

# **Probabilistic Robotics Course**

## **Dynamic Bayesian Networks (Filtering)**

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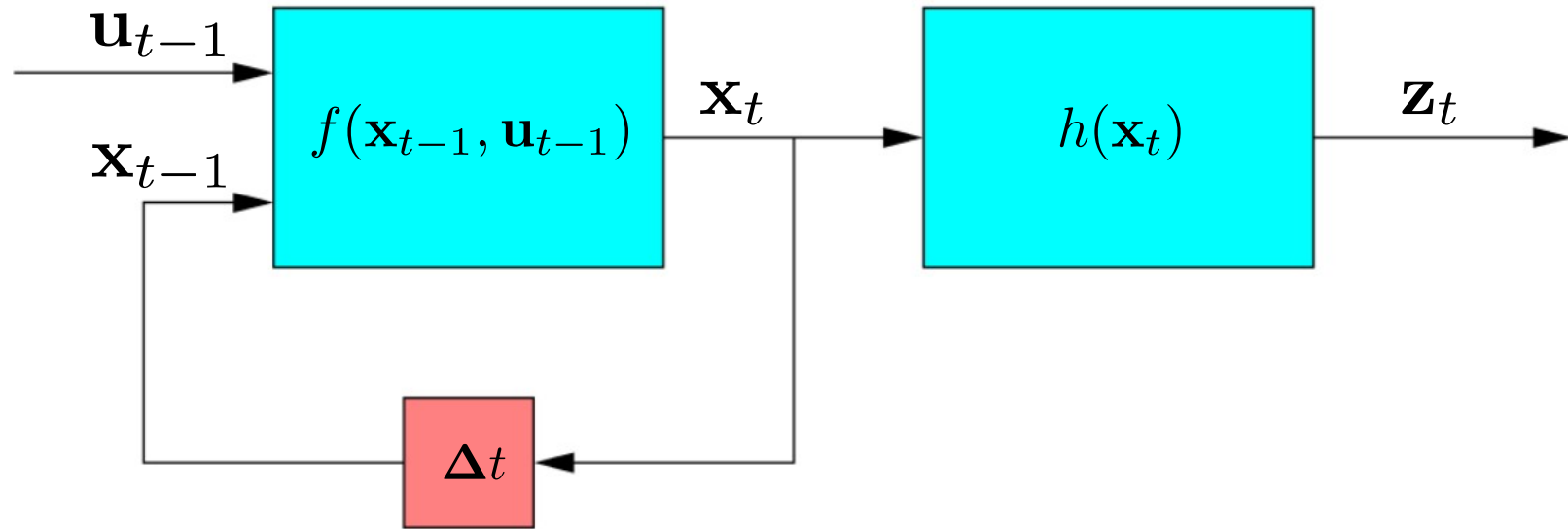
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# Overview

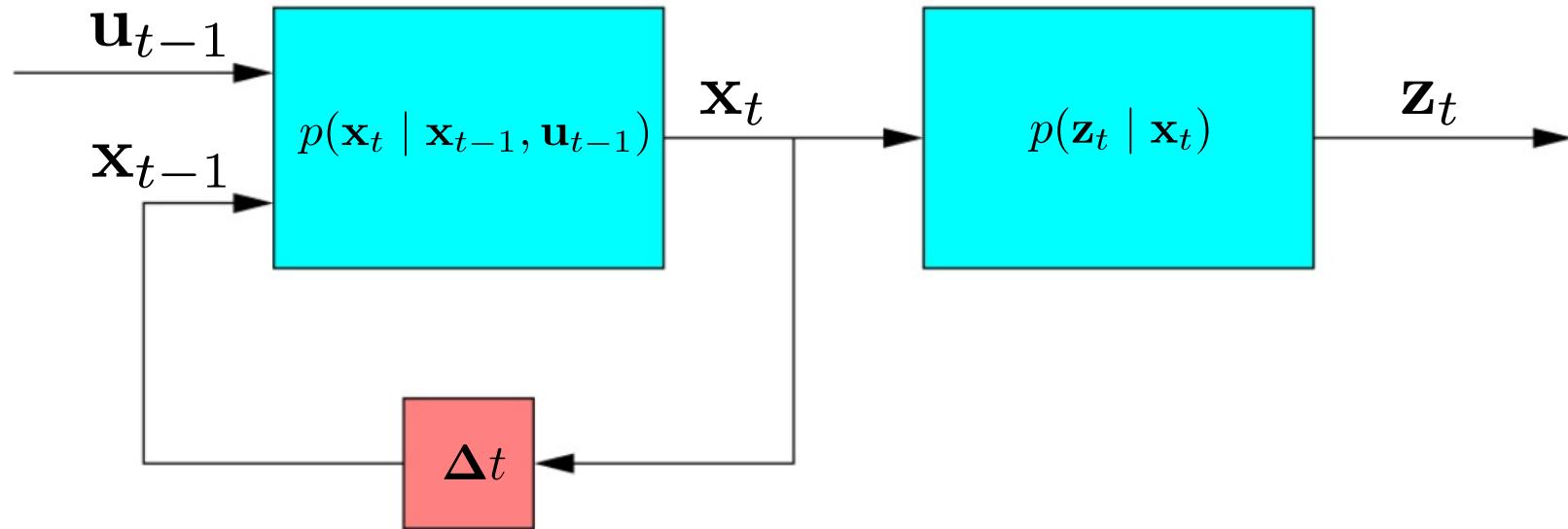
- Probabilistic Dynamic Systems
- Dynamic Bayesian Networks (DBN)
- Inference on DBN
- Recursive Bayes Equation

# Dynamic System Deterministic View



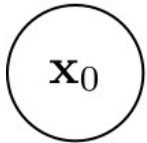
- $f(\mathbf{x}_{t-1}, \mathbf{u}_{t-1})$ : transition function
- $h(\mathbf{x}_t)$ : observation function
- $\mathbf{x}_{t-1}$ : previous state
- $\mathbf{x}_t$ : current state
- $\mathbf{z}_t$ : current observation
- $\mathbf{u}_{t-1}$ : previous control/action
- $\Delta t$ : delay

# Dynamic System Probabilistic View



- $p(\mathbf{x}_t \mid \mathbf{x}_{t-1}, \mathbf{u}_{t-1})$ : transition model
- $p(\mathbf{z}_t \mid \mathbf{x}_t)$ : observation model
- $\mathbf{x}_{t-1}$ : previous state
- $\mathbf{x}_t$ : current state
- $\mathbf{z}_t$ : current observation
- $\mathbf{u}_{t-1}$ : previous control/action
- $\Delta t$ : delay

# Evolution of a Dynamic System: State

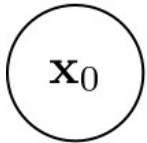


Let's start from a known initial state distribution  $p(\mathbf{x}_0)$ .

