Deforestation modelling using logistic regression and GIS

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ABSTRACT: A methodology has been used by means of which modellers and planners can quantify the certainty in predicting the location of deforestation. Geographic information system and logistic regression analyses were employed to predict the spatial distribution of deforestation and detects factors influencing forest degradation of Hyrcanian forests of western Gilan, Iran. The logistic regression model proposed that deforestation is a function of slope, distance to roads and residential areas. The coefficients for the explanatory variables indicated that the probability of deforestation is negatively related to slope, distance from roads and residential areas. Although the distance factor was found to be a contributor to deforestation, its effect is lower than that of slope. The correlates of deforestation may change over time, and so the spatial model should be periodically updated to reflect these changes. Like in any model, the quality may be improved by introducing the new variables that may contribute to explaining the spatial distribution of deforestation.

Keywords: manmade areas; physiographic factors; roads; probability; Hyrcanian forests

In the northern forests of Iran, deforestation and forest degradation are among the most important problems that have proved to be a prevailing factor for flooding, soil erosion, and in general for environment and humans (PIR BAVAGHAR et al. 2003). Therefore, detecting deforestation and identifying the factors influencing it are important, as this could be one stage in forest conservation, control of deforestation and is necessary in appropriate forest management planning (GRAINGER 1993; MAKINANO et al. 2010).

A spatial information system is a logical tool for monitoring and evaluating deforestation. The information of this system may offer a framework to develop a variety of powerful models, which could help managers to make decisions based on a methodologically robust basis (Felicisimo et al. 2002).

From a planning and management perspective, it is important to have a spatial view of where deforestation occurred, and its underlying drivers. One of the most important methods to detect the factors influencing deforestation and their spatial interaction is to model their influence on the landscape

using spatial data (Serneels, Lambin 2001; Laurance et al. 2002; Nagendra et al. 2003; Mertens et al. 2004; Etter et al. 2006).

Knowledge of the rate and extent of deforestation and its driving factors is necessary for environmental planners and managers (LUDEKE et al. 1990). To understand the deforestation process, determination and knowledge of the relationship between natural and manmade variables and deforestation are an essential step (LINKIE et al. 2004).

The changes in land use and land cover together with the influence of natural and anthropogenic factors have been intensively investigated (Felicisimo et al. 2002; Linkie et al. 2004; Ostapowicz 2005; Amini et al. 2009; Bagheri, Shataee 2010). However, there is also a need to conduct more studies, because the factors influencing deforestation are often site-specific (Geist, Lambin 2002; Linkie et al. 2004). For example, the factors influencing deforestation are different on various continents (Bawa, Dayanandan 1997) and, even when they are the same, they need not be equally important. Therefore, the study of deforestation on

a site-by-site basis is necessary which could only be possible by using the inexpensive geographic information system (GIS) software (LINKIE et al. 2004).

Several studies have attempted to understand a deforestation rate in the Hyrcanian forests (Rafieyan et al. 2003; Pir Bavaghar 2004; Salman Mahini et al. 2009; Bagheri, Shataee 2010). But there are just a few studies, trying to model these deforestations according to the factors influencing them.

All the forests of the Iranian territory became nationalized in 1962; therefore forests in Iran are basically state-owned. The population growth has increased needs for food and crop lands. Consequently, deforestation has occurred in these forests. Development of agricultural areas, i.e. converting forested areas to tea cultivation and rice fields, livestock grazing, urbanization, rural development, and expansion of the industrial areas, are the factors influencing to the largest extent deforestation in the northern forests of Iran. Therefore, the accessibility variables seem to be more important than other factors in the study area.

The objectives of this paper are to detect and analyze deforestation in watershed basin No. 28 in the Caspian forests. This research reveals if deforestations depend on the physiographic and socio-economic factors.

The above process is carried out under the hypothesis that the present deforestation is related to physiographic (elevation, slope, aspect) and surrogate socio-economic (distance to roads and residential areas) factors.

MATERIAL AND METHODS

Study area. Watershed basin No. 28, part of the Hyrcanian forests of Iran, is situated in eastern Gilan Province in the south-west of the Caspian Sea (Fig. 1).

