# CMPT 210: Probability and Computing

Lecture 15

Sharan Vaswani

March 3, 2023

#### Recap

**Expectation**/mean of a random variable R is denoted by  $\mathbb{E}[R]$  and "summarizes" its distribution. Formally,  $\mathbb{E}[R] := \sum_{\omega \in \mathcal{S}} \Pr[\omega] R[\omega]$ 

Example: When throwing a standard dice, if R is the random variable equal to the number on the dice.  $\mathbb{E}[R] = \sum_{i \in \{1,2,\dots,6\}} \frac{1}{6}[i] = \frac{7}{2}$ .

Alternate definition of expectation:  $\mathbb{E}[R] = \sum_{x \in \mathsf{Range}(R)} x \, \mathsf{Pr}[R = x].$ 

This definition does not depend on the sample space.

Example: If  $\mathcal{I}_A$  is the indicator random variable for event A, then  $\mathbb{E}[\mathcal{I}_A] = \Pr[\mathcal{I}_A = 1](1) + \Pr[\mathcal{I}_A = 0](0) = \Pr[A]$ . For  $\mathcal{I}_A$ , the expectation is equal to the probability that event A happens.

**Linearity of Expectation**: For n random variables  $R_1, R_2, \ldots, R_n$  and constants  $a_1, a_2, \ldots, a_n$ ,  $\mathbb{E}\left[\sum_{i=1}^n a_i R_i\right] = \sum_{i=1}^n a_i \mathbb{E}[R_i]$ .

1

#### Recap

If  $R \sim \text{Bernoulli}(p)$ ,  $\mathbb{E}[R] = p$ . Example: When tossing a coin, if R is the random variable equal to 1 if we get a heads.

If  $R \sim \text{Uniform}(\{v_1, \dots, v_n\})$ ,  $\mathbb{E}[R] = \frac{v_1 + v_2 + \dots + v_n}{n}$ . Example: When throwing an *n*-sided dice with numbers  $v_1, \dots v_n$ , if R is the random variable equal to the number.

If  $R \sim \text{Bin}(n, p)$ ,  $\mathbb{E}[R] = np$ . Example: When tossing n independent coins, if R is the random variable equal to the number of heads.

If  $R \sim \text{Geo}(p)$ ,  $\mathbb{E}[R] = \frac{1}{p}$ . Example: When tossing a coin repeatedly, if R is the random variable equal to the number of tosses required to get the first heads.

### Expectation - Examples - Coupon Collector Problem

 $\mathbf{Q}$ : In a game started by a coffee shop, each time we buy a coffee, we get a coupon. Each coupon has a color (amongst n different colors) and each time, the color of the coupon is selected uniformly at random from amongst the n colors. If we collect at least one coupon of each color, we can claim a free coffee. On average, how many coupons should we collect (coffees we should buy) to claim the prize?

Suppose we get the following sequence of coupons:

$$blue, green, green, red, blue, orange, blue, orange, gray$$

Let us partition this sequence into segments such that a segment ends when we collect a coupon of a new color we did not have before. For this example,

$$\underbrace{\textit{blue}}_{S_1}\underbrace{\textit{green}}_{S_2}\underbrace{\textit{green}, \textit{red}}_{S_3}\underbrace{\textit{blue}, \textit{orange}}_{S_4}\underbrace{\textit{blue}, \textit{orange}, \textit{gray}}_{S_5}$$

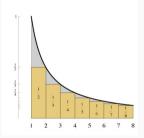
If the number of segments is equal to n, by definition, we will have collected coupons of the n different colors. Define  $X_k$  to be the random variable equal to the length of segment  $S_k$  and T to be the total number of coupons required to have at least one coupon per color.

## **Expectation - Examples - Coupon Collector Problem**

 $T=X_1+X_2+\ldots X_n$ . We wish to compute  $\mathbb{E}[T]$ . By linearity of expectation,  $\mathbb{E}[T]=\mathbb{E}[X_1]+\mathbb{E}[X_2]+\ldots+\mathbb{E}[X_n]$ .