# Evolution of a Dynamic System: Control



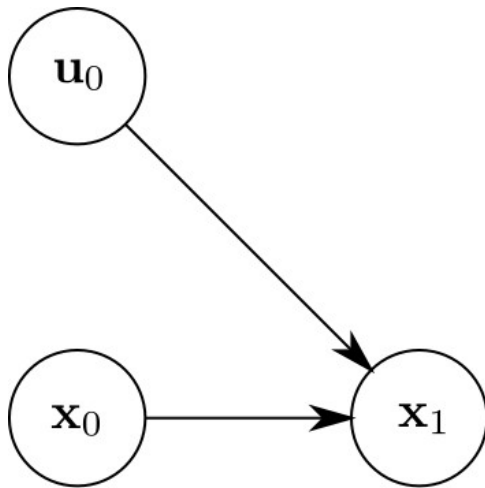
$\mathbf{u}_0$



$\mathbf{x}_0$

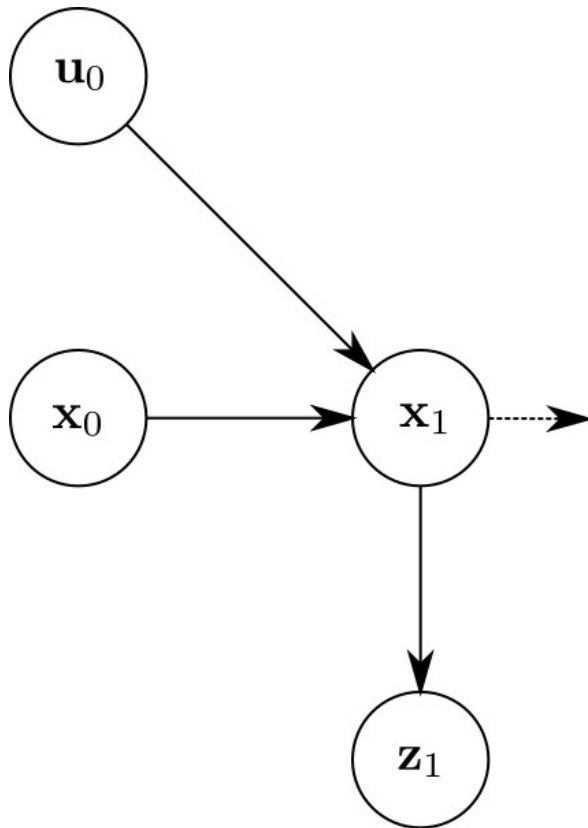
A control  $\mathbf{u}_0$  becomes available.

# Evolution of a Dynamic System: Transition



The transition model  $p(\mathbf{x}_t \mid \mathbf{x}_{t-1}, \mathbf{u}_{t-1})$  correlates the current state  $\mathbf{x}_1$  with the previous control  $\mathbf{u}_0$  and the previous state  $\mathbf{x}_0$ .

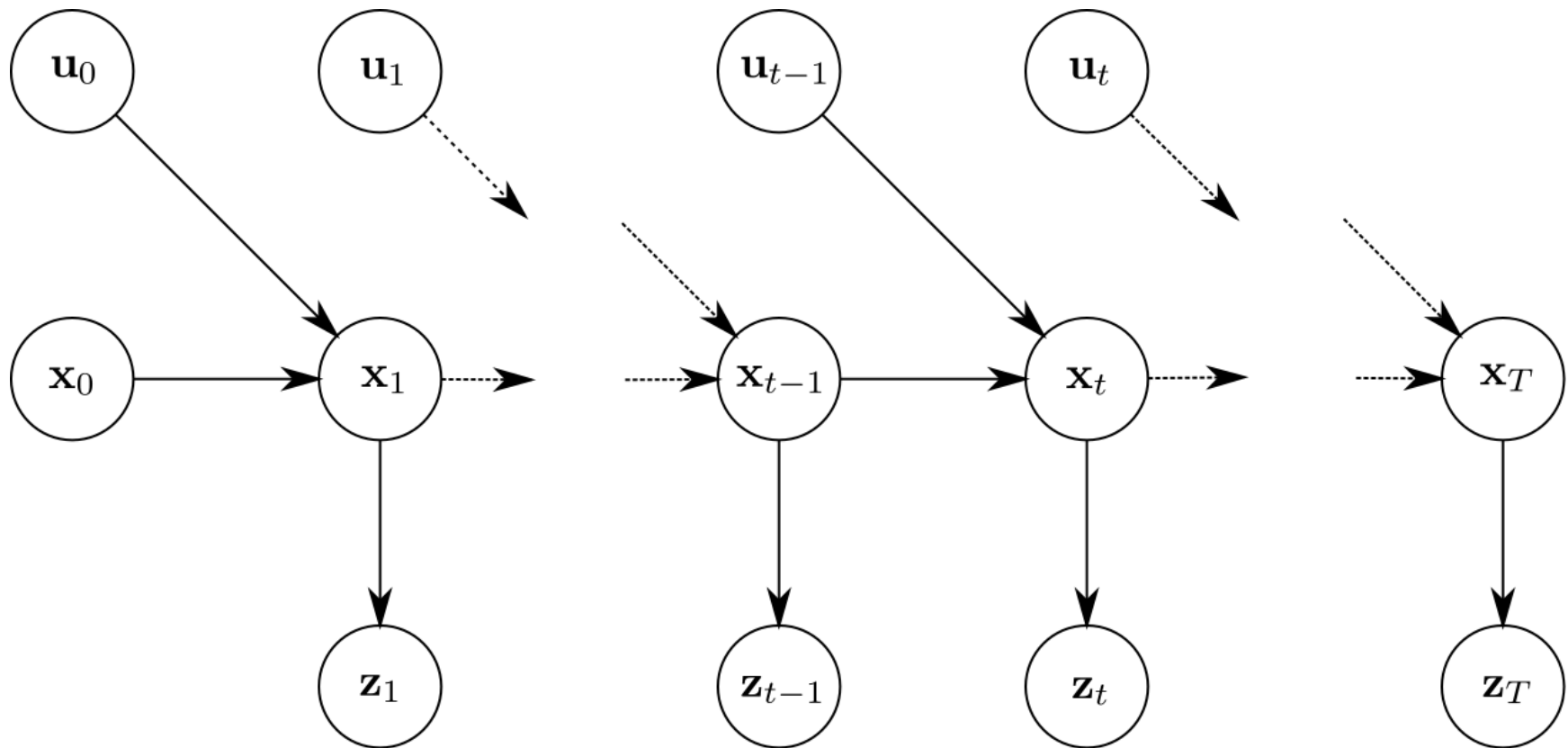
# Evolution of a Dynamic System: Observation



The observation model  $p(\mathbf{z}_t \mid \mathbf{x}_t)$  correlates the observation  $\mathbf{z}_1$  and the current state  $\mathbf{x}_1$ .

