## 9.7 Solutions to Chapter 7

## Exercise 7.1

Let us consider the following seasonal MA model:

$$X_t = Z_t + \Theta Z_{t-s}$$

The autocovariance function in lag h is

$$\gamma_X(h) = cov(X_{t+h}, X_t)$$

$$= cov(Z_{t+h} + \Theta Z_{t+h-s}, Z_t + \Theta Z_{t-s})$$

$$= \gamma_Z(h) + \Theta \gamma_Z(h-s) + \Theta \gamma_Z(h+s) + \Theta^2 \gamma_Z(h)$$
(1)

Note that

$$\gamma_Z(h) = \begin{cases} \sigma^2 & h = 0\\ 0 & h \neq 0 \end{cases}$$

Therefore,

$$\gamma_X(h) = \begin{cases}
\gamma_Z(0) + \Theta^2 \gamma_Z(0) & h = 0 \\
\Theta \gamma_Z(0) & h = \pm s \\
0 & h \neq 0, \pm s
\end{cases}$$

$$= \begin{cases}
\sigma^2 (1 + \Theta^2) & h = 0 \\
\Theta \sigma^2 & h = \pm s \\
0 & h \neq 0, \pm s
\end{cases}$$

The auto-correlation function is

$$\rho_X(h) = \frac{\gamma_X(h)}{\gamma_X(0)} = \begin{cases} 1 & h = 0\\ \frac{\Theta}{1 + \Theta^2} & h = \pm s\\ 0 & h \neq 0, \pm s \end{cases}$$

## Exercise 7.2

Let us consider the following seasonal AR model:

$$X_t = \Phi X_{t-s} + Z_t$$

For h = 0, we have

$$\begin{split} \gamma_X(0) &= var(X_t) \\ &= cov(X_t, X_t) \\ &= cov(\Phi X_{t-s} + Z_t, \Phi X_{t-s} + Z_t) \\ &= \Phi^2 cov(X_{t-s}, X_{t-s}) + 2\Phi cov(X_{t-s}, Z_t) + cov(Z_t, Z_t) \\ &= \Phi^2 \gamma_X(0) + \sigma^2 \qquad \text{(since } Z_t, X_{t'} \text{ are uncorrelated for } t > t') \end{split}$$

Therefore

$$\gamma_X(0) = \frac{\sigma^2}{1 - \Phi^2}$$

For h > 0,

$$\gamma_X(h) = cov(X_{t+h}, X_t) 
= cov(\Phi X_{t+h-s} + Z_{t+h}, X_t) 
= \Phi cov(X_{t+h-s}, X_t) + cov(Z_{t+h}, X_t) 
= \Phi \gamma_X(h-s) mtext{(since } Z_{t'} \text{ and } X_t \text{ are uncorrelated for } t' > t), mtext{(1)}$$

Note that, since  $\gamma_X$  is an even function, for  $s \neq 1$ , we have

$$\begin{cases} \gamma(1) = \Phi \, \gamma_X(1-s) = \Phi \, \gamma(s-1) \\ \gamma(s-1) = \Phi \, \gamma(s-1-s) = \Phi \, \gamma(-1) = \Phi \, \gamma(1) \end{cases}$$

Therefore, it is easy to see that

$$\gamma(1) = \Phi \gamma(s-1) = \Phi^2 \gamma(1)$$

and therefore,

$$(1 - \Phi^2)\gamma(1) = 0$$

We know that  $|\Phi| < 1$  and consequently  $1 - \Phi^2 \neq 0$ . So we can conclude that

$$\gamma(1) = 0$$

Using the same method, we can show that  $\gamma(h) = 0$  for  $h \neq ks$ ,  $k = 0, 1, 2, \dots$ From (1), we can show that, for h = ks:

$$\gamma_X(h) = \Phi \gamma_X(h - s)$$

$$= \Phi^2 \gamma_X(h - 2s)$$

$$\vdots$$

$$= \Phi^k \gamma_X(h - ks)$$

$$= \Phi^k \gamma_X(0)$$

$$= \Phi^k \frac{\sigma^2}{1 - \Phi^2}$$

Therefore,

$$\gamma_X(h) = \begin{cases} \frac{\sigma^2}{1 - \Phi^2} & h = 0\\ \frac{\Phi^k \sigma^2}{1 - \Phi^2} & h = ks, \quad k = 1, 2, \dots\\ 0 & h \neq ks, \quad k = 1, 2, \dots \end{cases}$$