Let us calculate  $\mathbb{E}[X_k]$ . If we are on segment k, we have seen k-1 colors before. Hence, the probability of seeing a new (one that we have not seen before) colored coupon in  $S_k$  is  $\frac{n-(k-1)}{n}$ .  $X_k \sim \text{Geo}\left(\frac{n-(k-1)}{n}\right)$ , and we know that  $\mathbb{E}[X_k] = \frac{n}{n-k+1}$ .

$$\mathbb{E}[T] = \sum_{k=1}^{n} \frac{n}{n-k+1} = n \left[ \frac{1}{n} + \frac{1}{n-1} + \dots + \frac{1}{1} \right]$$
$$\leq n \left[ 1 + \int_{1}^{n} \frac{dx}{x} \right] = n \left[ 1 + \ln(n) \right]$$

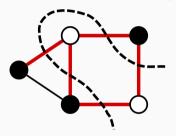


We also know that  $\mathbb{E}[T] \ge n \ln(n+1)$ . Hence,  $\mathbb{E}[T] = O(n \ln(n))$ , meaning that we need to buy  $O(n \ln(n))$  coffees to collect coupons of n colors and get a free coffee.



#### Max Cut

Given a graph  $G = (\mathcal{V}, \mathcal{E})$ , partition the graph's vertices into two complementary sets  $\mathcal{E}$  and  $\mathcal{T}$ , such that the number of edges between the set  $\mathcal{E}$  and the set  $\mathcal{T}$  is as large as possible.



Max Cut has applications to VLSI circuit design.

Equivalently, find a set  $\mathcal{U} \subseteq \mathcal{V}$  of vertices that solve the following

$$\max_{\mathcal{U} \subset \mathcal{V}} |\delta(\mathcal{U})| \text{ where } \delta(\mathcal{U}) := \{(u,v) \in \mathcal{E} | u \in \mathcal{U} \text{ and } v \notin \mathcal{U}\}$$

Here,  $\delta(\mathcal{U})$  is referred to as the "cut" corresponding to the set  $\mathcal{U}$ .

5

#### Max Cut

- ullet Max Cut is NP-hard (Karp, 1972), meaning that there is no polynomial (in  $|\mathcal{E}|$ ) time algorithm that solves Max Cut exactly.
- We want to find an approximate solution  $\mathcal{U}$  such that, if OPT is the size of the optimal cut, then,  $|\delta(\mathcal{U})| \geq \alpha \ OPT$  where  $\alpha \in (0,1)$  is the multiplicative approximation factor.
- Randomized algorithm that guarantees an approximate solution with  $\alpha = \frac{1}{2}$  with probability close to 1 (Erdos, 1967).
- Algorithm with  $\alpha = 0.878$ . (Goemans and Williamson, 1995).
- Under some technical conditions, no efficient algorithm has  $\alpha > 0.878$  (Khot et al, 2004).

We will use Erdos' randomized algorithm and first prove the result in expectation. We wish to prove that for  $\mathcal{U}$  returned by Erdos' algorithm,

$$\mathbb{E}[|\delta(\mathcal{U})|] \geq \frac{1}{2}\mathit{OPT}$$

**Algorithm**: Select  $\mathcal{U}$  to be a random subset of  $\mathcal{V}$  i.e. for each vertex v, choose v to be in the set  $\mathcal{U}$  independently with probability  $\frac{1}{2}$  (do not even look at the edges!).

#### Max Cut

**Claim**: For Erdos' algorithm,  $\mathbb{E}[|\delta(\mathcal{U})|] \geq \frac{1}{2}OPT$ .

**Proof**: For each edge  $(u, v) \in \mathcal{E}$ , let  $X_{u,v}$  be the indicator random variable equal to 1 iff the event  $E_{u,v} = \{(u, v) \in \delta(\mathcal{U})\}$  happens.

$$\mathbb{E}[|\delta(\mathcal{U})|] = \mathbb{E}\left[\sum_{(u,v)\in\mathcal{E}} X_{u,v}\right] = \sum_{(u,v)\in\mathcal{E}} \mathbb{E}\left[X_{u,v}\right] = \sum_{(u,v)\in\mathcal{E}} \Pr[E_{u,v}]$$
(Linearity of expectation, and Expectation of indicator r.v's.)

$$\begin{split} \Pr[E_{u,v}] &= \Pr[(u,v) \in \delta(\mathcal{U})] = \Pr\left[(u \in \mathcal{U} \cap v \notin \mathcal{U}) \cup (u \notin \mathcal{U} \cap v \in \mathcal{U})\right] \\ &= \Pr\left[(u \in \mathcal{U} \cap v \notin \mathcal{U})\right] + \Pr\left[(u \notin \mathcal{U} \cap v \in \mathcal{U})\right] \quad \text{(Union rule for mutually exclusive events)} \\ \Pr[E_{u,v}] &= \Pr[u \in \mathcal{U}] \, \Pr[v \notin \mathcal{U}] + \Pr[u \notin \mathcal{U}] \, \Pr[v \in \mathcal{U}] = \frac{1}{2} \, \frac{1}{2} + \frac{1}{2} \, \frac{1}{2} = \frac{1}{2}. \\ &\qquad \qquad \text{(Independent events)} \end{split}$$

$$\implies \mathbb{E}[|\delta(\mathcal{U})|] = \sum_{(u,v)\in\mathcal{E}} \Pr[E_{u,v}] = \frac{|\mathcal{E}|}{2} \ge \frac{\mathsf{OPT}}{2}.$$



### **Conditional Expectation**

Similar to probabilities, expectations can be conditioned on some event.