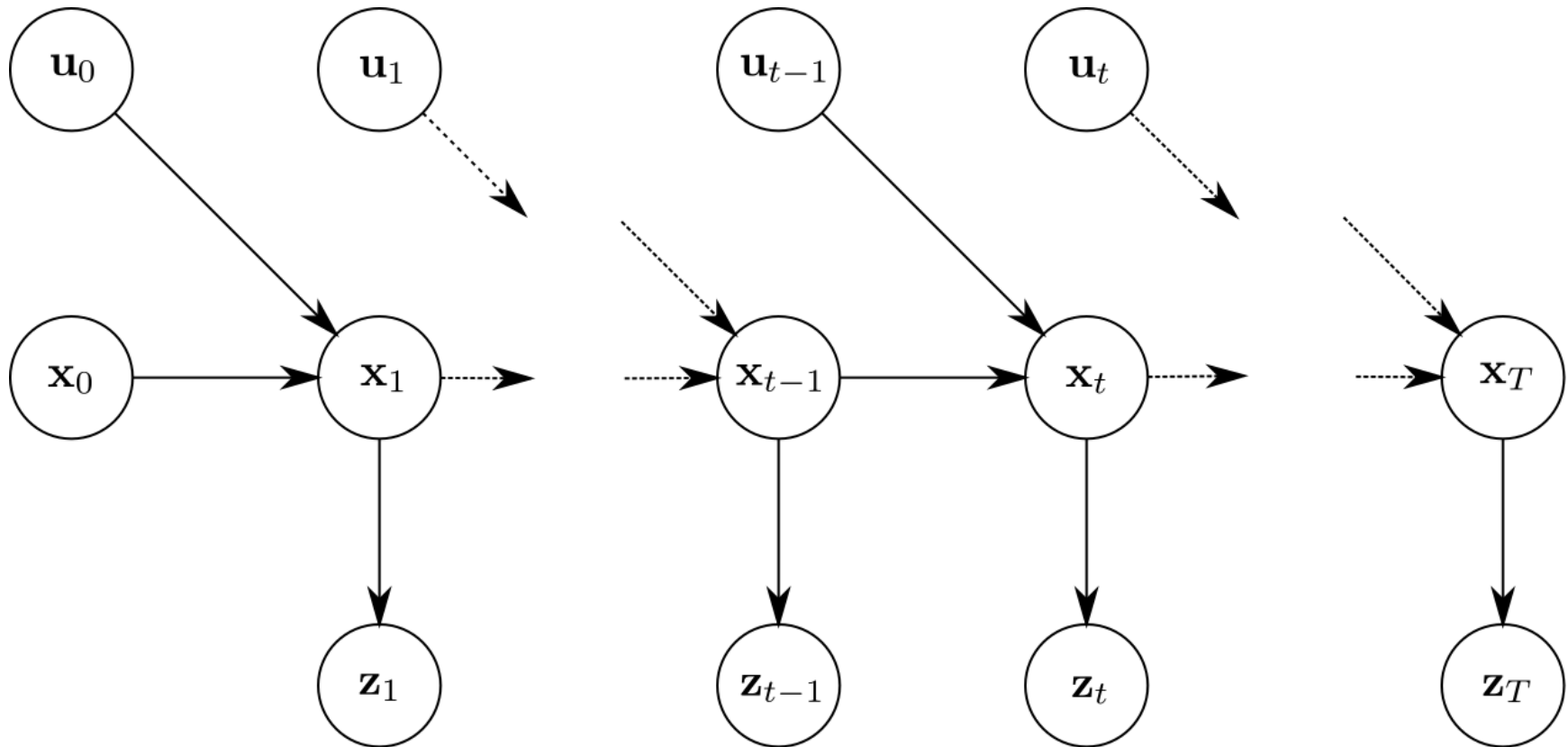


# Evolution of a Dynamic System



This leads to a recurrent structure, that depends on the *time  $t$* .

# Dynamic Bayesian Networks (DBN)



- Graphical representations of stochastic dynamic processes
- Characterized by a recurrent structure

# States in a DBN

The domain of the states  $\mathbf{x}_t$ , the controls  $\mathbf{u}_t$  and the observations  $\mathbf{z}_t$  are not restricted to be boolean or discrete.

Examples:

- Robot localization, with a laser range finder
  - States  $\mathbf{x}_t \in SE(2)$ , isometries on a plane
  - Observations  $\mathbf{z}_t \in \mathcal{R}^{\#beams}$ , laser range measurements
  - Controls  $\mathbf{u}_t \in \mathcal{R}^2$ , translational and rotational speed
- HMM (Hidden Markov Model)
  - States  $\mathbf{x}_t \in [X_1, \dots, X_{N_x}]$ , finite states
  - Observations  $\mathbf{z}_t \in [Z_1, \dots, Z_{N_z}]$ , finite observations
  - Controls  $\mathbf{u}_t \in [U_1, \dots, U_{N_u}]$ , finite controls

Inference in a DBN requires to design a data structure that can represent a *distribution* over states.

# Typical Inferences in a DBN

In a dynamic system, usually<sup>1</sup> we know:

- the observations  $\mathbf{z}_{1:T}$  made by the system, because we *measure* them.
- the controls  $\mathbf{u}_{0:T-1}$ , because we *issue* them

Typical inferences in a DBN:

name	query	known
Filtering	$p(\mathbf{x}_T   \mathbf{u}_{0:T-1}, \mathbf{z}_{1:T})$	$\mathbf{u}_{0:T-1}, \mathbf{z}_{1:T}$
Smoothing	$p(\mathbf{x}_t   \mathbf{u}_{0:T-1}, \mathbf{z}_{1:T}), 0 < t < T$	$\mathbf{u}_{0:T-1}, \mathbf{z}_{1:T}$
Max a Posteriori	$\operatorname{argmax}_{\mathbf{x}_{0:T}} p(\mathbf{x}_{0:T}   \mathbf{u}_{0:T-1}, \mathbf{z}_{1:T})$	$\mathbf{u}_{0:T-1}, \mathbf{z}_{1:T}$

