Incomplete Data Analysis

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Exercise

→ The following table shows a small artificial dataset with 4 missing values:

Subject number	Sex	Age(years)	Systolic blood pressure (mmHg)
1	Male	50	163.5
2	Male	41	126.4
3	Male	52	150.7
4	Male	58	190.4
5	Male	56	172.2
6	Male	45	NA
7	Male	42	136.3
8	Male	48	146.8
9	Male	57	162.5
10	Male	56	161.0
11	Male	55	148.7
12	Male	58	163.6
13	Female	57	NA
14	Female	44	140.6
15	Female	56	NA
16	Female	45	118.7
17	Female	48	NA
18	Female	50	104.6
19	Female	59	131.5
20	Female	55	126.9

Exercise

- (i) Carry out a complete case analysis to find the mean value of systolic blood pressure overall, and by sex. Also compute the associated standard error of the mean.
- (ii) Impute the missing values of systolic blood pressure by mean imputation. Use these filled in values to estimate the mean systolic blood pressure with corresponding standard error.
- (iii) Impute the missing values for systolic blood pressure by regression imputation, regressing on sex and age. Write down the regression equation used and the four imputed values obtained. Use these to estimate the mean systolic blood pressure and associated standard error.
- (iv) The same as in (iii) but now using stochastic regression imputation.
- (v) Suppose that hot deck imputation is to be used with strata defined by sex and age (≤ 50 years and > 50 years). Estimate the mean systolic blood pressure with associated standard error.

Exercise - solution

- \hookrightarrow Firstly, recall that the standard error of the mean is a measure of the precision of the mean.
- \hookrightarrow It depends on both the standard deviation and the sample size and it is defined as $\frac{\sigma}{\sqrt{n}}$.
- \hookrightarrow Since the standard deviation is seldom known in practice, we estimate it by the sample standard deviation, and thus the estimated standard error of the mean is given by $\frac{s}{\sqrt{n}}$, where $s^2 = \frac{1}{n-1} \sum_{i=1}^{n} (x_i \bar{x})^2$.
- → The mean SBP for men and for women are respectively 156.55 (5.27) and 124.46 (6.10).

 The values in brackets are the standard errors.

Exercise - solution

- Using mean imputation, the estimated SBP mean is 146.53 (we already knew this from the complete case analysis), while the associated standard error is 4.39.
- → Of course, the standard error of the mean based on mean imputation is smaller than that of the complete case analysis. Why? Have a look at the sample variance formula.

$$SBP = \beta_0 + \beta_1 Age + \beta_2 Sex + \varepsilon, \qquad \varepsilon \sim N(0, \sigma^2).$$

→ Here, gender is a dummy (binary) variable taking the value 0 if the subject is female and 1 if the subject is male.

Exercise - solution

 \hookrightarrow The expected values for SBP are then

$$E(\textit{SBP}) = \begin{cases} \beta_0 + \beta_1 \mathsf{Age} & \text{if the subject is female,} \\ (\beta_0 + \beta_2) + \beta_1 \mathsf{Age} & \text{if the subject is male.} \end{cases}$$

- We fit the regression model to the complete cases, obtain estimated regression coefficients and then impute the SBP values based on the regression equation.
- \hookrightarrow The estimated regression coefficients are $\widehat{\beta}_0=40.8784,\,\widehat{\beta}_1=1.6518,$ and $\widehat{\beta}_2=29.6318.$
- \hookrightarrow The imputed values are then

Subject number	Sex	Age	Imputed SBP
6	Male	45	144.84
13	female	57	135.03
15	female	56	133.38
17	Female	48	120.17

Exercise - solution

- → For the above regression imputation, the estimated mean SBP is 143.89 with a standard error of 4.65.
- → We will impute the SBP values based on stochastic regression imputation.
- → The steps are essentially the same as those in regression imputation, but we add a random term, normally distributed with mean zero and variance equal to the estimated variance of the residuals, which in this case is 13.34².

Subject number	Sex	Age	Imputed SBP
6	Male	45	133.90
13	female	57	126.75
15	female	56	148.38
17	Female	48	119.95

→ Of course, since we are adding a random term, different execution of the code will lead to different imputed values. The values obtained above were obtained using set.seed(1).

