.
- → As with the relative risk, OR = 1 is the null value, and it corresponds to the case where the odds of the outcome D are the same in both groups (exposed and unexposed), and it is again equivalent to no association (independence) between D and E.
- \hookrightarrow When OR > 1, there is a greater risk of *D* in the exposed group.
- \hookrightarrow The reverse is true when OR < 1, that is, there is a lower risk of D in the exposed group.

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Odds ratio

- → Let us revisit the example about infant mortality in the USA in 1991.
- → Let us first compute the odds of infant mortality in the unmarried mother group (exposed group)

$$\frac{16712/1213854}{1197142/1213854} = \frac{16712}{1197142}.$$

→ Similarly, the odds of infant mortality in the married mother (unexposed) group are

$$\frac{18784/2897205}{2878421/2897205} = \frac{18784}{2878421}.$$

 \hookrightarrow The odds ratio for infant mortality associated with an unmarried mother is

$$\mathsf{OR} = \frac{16712/1197142}{18784/2878421} = 2.14.$$



Odds ratio

- We will now re-do the calcution but now with low birth weight being the risk factor/exposure.

$$\frac{21054/292323}{271269/292323} = \frac{21054}{271269}.$$

→ Analogously, the odds of death in the normal birth weight group are

$$\frac{14442/3818736}{3804294/3818736} = \frac{14442}{3804294}.$$

→ Finally, the odds ratio for infant mortality associated with low birth weight is

$$\mathsf{OR} = \frac{21054/271269}{14442/3804294} = 20.4.$$



Odds ratio

→ So, let us compare what we have obtained for this example so far in terms of relative risks and odds ratios.

| | RR | OR |
|--|------|------|
| Exposure/risk factor: marital status | 2.12 | 2.14 |
| Exposure/risk factor: low birth weight | 19.0 | 20.4 |

- → The RRs and ORs for both exposures (mother's marital status and birth weight) are very similar.
- → But, is this always the case?

- \hookrightarrow Let us look formally at the relationship between the RR and the OR.
- \hookrightarrow By its definition, the odds ratio for D associated with E is

$$OR = \frac{\Pr(D \mid E)}{\Pr(\text{not } D \mid E)} \times \frac{\Pr(\text{not } D \mid \text{not } E)}{\Pr(D \mid \text{not } E)}$$

$$= \frac{\Pr(D \mid E)}{\Pr(D \mid \text{not } E)} \times \frac{\Pr(\text{not } D \mid \text{not } E)}{\Pr(\text{not } D \mid E)}$$

$$= RR \times \frac{\Pr(\text{not } D \mid \text{not } E)}{\Pr(\text{not } D \mid E)}.$$

Odds ratio

 \hookrightarrow Let us suppose first that RR > 1. By definition of relative risk, this means that

$$Pr(D \mid E) > Pr(D \mid \text{not } E),$$

which trivially implies that

$$1 - \Pr(D \mid E) < 1 - \Pr(D \mid \text{not } E),$$

and which, in turn, can be equivalently written as,

$$Pr(\text{not } D \mid E) < Pr(\text{not } D \mid \text{not } E),$$

thus implying that

$$\frac{\Pr(\text{not }D\mid \text{not }E)}{\Pr(\text{not }D\mid E)}>1.$$

- \hookrightarrow When RR < 1, a similar reasoning leads to the conclusion that OR < RR.
- → We thus arrive at the conclusion that the odds ratio is always farther away from one than
 the relative risk (except when the relative risk is one, in which case the odds ratio will also
 be one).
- \hookrightarrow How farther away the OR is from 1 compared to the RR will depend on both $\Pr(D \mid E)$ and $\Pr(D \mid \text{not } E)$, which are also the two probabilities involved in the computation of the risk ratio.
- \hookrightarrow When the risk of disease is low in both the exposed and unexposed groups, then $\Pr(\text{not }D\mid E)\approx 1$ and $\Pr(\text{not }E)$ not $E)\approx 1$ and so $R=\mathbb{R}$.