<sup>1</sup>usually does not mean always

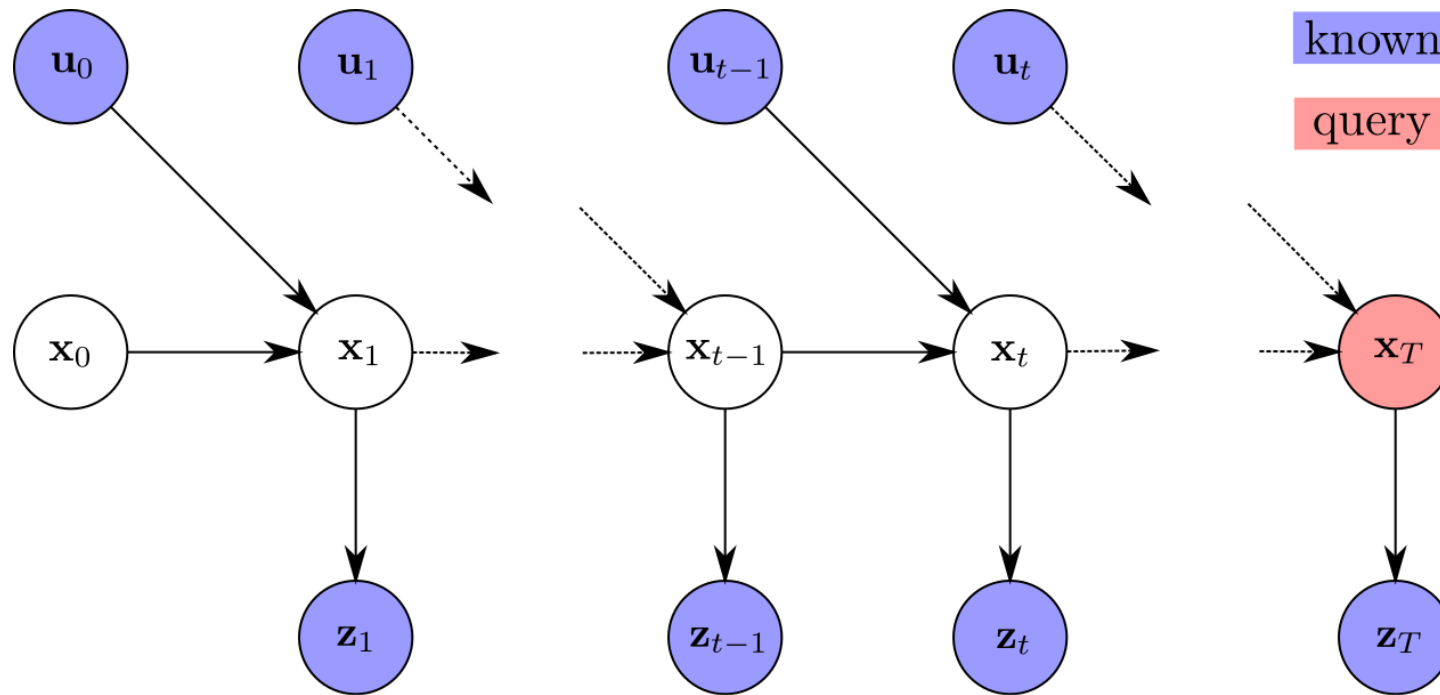
# Typical Inferences in a DBN

Using the traditional tools for Bayes Networks is not a good idea:

- too many variables (potentially infinite) render the solution intractable
- the domains are not necessarily discrete

However, we can exploit the recurrent structure to design procedures that take advantage of it

# DBN Inference: Filtering

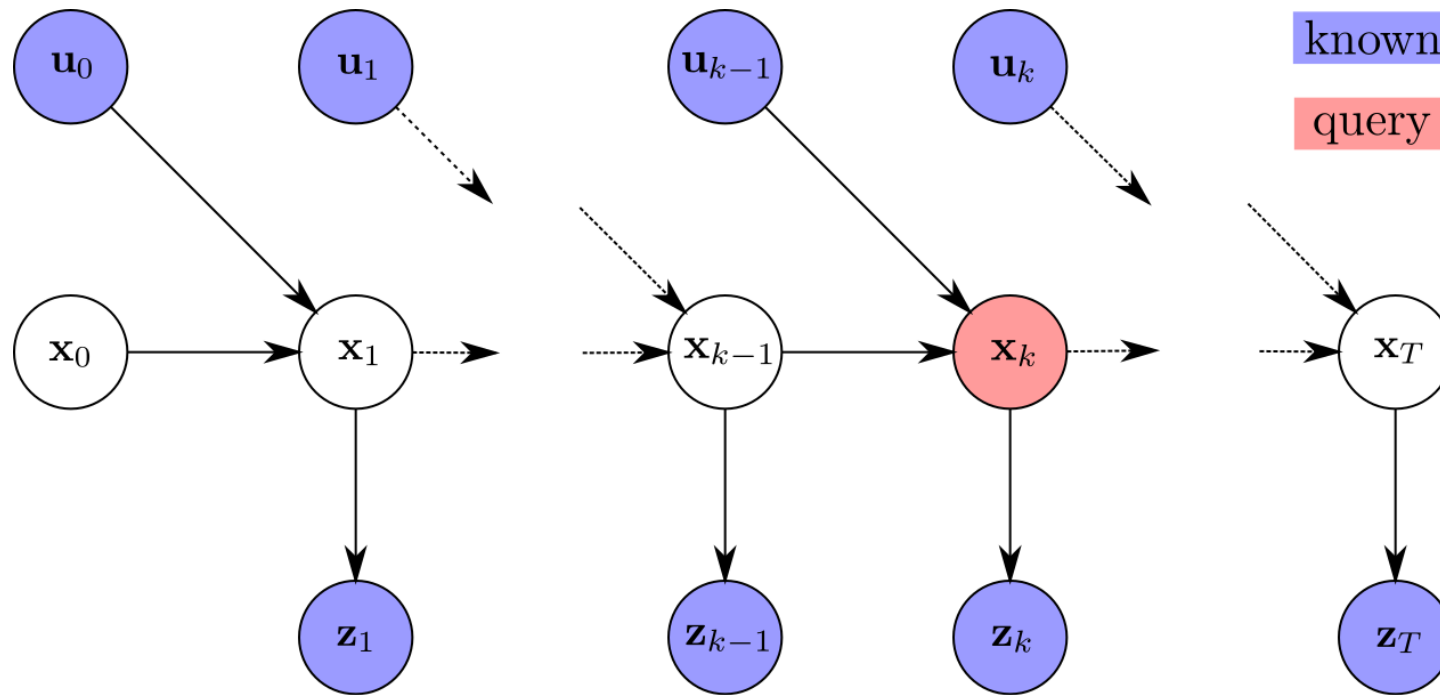


Given:

- the sequence of all observations  $\mathbf{z}_{1:T}$  up to the current time  $T$
- the sequence of all controls  $\mathbf{u}_{0:T-1}$

we want to compute the distribution over the current state  $p(\mathbf{x}_T | \mathbf{u}_{0:T-1}, \mathbf{z}_{1:T})$ .

