CMPT 210: Probability and Computing

Lecture 19

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Recap

Variance: Standard way to measure the deviation from the mean. For r.v. X,

$$Var[X] = \mathbb{E}[(X - \mathbb{E}[X])^2] = \sum_{x \in Range(X)} (x - \mu)^2 \Pr[X = x], \text{ where } \mu := \mathbb{E}[X].$$

Alternate Definition: $Var[X] = \mathbb{E}[X^2] - (\mathbb{E}[X])^2$.

If
$$X \sim \text{Ber}(p)$$
, $\text{Var}[X] = p(1-p)$.

If
$$X \sim \text{Uniform}(\{v_1, v_2, \dots v_n\})$$
, $\text{Var}[X] = \frac{[v_1^2 + v_2^2 + \dots v_n^2]}{n} - \left(\frac{[v_1 + v_2 + \dots v_n]}{n}\right)^2$.

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Recall that for a coin s.t. Pr[heads] = p, R is the r.v. equal to the number of coin tosses we need to get the first heads. Let A be the event that we get a heads in the first toss. Using the law of total expectation,

$$\mathbb{E}[R^2] = \mathbb{E}[R^2|A] \Pr[A] + \mathbb{E}[R^2|A^c] \Pr[A^c]$$

 $\mathbb{E}[R^2|A] = 1$ ($R^2 = 1$ if we get a heads in the first coin toss) and $\Pr[A] = p$. Hence,

$$\mathbb{E}[R^2] = (1)(p) + \mathbb{E}[R^2|A^c](1-p) \quad ; \quad \mathbb{E}[R^2|A^c] = \sum_{k=1} k^2 \Pr[R = k|A^c]$$

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$$\mathbb{E}[R^2] = (1)(p) + \mathbb{E}[R^2|A^c](1-p) \quad ; \quad \mathbb{E}[R^2|A^c] = \sum_{k=1}^{\infty} k^2 \Pr[R = k|A^c]$$

Note that
$$\Pr[R = k | A^c] = \Pr[R = k | \text{ first toss is a tails}] = (1 - p)^{k-2} p = \Pr[R = k - 1]$$

 $\implies \mathbb{E}[R^2 | A^c] = \sum_{k=1} k^2 \Pr[R = k - 1] = \sum_{t=0} (t+1)^2 \Pr[R = t]$ $(t := k - 1)$

Continuing from the previous slide,

$$\mathbb{E}[R^2|A^c] = \sum_{t=0}^{\infty} (t+1)^2 \Pr[R=t] = \sum_{t=0}^{\infty} t^2 \Pr[R=t] + 2 \sum_{t=0}^{\infty} t \Pr[R=t] + \sum_{t=0}^{\infty} \Pr[R=t] + \sum_{t=0}^{\infty} \Pr[R=t] + 2 \sum_{t=1}^{\infty} t \Pr[R=t] + \sum_{t=1}^{\infty} \Pr[R=t] + 2 \sum_{t=1}^{\infty} t \Pr[R=t] + 2 \sum_{t=1}^{$$

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$$= \sum_{t=1}^{\infty} t^2 \Pr[R=t] + 2 \sum_{t=1}^{\infty} t \Pr[R=t] + \sum_{t=1}^{\infty} \Pr[R=t] = \mathbb{E}[R^2] + 2\mathbb{E}[R] + 1$$

Putting everything together,

$$\mathbb{E}[R^2] = (1)(p) + (\mathbb{E}[R^2] + 2\mathbb{E}[R] + 1])(1-p) \implies p \,\mathbb{E}[R^2] = p + 2(1-p)\mathbb{E}[R] + (1-p)\mathbb{E}[1]$$

$$\implies p \,\mathbb{E}[R^2] = p + \frac{2(1-p)}{p} + (1-p) \qquad (\mathbb{E}[R] = \frac{1}{p}, \,\mathbb{E}[1] = 1)$$

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$$\sigma_X := \sqrt{\mathsf{Var}[X]} = \sqrt{\mathbb{E}[X^2] - (\mathbb{E}[X])^2}$$

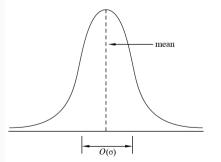
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Standard deviation for a "bell"-shaped distribution indicates how wide the "main part" of the distribution is.