The study region covers approximately 24,000 ha. Elevation ranges from 0 to 2,900 m a.s.l. and slope varies from 0 to 250%. The study area is flat in the north and rugged mountains cover southern parts. This region has a temperate climate and precipitation (the mean annual precipitation is about 1,400 mm) is distributed throughout the year (SAGHEB TALEBI et al. 2003).

Hyrcanian forests are mixed and uneven-aged deciduous forests. Spot cutting in limited areas has been applied in these forests.

Research data. Data used in this study were:

- Digital 1:25,000 thematic-topographic maps produced in 1982, which were extracted based on 1:20,000 aerial photos acquired in 1967 according to the Iranian Forests, Rangelands and Watershed Organization (FRWO) order and National Cartographic Centre (NCC) supervision. These maps have 63 different data layers including: residential areas, roads, railways, forests, ranges, gardens, contour lines, etc.
- Digital thematic-topographic maps dated 2001, which were generated based on 1:20,000 aerial photos dated 1994. These maps have 63 different data layers that are the same as in previous maps.

Methodology. Deforestation mapping. The layer of forest classes was extracted from both 1967 and 1994 digital maps and then the values of 0 and 1 were labelled to non-forest and forest areas, respectively, in ArcGIS 9.3 software. This process was done for all the map sheets covering the watershed. By comparing forest maps related to the start and end of the period (1967 and 1994), deforestation maps were obtained. The maps were exported to Idrisi Selva software, in 30-m raster-grid format. The flow chart of the methodology and processing steps carried out in this study is shown in Fig. 2.

Explanatory variables. After extracting contour lines from the 3D digital maps of 1:25,000 scale, a digital elevation model having a spatial resolution

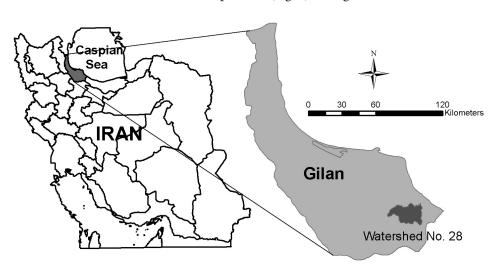
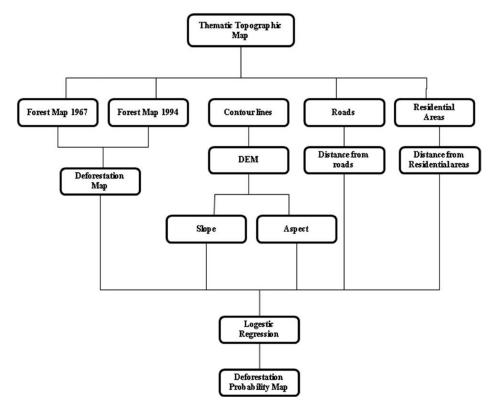


Fig. 1. The location of the study area in Iran (left) and in Gilan (right)

Fig. 2. Flowchart of the methods and materials



of 30 m was created. Slope and aspect layers were calculated using a digital elevation model (DEM). According to BEERS et al. (1966) a cosine function (Eq. 1) was applied to transform the aspect into a number ranging from 0 (southwest-facing) to 2 (northeast-facing) to create a more direct measure of radiation load for statistical analysis. Distances to the nearest road and settlement were calculated as a series of buffers of 200 m expanding from each road segment and settlement centre, respectively. Road network and human settlements were derived from the digital map dated 1967. These maps were converted from vector to raster format with 30-m grid cells.

$$Cos (45 - Aspect) + 1 \tag{1}$$

Datasets for modelling and validation. Production and validation of the logistic regression model were performed by a sampling of the geographical and environmental space of the study region. To avoid biases, the samples were balanced to have the same number of positive (deforested) and negative (non-deforested) cases (Felicisimo et al. 2002). 100 points were selected in areas presenting deforestation between 1967 and 1994, and 100 points in areas that remained forested over the same period. These points should be separated by at least 1,000 m to reduce the effects of spatial autocorrelation. In this study a sepa-

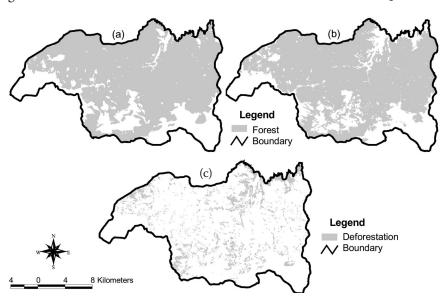


Fig. 3. Forest maps of watershed basin No. 28 for 1967 (a), 1994 (b), and the deforestation map in this period (c)

ration distance of 500 m (at least 500 m) was used instead, because of the relatively limited area in extent (Linkie et al. 2004). Each observation is a 30-m grid cell containing either deforestation or not. 80% of observations were allocated to establish the model, and 20% of them were allocated to the validation model. The model has been calibrated using a separate dataset, to reduce the likelihood of overestimating the predictive ability of the model (Chatfield 1995; Wilson et al. 2005).