- → We have seen before that the relative risk is not symmetric in the roles of D and E, that is, the relative risk for D associated with E does not (need to) coincide with the relative risk for E associated with D.
- The odds ratio enjoys the property of being symmetric with respect to the roles of *D* and *E*, that is, reversing the roles of *D* and *E* makes no difference in its computation.
- → As we will see in one of the next lectures, this property will be key when estimating the association between D and E in certain study designs (case-control studies).
- \hookrightarrow In what follows, let $OR_{D|E}$ denote the odds ratio for D associated with E and $OR_{E|D}$ denote the odds ratio for E associated with D.

$$\begin{split} \operatorname{OR}_{D|E} &= \frac{\operatorname{Pr}(D \mid E)}{\operatorname{Pr}(\operatorname{not} D \mid E)} \middle/ \frac{\operatorname{Pr}(D \mid \operatorname{not} E)}{\operatorname{Pr}(\operatorname{not} D \mid \operatorname{not} E)} \\ &= \frac{\operatorname{Pr}(D\&E) / \operatorname{Pr}(E)}{\operatorname{Pr}(\operatorname{not} D\&E) / \operatorname{Pr}(E)} \middle/ \frac{\operatorname{Pr}(D\&\operatorname{not} E) / \operatorname{Pr}(\operatorname{not} E)}{\operatorname{Pr}(\operatorname{not} D\&\operatorname{not} E) / \operatorname{Pr}(\operatorname{not} E)} \\ &= \frac{\operatorname{Pr}(D\&E)}{\operatorname{Pr}(\operatorname{not} D\&E)} \middle/ \frac{\operatorname{Pr}(D\&\operatorname{not} E)}{\operatorname{Pr}(\operatorname{not} D\&\operatorname{not} E)} \\ &= \frac{\operatorname{Pr}(D\&E)}{\operatorname{Pr}(D\&\operatorname{not} E)} \middle/ \frac{\operatorname{Pr}(\operatorname{not} D\&\operatorname{not} E)}{\operatorname{Pr}(\operatorname{not} D\&\operatorname{not} E)} \\ &= \frac{\operatorname{Pr}(D\&E) / \operatorname{Pr}(D)}{\operatorname{Pr}(D\&\operatorname{not} E) / \operatorname{Pr}(\operatorname{not} D)} \middle/ \frac{\operatorname{Pr}(\operatorname{not} D\&\operatorname{not} E) / \operatorname{Pr}(\operatorname{not} D)}{\operatorname{Pr}(\operatorname{not} E \mid D)} \\ &= \frac{\operatorname{Pr}(E \mid D)}{\operatorname{Pr}(\operatorname{not} E \mid D)} \middle/ \frac{\operatorname{Pr}(E \mid \operatorname{not} D)}{\operatorname{Pr}(\operatorname{not} E \mid \operatorname{not} D)} \\ &= \operatorname{OR}_{E|D} \end{split}$$

Association is not causation!

- → They do not imply (without further considerations) any causal effect, that is, we cannot conclude that the risk factor causes the health outcome.
- Some third factor (known as confounding variable) might be responsible for the relationship between the exposure and the outcome under study (i.e., the groups receiving and not receiving the exposure may be different from one another in some other important variable that is also related to the outcome).
- → For instance, in the running example of mother's marital status and infant mortality, the apparent association captured by the RR and OR could be due to another factor(s) that is (are) related both to marital status and infant mortality.
- → We will learn how to deal with confounding variables in a couple of weeks.

Extra material: vaccine efficacy (and its connection to the relative risk...)

→ As can be read here:

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https://www.thelancet.com/journals/lanmic/article/
PIIS2666-5247(21)00069-0/fulltext
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vaccine efficacy is generally reported as a relative risk reduction (RRR), which is defined as 1 - RR.

- → Here the exposure is beneficial and the exposed group consists of those who were vaccinated. The unexposed group then consists of those who were not vaccinated. The outcome D is the getting infected with the virus SARS-CoV-2.
- → This blog entry nicely illustrates the calculation of the efficacy of the Pfizer and Moderna (for COVID-19) vaccines

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https://towardsdatascience.com/pfizer-and-moderna-vaccine-efficacy-calculated-from-data-9566897173c
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