

# A Cautionary Note on Estimating Effect Sizes

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## Abstract

An increasingly popular approach to statistics is to focus on estimation and to forgo hypothesis testing altogether. Through an example, we show that estimates and confidence of effect sizes are overestimated the null is ignored, and when it is a plausible description of the data. We illustrate how this overestimation can be avoided by incorporating the plausibility of the null into the estimation process.

Consider the following scenario: Your colleague has just conducted an experiment for a Registered Report. The analysis yields  $p < 0.05$  and your colleague believes that the null hypothesis can be rejected. In line with recommendations both old (e.g., Grant, 1962; Loftus, 1996) and new (e.g., Harrington et al., 2019; Cumming, 2014) you convince your colleague that it is better to replace the  $p$ -value with an estimate of the effect size and a 95% confidence interval (but see Morey, Hoekstra, Rouder, Lee, & Wagenmakers, 2016). You also manage to convince your colleague to plot the data. Instead of simply reporting  $p < 0.05$ , the statistical analysis in the report is now more informative. The result is shown in Figure 1. In the text of the paper, the result is summarized as Cohen's  $d = 0.30$ ,  $CI = [0.02, 0.58]$ , in line with guidelines for reporting statistics (e.g., the guidelines of the Psychonomic Society<sup>1</sup>, or those of Psychological Science<sup>2</sup>). Given the results shown in Figure 1, what is a reasonable point estimate of the effect size? A straightforward answer is “0.30” which makes intuitive sense from an estimation perspective. However, your colleague now tells you about the nature of the experiment: plants grow faster when you talk to them.<sup>3</sup> Suddenly, a population effect size of “0” also appears plausible. Any

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<sup>1</sup><https://www.springer.com/psychology?SGWID=0-10126-6-1390050-0>

<sup>2</sup>[https://www.psychologicalscience.org/publications/psychological\\_science/ps-submissions#STAT](https://www.psychologicalscience.org/publications/psychological_science/ps-submissions#STAT)

<sup>3</sup>Specifically, imagine your colleague took 100 plants and measured their growth three times during two weeks. The first week 50 plants were randomly selected and spoken to while the other served as control. The next week, the roles reversed and the previously spoken to plants served as controls while the control plants were now talked to. The quantity of interest is the difference in growth between the weeks. This example is inspired by (Berger & Delampady, 1987).

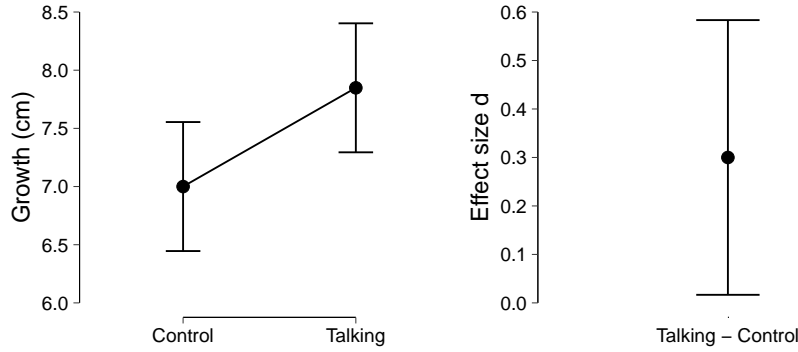


Figure 1: The left panel shows a descriptives plot with the mean and 95% confidence interval of the simulated plant growth. The right panel shows an estimate of the effect size, Cohen’s  $d$ , and a 95% confidence interval.

observed difference may merely be due to the individual differences between the plants.<sup>4</sup>

## When are Effect Sizes Overestimated?

Point estimates and confidence intervals are based on the alternative hypothesis and tend to overestimate effect sizes. This overestimation is caused by the strong assumption that the null or perinull hypothesis is irrelevant. However, the null or perinull can have high plausibility after seeing the data, thus, there can be evidence for the absence of an effect. This happens if the alternative is a-priori unlikely (e.g., for a null hypothesis like “Talking to plants does not affect their growth.”), or when the data are so uninformative that after seeing the data there is substantial uncertainty about which model best describes the data. If the null hypothesis still has a high plausibility after seeing the data, it is obvious that it cannot be ignored, but that is exactly what is done when estimates are only based on the alternative hypothesis. As a result, the estimates are overconfident and, because the null hypothesis would shrink the estimates towards zero, overestimated.