# DBN Inference: Smoothing



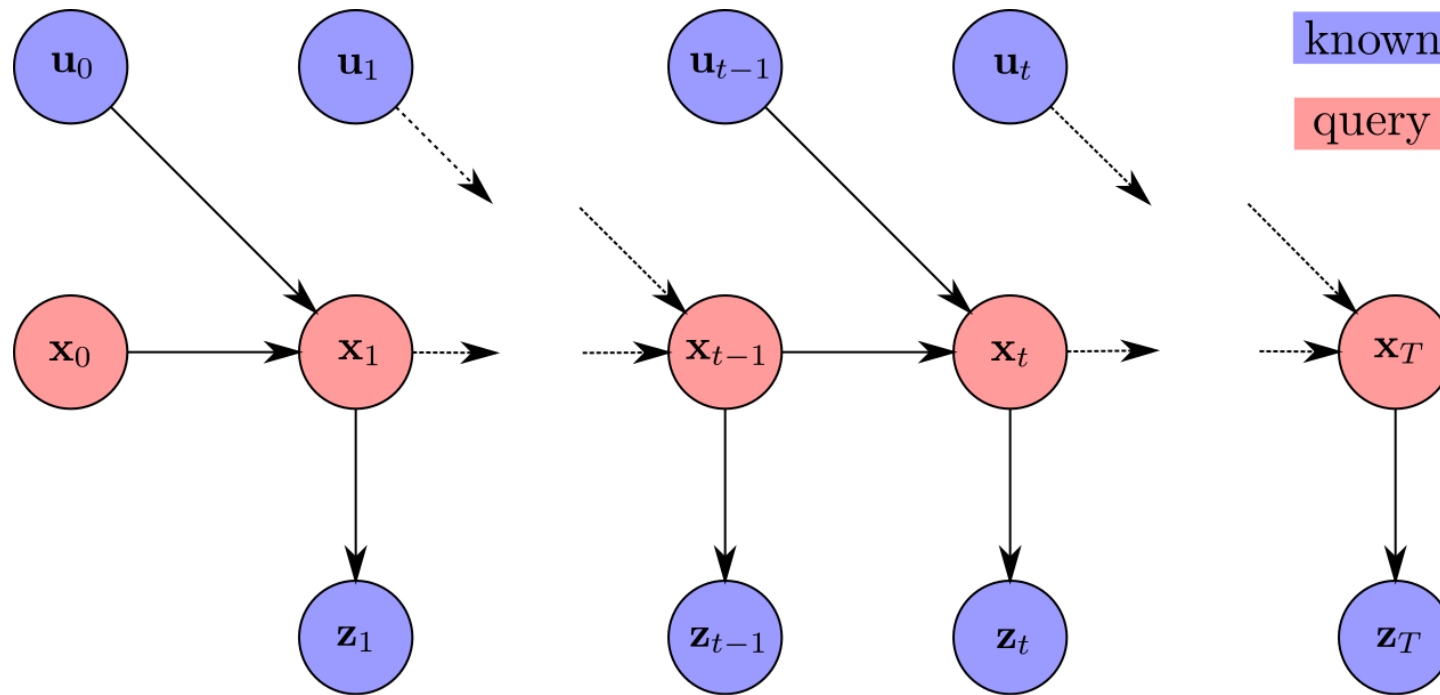
Given:

- the sequence of all observations  $\mathbf{z}_{1:T}$  up to the current time  $T$
- the sequence of all controls  $\mathbf{u}_{0:T-1}$

we want to compute the distribution over a past state  $p(\mathbf{x}_k | \mathbf{u}_{0:T-1}, \mathbf{z}_{1:T})$ .

Knowing also the controls  $\mathbf{u}_{0:T-1}$  and the observations  $\mathbf{z}_{1:T}$  *after* time  $k$ , leads to more accurate estimates than pure filtering.

# DBN Inference: Maximum a Posteriori



Given:

- the sequence of all observations  $\mathbf{z}_{1:T}$  up to the current time  $T$
- the sequence of all controls  $\mathbf{u}_{0:T-1}$

we want to find the most likely *trajectory* of states  $\mathbf{x}_{0:T}$ .

In this case we are not seeking for a distribution.  
Just the most likely *sequence*.



# DBN Inference: Belief

- Algorithms for performing inference on a DBN keep track of the *estimate* of a distribution of states.
- This distribution should be stored in an appropriate data structure.
- The structure depends on:
  - the knowledge of the characteristics of the distribution (e.g. Gaussian)
  - the domain of the state variables (e.g. continuous vs discrete)

When we write  $b(\mathbf{x}_t)$  we mean our current belief of  $p(\mathbf{x}_t|\dots)$

The algorithms for performing inference on a DBN work by updating a belief.

# DBN Inference: Belief

- In the simple case of a system with discrete state  $\mathbf{x} \in \{X_{1:n}\}$ , the belief can be represented through an array  $\mathbf{x}$  of float values. Each cell of the array  $\mathbf{x}[i] = p(\mathbf{x} = X_i)$  contains the probability of that state
- If our system has a continuous state and we know it is distributed according to a Gaussian, we can represent the belief through its parameters (mean and covariance matrix)
- If the state is continuous but the distribution is unknown, we can use some approximate representation (e.g. weighed samples of state values).

# Filtering: Bayes Recursion

We want to compute:  $p(\mathbf{x}_T | \mathbf{u}_{0:T-1}, \mathbf{z}_{1:T})$

We know:

- the observations  $\mathbf{z}_{1:T}$
- the controls  $\mathbf{u}_{0:T-1}$
- $p(\mathbf{x}_t | \mathbf{x}_{t-1}, \mathbf{u}_{t-1})$ : the transition model. It is a function that, given the previous state  $\mathbf{x}_{t-1}$  and control  $\mathbf{u}_{t-1}$ , tells us how likely it is to land in state  $\mathbf{x}_t$ .
- $p(\mathbf{z}_t | \mathbf{x}_t)$ : the observation model. It is a function, that given the current state  $\mathbf{x}_t$ , tells us how likely it is to observe  $\mathbf{z}_t$ .
- $b(\mathbf{x}_{t-1})$ , which is our belief about the previous state  
$$p(\mathbf{x}_{t-1} | \mathbf{u}_{0:t-2}, \mathbf{z}_{1:t-1})$$

# Filtering: Bayes Rule

$$p(\mathbf{x}_T | \mathbf{u}_{0:T-1}, \mathbf{z}_{1:T}) = \quad (1)$$

- splitting  $\mathbf{z}_t$ :

$$= p(\underbrace{\mathbf{x}_t}_A | \underbrace{\mathbf{z}_t}_B, \underbrace{\mathbf{u}_{0:t-1}, \mathbf{z}_{1:t-1}}_C) \quad (2)$$

- recall the conditional Bayes rule  $p(A|B, C) = \frac{p(B|A, C)p(A|C)}{p(B|C)}$

$$= \frac{p(\underbrace{\mathbf{z}_t}_B | \underbrace{\mathbf{x}_t}_A, \underbrace{\mathbf{u}_{0:t-1}, \mathbf{z}_{1:t-1}}_C) p(\underbrace{\mathbf{x}_t}_A | \underbrace{\mathbf{u}_{0:t-1}, \mathbf{z}_{1:t-1}}_C)}{p(\underbrace{\mathbf{z}_t}_B | \underbrace{\mathbf{u}_{0:t-1}, \mathbf{z}_{1:t-1}}_C)} \quad (3)$$

# Filtering: Denominator

- let the denominator

$$\eta_t = 1/p(\mathbf{z}_t \mid \mathbf{u}_{0:t-1}, \mathbf{z}_{1:t-1}) \quad (4)$$

Note that  $\eta_t$  does not depend on the state  $\mathbf{x}$ , thus to the extent of our computation is just a normalizing constant.