Modelling the spatial distribution of deforestation using logistic regression. To investigate the correlations that exist between a dichotomous dependent variable (deforested/non-deforested) and independent variables which cannot be assumed to satisfy the required assumptions of discriminant analysis (normality assumption), a logistic regression model has been used (Ludeke et al. 1990). In logistic regression, a dependent variable transforms into a logit variable (the natural log of the odds of the dependent variable occurring or not), and then based on the independent variables, maximum likelihood estimation is applied to estimate the probability of occurrence of a certain event (deforestation) (Rueda 2010).

The dependent variable used for calibrating the model was derived from the analysis of two datasets of forest maps. This information was extracted from 30-m grid cells of forest layers, using Idrisi Selva software. The extraction information was exported into SPSS 16 software for further analysis. A multivariate, spatially explicit model of the deforestation was developed using the logistic regression (Schneider, Pontius 2001; Sernels, Lambin 2001; Wilson et al. 2005). The model was used to determine the variables that explain the spatial distribution of deforestation.

A logistic model was developed based on the binary response variable (one – deforestation; zero – non-deforestation) (RIVERA et al. 2012) and the explanatory variables (elevation, aspect, slope, distance to roads and residential areas). Before developing the model, the explanatory variables were standardized by dividing values by their root-mean-square because of the easier comparison of the relative effect of each variable (ETTER et al. 2006). The logistic function gives the probability of forest loss as a function of the explanatory variables.

The logistic function (Eq. 2) results bounded between 0 and 1 as follows:

$$p = E(Y) = \frac{\exp(\beta_0 + \beta_1 X_1 + \beta_2 X_2 + \dots + \beta_i X_i)}{1 + \exp(\beta_0 + \beta_1 X_1 + \beta_2 X_2 + \dots + \beta_i X_i)}$$
(2)

where:

p − probability of deforestation in the cell,

E(Y) – expected value of the binary dependent variable Y,



- constant to be estimated,

- predicted coefficient of each independent variable X_i (SCHNEIDER, PONTIUS 2001).

The amount of the contribution of each factor to deforestation is described by the regression coefficients. We could transform the logistic function into a linear response with the following transformation:

$$p' = \log_e(p/1 - p) \tag{3}$$

hence

$$p' = (\beta_0 + \beta_1 X_1 + \beta_2 X_2 + \beta_3 X_3) \tag{4}$$

This transformation which allows linear regression to estimate each β_i is called a logit or logistic transformation (transformation from Eq. 3 to Eq. 4). The final result is a probability score (p) for each cell (Schneider, Pontius 2001). Notice that the logit transformation of dichotomous data ensures that the dependent variable of the regression is continuous, and the new dependent variable (logit transformation of the probability) is unbounded. Furthermore, it ensures that the predicted probability will be continuous within the range from 0 to 1. The final step is the classification of these results.

Validation of the logistic regression model. To indicate the effectiveness and soundness of the model, statistical tests of individual predictors including goodness-of-fit statistic, and validations of predicted probabilities have been accomplished (PENG et al. 2002). At first, R square of the model was calculated. In this method, R^2 is called pseudo R^2 because it is not computed in the same way as the regular regression R^2 (LUDEKE et al. 1990). This R square indicates the fitness of the model, but does not give as much information as the regular regression R^2 about the scatter of the data around the fitted line (Ludeke et al. 1990). The value of R^2 is low in logistic regression models because of the binary response variable (B10 et al. 1998). For a very good fit of a logistic regression model, R^2 should have values between 0.2 and 0.4 (WILSON et al. 2005). The statistical significance of individual regression coefficients (β_i) was tested using the Wald chi-square statistic. For testing goodness-of-fit, the Hosmer-Lemeshow (H-L) test was used in this study. According to the Hosmer-Lemeshow test, the observations are grouped into deciles of risk according to a comparison of the observed probability with the expected probability within each decile. The area under the Receiver Operator Characteristic Curve (ROC) is usually used as discrimination ability. ROC was calculated by comparing the predictions of deforestation with the actual