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Proof:

$$Var[aR + b] = \mathbb{E}[(aR + b)^{2}] - (\mathbb{E}[aR + b])^{2} = \mathbb{E}[a^{2}R^{2} + 2abR + b^{2}] - (\mathbb{E}[aR] + \mathbb{E}[b])^{2}$$

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Similarly, for the standard deviation,

$$\sigma_{aR+b} = \sqrt{\operatorname{Var}[aR+b]} = \sqrt{a^2\operatorname{Var}[R]} = |a| \sigma_R$$

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Note the difference from the property of expectation,

$$\mathbb{E}[aR+b]=a\mathbb{E}[R]+b$$

Recall that for r.v's R and S, $\mathbb{E}[R+S] = \mathbb{E}[R] + \mathbb{E}[S]$. In general, such a property is not true for the variance, i.e. variance of a sum is not necessarily equal to the sum of the variances.

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$$Var[R + S] = \mathbb{E}[(R + S)^{2}] - (\mathbb{E}[R + S])^{2} = \mathbb{E}[R^{2} + S^{2} + 2RS] - (\mathbb{E}[R] + \mathbb{E}[S])^{2}$$
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Pairwise Independence: Random variables $R_1, R_2, R_3, \ldots R_n$ are *pairwise* independent if for any pair R_i and R_j , for $x \in \text{Range}(R_i)$ and $y \in \text{Range}(R_j)$, events $\Pr[R_i = x]$ and $\Pr[R_j = y]$ are pairwise independent implying that $\Pr[(R_i = x) \cap (R_j = y)] = \Pr[R_i = x] \Pr[R_j = y]$.

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$$\begin{aligned} \textit{Proof} : \mathsf{Var}[R_1 + R_2 + \dots R_n] &= \mathbb{E}[(R_1 + R_2 + \dots R_n)^2] - (\mathbb{E}[R_1 + R_2 + \dots R_n])^2 \\ &= \sum_{i=1}^n [\mathbb{E}[R_i^2] - (\mathbb{E}[R_i])^2] + 2 \sum_{i,j|1 \leq i < j \leq n} [\mathbb{E}[R_i R_j] - \mathbb{E}[R_i] \, \mathbb{E}[R_j]] \\ &\Longrightarrow \mathsf{Var}[R_1 + R_2 + \dots R_n] &= \sum_{i=1}^n \mathsf{Var}[R_i] \end{aligned} \qquad (\mathsf{Since the r.v's are pairwise independent})$$

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Proof:
$$Var[R_1 + R_2 + \dots R_n] = \mathbb{E}[(R_1 + R_2 + \dots R_n)^2] - (\mathbb{E}[R_1 + R_2 + \dots R_n])^2$$

$$= \sum_{i=1}^n [\mathbb{E}[R_i^2] - (\mathbb{E}[R_i])^2] + 2 \sum_{i,j|1 \le i < j \le n} [\mathbb{E}[R_i R_j] - \mathbb{E}[R_i] \mathbb{E}[R_j]]$$

$$\implies \operatorname{Var}[R_1 + R_2 + \dots R_n] = \sum_{i=1}^n \operatorname{Var}[R_i]$$
 (Since the r.v's are pairwise independent)

Importantly, we do not require the r.v's to be mutually independent. Mutual independence \Rightarrow pairwise independence, but pairwise independence \Rightarrow mutual independence.

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Hence,

$$R = R_1 + R_2 + \ldots + R_n \implies \mathsf{Var}[R] = \mathsf{Var}[R_1 + R_2 + \ldots + R_n]$$

Since R_1, R_2, \ldots, R_n are mutually independent indicator random variables,

$$Var[R] = Var[R_1] + Var[R_2] + \ldots + Var[R_n]$$

Since the variance of an indicator (Bernoulli) r.v. is p(1-p),

$$Var[R] = n p (1 - p).$$



Q: In a class of *n* students, what is the probability that two students share the same birthday? Assume that (i) each student is equally likely to be born on any day of the year, (ii) no leap years and (iii) student birthdays are independent of each other.