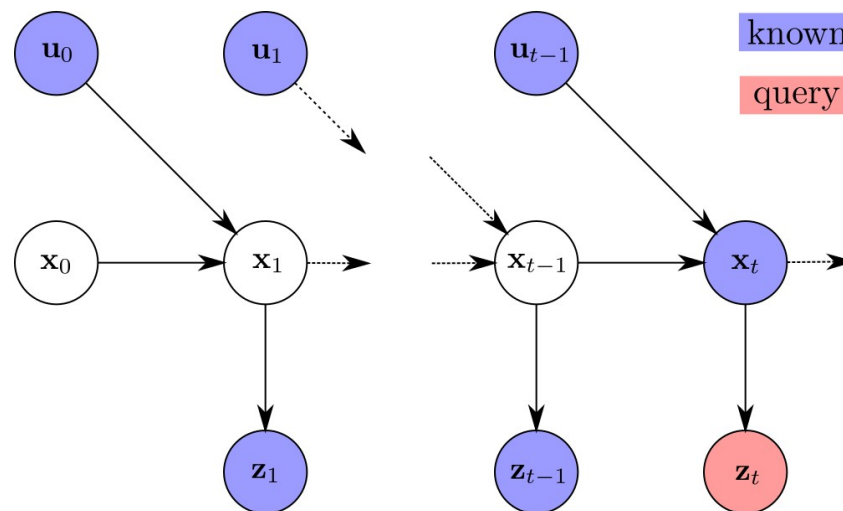
We will come back to the denominator later.

# Filtering: Observation model

- our filtering equation becomes:

$$\eta_t p(\mathbf{z}_t \mid \mathbf{x}_t, \mathbf{u}_{0:t-1}, \mathbf{z}_{1:t-1}) p(\mathbf{x}_t \mid \mathbf{u}_{0:t-1}, \mathbf{z}_{1:t-1}) \quad (5)$$

Note that  $p(\mathbf{z}_t \mid \mathbf{x}_t, \mathbf{u}_{0:t-1}, \mathbf{z}_{1:t-1})$  means this:



- if we know  $\mathbf{x}_t$ , we do not need to know  $\mathbf{u}_{0:t-1}, \mathbf{z}_{1:t-1}$  to predict  $\mathbf{z}_t$ , since the state  $\mathbf{x}_t$  encodes all the knowledge about the past (Markov assumption):

$$p(\mathbf{z}_t \mid \mathbf{x}_t, \mathbf{u}_{0:t-1}, \mathbf{z}_{1:t-1}) = p(\mathbf{z}_t \mid \mathbf{x}_t) \quad (6)$$

# Filtering: Transition Model

- thus, our current equation is:

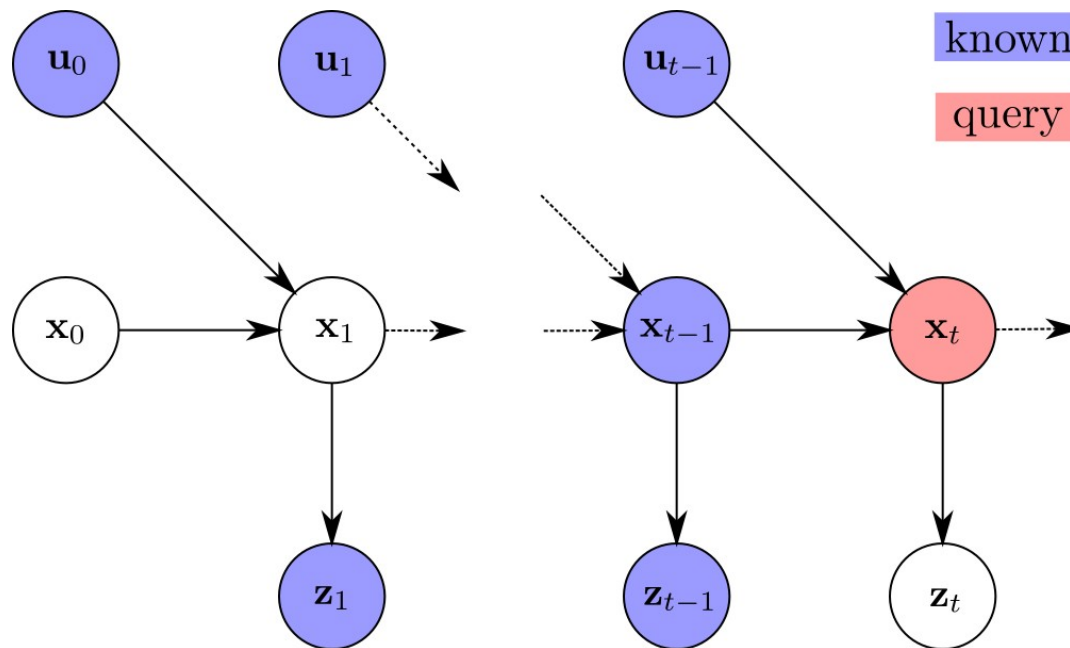
$$p(\mathbf{x}_t \mid \mathbf{u}_{0:t-1}, \mathbf{z}_{1:t}) = \eta_t p(\mathbf{z}_t \mid \mathbf{x}_t) p(\mathbf{x}_t \mid \mathbf{u}_{0:t-1}, \mathbf{z}_{1:t-1}) \quad (7)$$

Still the second part of the equation is obscure.

Our task is to manipulate it, to get something that matches our preconditions.