ones (WILSON et al. 2005; ETTER et al. 2006). These values range from 0.5 to 1.0. The value above 0.7 indicates an accurate model fit, above 0.9 indicates a highly accurate model (LINKIE et al. 2004) and the value of 0.5 indicates a random model.

RESULTS

The forest maps of watershed basin No. 28 for 1967 and 1994 are depicted in Fig. 3. Approximately 12% (2,902 ha) of the total area was deforested during the 27-year period. Therefore, the mean annual deforestation rate was 0.44%.

The coefficients and the value of the area under the ROC curve are listed in Table 1. During this period the probability of deforestation was significantly and negatively determined by slope (Wald = 7.057, df = 1), distance to roads (Wald = 4.295, df = 1), and distance to residential areas (Wald = 13.651, df = 1). The logistic regression model proposed that deforestation is a function of slope, distance to roads and residential areas (Table 1). The coefficients for the explanatory variables indicated that the probability of deforestation is negatively related to slope, distance from roads and residential areas (Table 1). Although the distance factor was found to be a contributor to deforestation, their effects are lower than those of slope.

Although the logistic regression goodness of fit measured by the Nagelkerke R^2 statistic is low, the significant Chi-square value (54.12, df = 3, P < 0.001) and high correct classification percentage (72.5%) indicate the perfect fit of the model in explaining the relationship between independent and dependent variables (Table 1).

The area under the ROC curve of the model was 0.807, so this model has a good discrimination ability (LINKIE et al. 2004).

The best-fit model of deforestation (Table 1) was used to predict the probability of deforestation of the remaining areas of the watershed. The probability of

Table 1. Result of the logit analysis

Variable	В	S.E	Wald	Sig	Exp (B)
St. Build	- 0.051	0.014	13.651	0.000	0.950
St. Road	-0.012	0.006	4.295	0.038	0.998
St. Slope	-0.183	0.069	7.057	0.008	0.833
Constant	2.683	0.565	22.567	0.000	14.629

Chi-square value = 54.125; Nagelkerke R square = 0.383; ROC = 0.807, SE = 0.031, Sig = 0.000; Hosmer & Lemeshow test Chi-square = 7.433, Sig = 0.491; Correct classification = 72.5%, RMSE% = 23%, St. Slope, St. Build & St. Road: Standardized value of slope, distance from residential areas and roads

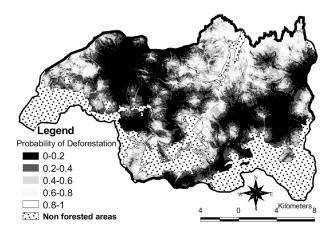


Fig. 4. Probability of deforestation (lighter areas have a higher probability of deforestation and are located in a relatively flat terrain, these areas also have shorter distances to roads and residential areas)

deforestation ranged from zero to 0.94 (Fig. 4). At the final step a classified map was generated based on the predicted probability.

DISCUSSION

The deforestation rate of the study area was 0.44% per year (2,902 ha). This study has shown the value of producing site-specific models based on logistic regression that can be used in forest management.

The deforestation models developed in this study can be used to predict future patterns of deforestation and to identify where to focus stronger protection for the best results. This analysis found that the spatial pattern of forest loss was dependent on several physiographic and anthropogenic factors and that the logistic regression models could be used to accurately predict future deforestation trends (Linkie et al. 2004). One of these factors was slope, which was important during this period because the areas with steep slopes tended to include more rugged terrain and are further from existing deforestation fronts. This may also partly explain why low slope forests are the most threatened forest type (LINKIE et al. 2004). We conclude that the areas with lower slopes are more accessible, and more suitable for agricultural activities that are the most important factors causing deforestation.

As expected from previous studies (PIR BA-VAGHAR 2004; WILSON et al. 2005; AMINI et al. 2009; BAGHERI, SHATAEE 2010), the position of roads and residential areas were important in determining deforestation patterns.