## A Spike-and-Slab Perspective

Here, we illustrate the overestimation and a remedy against it by reanalyzing the hypothetical data from Figure 1.<sup>5</sup> We consider the spike-and-slab model, which consists of two components. The first component corresponds to the position that talking to plants does not make them grow faster or slower ( $d = 0$ ). The second component corresponds to the position that speaking to plants does influence their growth ( $d \neq 0$ ). Here we view the spike-and-slab as a single model, although it can also be viewed as a form of Bayesian model averaging (see Box 1 for more details). Typically, an estimate of the effect size is solely based

<sup>4</sup>Unless you talk out loud, with consumption, and the plant is near.

<sup>5</sup>R code for the analysis is available at <https://osf.io/uq8st/>.

on the second component, which yields a point estimate and an uncertainty interval (for frequentists,  $d = 0.30$ , 95% CI: [0.02, 0.58]; for Bayesians  $d = 0.29$ , 95% CRI: [0.02, 0.57]). The spike-and-slab model also considers the possibility that an effect can be exactly zero and thus the estimate is an average of the two components and thus shrunk towards zero.<sup>6</sup> Figure 2 contrasts inference based on one component with inference based on both components by showing their 95% credible intervals and posterior means. The posterior mean and 95%

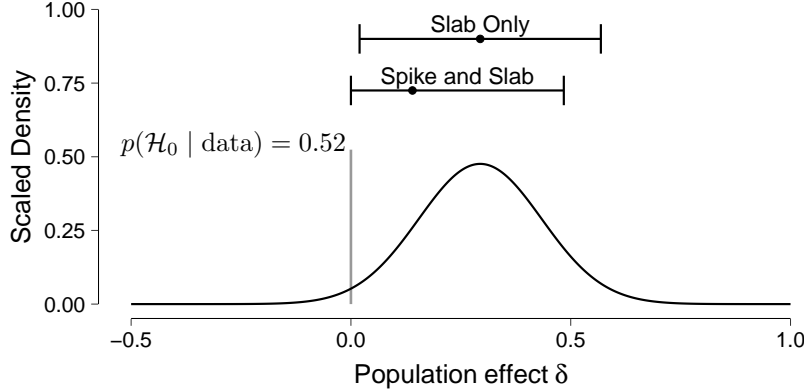


Figure 2: Visualization of the results from the spike-and-slab model. The black line represents the posterior distribution of effect size given the slab (i.e., the effect is non-zero). The posterior is scaled so that its mode equals the posterior probability of the alternative model. The gray line represents the posterior probability of the spike (i.e., the effect is exactly 0). The error bars and dots above the density show 95% credible intervals and the posterior mean for a posterior based on only the slab and for a posterior based on the spike-and slab.

credible interval of the spike-and-slab model are shrunk towards 0 compared to the results based on only the slab (0.14 (95% CRI: [0.00, 0.48]) vs. 0.29 (95% CRI: [0.02, 0.57])). This shrinkage is caused by the probability that the effect is absent, i.e., the null hypothesis is non-negligible. Note that although 0 appears to be included in the credible interval, this cannot be interpreted directly as evidence against an effect (or significance testing). Instead, the spike-and-slab posterior contains the evidence for there being an effect, combined with the evidence of the size of the effect across the two components. Note that as the posterior probability of the null decreases, the spike-and-slab results approach those of the alternative model.

## Discussion

We argue that estimates of effect sizes based only on the alternative hypothesis tend to be overconfident, in particular when a null or perinull hypothesis

<sup>6</sup>For the spike-and-slab model, the posterior distribution is constructed in the following manner:  $p(\delta|\text{data}) = 1\{\delta = 0\}pr(\mathcal{M}_0|\text{data}) + p(\delta|\text{data}, \mathcal{M}_1)pr(\mathcal{M}_1|\text{data})$ . Here,  $1\{\delta = 0\}$  is the Dirac delta function which represents the spike under the null,  $pr$  denotes probability of a model, and  $p$  denotes density related to the magnitude of the effect. Posterior model probabilities are obtained using prior model probabilities of  $1/2$ .

may also describe the data well. Consequently, point estimates and confidence intervals based solely on the alternative overestimate effect sizes. A solution for this overestimation is the spike-and-slab model which explicitly considers the possibility that an effect is exactly 0 (as is also advocated for by [Kiers & Tendeiro, 2019](#); [Rouder, Haaf, & Vandekerckhove, 2018](#)). Although this idea is not new, the influence of the null is still too often ignored in practice.