# Filtering: Transition Model

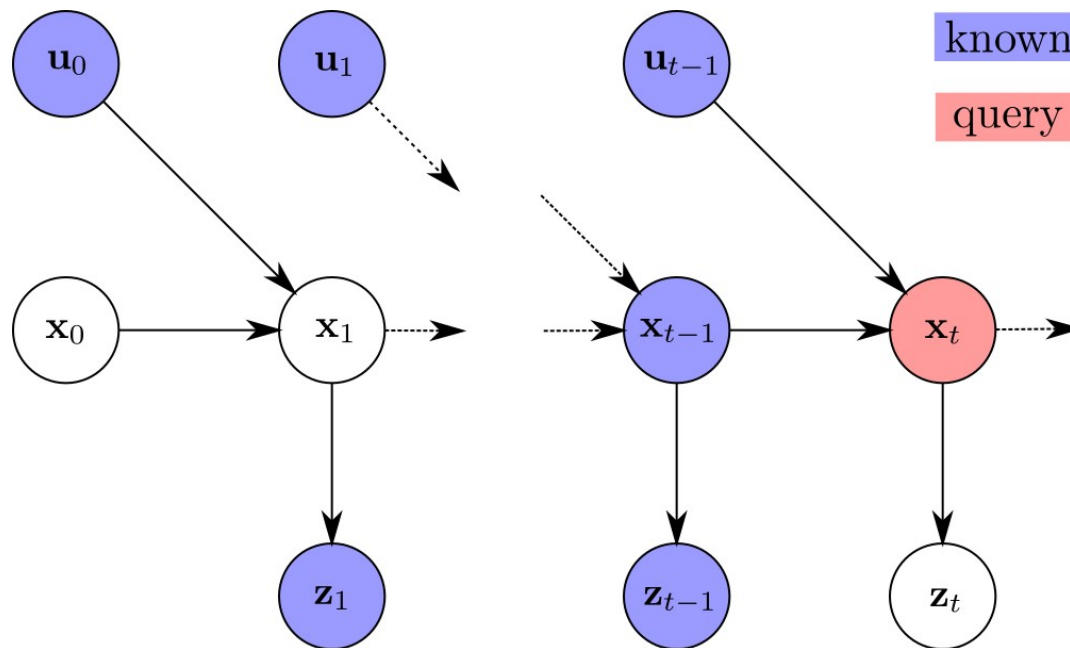
Knowing  $\mathbf{x}_{t-1}$  would make our life much easier, as we could repeat the trick done for the observation model:





# Filtering: Transition Model

Knowing  $\mathbf{x}_{t-1}$  would make our life much easier, as we could repeat the trick done for the observation model:



■ thus:

$$p(\mathbf{x}_t \mid \mathbf{x}_{t-1}, \mathbf{u}_{0:t-1}, \mathbf{z}_{1:t-1}) = p(\mathbf{x}_t \mid \mathbf{x}_{t-1}, \mathbf{u}_{t-1}) \quad (8)$$

# Filtering: Transition Model

The sad truth is that we do not have  $\mathbf{x}_{t-1}$ , however:

- recalling the probability identities:

*marginalization:* 
$$p(A|C) = \sum_B p(A, B|C) \quad (9)$$

*chain rule:* 
$$p(A, B|C) = p(A|B, C)p(B|C) \quad (10)$$

- by combining the two above we obtain:

$$p(A|C) = \sum_B p(A|B, C)p(B|C) \quad (11)$$

# Filtering: Transition Model

- let's look again at our problematic equation, and put some letters

$$\begin{aligned} & p(\underbrace{\mathbf{x}_t}_A \mid \underbrace{\mathbf{u}_{0:t-1}, \mathbf{z}_{1:t-1}}_C) = \\ & \sum_{\mathbf{x}_{t-1}} p(\underbrace{\mathbf{x}_t}_A \mid \underbrace{\mathbf{x}_{t-1}}_B, \underbrace{\mathbf{u}_{0:t-1}, \mathbf{z}_{1:t-1}}_C) p(\underbrace{\mathbf{x}_{t-1}}_B \mid \underbrace{\mathbf{u}_{0:t-1}, \mathbf{z}_{1:t-1}}_C) \end{aligned}$$

- putting in the result of Eq. (8), we highlight the transition model as:

$$= \sum_{\mathbf{x}_{t-1}} p(\mathbf{x}_t \mid \mathbf{x}_{t-1}, \mathbf{u}_{t-1}) p(\mathbf{x}_{t-1} \mid \mathbf{u}_{0:t-1}, \mathbf{z}_{1:t-1}) \quad (12)$$

$$p(A|C) = \sum_B p(A|B, C)p(B|C)$$

# Filtering: Wrapup

- after our efforts, we figure out that the recursive filtering equation is the following:

$$p(\mathbf{x}_t \mid \mathbf{u}_{0:t-1}, \mathbf{z}_{1:t}) = \quad (13)$$

$$\eta_t p(\mathbf{z}_t \mid \mathbf{x}_t) \sum_{\mathbf{x}_{t-1}} p(\mathbf{x}_t \mid \mathbf{x}_{t-1}, \mathbf{u}_{t-1}) p(\mathbf{x}_{t-1} \mid \mathbf{u}_{0:t-1}, \mathbf{z}_{1:t-1})$$

Yet, if in the last term of the product in the summation, we would not have a dependency from  $\mathbf{u}_{t-1}$ , we would have a *recursive* equation.

Luckily we have:

$$p(\mathbf{x}_{t-1} \mid \mathbf{u}_{0:t-1}, \mathbf{z}_{1:t-1}) = p(\mathbf{x}_{t-1} \mid \mathbf{u}_{0:t-2}, \mathbf{z}_{1:t-1}) \quad (14)$$

Since the last control has no influence on  $\mathbf{x}_{t-1}$ , if we don't know  $\mathbf{x}_t$ .

# Filtering: Wrapup

- we can finally write the recursive equation of filtering as:

$$\overbrace{p(\mathbf{x}_t \mid \mathbf{u}_{0:t-1}, \mathbf{z}_{1:t})}^{b(\mathbf{x}_t)} = \quad (15)$$

$$\eta_t p(\mathbf{z}_t \mid \mathbf{x}_t) \sum_{\mathbf{x}_{t-1}} p(\mathbf{x}_t \mid \mathbf{x}_{t-1}, \mathbf{u}_{t-1}) \underbrace{p(\mathbf{x}_{t-1} \mid \mathbf{u}_{0:t-2}, \mathbf{z}_{1:t-1})}_{b(\mathbf{x}_{t-1})}$$

During the estimation, we do not have the true distribution, but rather the beliefs *estimate*.