Deforestation has been seen to have a negative relationship with slope, distance from roads and residential areas. Negative coefficients of these factors indicate that higher values are associated with lower probabilities of deforestation (WILSON et al. 2005; Amini et al. 2009; Bagheri, Shataee 2010). In general, a steep slope limits deforestation due to difficulties associated with transportation. The spatial patterns of deforestation across this area highlighted the critical role of accessibility, with the importance of distance to roads and residential areas. These results reflect the findings of other deforestation assessments (Ludek et al. 1990; Linkie et al. 2004, 2010; Amini et al. 2009; Bagheri, Shataee 2010). It was found that 83% of deforestation occurred within a 2-km distance from roads. Similarly, LUDEKE et al. (1990) also found that deforestation decreased rapidly with a distance from roads and there was a steep drop in the percentage area deforested beyond 2 km from access routes. WILSON et al. (2005) also mentioned that 90% of the deforested area is within 2.5 kilometres from roads.

The validation analysis showed that the explanatory variables included in the model had a sufficient explanatory power to discriminate between deforested and non-deforested areas.

The correlates of deforestation may change over time and so the spatial model should be periodically updated to reflect these changes. Like in any model, the quality may be improved by introducing the new variables that may contribute to explaining the spatial distribution of deforestation.

The results of this analysis are based on the assumption that the existing forest maps are accurate. Furthermore, the accuracy of this assessment relies on the quality and accuracy of the maps of the explanatory variables included in the model. These maps are the most detailed and comprehensive presently available for these forests.

CONCLUSIONS

This study fulfilled its aim by predicting spatial patterns of deforestation in the northern forests of Iran and understanding the underlying drivers. Deforestation is indeed as interplay between several factors. Accessibility was found to be an important variable for explaining the patterns of deforestation observed in the study area. The results indicated that slope, major roads, and residential areas have a strongly significant correlation with deforestation. So the results highlighted the critical role of accessibility. The results did not indicate a significant relationship with aspect and elevation. The results also showed the utility of a statistical modelling ap-

proach to analyse and predict deforestation. The logistic regression goodness of fit is low, suggesting that missing variables such as livestock's role, might further explain differences between low and high deforestation. In spite of this, the modelling approach developed by this study would benefit conservation planning (WILSON et al. 2005; SMITH et al. 2008; LINKIE et al. 2010).

References

Amini M.R., Shataee Sh., Moaieri M.H., Ghazanfari H. (2009):
Deforestation modeling and investigation on related physiographic and human factors using satellite images and GIS (Case study: Armardeh forests of Baneh). Iranian Journal of Forest and Poplar Research, 17: 431–443. (in Persian)

Bagheri R., Shataee Sh. (2010): Modeling forest area decreases using logistic regression (Case study: Chehl-Chay catchment, Golestan province). Iranian Journal of Forest, 2: 243–252. (in Persian)

Beers T.W., Press P.E., Wensel L.C. (1996): Aspect transformation in site productivity research. Journal of Forestry, 64: 691–692. Bio A.M.F., Alkemende R., Barendregt A. (1998): Determining alternative models for vegetation response analysis: a non-parametric approach. Journal of Vegetation Science, 9: 5–16. Chatfield C. (1995): Model uncertainty, data mining and statistical inference. Journal of the Royal Statistical Society, 158: 419–466.

Etter A., McAlpine C., Wilson K., Phinn S., Possingham H. (2006): Regional patterns of agricultural land use and deforestation in Colombia. Agriculture, Ecosystems & Environment. 114: 369–386.

Felicíslmo A.M., Francés E., Fernández J.M., González-Díez A., Varas J. (2002): Modeling the potential distribution of forests with a GIS. Photogrammetric Engineering and Remote Sensing, 68: 455–461.

Geist H.J., Lambin E.F. (2002): Proximate causes and underlying driving forces of tropical deforestation. Bioscience, 52: 143–150.

Laurance W.F., Albernaz A.K., Schroth G., Fearnside P.M., Bergen S., Venticinque E.M., Da Costa C. (2002): Predictors of deforestation in the Brazilian Amazon. Journal of Biogeography, 29: 737–748.