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For d := 365 (since no leap years),

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Define M to be the number of pairs of students with matching birthdays. For a fixed ordering of the students, let $X_{i,j}$ be the indicator r.v. corresponding to the event $E_{i,j}$ that the birthdays of students i and j match. Hence,

$$M = \sum_{i,j|1 \le i < j \le n} X_{i,j} \implies \mathbb{E}[M] = \mathbb{E}[\sum_{i,j|1 \le i < j \le n} X_{i,j}] = \sum_{i,j|1 \le i < j \le n} \mathbb{E}[X_{i,j}] = \sum_{i,j|1 \le i < j \le n} \Pr[E_{i,j}]$$
(Linearity of expectation)

For a pair of students i, j, let B_i be the r.v. equal to the day of student i's birthday. Range (B_i) = $\{1, 2, \ldots, d\}$. For all $k \in [d]$, $\Pr[B_i = k] = 1/d$ (each student is equally likely to be born on any day of the year).

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$$E_{i,j} = (B_i = 1 \cap B_j = 1) \cup (B_i = 2 \cap B_j = 2) \cup \dots$$

$$\implies \Pr[E_{i,j}] = \sum_{k=1}^d \Pr[B_i = k \cap B_j = k] = \sum_{k=1}^d \Pr[B_i = k] \Pr[B_j = k] = \sum_{k=1}^d \frac{1}{d^2} = \frac{1}{d}$$
(student birthdays are independent of each other)

$$\implies \mathbb{E}[M] = \sum_{i,j|1 \le i < j \le n} \Pr[E_{i,j}] = \frac{1}{d} \sum_{i,j|1 \le i < j \le n} (1) = \frac{1}{d} [(n-1) + (n-2) + \ldots + 1] = \frac{n(n-1)}{2d}$$

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Hence, in our class of 75 students, on average, there are $\frac{(21)(41)}{365} = 7.60$ students with matching birthdays.

Q: Are the $X_{i,j}$ r.v's mutually independent?

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No, because if
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 and $X_{j,k} = 1$, then, $\Pr[X_{i,k} = 1 | X_{j,k} = 1 \cap X_{i,j} = 1] = 1 \neq \frac{1}{d} = \Pr[X_{i,k} = 1].$

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Yes, because for all i, j and i', j' (where $i \neq i'$), $\Pr[X_{i,j} = 1 | X_{i',j'} = 1] = \Pr[X_{i,j} = 1]$ because if students i' and j' have matching birthdays, it does not tell us anything about whether i and j have matching birthdays.

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$$\mathsf{Var}[M] = \mathsf{Var}[\sum_{i,j | 1 \leq i < j \leq n} X_{i,j}]$$

Since $X_{i,j}$ are pairwise independent, the variance of the sum is equal to the sum of the variance.

$$\implies \mathsf{Var}[M] = \sum_{i,j|1 \le i < j \le n} \mathsf{Var}[X_{i,j}] = \sum_{i,j|1 \le i < j \le n} \frac{1}{d} \left(1 - \frac{1}{d}\right) = \frac{1}{d} \left(1 - \frac{1}{d}\right) \frac{n(n-1)}{2}$$

$$(\mathsf{Since}\ X_{i,j}\ \mathsf{is\ an\ indicator\ (Bernoulli)}\ \mathsf{r.v.\ and\ } \mathsf{Pr}[X_{i,j} = 1] = \frac{1}{d})$$

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Hence, in our class of 75 students, the standard deviation for the matching birthdays is equal to $\sqrt{\frac{(37)(75)}{365}} \frac{364}{365} \approx 2.75$.



For two random variables R and S, the covariance between R and S is defined as:

$$\mathsf{Cov}[R,S] := \mathbb{E}[(R - \mathbb{E}[R]) \, (S - \mathbb{E}[S])] = \mathbb{E}[RS] - \mathbb{E}[R] \, \mathbb{E}[S]$$

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Covariance generalizes the notion of variance to multiple random variables.