This approach may contrast with the popular estimation mindset, where it is argued that statistical significance should be abandoned in favor of estimation ([McShane, Gal, Gelman, Robert, & Tackett, 2019](#); [Valentine, Aloe, & Lau, 2015](#); [Cumming, 2014](#)). Some may argue that all null hypotheses are false ([Cohen, 1990](#); [Meehl, 1978](#)) and therefore there is no need to consider the component that states that an effect is exactly 0. However, there are statistical motivations to consider a point null ([Haaf, Ly, & Wagenmakers, 2019](#); [Berger & Delampady, 1987](#)) and several large-scale replications studies have demonstrated that a near-zero effect size is reasonable in practice, i.e., when the null is a plausible description of the data (e.g., see the meta-analyses conducted by [Klein et al., 2018](#); [Camerer et al., 2018](#); [Nosek & Lakens, 2014](#)).

### When (not) to consider the spike

When there is overwhelming evidence that an effect is non-zero, the results from a spike-and-slab model become virtually identical to those of a slab only model. Thus, in some situations, it can be argued that it is justified to ignore the spike and that the use of the spike-and-slab model is needlessly complicated. There are two main reasons for ignoring the spike. First, if the effect is known to be large then the posterior probability of the spike will be negligible. Second, if the presence of the effect is established, for instance by prior research, it is sensible to discard the spike a-priori. Yet, the spike does not hurt in either of these situations.

### Conclusion

In sum, we argue that descriptions of effect sizes based on only the alternative hypothesis are overconfident and as a consequence overestimate effect sizes. A remedy for this overestimation is the spike-and-slab model which explicitly accounts for the possibility of the null of no effect. Although this idea is not new, it remains underutilized in practice, and we hope this paper brings more attention to the spike-and-slab approach.

## References

- Amrhein, V., Greenland, S., & McShane, B. (2019). *Scientists rise up against statistical significance*. Nature Publishing Group.
- Berger, J. O., & Delampady, M. (1987). Testing precise hypotheses. *Statistical Science*, 2, 317–352.
- Camerer, C. F., Dreber, A., Holzmeister, F., Ho, T.-H., Huber, J., Johannesson, M., ... Wu, H. (2018). Evaluating replicability of social science experiments in *Nature* and *Science*. *Nature Human Behaviour*, 2, 637–644.
- Cohen, J. (1990). Things I have learned (thus far). *American Psychologist*, 45, 1304–1312.