Linkie M., Smith R.J., Leader-Williams N. (2004): Mapping and predicting deforestation patterns in the lowlands of Sumatra. Biodiversity Conservation, 13: 1809–1818.

Linkie M., Rood E., Smith R.J. (2010): Modelling the effectiveness of enforcement strategies for avoiding tropical deforestation in Kerinci SeblatNational Park, Sumatra. Biodiversity Conservation, 19: 973–984.

Ludeke A.K. (1990): An analysis of antropogenic deforestation using logistic regression and GIS. Journal of Environmental Management, 31: 247–259.

- Makinano M.M., Santillan J.R., Paringit E.C. (2010): Detection and analysis of deforestation in cloud-contaminated Landsat images: A case of two Philippine provinces with history of forest resource utilization. In: Proceeding of the 31st Asian Conference on Remote Sensing (ACRS 2010): Remote Sensing for Global Change and Sustainable Development, Hanoi, Nov 1–5, 2010: 44–52.
- Mertens B., Lambin E. (1999): Modelling land cover dynamics: integration of fine-scale land cover data with landscape attributes. International Journal of Applied Earth Observation and Geoinformation, 1: 48–52.
- Nagendra H., Southworth J., Tucker C.J. (2003): Accessibility as a determinant of landscape transformation in western Hondrus: linking pattern and process. Landscape Ecology, 18: 141–158.
- Ostapowicz K. (2005): Model of forests spatial distribution in the western part of the Karpaty Mts. In: 8th AGILE Conference on Geographic Information Science, Estoril, May 26–28, 2005: 611–617.
- Peng C.J., Lee K.L., Ingersoll G.M. (2002): An introduction to Logistic regression analysis and reporting. The Journal of Educational Research, 96: 3–14.
- Pir Bavaghar M., Darvishsefat A.A., Namiranian M. (2003): The study of spatial distribution of forest changes in the northern forests of Iran. In: Proceedings of the Map Asia Conference, Kuala Lumpur, Oct 14–15, 2003: 1–6.
- Pir Bavaghar M. (2004): Forest Area Change Detection Related To Topographic Factors and Residential Areas (Case Study: Eastern Forests of Gilan Province). [MSc. Thesis.] Karaj, University of Tehran: 110.
- Rafieyan O., Darvishsefat A.A., Namiranian M. (2003): Forest area change detection using ETM+ data in northern forest of Iran. In: The First International Conference on Environmental Research and Assessment, Bucharest, Mar 23–27, 2003: 1–4.

- Rivera S., Martinez de Anguita P., Ramsey R.D., Crowl T.A. (2012): Spatial modeling of tropical deforestation using socioeconomic and biophysical data. Small-scale Forestry, 12: 321–334.
- Rueda X. (2010): Understanding deforestation in the southern Yucatan: insights from a sub-regional, multi-temporal analysis. Regional Environmental Change, 10:175–189.
- Sagheb Talebi Kh., Sajedi T., Yazdian F. (2003): Forests of Iran, Research Institute of Forests and Rangelands, Forest Research Devision, Iran: 28.
- Salman Mahini A., Feghhi J., Nadali A., Riazi B. (2009): Tree cover change detection through artificial neural network classification using Landsat TM and ETM+ images (case study: Golestan Province, Iran). Iranian Journal of Forest and Poplar Research, 16: 495–505.
- Schneider L.C., Pontius R.G. (2001): Modelling land-use changes in the Ipswich watershed, Massachusetts, USA. Agriculture, Ecosystems & Environment, 85: 83–94.
- Serneels S., Lambin E.F. (2001): Proximate causes of land use change in Narok District, Kenya: a spatial statistical model. Agriculture, Ecosystems & Environment, 85: 65–81.
- Smith R.J., Easton J., Nhancale B.A., Armstrong A.J., Culverwell J., Dlamini S., Goodman P.S., Loffler L., Matthews W.S., Monadjem A., Mulqueeny C.M., Ngwenya P., Ntumi C.P., Soto B., Leader-Williams N. (2008): Designing a transfrontier conservation landscape for the Maputaland centre of endemism using biodiversity, economic and threat data. Biological Conservation, 141: 2127–2138.
- Wilson K., Newton A., Echeverria C., Weston Ch., Burgman M. (2005): A vulnerability analysis of the temperate forests of south central Chile. Biological Conservation, 122: 9–21.

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