Exercise - solution

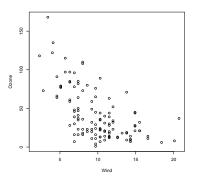
- \hookrightarrow The estimated mean is thus 143.67 (4.71).
- \hookrightarrow We now turn our attention to hot deck imputation. We define strata based on gender and age (\le 50 and > 50).
- → For instance, we see that subject number 6 is male and is 45 years old. Thus, possible candidates to give him his SBP value are subjects number 1, 2, 7, and 8. We now just need to pick one of these subjects in a random fashion. Using the seed equal to 1, we have obtained subject number 8 to be the donor.
- \hookrightarrow We proceed in the same fashion for the other subjects.

Cautionary note

- So far, we have used normal linear regression models, which assume, normality of the error terms, homoscedasticity (i.e., the variance of the error terms is constant, or equivalently, does not depend on the regressors), independence of the error terms, and linearity of the predictor.
- → All these assumptions can be checked informally using diagnostic plots (c.f. your favourite regression book).
- → We give a specific example concerning the linearity of the predictor. Let us use the airquality dataset again.

Cautionary note

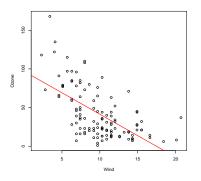
→ Below we have a scatterplot of the variables wind and ozone (which has missing values, but this is not the point here).



 \hookrightarrow The relationship does not seem to be linear, as the plot in the next slide indicates.

Cautionary note

 \hookrightarrow Scatterplot with the fitted regression line superimposed.

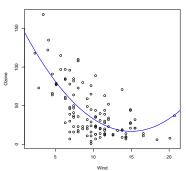


 \hookrightarrow The nonlinear curvature might suggest the inclusion of a quadratic term.

Cautionary note

Ozone =
$$\beta_0 + \beta_1 \text{Wind} + \beta_2 \text{Wind}^2 + \epsilon$$
, $\epsilon \sim \text{N}(0, \sigma^2)$.

⇒ Below, we show the scatterplot along with the quadratic fit, which seems much better.



Illustrative simulation study

- → To better illustrate the performance of some of the traditional methods to handle missing data I have conducted a Monte Carlo simulation study.
- → A Monte Carlo simulation study generates a large number of samples (e.g., 1000) from a population with a specified set of parameter values.
- Estimating a statistical model on each sample and saving the resulting parameter estimates creates an empirical sampling distribution for each model parameter.
- → The difference between the average parameter estimate and the true population parameter
 is of particular importance because it quantifies the bias.
- ∴ You uploaded on Learn three interesting papers (in my opinion!) about simulation studies.

Illustrative simulation study

→ I have generated 1000 datasets, each consisting of 200 observations, from a bivariate normal distribution with the following structure

$$\boldsymbol{\mu} = \begin{bmatrix} \mu_1 \\ \mu_2 \end{bmatrix} = \begin{bmatrix} 0 \\ 0 \end{bmatrix}, \qquad \boldsymbol{\Sigma} = \begin{bmatrix} \sigma_1^2 & \sigma_{1,2} \\ \sigma_{2,1} & \sigma_2^2 \end{bmatrix} = \begin{bmatrix} 1 & 0.5 \\ 0.5 & 1 \end{bmatrix}, \qquad \rho = 0.5.$$

$$\begin{bmatrix} Y_1 \\ Y_2 \end{bmatrix} \sim \mathsf{N}_2(\boldsymbol{\mu}, \boldsymbol{\Sigma})$$

→ We will impose missingness on Y₂ only, in a similar fashion as we did in Exercise 3 of Workshop 1.

Illustrative simulation study

→ The first simulation created MCAR data by imposing

$$Pr(R_2 = 1 \mid Y_1, Y_2) = 0.5.$$

$$\Pr(R_2 = 1 \mid Y_1, Y_2) = \frac{e^{\beta_0 + \beta_1 Y_1}}{1 + e^{\beta_0 + \beta_1 Y_1}}, \quad \beta_0 = 1.5, \quad \beta_1 = 3.$$

Finally, the third simulation study created MNAR data by imposing

$$\Pr(R_2 = 1 \mid Y_1, Y_2) = \frac{e^{\beta_0 + \beta_1 Y_1 + \beta_2 Y_2}}{1 + e^{\beta_0 + \beta_1 Y_1 + \beta_2 Y_2}}, \quad \beta_0 = 1.5, \quad \beta_1 = 3, \quad \beta_2 = 5.$$