## Exercise 7.3

$$\gamma_{X}(h) = cov(X_{t+h}, X_{t}) 
= cov(\Phi X_{t+h-12} + Z_{t+h} + \theta Z_{t+h-1}, X_{t}) 
= \Phi \underbrace{cov(X_{t+h-12}, X_{t})}_{\gamma_{X}(h-12)} + cov(Z_{t+h}, X_{t}) + \theta cov(Z_{t+h-1}, X_{t}) 
\qquad \gamma_{X}(h-12) 
= \begin{cases}
\Phi \gamma_{X}(12) + cov(Z_{t}, X_{t}) + \theta cov(Z_{t-1}, X_{t}) & h = 0 
\Phi \gamma_{X}(11) + cov(Z_{t+1}, X_{t}) + \theta cov(Z_{t}, X_{t}) & h = 1 
\Phi \gamma_{X}(0) + cov(Z_{t+12}, X_{t}) + \theta cov(Z_{t+11}, X_{t}) & h = 12 
\Phi \gamma_{X}(h-12) & \text{otherwise}
\end{cases}$$
(A)

Note that

$$\bullet \ cov(X_t, Z_{t'}) = 0 \quad \text{if } t' > t \tag{1}$$

•

$$cov(Z_t, X_t) = cov(Z_t, \Phi X_{t-12} + Z_t + \theta Z_{t-1})$$
  
= 0 + cov(Z\_t, Z\_t) + 0 (using (1)) = var(Z\_t) = \sigma^2 (2)

•

$$cov(Z_{t-1}, X_t) = cov(Z_{t-1}, \Phi X_{t-12} + Z_t + \theta Z_{t-1})$$

$$= \Phi cov(Z_{t-1}, X_{t-12}) + 0 + \theta cov(Z_{t-1}, Z_{t-1})$$

$$= \theta \sigma^2 \quad \text{(since } Z_t \text{ are WN, using (1) and (2))}$$
(3)

Consequently, we can rewrite (A) as follows:

$$\gamma_X(h) = \begin{cases} \Phi \, \gamma_X(12) + cov(Z_t, X_t) + \theta \, cov(Z_{t-1}, X_t) & h = 0 \\ \Phi \, \gamma_X(11) + cov(Z_{t+1}, X_t) + \theta \, cov(Z_t, X_t) & h = 1 \\ \Phi \, \gamma_X(0) + cov(Z_{t+12}, X_t) + \theta \, cov(Z_{t+11}, X_t) & h = 12 \\ \Phi \, \gamma_X(h - 12) & \text{otherwise} \end{cases}$$

$$= \begin{cases} \Phi \, \gamma_X(12) + \sigma^2 + \theta \cdot \theta \sigma^2 = \Phi \gamma_X(12) + (1 + \theta^2) \sigma^2 & h = 0 \\ \Phi \, \gamma_X(11) + 0 + \theta \sigma^2 & h = 1 \\ \Phi \, \gamma_X(0) + 0 + 0 & h = 12 \\ \Phi \, \gamma_X(h - 12) & \text{otherwise} \end{cases}$$

Since

$$\gamma_X(0) = \Phi \gamma_X(12) + \sigma^2 + \theta^2 \sigma^2$$

and, we have

$$\gamma_X(12) = \Phi \, \gamma_X(0)$$

Therefore, we can conclude that:

$$\gamma_X(0) = \Phi^2 \gamma_X(0) + \sigma^2 (1 + \theta^2) \quad \Rightarrow \quad \gamma_X(0) = \frac{\sigma^2 (1 + \theta^2)}{1 - \Phi^2}$$

Since  $\gamma_X(h) = \Phi \gamma_X(h-12)$ , we can easily show that

$$\gamma_X(12k) = \Phi^k \gamma_X(0)$$

Similarly, since

$$\gamma_X(1) = \Phi \gamma_X(11) + \theta \sigma^2$$
 and  $\gamma_X(h) = \Phi \gamma_X(h-12)$  for  $h \neq 0, 1, 12$ 

we can conclude that

$$\gamma_X(1) = \frac{\theta}{1 - \Phi^2} \sigma^2$$

and

$$\gamma_X(12k \pm 1) = \Phi^k \gamma_X(1) = \Phi^k \cdot \frac{\theta}{1 - \Phi^2} \sigma^2 = \frac{\Phi^k \theta}{1 + \theta^2} \gamma_X(0)$$

Therefore, for the auto-correlation function we have:

$$\rho_X(h) = \frac{\gamma_X(h)}{\gamma_X(0)} = \begin{cases} 1 & h = 0\\ \Phi^k & h = 12k\\ \frac{\Phi^k \theta}{1 + \theta^2} & h = 12k \pm 1\\ 0 & \text{otherwise} \end{cases}$$