CMPT 210: Probability and Computing

Lecture 13

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Recap

Random variable: A random "variable" R on a probability space is a total function whose domain is the sample space S. The codomain is denoted by V (usually a subset of the real numbers), meaning that $R: S \to V$.

Example: Suppose we toss three independent, unbiased coins. In this case, $\mathcal{S} = \{HHH, HHT, HTH, HTT, THH, THT, TTH, TTT\}$. C is a random variable equal to the number of heads that appear such that $C: \mathcal{S} \to \{0,1,2,3\}$. C(HHT) = 2. An random variable partitions the sample space into several blocks. For r.v. R, for all $i \in \text{Range}(R)$, the event $[R=i] = \{\omega \in \mathcal{S} | R(\omega) = i\}$. For any r.v. R, $\sum_{i \in \text{Range}(R)} \Pr[R=i] = 1$.

Example: For the above r.v.
$$C$$
, $[C = 2] = \{HHT, HTH, THH\}$ and $Pr[C = 2] = \frac{3}{8}$. $\sum_{i \in Range(C)} Pr[C = i] = Pr[C = 0] + Pr[C = 1] + Pr[C = 2] + Pr[C = 3] = \frac{1}{8} + \frac{3}{8} + \frac{3}{8} + \frac{3}{8} = 1$.

Recap

Indicator Random Variable: An indicator random variable corresponding to an event E is denoted as \mathcal{I}_E and is defined such that for $\omega \in E$, $\mathcal{I}_E[\omega] = 1$ and for $\omega \notin E$, $\mathcal{I}_E[\omega] = 0$.

Example: When throwing two dice, if E is the event that both throws of the dice result in a prime number, then $\mathcal{I}_E((2,4)) = 0$ and $\mathcal{I}_E((2,3)) = 1$.

Probability density function (PDF): Let R be a r.v. with codomain V. The probability density function of R is the function $PDF_R : V \to [0,1]$, such that $PDF_R[x] = Pr[R = x]$ if $x \in Range(R)$ and equal to zero if $x \notin Range(R)$.

Cumulative distribution function (CDF): The cumulative distribution function of R is the function $CDF_R : \mathbb{R} \to [0,1]$, such that $CDF_R[x] = Pr[R \le x]$.

Importantly, neither PDF_R nor CDF_R involves the sample space of an experiment.

Example: If we flip three coins, and C counts the number of heads, then $PDF_C[0] = Pr[C = 0] = \frac{1}{8}$, and $CDF_C[2.3] = Pr[C \le 2.3] = Pr[C = 0] + Pr[C = 1] + Pr[C = 2] = \frac{7}{8}$.

Bernoulli Distribution

Canonical Example: We toss a biased coin such that the probability of getting a heads is p. Let R be the random variable such that R=1 when the coin comes up heads and R=0 if the coin comes up tails. R follows the Bernoulli distribution.

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PDF_R for Bernoulli distribution: $f: \{0,1\} \to [0,1]$ meaning that Bernoulli random variables take values in $\{0,1\}$. It can be fully specified by the "probability of success" (of an experiment) p (probability of getting a heads in the example). Formally, PDF_R is given by:

$$f(1) = p$$
 ; $f(0) = q := 1 - p$.

In the example, Pr[R = 1] = f(1) = p = Pr[event that we get a heads].

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In the example, Pr[R = 1] = f(1) = p = Pr[event that we get a heads].

 CDF_R for Bernoulli distribution: $F: \mathbb{R} \to [0,1]$:

$$F(x) = 0$$
 (for $x < 0$)
= 1 - p (for $0 \le x < 1$)
= 1 (for $x \ge 1$)

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PDF_R for Uniform distribution: $f: V \to [0,1]$ such that for all $v \in V$, f(v) = 1/|v|. In the example, $f(1) = f(2) = \ldots = f(6) = \frac{1}{6}$.

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 CDF_R for Uniform distribution: For n elements in V arranged in increasing order – (v_1, v_2, \ldots, v_n) , the CDF is:

$$F(x) = 0$$
 (for $x < v_1$)
 $= k/n$ (for $v_k \le x < v_{k+1}$)
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Q: If X has a Bernoulli distribution, when is X also uniform? Ans: When p = 1/2

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$$E_k = (A_1 \cap A_2 \dots A_k \cap A_{k+1}^c \cap A_{k+2}^c \cap \dots \cap A_n^c) \cup (A_1^c \cap A_2 \dots A_k \cap A_{k+1} \cap A_{k+2}^c \cap \dots \cap A_n^c) \cup \dots$$

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$$Pr[E_{k}] = Pr[(A_{1} \cap A_{2} \dots A_{k} \cap A_{k+1}^{c} \cap A_{k+2}^{c} \cap \dots \cap A_{n}^{c})] + Pr[A_{1}^{c} \cap A_{2} \dots A_{k} \cap A_{k+1} \cap \dots \cap A_{n}^{c}] + \dots$$

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 (Independence of tosses)

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. For $k \in \{0, 1, ..., n\}$, $f(k) = \binom{n}{k} p^k (1-p)^{n-k}$.