Illustrative simulation study - MCAR results

	Estimate	Population parameter	CC	MI	RI	SRI
Ī	μ_1	0	-0.0056	-0.0029	-0.0029	-0.0029
	μ_2	0	0.0007	0.0007	0.0021	-0.0005
	$\frac{\mu_2}{\sigma_1^2}$	1	0.9992	1.0006	1.0006	1.0006
	σ_2^2	1	1.0016	0.499	0.6270	1.0036
	ρ	0.5	0.4962	0.3499	0.6279	0.4971

Illustrative simulation study - MAR results

	Estimate	Population parameter	CC	MI	RI	SRI
-	μ_1	0	0.4671	-0.0029	-0.0029	-0.0029
	μ_2	0	0.2352	0.2352	0.0007	0.0024
	$\mu_2 \ \sigma_1^2$	1	0.6008	1.0006	1.0006	1.0006
	σ_2^2	1	0.9000	0.5984	0.7539	1.0042
	ρ	0.5	0.4060	0.2569	0.5715	0.4935

Illustrative simulation study - MNAR results

	Estimate	Population parameter	CC	MI	RI	SRI
-	μ_1	0	0.5064	-0.0029	-0.0029	-0.0029
	μ_2	0	0.6030	0.6030	0.5286	0.5289
	$\mu_2 \ \sigma_1^2$	1	0.6592	1.0006	1.0006	1.0006
	σ_2^2	1	0.5212	0.3025	0.3204	0.5337
	ρ	0.5	0.1626	0.1006	0.2531	0.1963

Imputation of several missing variables - routine multivariate imputation

- → When presenting the conditional mean imputation and stochastic regression imputation methods, we assumed that only one variable was subject to missingness.
- → We now put ourselves in the situation where several variables have missing values.
- One possibility, directly extending what we have seen before, is to specify a multivariate regression model. However, the main drawback of this approach is that it is not trivial to set up a reasonable multivariate regression model.
- → Usually, a multivariate normal or t distribution is used for continuous variables, and a multinomial distribution for discrete outcomes.
- → Positive point is that software exist to fit such models.

Imputation of several missing variables - iterative regression imputation

- → A different way to extend the univariate methods seen before is to apply them iteratively to the variables with missingness.
- \hookrightarrow Supposing the variables with missingness are $Y_{(1)}, Y_{(2)}, \ldots, Y_{(k)}$. Suppose further that Y_{obs} denotes the fully observed variables (i.e., no missingness).
- → Iterative regression imputation consists of the following steps:
 - Imputing all missing values in $Y_{(1)}$, $Y_{(2)}$, ..., $Y_{(k)}$ using some crude imputation method (e.g., mean imputation).
 - 2 Impute $Y_{(1)}$ given $Y_{(2)}, \ldots, Y_{(k)}$ and Y_{obs} (using for instance stochastic regression imputation).
 - 3 Impute $Y_{(2)}$ given $Y_{(1)}, Y_{(3)}, \dots, Y_{(k)}$ and Y_{obs} (using the newly imputed values for $Y_{(1)}$).
 - 4 ...
 - Impute $Y_{(k)}$ given $Y_{(1)}, Y_{(2)}, \dots, Y_{(k-1)}$ and Y_{obs} (using the newly imputed values for $Y_{(1)}, Y_{(2)}, \dots, Y_{(k-1)}$).



Imputation of several missing variables – iterative regression imputation

- → The disadvantage is that the user needs to be more careful in this setting in order to ensure that the separate regression models are consistent with each other.
- → For instance, on a survey, it would not make sense to impute age based on income but then later ignoring age when imputing income.

Summary

- → Best approach is possibly stochastic regression imputation.
- \hookrightarrow Can be reasonable with only a mild percentage of missing values (e.g., < 5%).
- → We should be aware that all methods, although to different extents, results in overly precise estimates (i.e., standard errors too small):
 - → Analyses after single imputation do not know that some of the values have been imputed.
 - → Simply treats imputed values as if they were observed.
 - \hookrightarrow Does not take into account the uncertainty in the imputed values.