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The covariance between two r.v's is symmetric i.e. Cov[R, S] = Cov[S, R].

For two arbitrary (not necessarily independent) r.v's, $\it R$ and $\it S$,

$$\mathsf{Var}[R+S] = \mathsf{Var}[R] + \mathsf{Var}[S] + 2\,\mathsf{Cov}[R,S]$$

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Recall from Lecture 17, Slide 7, where we showed that,

$$\mathsf{Var}[R+S] = \mathsf{Var}[R] + \mathsf{Var}[S] + 2(\mathbb{E}[RS] - \mathbb{E}[R]\,\mathbb{E}[S]) = \mathsf{Var}[R] + \mathsf{Var}[S] + 2\,\mathsf{Cov}[R,S].$$

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Covariance¹

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Generalization to multiple random variables $R_1, R_2, \dots R_n$ (Recall from Lecture 17, Slide 8):

$$\operatorname{Var}\left[\sum_{i=1}^{n} R_{i}\right] = \sum_{i=1}^{n} \operatorname{Var}[R_{i}] + 2 \sum_{1 \leq i < j \leq n} \operatorname{Cov}[R_{i}, R_{j}]$$

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We know that $Cov[X, Y] = \mathbb{E}[XY] - \mathbb{E}[X]\mathbb{E}[Y]$. Note that $X = \mathcal{I}_A$ and $Y = \mathcal{I}_B$. We can conclude that $XY = \mathcal{I}_{A \cap B}$ since XY = 1 iff both events A and B happen.

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$$\implies \mathbb{E}[X] = \Pr[A] ; \mathbb{E}[Y] = \Pr[B]; \mathbb{E}[XY] = \Pr[A \cap B]$$

$$\implies \operatorname{Cov}[X, Y] = \mathbb{E}[XY] - \mathbb{E}[X] \mathbb{E}[Y] = \Pr[A \cap B] - \Pr[A] \Pr[B]$$

If $Cov[X, Y] > 0 \implies Pr[A \cap B] > Pr[A] Pr[B]$. Hence,

$$\Pr[A|B] = \frac{\Pr[A \cap B]}{\Pr[B]} > \frac{\Pr[A]\Pr[B]}{\Pr[B]} = \Pr[A]$$

If Cov[X,Y] > 0, it implies that Pr[A|B] > Pr[A] and hence, the probability that event A happens increases if B is going to happen/has happened. Similarly, if Cov[X,Y] < 0, Pr[A|B] < Pr[A]. In this case, if B happens, then the probability of event A decreases.

The correlation between two r.v's R_1 and R_2 is defined as:

$$\mathsf{Corr}[R_1,R_2] = \frac{\mathsf{Cov}[R_1,R_2]}{\sqrt{\mathsf{Var}[R_1]\,\mathsf{Var}[R_2]}}$$

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If R_1 and R_2 are independent, $Cov[R_1, R_2] = 0$ and $Corr[R_1, R_2] = 0$.

If
$$R_1 = -R_2 = R$$
, then,

$$\begin{aligned} \operatorname{Corr}[R,-R] &= \frac{\operatorname{Cov}[R,-R]}{\sqrt{\operatorname{Var}[R]\operatorname{Var}[-R]}} = \frac{\operatorname{Cov}[R,-R]}{\sqrt{\operatorname{Var}[R](-1)^2\operatorname{Var}[R]}} = \frac{\operatorname{Cov}[R,-R]}{\operatorname{Var}[R]} \\ &= \frac{\mathbb{E}[-R^2] - \mathbb{E}[R]\,\mathbb{E}[-R]}{\operatorname{Var}[R]} = \frac{-\mathbb{E}[R^2] + \mathbb{E}[R]\,\mathbb{E}[R]}{\operatorname{Var}[R]} = \frac{-\operatorname{Var}[R]}{\operatorname{Var}[R]} = -1 \end{aligned}$$

