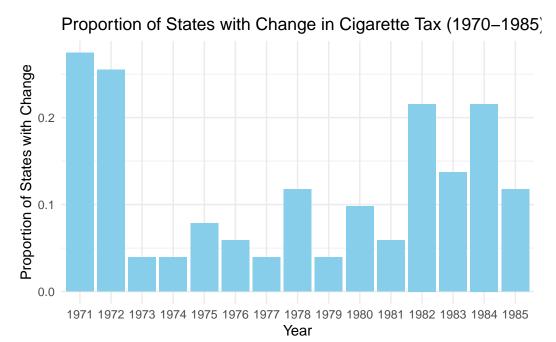
Homework 3

Varun Saxena

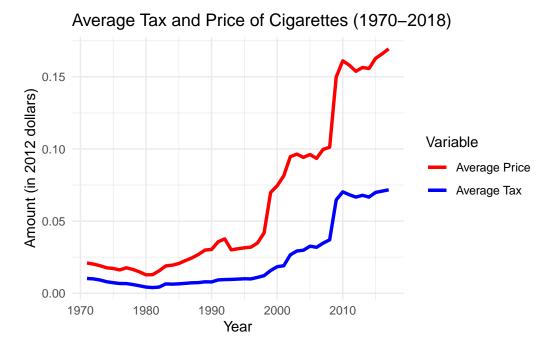
Loading required package: pacman

 $Repository\ https://github.com/varunsaxena2/saxena-v-hwk3-2/upload$

1. Here is the bar graph showing the proportion of states with a change in their cigarette tax in each year from 1970 to 1985.

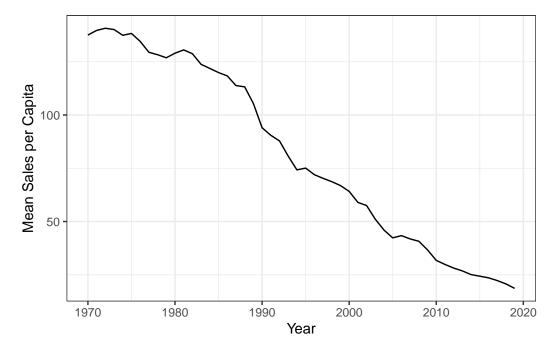


2. Here is the line graph showing the average tax (in 2012 dollars) on cigarettes and the average price of a pack of cigarettes from 1970 to 2018.

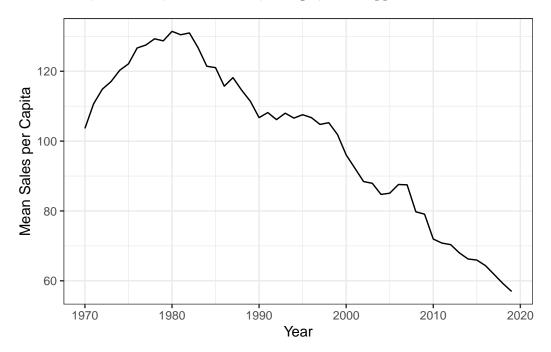


3.

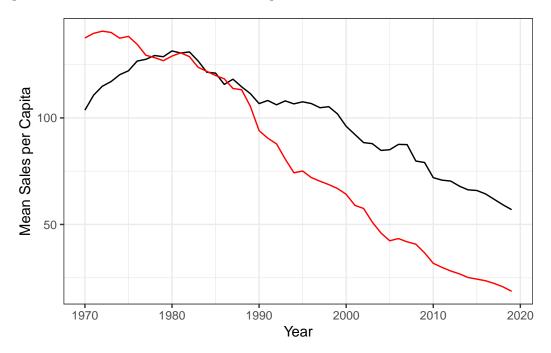
The 5 states with the highest increases in cigarette prices (in dollars) over the time period are New York, District of Columbia, Connecticut, Rhode Island, Massachusetts.



4. The 5 states with the lowest increases in cigarette prices (in dollars) over the time period are Missouri, Tennessee, North Dakota, Georgia, Mississippi.



5. This plot displays the previous two graphs overlayed ontop of one another. The five highest price states is in red and the five lowest price states is in black.



From this, we can gather that cigarettes follow normal economic trends, with higher taxes correlating with lower sales due to a price increase. Initially this trend is flipped, but over time, the gap in sales per capita grows, indicating consumer sensitivity to price changes.