- Eq. (16) tells us how to update a current belief once new observations/controls become available:

$$b(\mathbf{x}_t) = \eta_t p(\mathbf{z}_t \mid \mathbf{x}_t) \sum_{\mathbf{x}_{t-1}} p(\mathbf{x}_t \mid \mathbf{x}_{t-1}, \mathbf{u}_{t-1}) b(\mathbf{x}_{t-1}) \quad (16)$$

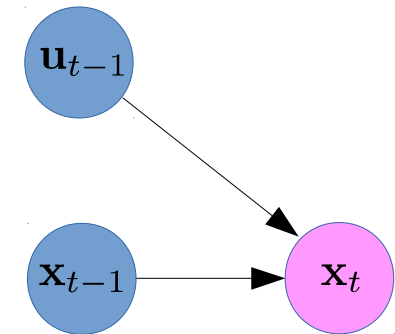
# Normalizer: $\eta_t$

The *normalizer*  $\eta_t$  is just a constant ensuring that  $b(\mathbf{x}_t)$  is still a probability distribution:

$$\eta_t = \frac{1}{\sum_{\mathbf{x}_t} p(\mathbf{z}_t \mid \mathbf{x}_t) \sum_{\mathbf{x}_{t-1}} p(\mathbf{x}_t \mid \mathbf{x}_{t-1}, \mathbf{u}_{t-1}) b(\mathbf{x}_{t-1})} \quad (17)$$

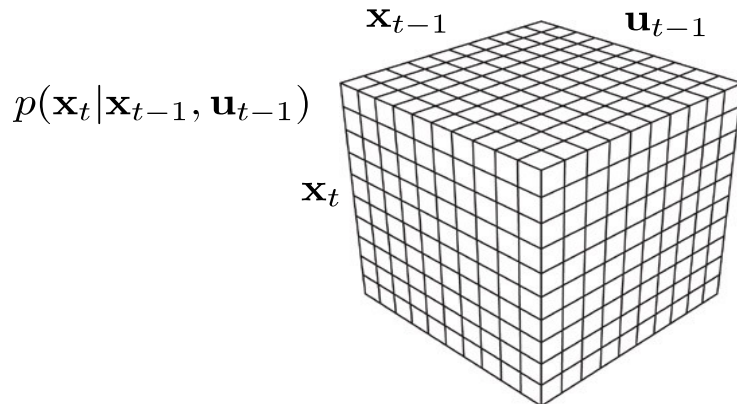
# Filtering: Alternative Formulation

**Predict:** incorporate in the last belief  $b_{t-1|t-1}$  the most recent control  $\mathbf{u}_{t-1}$ .



Ingredients:

- Transition model

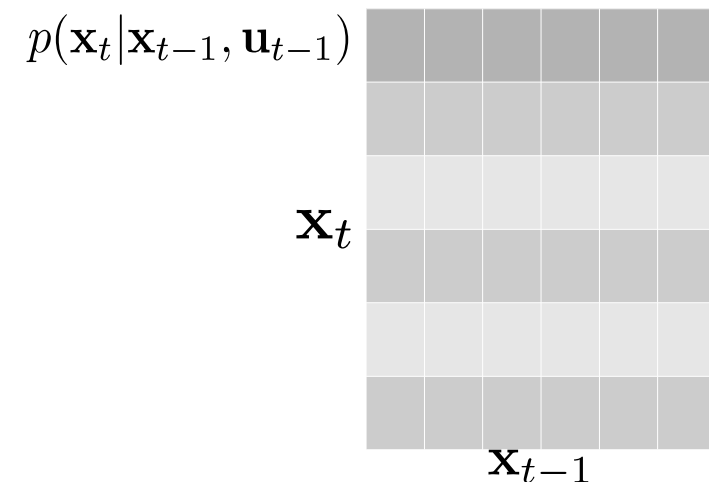


- Prior belief



- Control  $\mathbf{u}_{t-1}$

The control is known, so we can work with a "2D" distribution selected according to the current  $\mathbf{u}_{t-1}$ .



# Filtering: Alternative Formulation

## Predict:

- From the transition model and the last state, compute the following joint distribution through *chain rule*:

$$p(\mathbf{x}_t, \mathbf{x}_{t-1} | t-1) = p(\mathbf{x}_t | \mathbf{x}_{t-1}, \mathbf{u}_{t-1}) \underbrace{p(\mathbf{x}_{t-1} | t-1)}_{b_{t-1|t-1}}$$

- From the joint, remove  $\mathbf{x}_{t-1}$  through *marginalization*:

$$\underbrace{p(\mathbf{x}_t | t-1)}_{b_{t|t-1}} = \sum_{\mathbf{x}_{t-1}} p(\mathbf{x}_t, \mathbf{x}_{t-1} | t-1)$$

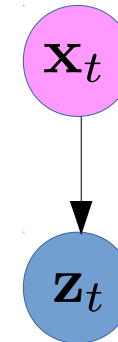
- Programmatically (discrete case)

```
BeliefType b_pred = BeliefType::Zero;
for (x_i : X)
    for (x_j: X)
        b_pred[x_j] += b[x_i]*transitionModel(x_j,x_i,u);
```



# Filtering: Alternative Formulation

**Update:** incorporate in the predicted belief  $b_{t|t-1}$  the new measurement  $\mathbf{z}_t$

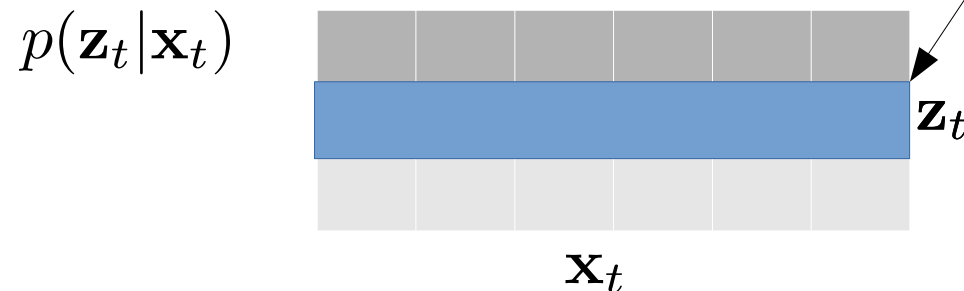


## Ingredients

- Predicted belief



- Observation model



we focus on the known measurement, the rest is irrelevant

- Known measurement  $\mathbf{z}_t$

# Filtering: Alternative Formulation

**Update:** from the predicted belief  $b_{t|t-1}$ , compute the joint distribution that predicts the observation.

- Joint over state and measurement (*chain rule*):

$$p(\mathbf{x}_t, \mathbf{z}_t | t) = p(\mathbf{z}_t | \mathbf{x}_t) p(\mathbf{x}_t, |t - 1)$$

- *Condition* on the actual measurement:

$$\underbrace{p(\mathbf{x}_t | t)}_{b_{t|t}} = \frac{p(\mathbf{x}_t, \mathbf{z}_t | t)}{p(\mathbf{z}_t | t)}$$

- Programmatically (discrete case)

```
float normalizer=0;
for (x_i : X) {
    b[x_i] = b_pred[x_i] * observationModel(z,x_i);
    normalizer += b[x_i];
}
b *= 1./normalizer;
```