For random variable R, the expected value of R conditioned on an event A is given by:

$$\mathbb{E}[R|A] = \sum_{x \in \mathsf{Range}(R)} x \; \mathsf{Pr}[R = x|A]$$

 $\mathbf{Q}$ : If we throw a standard dice and define R to be the random variable equal to the number that comes up, what is the expected value of R given that the number is at most 4?

Let A be the event that the number is at most 4.

$$\begin{aligned} \Pr[R = 1|A] &= \frac{\Pr[(R = 1) \cap A]}{\Pr[A]} = \frac{\Pr[R = 1]}{\Pr[A]} = \frac{1/6}{4/6} = 1/4. \\ \Pr[R = 2|A] &= \Pr[R = 3|A] = \Pr[R = 4|A] = \frac{1}{4} \text{ and } \Pr[R = 5|A] = \Pr[R = 6|A] = 0. \end{aligned}$$

$$\mathbb{E}[R|A] = \sum_{x \in \{1,2,3,4\}} x \Pr[R = x|A] = \frac{1}{4}[1+2+3+4] = \frac{5}{2}.$$

Q: What is the expected value of R given that the number is at least 4?

### Law of Total Expectation

If R is a random variable  $S \to V$  and events  $A_1, A_2, \ldots A_n$  form a partition of the sample space i.e. for all  $i, j, A_i \cap A_j = \emptyset$  and  $A_1 \cup A_2 \cup \ldots \cup A_n = S$ , then,

$$\mathbb{E}[R] = \sum_{i} \mathbb{E}[R|A_{i}] \, \mathsf{Pr}[A_{i}] \, .$$

Proof:

$$\mathbb{E}[R] = \sum_{x \in \mathsf{Range}(R)} x \, \mathsf{Pr}[R = x] = \sum_{x \in \mathsf{Range}(R)} x \, \sum_{i} \mathsf{Pr}[R = x|A_{i}] \, \mathsf{Pr}[A_{i}]$$

$$= \sum_{i} \mathsf{Pr}[A_{i}] \sum_{x \in \mathsf{Range}(R)} x \, \mathsf{Pr}[R = x|A_{i}]$$

$$\implies \mathbb{E}[R] = \sum_{i} \mathsf{Pr}[A_{i}] \, \mathbb{E}[R|A_{i}].$$

#### Conditional Expectation - Examples

Q: Suppose that 49.6% of the people in the world are male and the rest female. If the expected height of a randomly chosen male is 5 feet 11 inches, while the expected height of a randomly chosen female is 5 feet 5 inches, what is the expected height of a randomly chosen person?

Define H to be the random variable equal to the height (in feet) of a randomly chosen person. Define M to be the event that the person is male and F the event that the person is female. We wish to compute  $\mathbb{E}[H]$  and we know that  $\mathbb{E}[H|M] = 5 + \frac{11}{12}$  and  $\mathbb{E}[H|F] = 5 + \frac{5}{12}$ . Pr[M] = 0.496 and Pr[F] = 1 - 0.496 = 0.504. Hence,  $\mathbb{E}[H] = \mathbb{E}[H|M] \Pr[M] + \mathbb{E}[H|F] \Pr[F] = \frac{71}{13}(0.496) + \frac{65}{12}(0.504)$ .

10



### Randomized Quick Select

Given an array A of n distinct numbers, return the  $k^{th}$  smallest element in A for  $k \in [1, n]$ .

#### Algorithm Randomized Quick Select

- 1: function QuickSelect(A, k)
- 2: If Length(A) = 1, return A[1].
- 3: Select  $p \in A$  uniformly at random.
- 4: Construct sets Left :=  $\{x \in A | x < p\}$  and Right :=  $\{x \in A | x > p\}$ .
- 5: r = |Left| + 1 {Element p is the  $r^{th}$  smallest element in A.}
- 6: if k = r then
- 7: return p
- 8: else if k < r then
- 9: QuickSelect(Left, k)
- 10: **else**
- 11: QuickSelect(Right, k r)
- 12: end if

#### Randomized QuickSelect

If  $A = \{2, 7, 0, 1, 3\}$  and we wish to find the  $2^{nd}$  smallest element meaning that k = 2. According to the algorithm,  $p \sim \text{Uniform}(A)$ . Say p = 3.