6. This model regresses ln(sales) on ln(price).

Call:

lm(formula = ln_sales ~ ln_price_2012, data = restricted_cig.data)

Residuals:

Min 1Q Median 3Q Max -0.68212 -0.08585 -0.00291 0.08552 0.83479

Coefficients:

Estimate Std. Error t value Pr(>|t|)
(Intercept) 3.19727 0.07701 41.52 <2e-16 ***
ln_price_2012 -0.40701 0.01933 -21.05 <2e-16 ***
--Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1

Residual standard error: 0.188 on 1018 degrees of freedom Multiple R-squared: 0.3033, Adjusted R-squared: 0.3026 F-statistic: 443.1 on 1 and 1018 DF, p-value: < 2.2e-16

The negative coefficient indicates that increases in price correlate with decreases in sales, indicating elastic demand.

7. This instrumental variable regression regresses $\ln(\text{sales})$ on $\ln(\text{price})$ with $\ln(\text{tax})$ serving as the instrumental variable.

The results of the study indicate that when tax is used as an instrumental variable to reduce endogeneity effects, the result is the same as the previous model, with a negative association between sales and price.

8. This model regresses ln(sales) on ln(tax)

Call:

lm(formula = ln_sales ~ ln_tax_2012, data = restricted_cig.data)

Residuals:

Min 1Q Median 3Q Max -0.78488 -0.08940 0.00176 0.09108 0.94920

Coefficients:

```
Estimate Std. Error t value Pr(>|t|)
(Intercept) 3.64736    0.07946    45.90    <2e-16 ***
ln_tax_2012 -0.23091    0.01568 -14.72    <2e-16 ***
---
Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
```

Residual standard error: 0.2045 on 1018 degrees of freedom Multiple R-squared: 0.1756, Adjusted R-squared: 0.1748 F-statistic: 216.8 on 1 and 1018 DF, p-value: < 2.2e-16

The negative estimator coefficient indicates that an increase in tax is associated with a decrease in sales, indicating the presence of price elasticity of demand.

9. These three models regress $\ln(\text{sales})$ on $\ln(\text{price})$, $\ln(\text{price})$ with $\ln(\text{tax})$ as an IV, and $\ln(\text{tax})$, respectively.

```
Call:
lm(formula = ln_sales ~ ln_price_2012, data = recent_cig.data)
Residuals:
     Min
               1Q
                    Median
                                 3Q
                                         Max
-0.92230 -0.17004 0.00664 0.17869 1.10282
Coefficients:
              Estimate Std. Error t value Pr(>|t|)
(Intercept)
               2.92463
                          0.03338
                                    87.63
                                            <2e-16 ***
ln_price_2012 -0.49841
                          0.01235 -40.37
                                            <2e-16 ***
Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
Residual standard error: 0.296 on 1273 degrees of freedom
Multiple R-squared: 0.5614,
                               Adjusted R-squared: 0.5611
F-statistic: 1630 on 1 and 1273 DF, p-value: < 2.2e-16
TSLS estimation, Dep. Var.: ln_sales, Endo.: ln_price_2012, Instr.: ln_tax_2012
Second stage: Dep. Var.: ln_sales
Observations: 1,275
Standard-errors: IID
                   Estimate Std. Error t value Pr(>|t|)
(Intercept)
                              0.034905 80.2675 < 2.2e-16 ***
                   2.801696
fit_ln_price_2012 -0.545355
                              0.012944 -42.1329 < 2.2e-16 ***
Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
RMSE: 0.297451
                Adj. R2: 0.556115
F-test (1st stage), ln_price_2012: stat = 14,664.4, p < 2.2e-16, on 1 and 1,273 DoF.
                       Wu-Hausman: stat = 191.5, p < 2.2e-16, on 1 and 1,272 DoF.
Call:
lm(formula = ln_sales ~ ln_tax_2012, data = recent_cig.data)
Residuals:
```

Max

3Q

Min

1Q

Median

-0.88265 -0.14516 0.00644 0.14463 1.15191

Coefficients:

Estimate Std. Error t value Pr(>|t|)
(Intercept) 2.744539 0.033594 81.70 <2e-16 ***
ln_tax_2012 -0.388772 0.008558 -45.43 <2e-16 ***
--Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1

Residual standard error: 0.2761 on 1273 degrees of freedom Multiple R-squared: 0.6185, Adjusted R-squared: 0.6182 F-statistic: 2064 on 1 and 1273 DF, p-value: < 2.2e-16 10. The result of these regressions are that the price elasticity of demand is greater in magnitude across all three regressions as indicated by the estimator coefficient. This would indicate that in more recent years, consumers are more sensitive to price and tax changes. A potential reason for this increased sensitivity to price changes is that there are more alternatives to cigarettes, with e-cigarettes and other such products. The addictive nature of cigarettes would render it an inelastic good, but the presence of alternative goods makes this good more elastic to price changes.