- Cumming, G. (2014). The new statistics: Why and how. *Psychological Science*, 25, 7–29.
- Grant, D. A. (1962). Testing the null hypothesis and the strategy and tactics of investigating theoretical models. *Psychological Review*, 69, 54–61.
- Grünwald, P., Van Ommen, T., et al. (2017). Inconsistency of bayesian inference for misspecified linear models, and a proposal for repairing it. *Bayesian Analysis*, 12(4), 1069–1103.
- Haaf, J. M., Ly, A., & Wagenmakers, E.-J. (2019). Retire significance, but still test hypotheses. *Nature*, 567, 461.
- Haldane, J. B. S. (1932). A note on inverse probability. *Mathematical Proceedings of the Cambridge Philosophical Society*, 28, 55–61.
- Harrington, D., D’Agostino Sr, R. B., Gatsonis, C., Hogan, J. W., Hunter, D. J., Normand, S.-L. T., . . . Hamel, M. B. (2019). *New guidelines for statistical reporting in the journal*. Mass Medical Soc.
- Iverson, G. J., Wagenmakers, E.-J., & Lee, M. D. (2010). A model averaging approach to replication: The case of  $p_{rep}$ . *Psychological Methods*, 15, 172–181.
- Jevons, W. S. (1874/1913). *The principles of science: A treatise on logic and scientific method*. London: MacMillan.
- Kiers, H., & Tendeiro, J. (2019). With Bayesian estimation one can get all that Bayes factors offer, and more. *manuscript submitted for publication*. Retrieved from [psyarxiv.com/zbpmy](https://psyarxiv.com/zbpmy)
- Klein, R., Vianello, M., Hasselman, F., Adams, B., Adams, R., Alper, S., . . . Nosek, B. (2018). Many Labs 2: Investigating variation in replicability across sample and setting. *Advances in Methods and Practices in Psychological Science*, 1, 443–490.
- Loftus, G. R. (1996). Psychology will be a much better science when we change the way we analyze data. *Current Directions in Psychological Science*, 5, 161–171.
- McShane, B. B., Gal, D., Gelman, A., Robert, C., & Tackett, J. L. (2019). Abandon statistical significance. *The American Statistician*, 73(sup1), 235–245.
- Meehl, P. E. (1978). Theoretical risks and tabular asterisks: Sir Karl, Sir Ronald, and the slow progress of soft psychology. *Journal of Consulting and Clinical Psychology*, 46, 806–834.
- Minka, T. P. (2000). Bayesian model averaging is not model combination. Available electronically at <http://www.stat.cmu.edu/minka/papers/bma.html>, 1–2.
- Morey, R. D., Hoekstra, R., Rouder, J. N., Lee, M. D., & Wagenmakers, E.-J. (2016). The fallacy of placing confidence in confidence intervals. *Psychonomic Bulletin & Review*, 23, 103–123.
- Nosek, B. A., & Lakens, D. (2014). Registered reports: A method to increase the credibility of published results. *Social Psychology*, 45, 137–141.
- Rouder, J. N., Haaf, J. M., & Vandekerckhove, J. (2018). Bayesian inference for psychology, part IV: Parameter estimation and Bayes factors. *Psychonomic Bulletin & Review*, 25, 102–113.
- Valentine, J. C., Aloe, A. M., & Lau, T. S. (2015). Life after nhst: How to describe your data without “p-ing” everywhere. *Basic and Applied Social Psychology*, 37(5), 260–273.

- Wrinch, D., & Jeffreys, H. (1921). On certain fundamental principles of scientific inquiry. *Philosophical Magazine*, 42, 369–390.
- Zellner, A., & Siow, A. (1980). Posterior odds ratios for selected regression hypotheses. In J. M. Bernardo, M. H. DeGroot, D. V. Lindley, & A. F. M. Smith (Eds.), *Bayesian statistics* (pp. 585–603). Valencia: University Press.
- Zellner, A., & Vandaele, W. (1975). Bayes–Stein estimators for k-means, regression and simultaneous equation models. In S. E. Fienberg & A. Zellner (Eds.), *Studies in Bayesian econometrics and statistics in honor of Leonard J. Savage* (pp. 627–653). Amsterdam, The Netherlands: North-Holland Publishing Company.

#### Box 1: The Spike-and-Slab as Bayesian model averaging

The spike-and-slab can be seen as a single model that consists of two components: the slab, which accounts for the possibility that the effect is non-zero, and the spike, which accounts for an effect of exactly 0. However, the spike-and-slab can also be seen as a form of Bayesian model averaging. From that perspective, the spike and the slab are two individual models. The slab represents the unconstrained model that freely estimates effect size, and the spike represents the constrained model where the effect size is fixed to 0. Next, the results for each model are weighted by the posterior model probabilities and averaged, so that inference can be made using results that are averaged over the models considered. It has been shown repeatedly that averaging over the models considered provides the best predictive performance (Zellner & Vandaele, 1975, pp. 640–641, as described in Zellner & Siow, 1980, p. 600–601; Haldane, 1932, p. 57, Iverson, Wagenmakers, & Lee, 2010, Rouder et al., 2018), and conceptually similar ideas date back much further (Wrinch & Jeffreys, 1921, p. 387, Jevons, 1874/1913). Note that these two perspective—a two-component model or averaging of two models—differ in semantics but are mathematically equivalent. Here we view the spike-and-slab as a single model because we focus on estimation.