Then after step 1, Left =  $\{2,0,1\}$  and Right =  $\{7\}$ . r := |Left| + 1 = 3 + 1 = 4. Since r > k, we recurse on the left-hand side by calling the algorithm on  $\{2,0,1\}$  with k=2.

 $p \sim \text{Uniform}(\{2,0,1\})$ . Say p=1. After step 2, Left  $= \{0\}$  and Right  $= \{2\}$ . r := |Left| + 1 = 1 + 1 = 2. Since r = k, we terminate the recursion and return p=1 as the second-smallest element in A.

Q: Run the algorithm if p = 0 in the first step?

Q: Run the algorithm if p = 1 in the first step?

#### Randomized Quick Select – Analysis

**Alternate way**: Sort the elements in A and return the  $k^{th}$  element in the sorted list. Uses  $O(n \log(n))$  comparisons.

Q: Can Randomized Quick Select do better – what is the maximum number of comparisons required by Randomized Quick Select in the worst-case?

In the worst case, Randomized Quick Select is worse than the naive strategy of sorting and returning the  $k^{th}$  element. What about the average (over the pivot selection) case?

**Claim**: For any array A with n distinct elements, and for any  $k \in [n]$ , Randomized Quick Select performs fewer than 8n comparisons in expectation.

In order to prove this claim, we will need to prove the following lemma.

### Randomized Quick Select – Analysis

**Lemma**: The child sub-problem's array (either Left or Right) after the partitioning (in Line 4 of the algorithm) has expected size smaller than  $\frac{7n}{8}$ .

*Proof*: Let us define a "good" event  $\mathcal{E}$  that the randomly chosen pivot splits the array roughly in half. Formally, if n is the length of the array, then  $\mathcal{E}$  is the event that  $r \in \left(\frac{n}{4}, \frac{3n}{4}\right]$  (for simplicity, let us assume that n is divisible by 4.) Since r is chosen randomly,  $\Pr[\mathcal{E}] = \frac{3n/4 - n/4}{n} = \frac{1}{2}$ .

Recall that |Left| = r - 1 and |Right| = n - r. Hence if event  $\mathcal E$  happens, then  $|\text{Left}| < \frac{3n}{4}$  and  $|\text{Right}| < \frac{3n}{4}$ . Hence,  $|\text{Child}| < \frac{3n}{4}$ . If event  $\mathcal E$  does not happen, in the worst-case, |Child| < n. By using the law of total expectation,

$$\begin{split} \mathbb{E}[|\mathsf{Child}|] &= \mathbb{E}[|\mathsf{Child}|\,|\mathcal{E}]\,\mathsf{Pr}[\mathcal{E}] + \mathbb{E}[|\mathsf{Child}|\,|\mathcal{E}^c]\,\mathsf{Pr}[\mathcal{E}^c] \\ &< \frac{3n}{4}\frac{1}{2} + (n)\frac{1}{2} = \frac{7n}{8}. \end{split}$$

Hence on average, the size of the child sub-problem is smaller than  $\frac{7n}{8}$ , proving the lemma.

# Randomized Quick Select - Analysis

In order to upper-bound the total number of comparisons, we use the Lemma with an induction on n. Recall that we need to prove that Randomized Quick Select requires fewer than 8n comparisons in expectation.

**Base case**: If n = 1, then we require 0 < 8 comparisons. Hence the base case is satisfied.

**Inductive Step**: Assume that for m < n,

 $\mathbb{E}[\mathsf{Total} \ \mathsf{number} \ \mathsf{of} \ \mathsf{comparisons} \ \mathsf{for} \ \mathsf{size} \ \mathit{m} \ \mathsf{array}] < 8 \ \mathit{m}.$ 

 $\mathbb{E}[\text{Total number of comparisons for size } n \text{ array}]$ 

$$\mathbb{E}[(n-1)+\mathsf{Total}$$
 number of comparisons in child sub-problem]

$$=(n-1)+\mathbb{E}[\text{Total number of comparisons in child sub-problem}]$$
 (Linearity of expectation)

$$<(n-1)+8\mathbb{E}[|\mathsf{Child}|]$$
 (Induction hypothesis)

$$<(n-1)+8\frac{7n}{8}<8n.$$
 (Lemma)

Hence, for any  $k \in [n]$ , on average, Randomized Quick Select requires fewer than 8n comparisons, even though it might require  $O(n^2)$  comparisons in the worst-case.