Proof: Let E_k be the event we get k heads. Let A_i be the event we get a heads in toss i.

$$E_{k} = (A_{1} \cap A_{2} \dots A_{k} \cap A_{k+1}^{c} \cap A_{k+2}^{c} \cap \dots \cap A_{n}^{c}) \cup (A_{1}^{c} \cap A_{2} \dots A_{k} \cap A_{k+1} \cap A_{k+2}^{c} \cap \dots \cap A_{n}^{c}) \cup \dots$$

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$$= Pr[A_{1}] Pr[A_{2}] Pr[A_{k}] Pr[A_{k+1}^{c}] Pr[A_{k+2}^{c}] \dots Pr[A_{n}^{c}] + \dots \quad \text{(Independence of tosses)}$$

$$= p^{k} (1 - p)^{n-k} + p^{k} (1 - p)^{n-k} + \dots$$

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Proof: Let E_k be the event we get k heads. Let A_i be the event we get a heads in toss i.

$$\begin{split} E_k &= (A_1 \cap A_2 \dots A_k \cap A_{k+1}^c \cap A_{k+2}^c \cap \dots \cap A_n^c) \cup (A_1^c \cap A_2 \dots A_k \cap A_{k+1} \cap A_{k+2}^c \cap \dots \cap A_n^c) \cup \dots \\ \Pr[E_k] &= \Pr[(A_1 \cap A_2 \dots A_k \cap A_{k+1}^c \cap A_{k+2}^c \cap \dots \cap A_n^c)] + \Pr[A_1^c \cap A_2 \dots A_k \cap A_{k+1} \cap \dots \cap A_n^c)] + \dots \\ &= \Pr[A_1] \Pr[A_2] \Pr[A_k] \Pr[A_{k+1}^c] \Pr[A_{k+2}^c] \dots \Pr[A_n^c] + \dots \\ &= \Pr[A_1] \Pr[A_2] \Pr[A_k] \Pr[A_{k+1}^c] \Pr[A_{k+2}^c] \dots \Pr[A_n^c] + \dots \\ &= \Pr[E_k] = \binom{n}{k} p^k (1-p)^{n-k} \end{split}$$
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(Number of terms = number of ways to choose the k tosses that result in heads = $\binom{n}{k}$)

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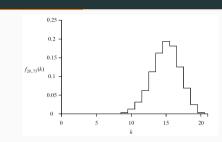
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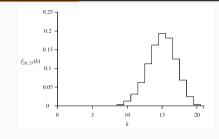
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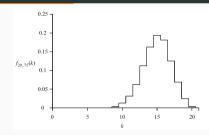


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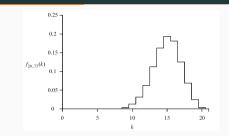
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Q: Prove that $\sum_{k \in \mathsf{Range}(\mathsf{R})} \mathsf{PDF}_R[k] = 1$. By the Binomial Theorem, $\sum_{k \in \mathsf{Range}(\mathsf{R})} \mathsf{PDF}_R[k] = \sum_{k=0}^n \binom{n}{k} p^k (1-p)^{n-k} = (p+1-p)^n = 1$.

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 CDF_R for Binomial distribution: $F: \mathbb{R} \to [0,1]$:

$$F(x) = 0$$

$$= \sum_{i=0}^{k} {n \choose i} p^{i} (1-p)^{n-i}$$

$$= 1.$$
(for $k \le x < k+1$)
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Canonical Example: We toss a biased coin independently multiple times. The probability of getting a heads is p. Let R be the random variable equal to the number of tosses needed to get the first heads. R follows the geometric distribution.

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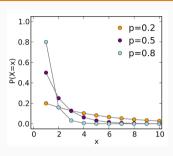
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Q: Prove that $\sum_{k \in \mathsf{Range}(\mathsf{R})} \mathsf{PDF}_R[k] = 1$.

By the sum of geometric series, $\sum_{k \in \mathsf{Range}(R)} \mathsf{PDF}_R[k] = \sum_{k=1}^\infty (1-p)^{k-1} p = \frac{p}{1-(1-p)} = 1$.

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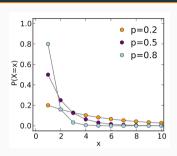


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